

# Introduction to Bayesian Inference & MCMC in RevBayes



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**TRACY HEATH**

Ecology, Evolution, & Organismal Biology  
Iowa State University

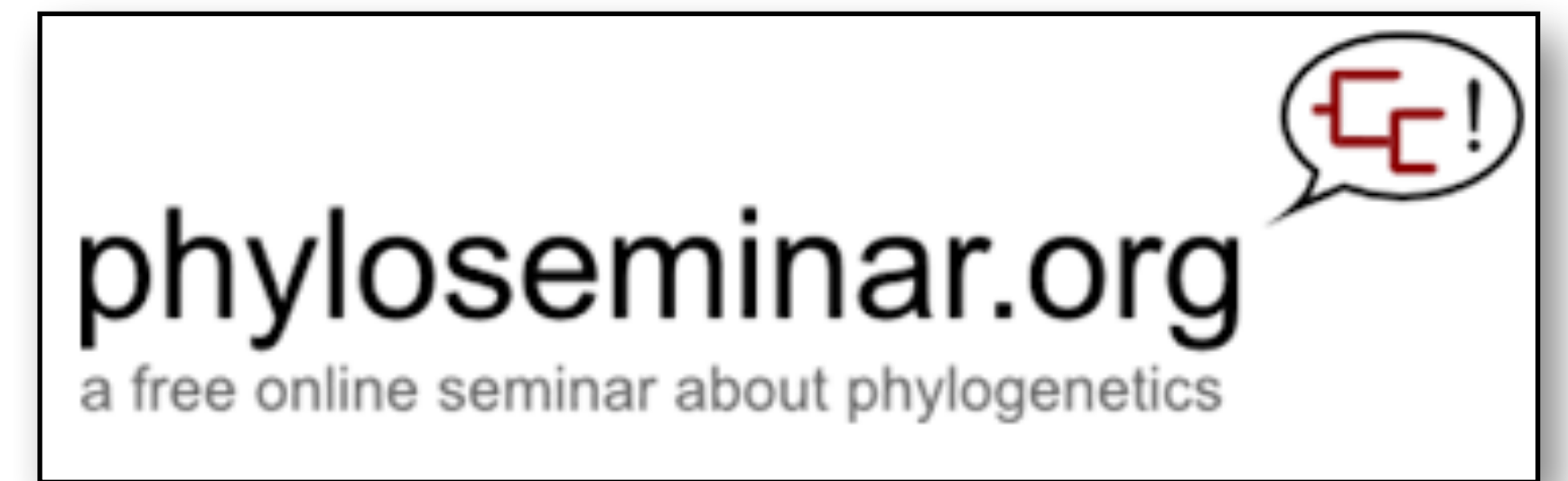




# Video Lectures

## Paul Lewis's Primer on Phylogenetics

- Trees & Likelihood 
- Substitution Models 
- Bayesian Statistics & MCMC 
- Bayesian Phylogenetics 

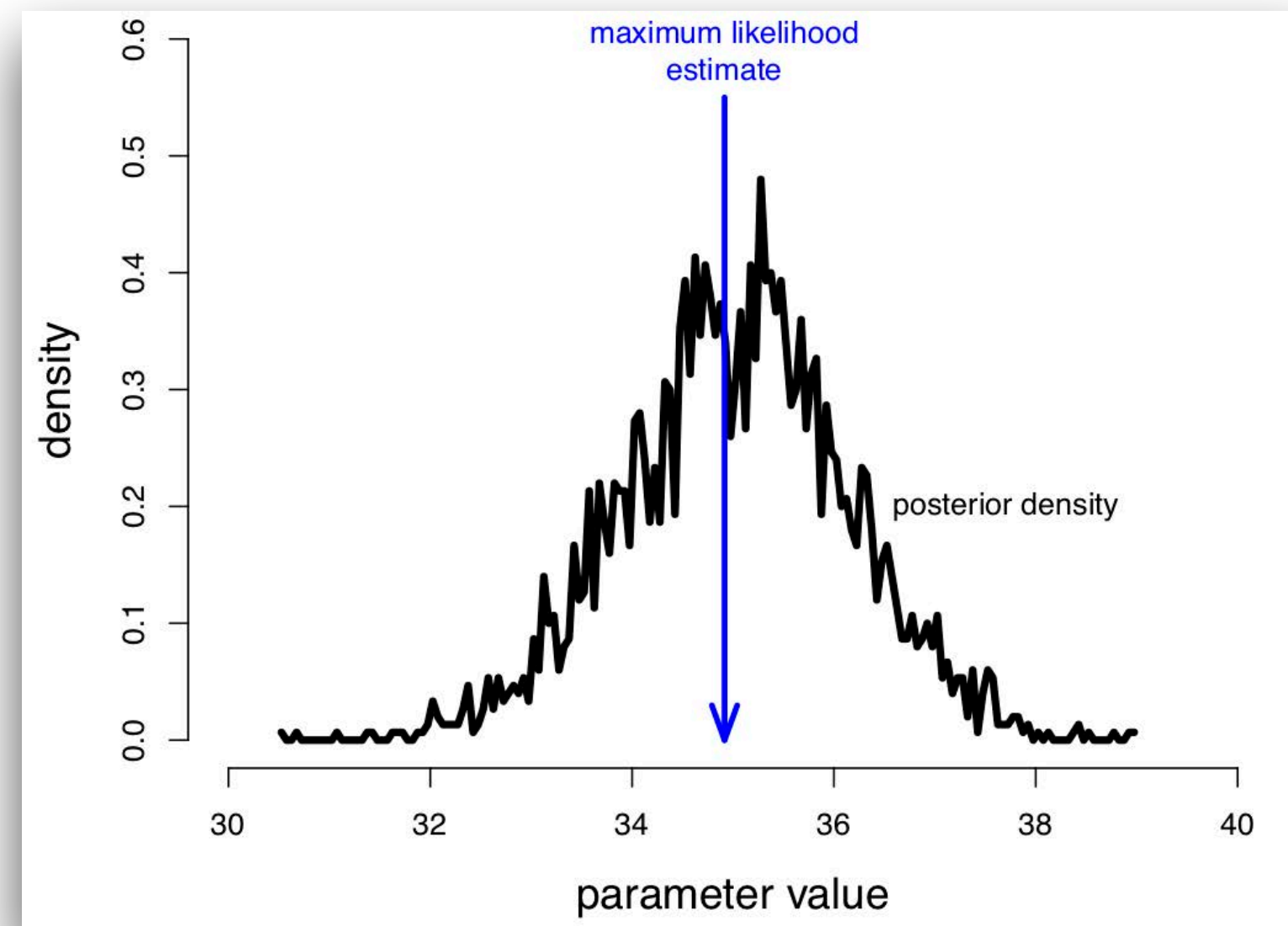




# Bayesian or Maximum Likelihood

- estimates  $\Pr(\theta | \mathbf{X})$
- estimates a **distribution**
- parameters are random variables
- average over nuisance parameters

Bayesian



- estimates  $\Pr(\mathbf{X} | \theta)$
- **point** estimate
- parameters are fixed/unknown
- optimize nuisance parameters

Maximum Likelihood



# Bayes Rule

POSTERIOR PROBABILITY

LIKELIHOOD

PRIOR PROBABILITY

$$\Pr(\theta \mid D) = \frac{\Pr(D \mid \theta) \Pr(\theta)}{\sum_{\theta} \Pr(D \mid \theta) \Pr(\theta)}$$

MARGINAL PROBABILITY OF THE DATA

The diagram illustrates Bayes Rule with the following components and labels:

- POSTERIOR PROBABILITY:** Points to the term  $\Pr(\theta \mid D)$  in the numerator of the left side of the equation.
- LIKELIHOOD:** Points to the term  $\Pr(D \mid \theta)$  in the numerator of the right side of the equation.
- PRIOR PROBABILITY:** Points to the term  $\Pr(\theta)$  in the numerator of the right side of the equation.
- MARGINAL PROBABILITY OF THE DATA:** Points to the denominator term  $\sum_{\theta} \Pr(D \mid \theta) \Pr(\theta)$ .



# Bayesian Inference

Estimate the probability of a hypothesis (model) conditional on observed data

The probability represents a **researcher's degree of belief**

Bayes Rule (also called Bayes Theorem) specifies the conditional probability of the hypothesis given the data



# Bayes Rule

the posterior probability of a discrete parameter  $\delta$  conditional on the data  $D$  is

$$\Pr(\delta \mid D) = \frac{\Pr(D \mid \delta) \Pr(\delta)}{\sum_{\delta} \Pr(D \mid \delta) \Pr(\delta)}$$



the likelihood marginalized over all possible values of  $\delta$



# Bayes Rule

the posterior probability of a discrete parameter  $\theta$  conditional on the data  $D$  is

$$f(\theta \mid D) = \frac{f(D \mid \theta)f(\theta)}{\int_{\theta} f(D \mid \theta)f(\theta)}$$



the likelihood marginalized over all possible values of  $\theta$



# Priors

Prior distributions are an important part of Bayesian statistics

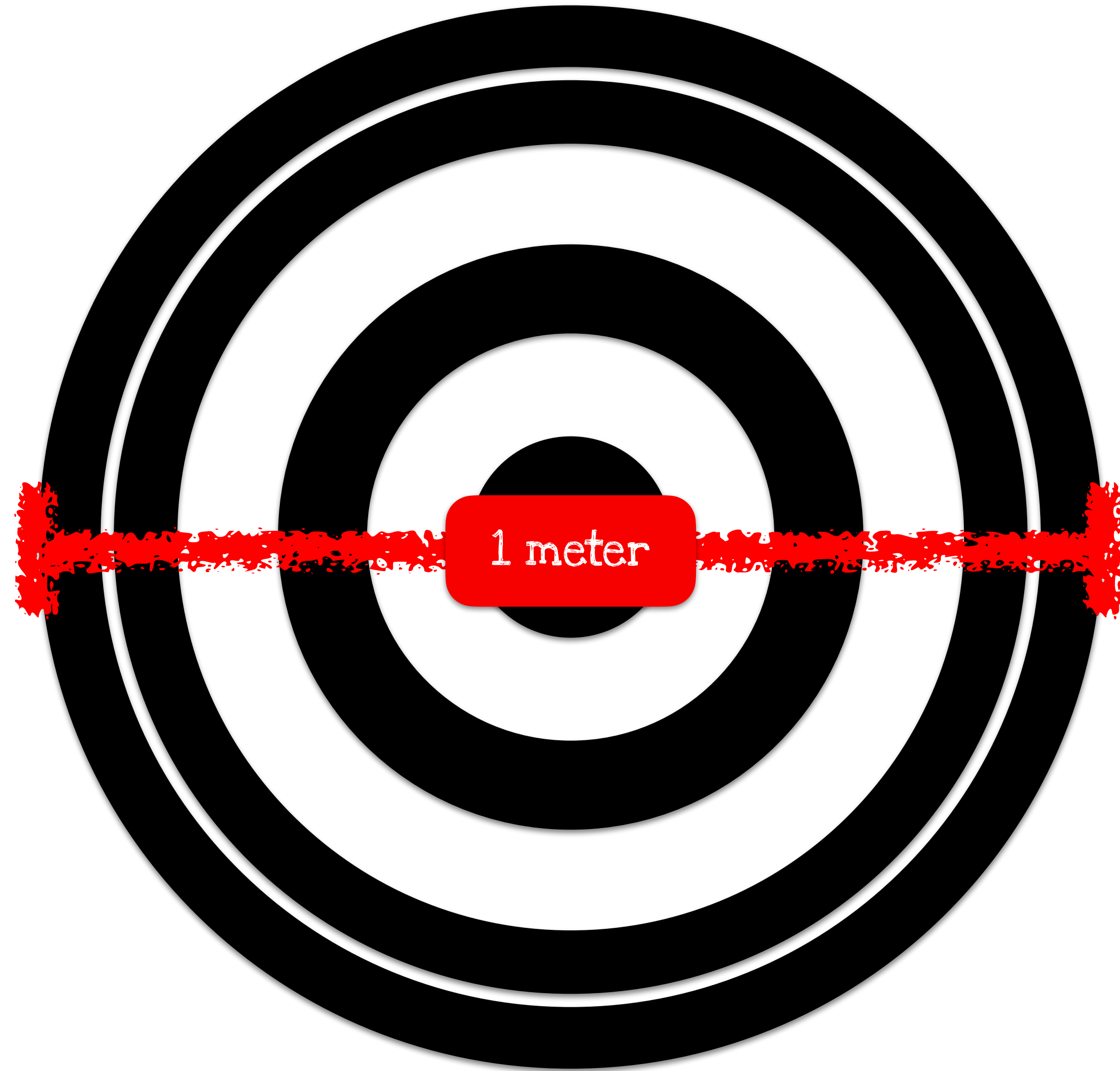
The distribution of  $\theta$  before any data are collected is the prior

$$f(\theta)$$

The prior describes your uncertainty in the parameters of your model



# Priors: Archery





# Priors: Archery



In this example we want to assess an archer's accuracy at hitting the bullseye

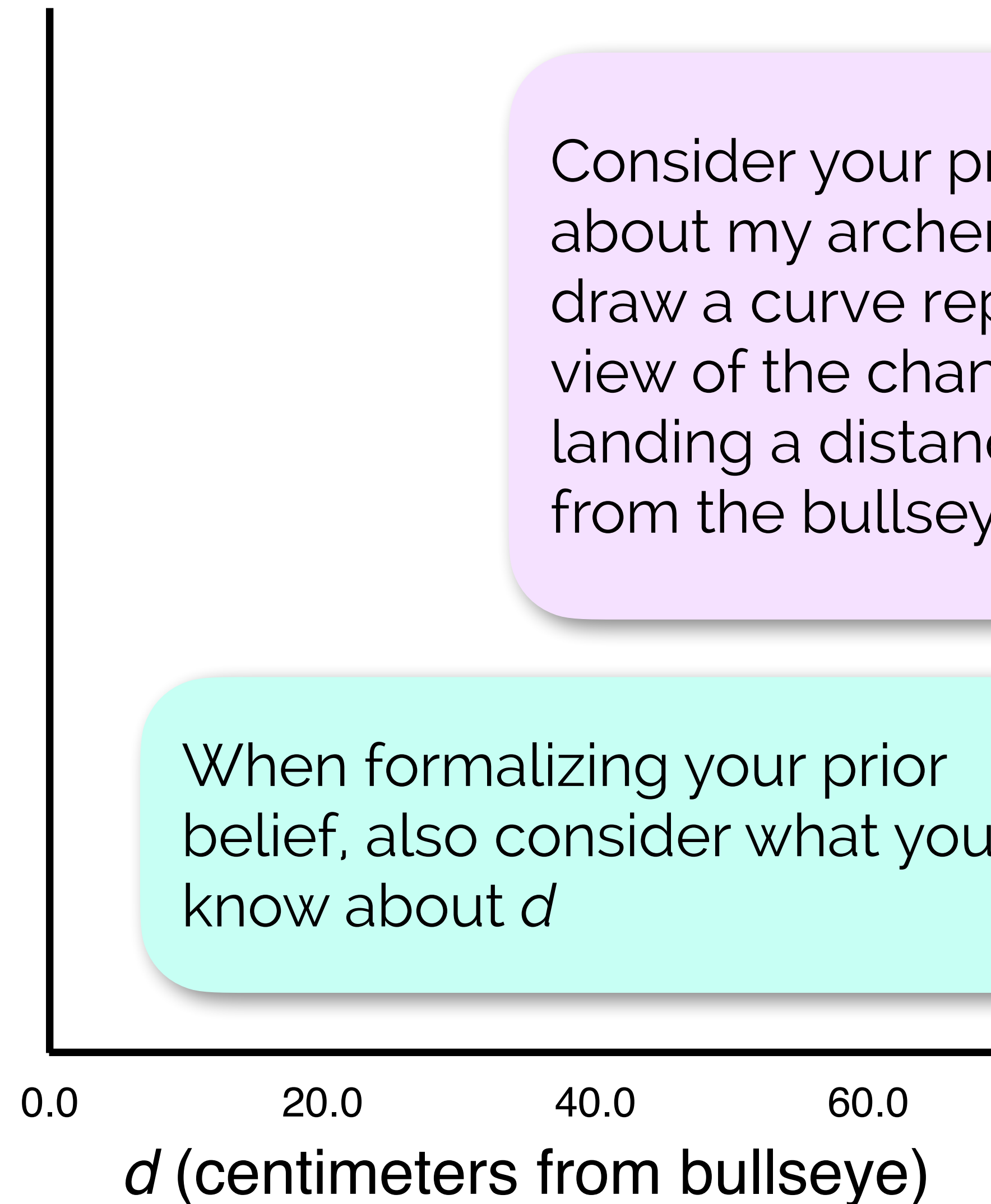
To quantify this, we will measure the distance  $d$  from the center of the target (in centimeters)

$d$  is an absolute value





# Priors: Archery

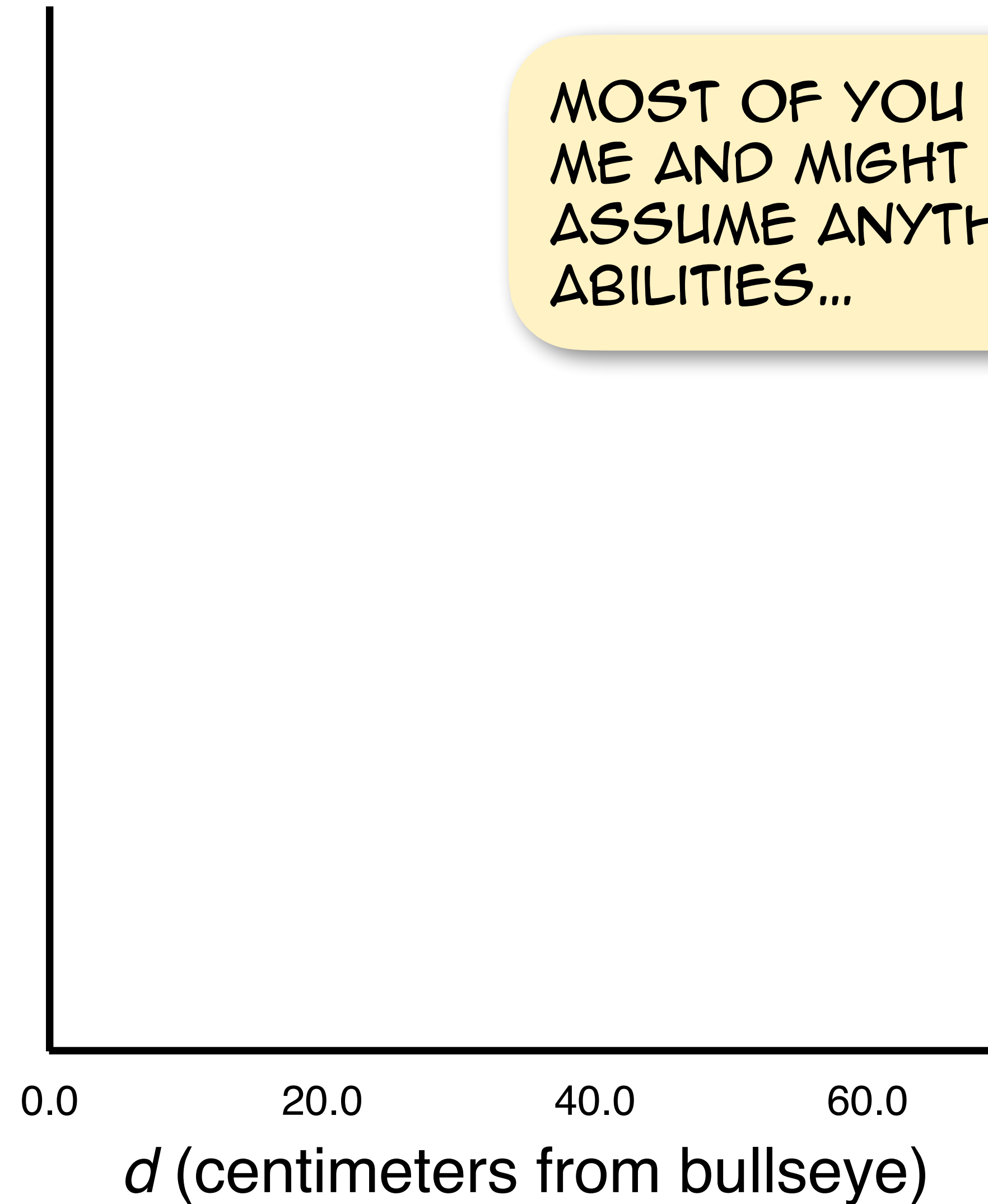




# Priors: Archery



MOST OF YOU DON'T KNOW  
ME AND MIGHT NOT WANT TO  
ASSUME ANYTHING ABOUT MY  
ABILITIES...

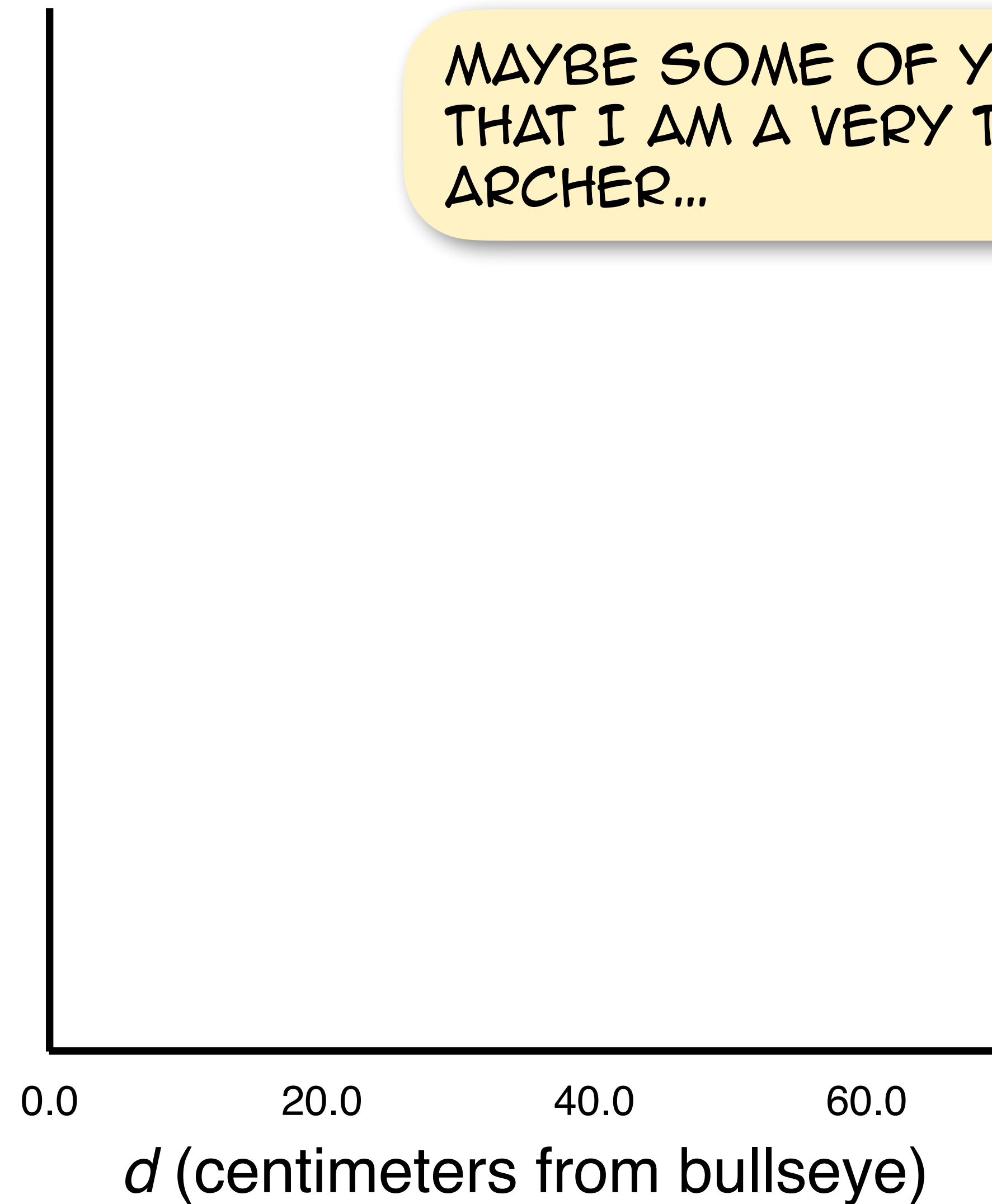




# Priors: Archery



MAYBE SOME OF YOU ASSUME  
THAT I AM A VERY TALENTED  
ARCHER...





# Priors: Archery

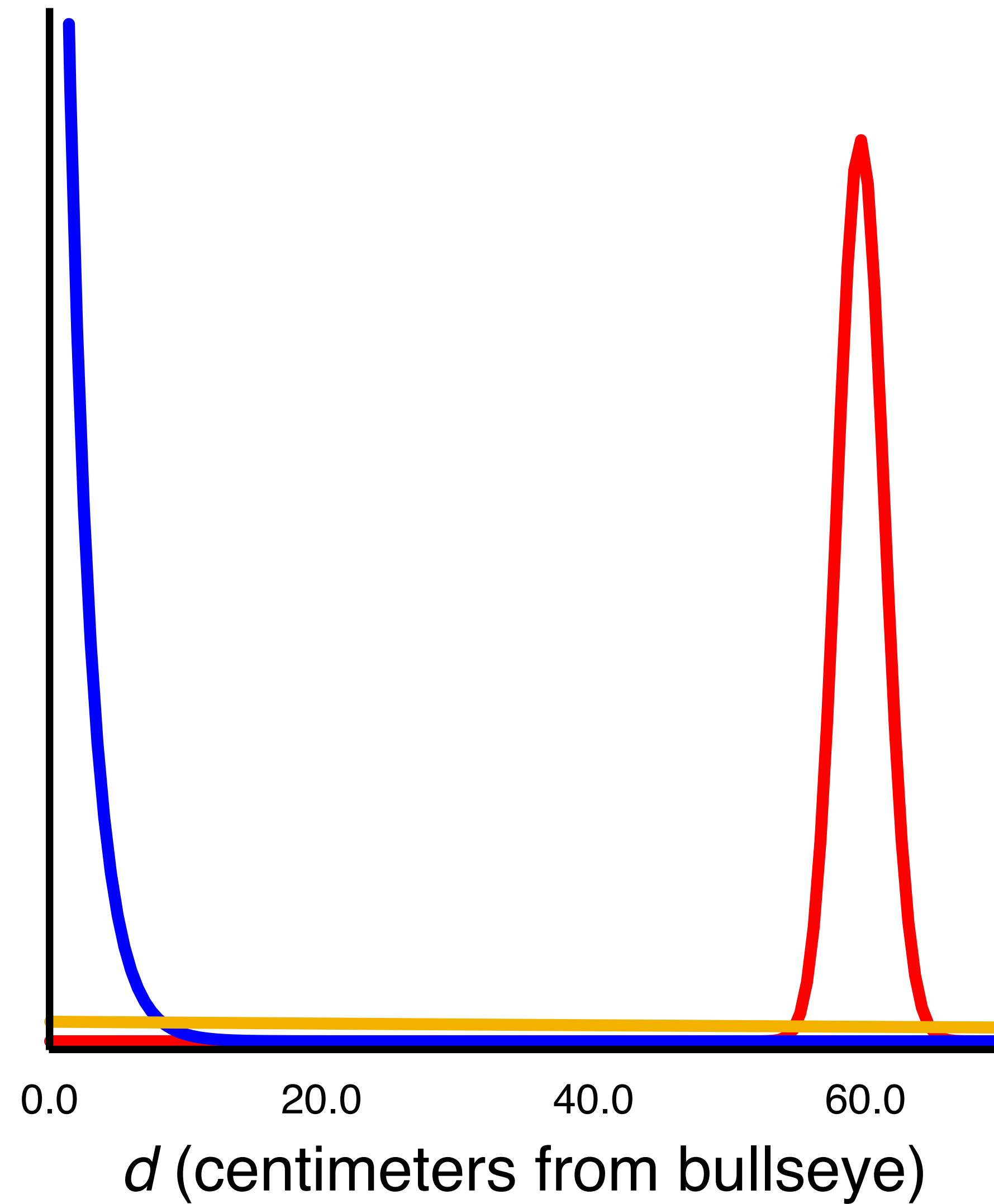


MAYBE SOME OF YOU THINK I MIGHT  
BE A TALENTED ARCHER AND THERE IS  
SOMETHING WRONG WITH MY BOW...



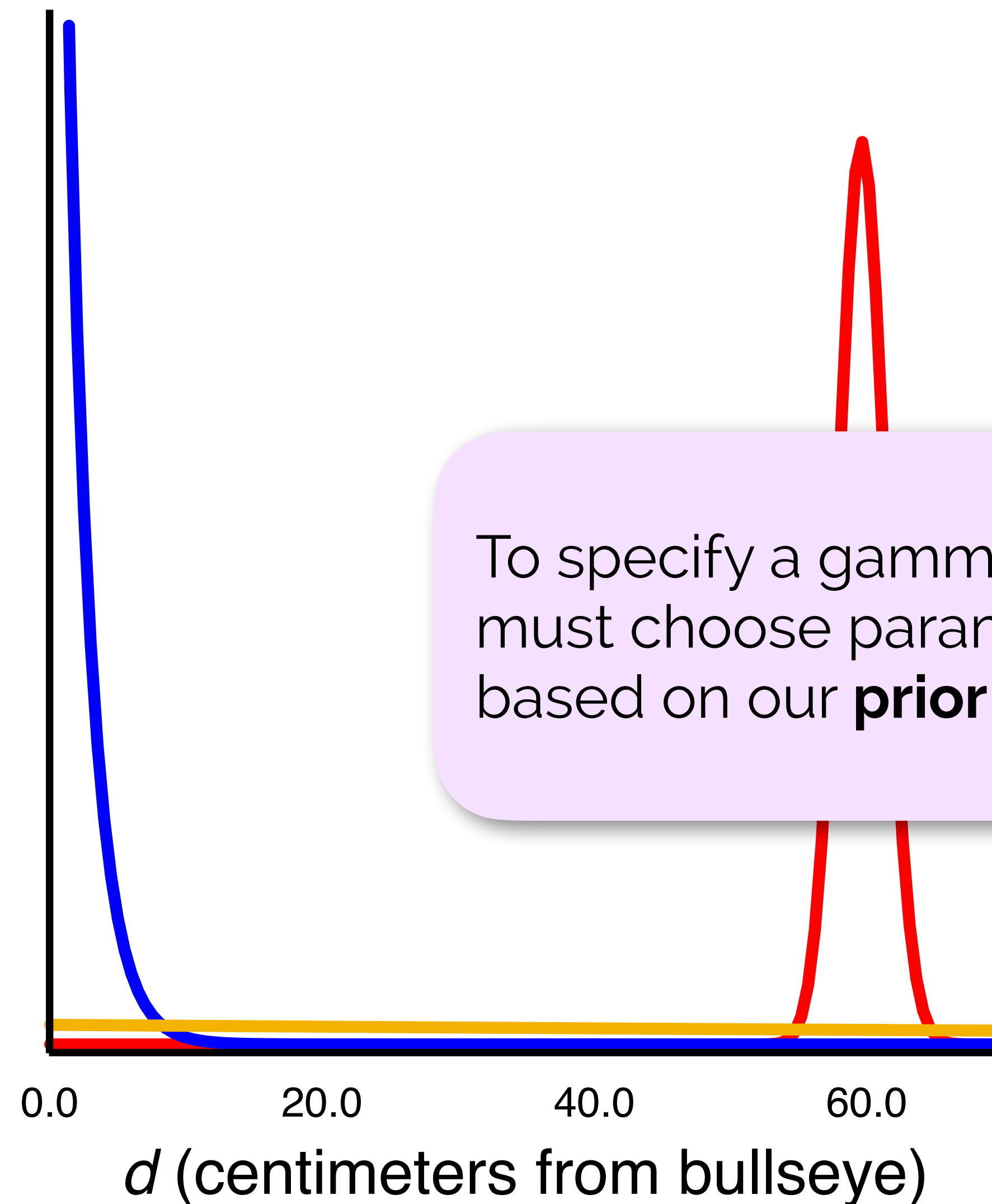


# Priors: Archery





# Priors: Archery



Each of these prior densities can be defined using a gamma distribution.

$$d \sim \text{Gamma}(\alpha, \beta)$$

$$f(d \mid \alpha, \beta) = \frac{1}{\Gamma(\alpha)\beta^\alpha} d^{\alpha-1} e^{-\frac{d}{\beta}}$$



# Priors: Archery

Let's assume that I will consistently miss the target

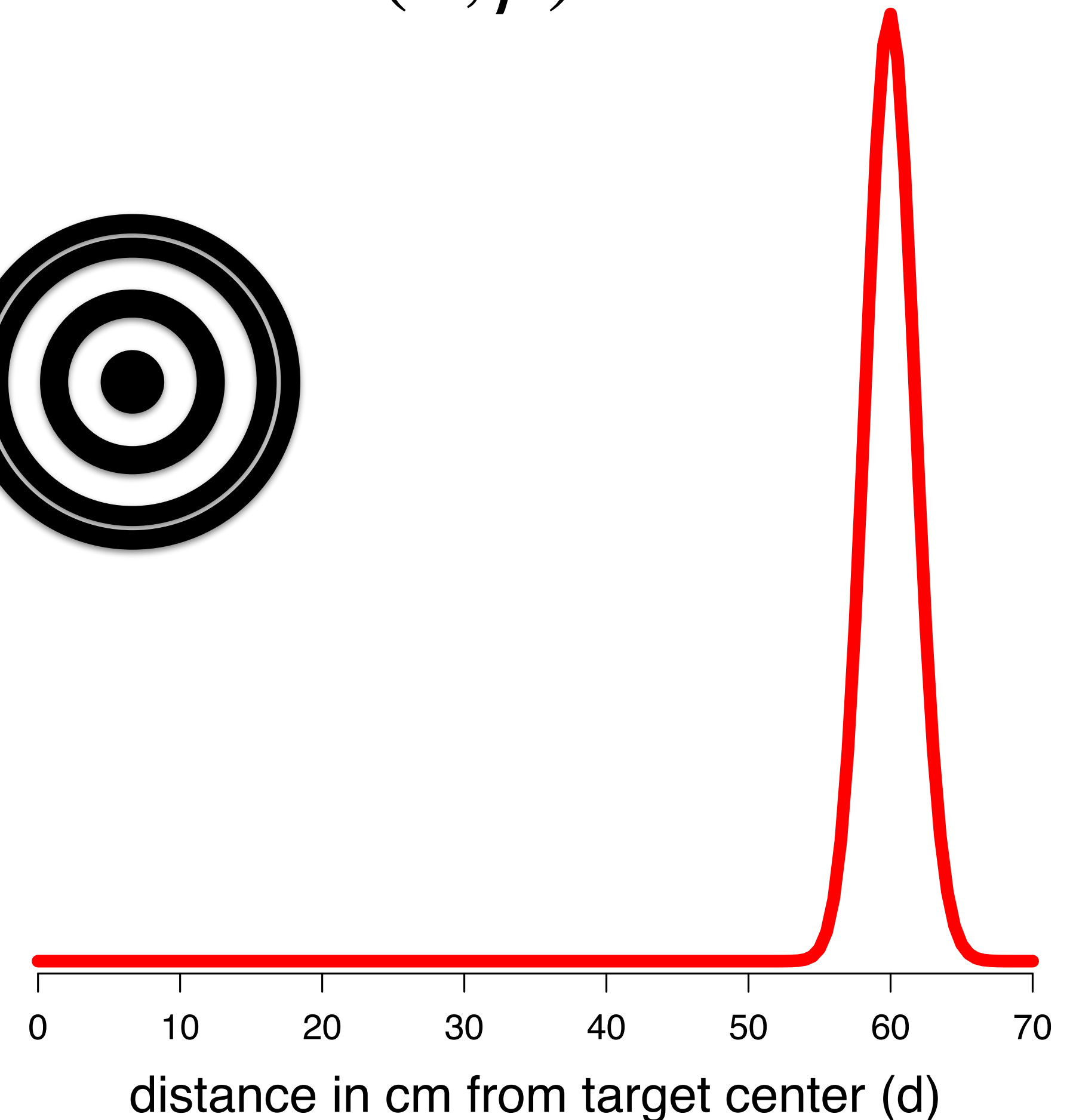
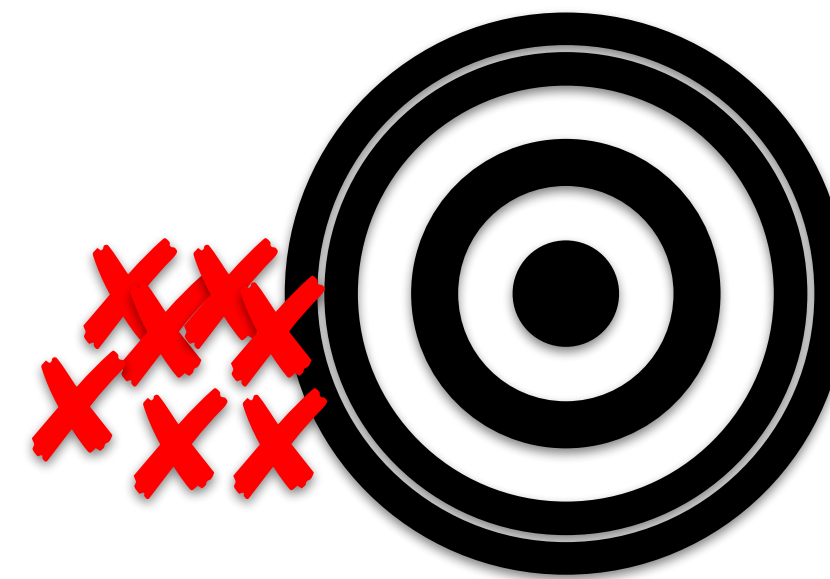
This is a gamma distribution with a mean ( $m$ ) of 60 and a variance ( $v$ ) of 3

**mean** = accuracy

**variance** = precision



$$d \sim \text{Gamma}(\alpha, \beta)$$





# Priors: Archery

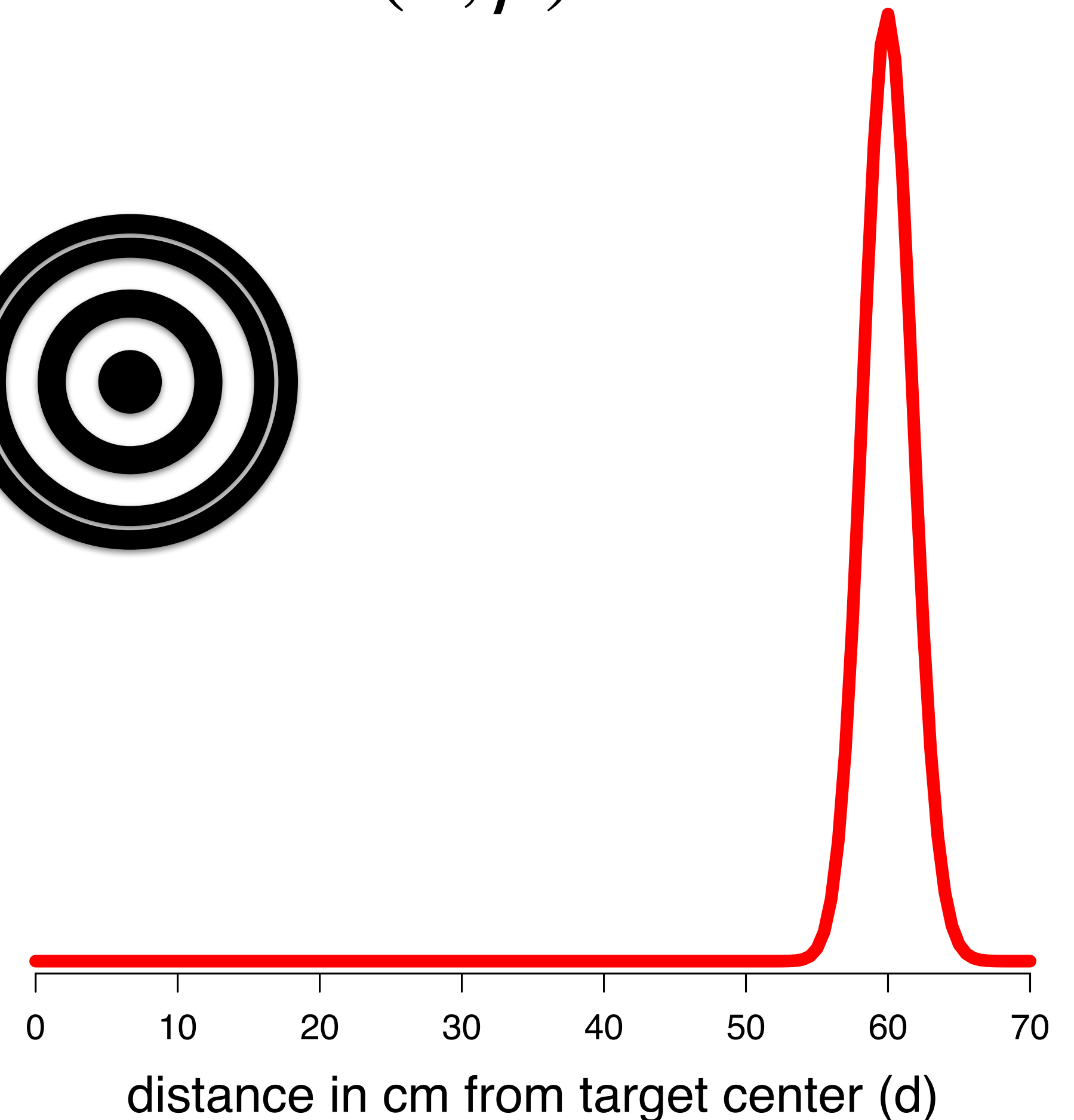
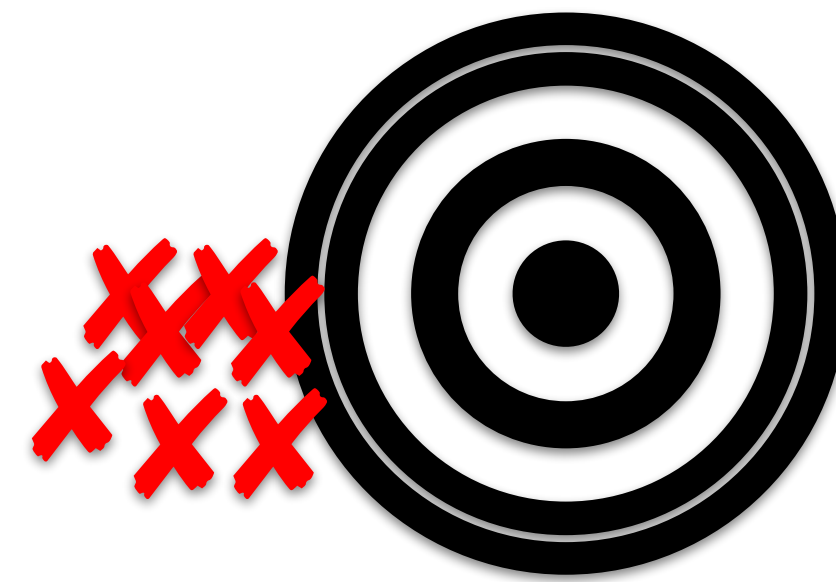
If we have prior knowledge of the mean and variance of the gamma distribution, we can compute the shape and rate parameters

$$m = \frac{\alpha}{\beta}, \alpha = \frac{m^2}{v}$$

$$v = \frac{\alpha}{\beta^2}, \beta = \frac{m}{v}$$



$$d \sim \text{Gamma}(\alpha, \beta)$$





# Priors: Archery

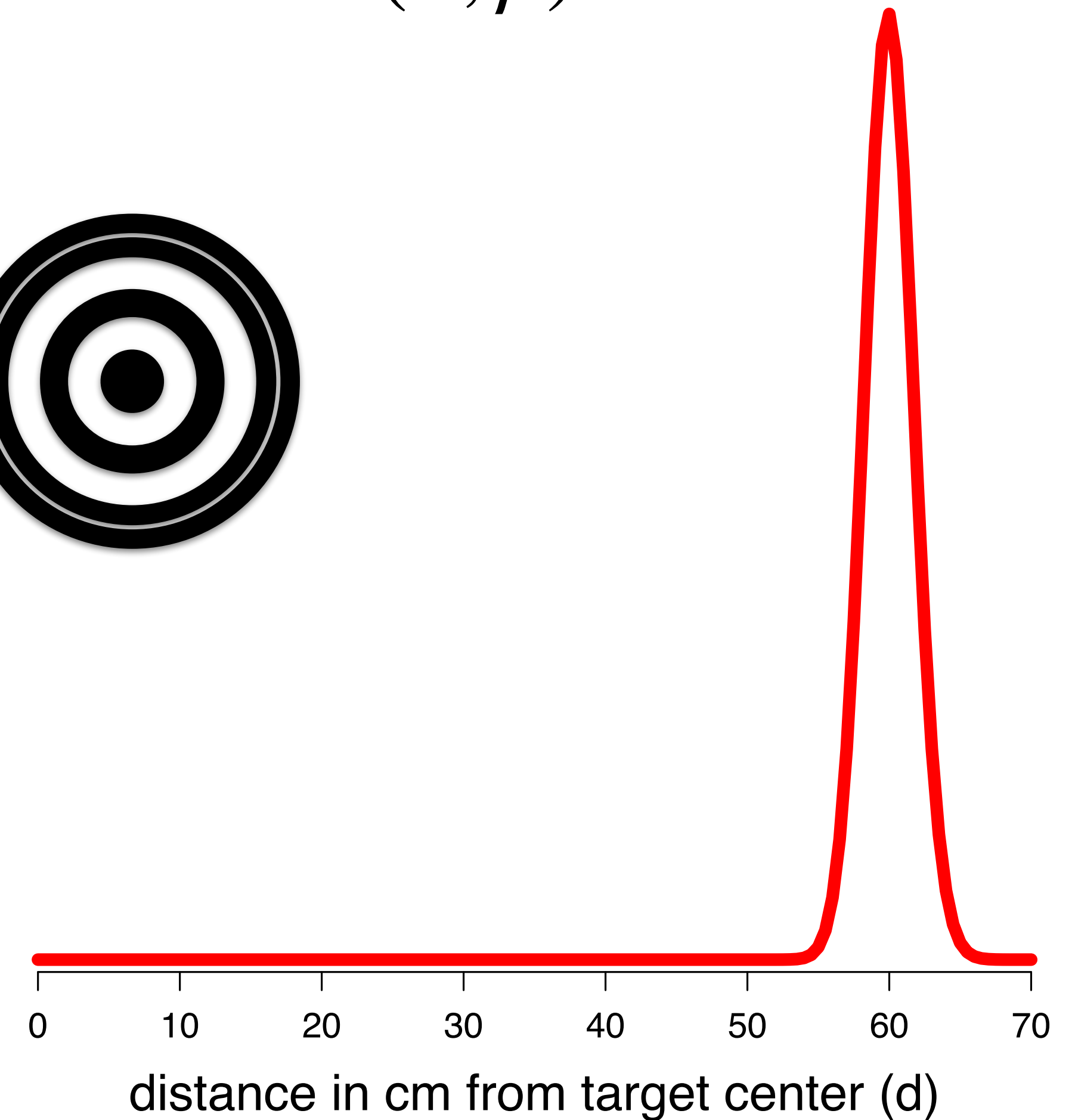
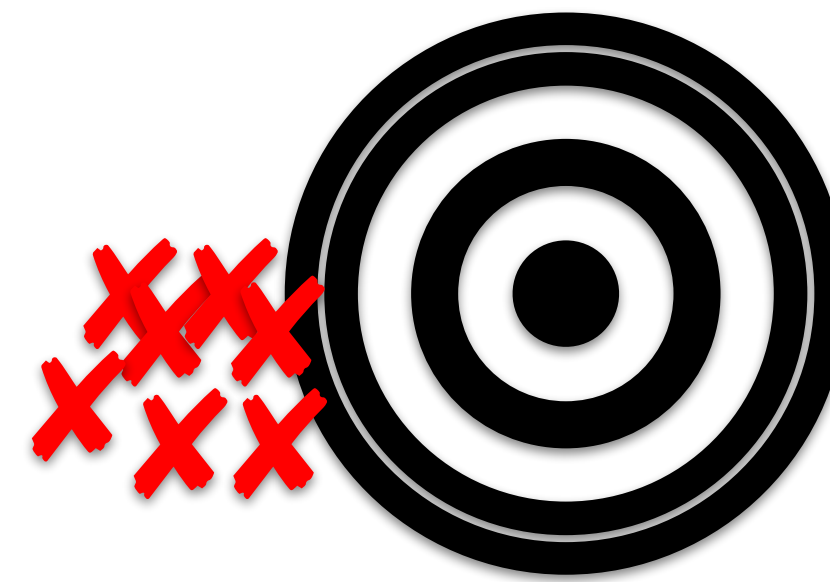
$$m = 60, \quad v = 3$$

$$\alpha = \frac{60^2}{3} = 1200$$

$$\beta = \frac{60}{3} = 20$$



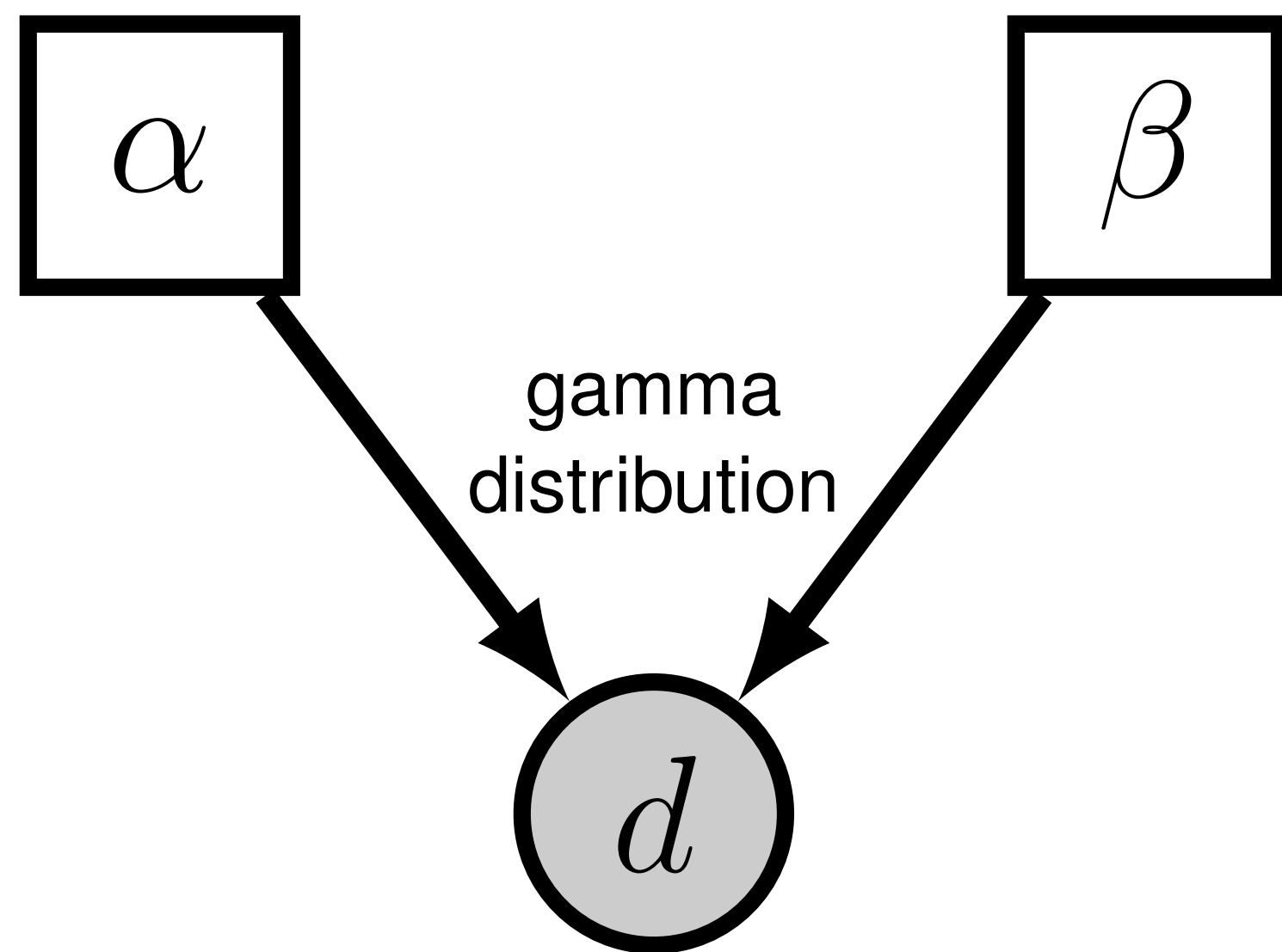
$$d \sim \text{Gamma}(\alpha, \beta)$$



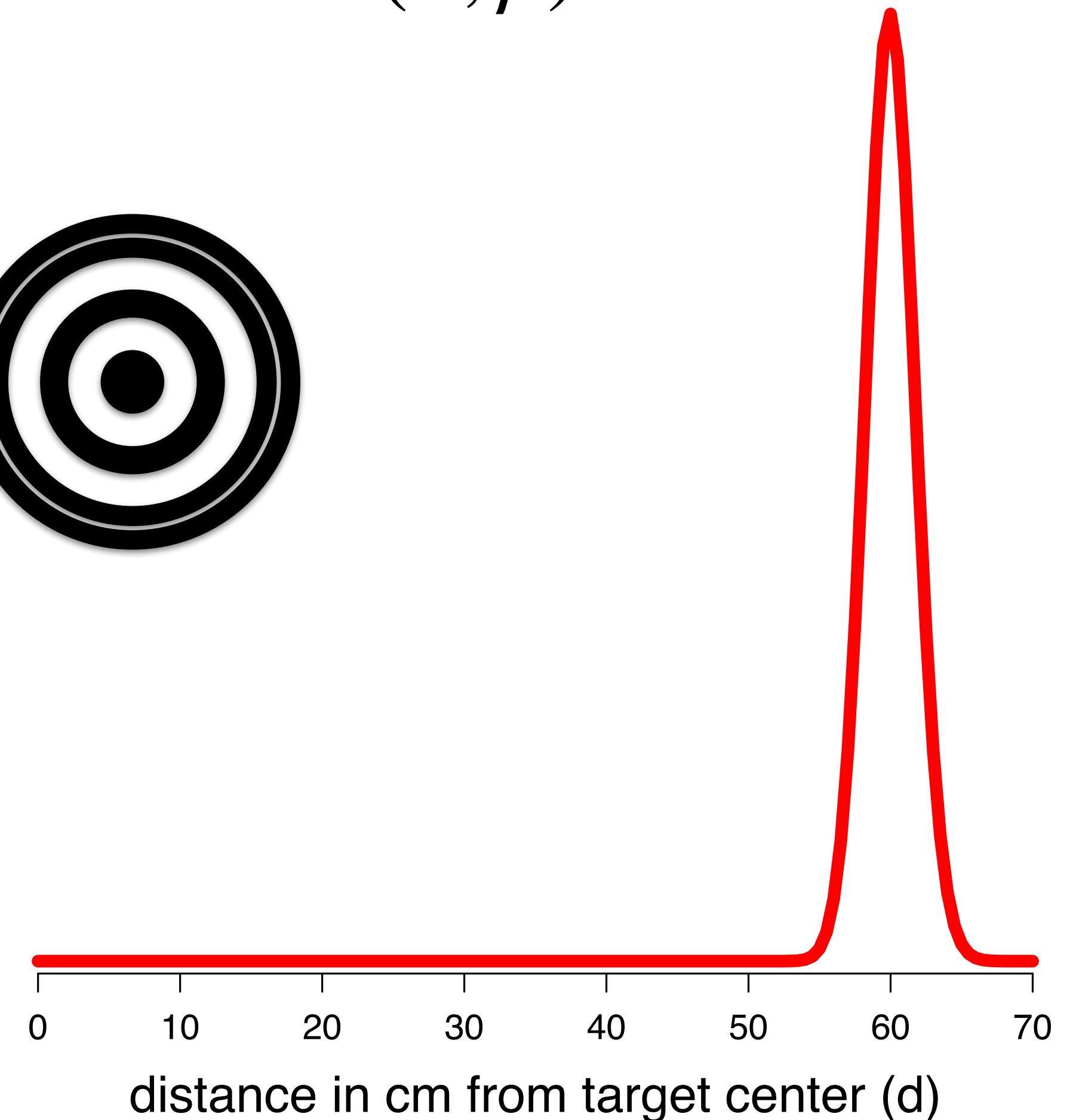
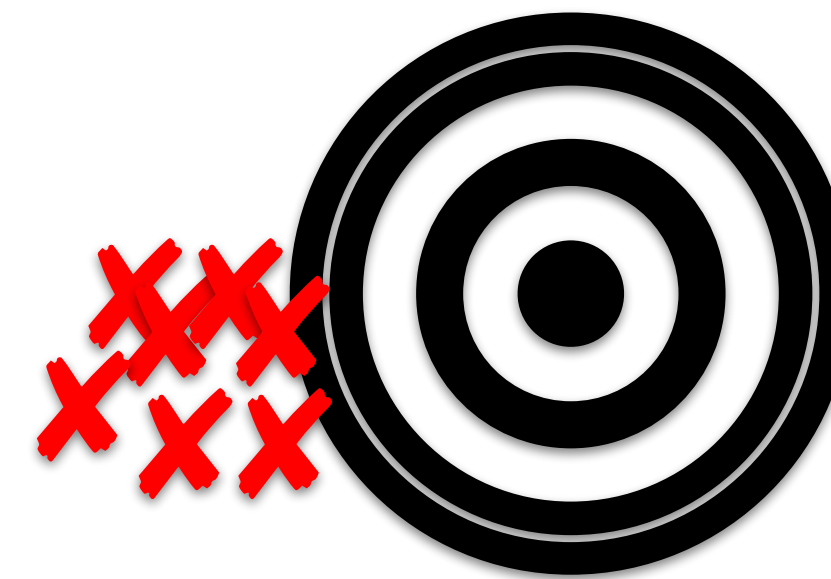


# Priors: Archery

Another way of expressing this distribution is with a probabilistic graphical model



$$d \sim \text{Gamma}(\alpha, \beta)$$



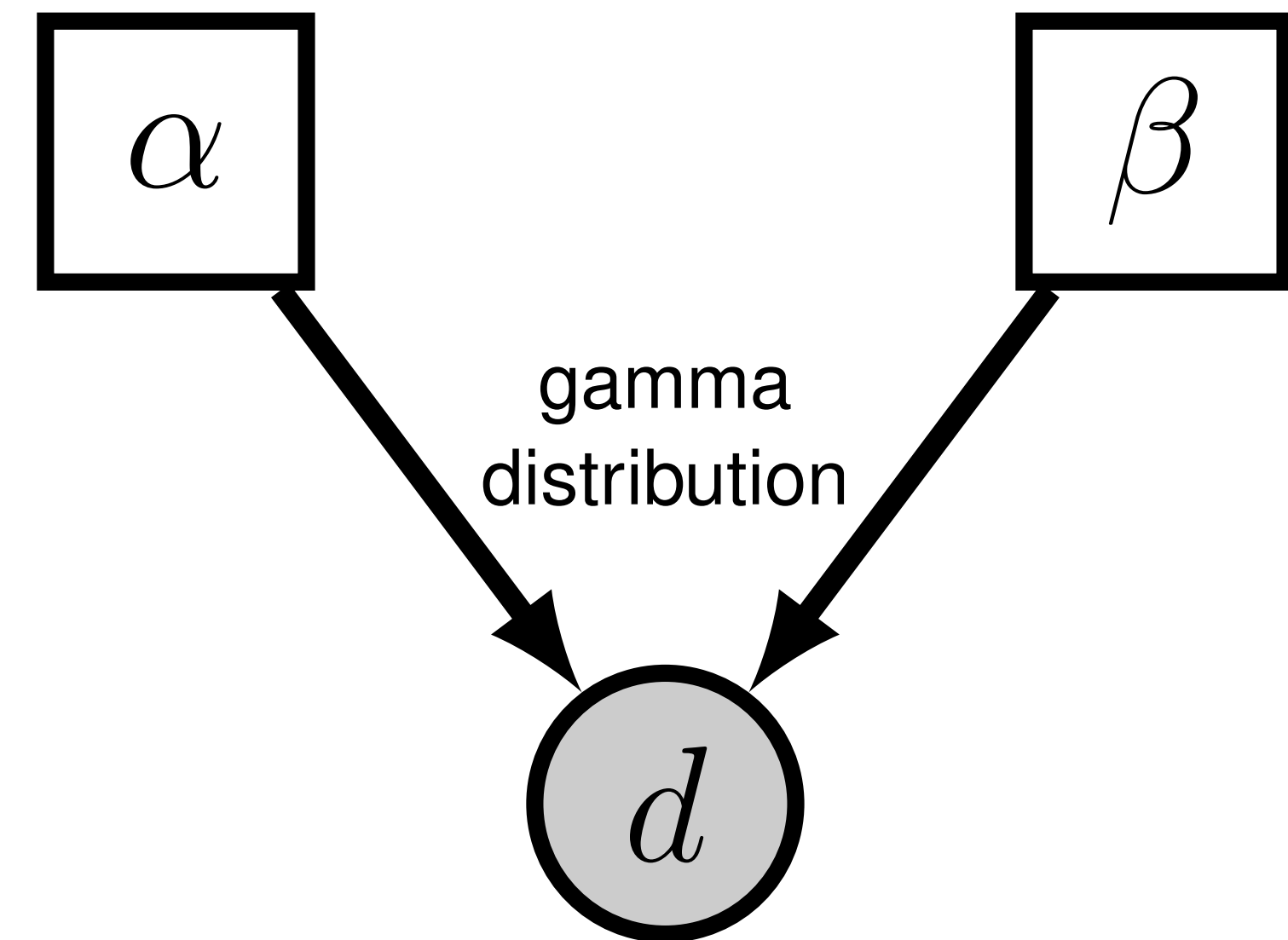


# Priors: Archery



$$d \sim \text{Gamma}(\alpha, \beta)$$

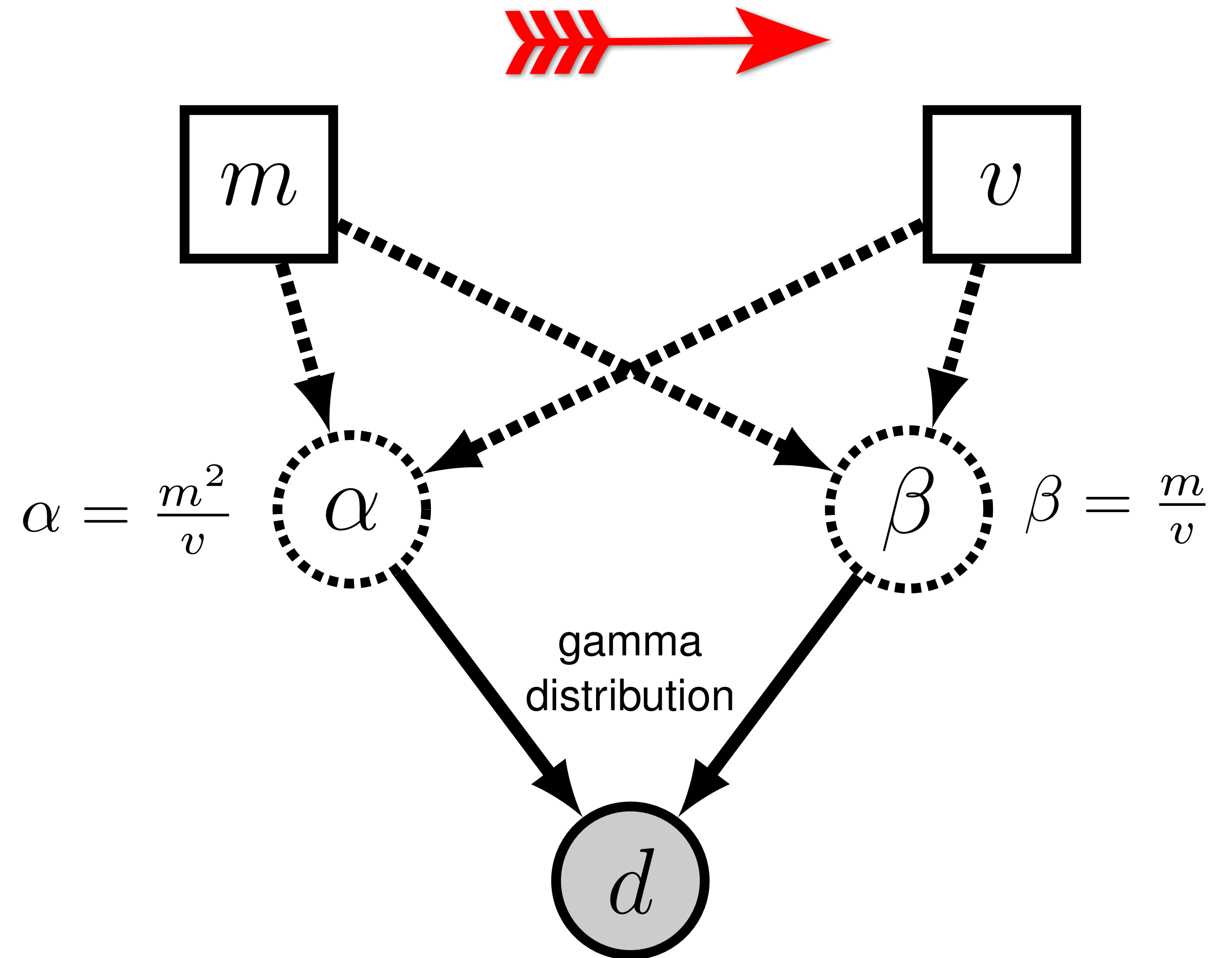
This shows that our observed datum ( $d =$  a single observed shot) is conditionally dependent on the shape ( $\alpha$ ) and rate ( $\beta$ ) of the gamma distribution





# Priors: Archery

We can parameterize the model using the mean ( $m$ ) and variance ( $v$ ), where  $\alpha$  and  $\beta$  are computed using  $m$  and  $v$



We may have more intuition about the mean and variance than we do about the shape and rate.

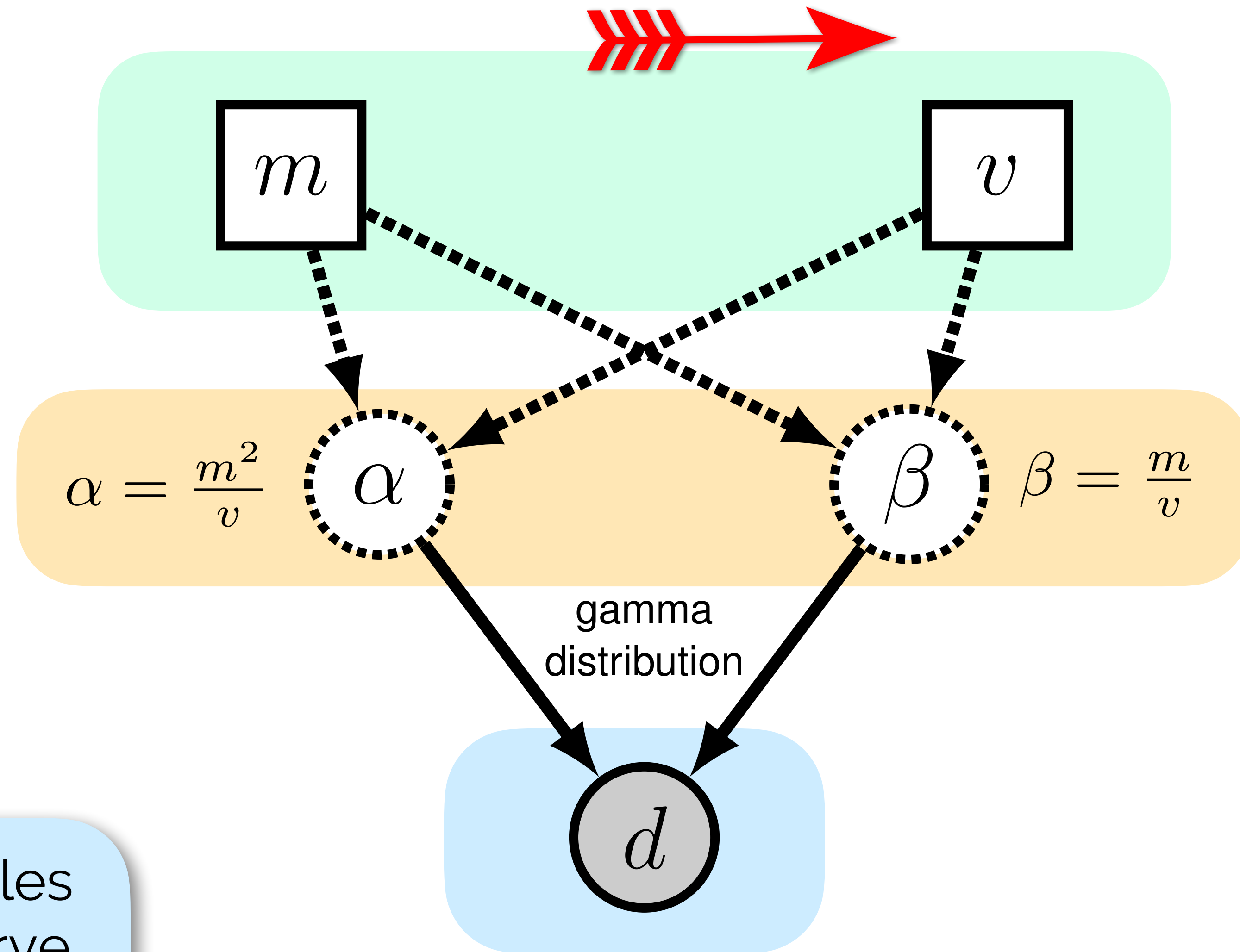


# Priors: Archery

Constant nodes represent a fixed value that is asserted or known

Deterministic nodes represent unknown random variable whose values are determined by other nodes

Stochastic nodes are random variables generated by the model. If we observe the value of a stochastic node, we fix it to that value



This graphical model has 3 types of nodes

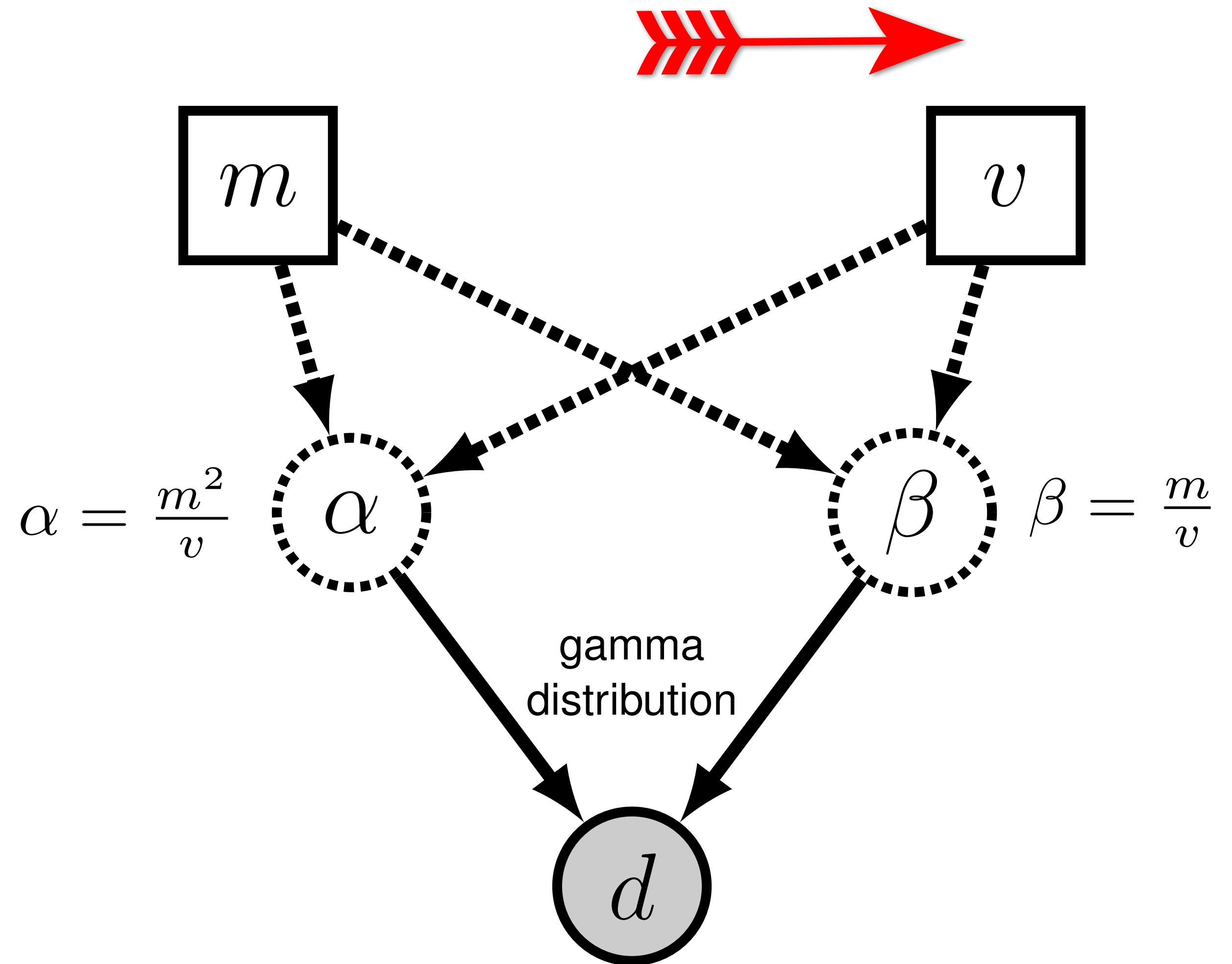


# Priors: Archery

If we set  $m$  and  $v$  to values corresponding to our assumed model, then we can calculate the likelihood of any observed shot

$$f(d \mid \alpha, \beta) = \frac{1}{\Gamma(\alpha)\beta^\alpha} d^{\alpha-1} e^{-\frac{d}{\beta}}$$

$$f(d = 39.76 \mid \alpha = 1200, \beta = 20) = 7.89916e - 40$$



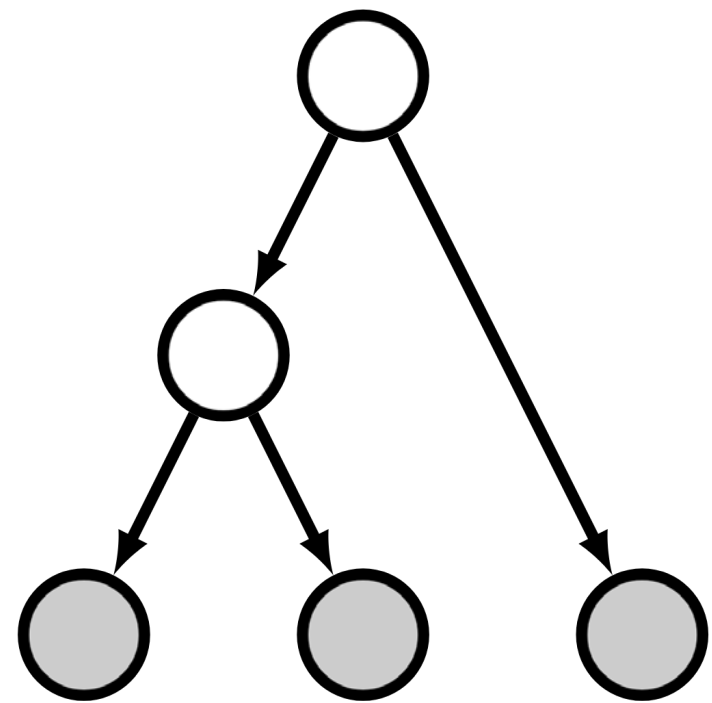


# RevBayes Demo: Archery

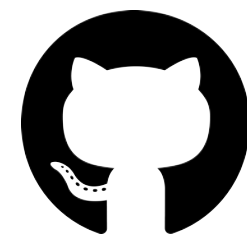


Bayesian Inference of Evolutionary Parameters

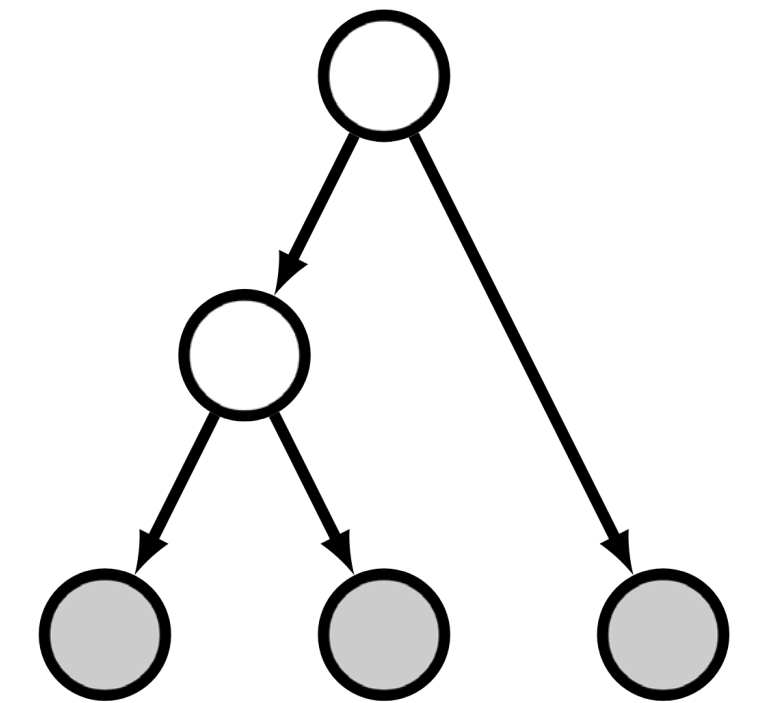
Fully integrative Bayesian inference of phylogenetic parameters using **probabilistic graphical models** and an **interpreted language**



<http://revbayes.com>



<https://github.com/revbayes>



Höhna et al. 2016. RevBayes: Bayesian phylogenetic inference using graphical models and an interactive model-specification language. *Systematic Biology*. (doi: 10.1093/sysbio/syw021)

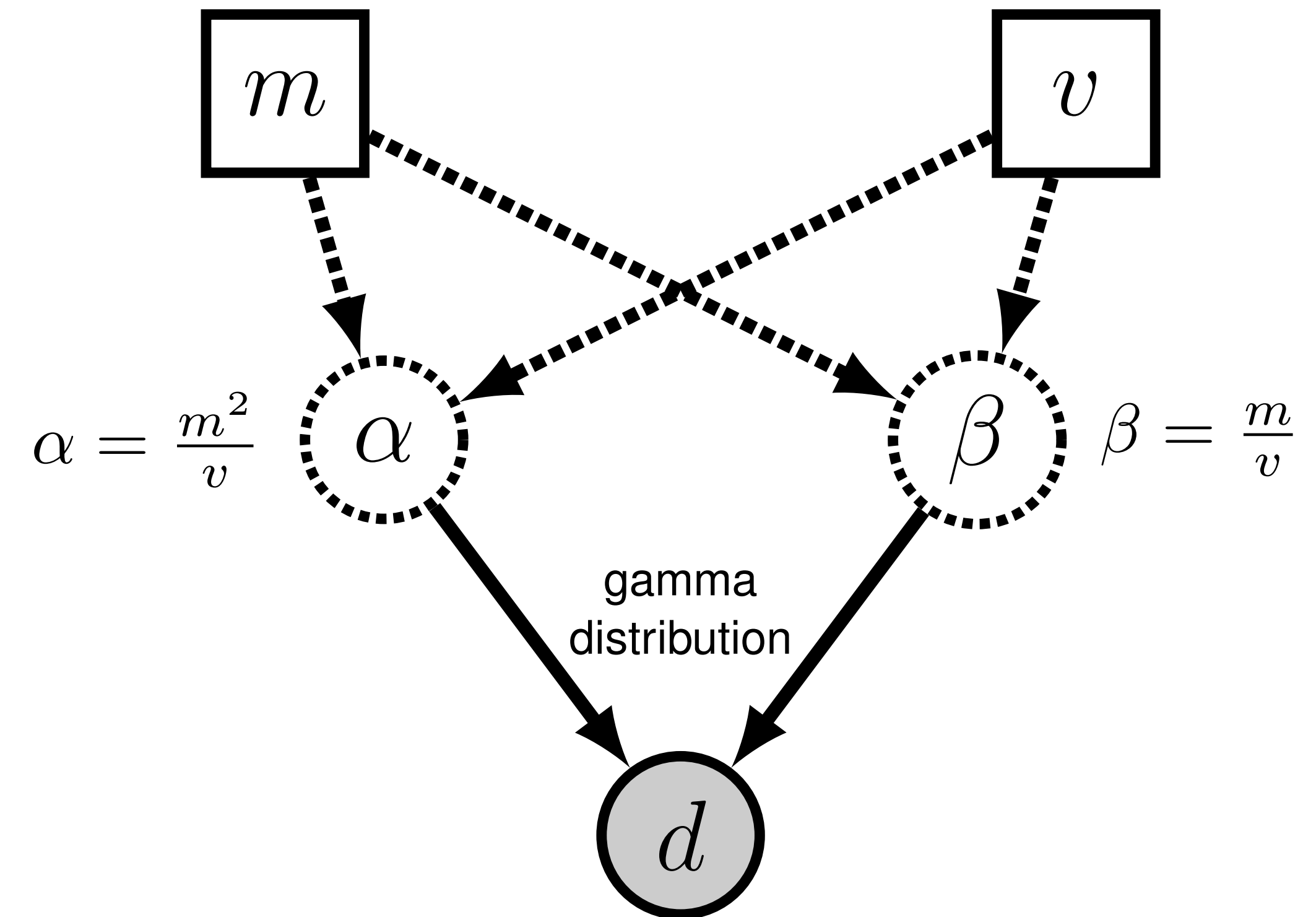


# Graphical Models in RevBayes



Graphical models provide a way to visually and computationally represent complex, parameter-rich probabilistic models

We can depict the conditional dependence structure of various parameters and other variables



Höhna et al. 2014. Probabilistic graphical model representation in Phylogenetics. *Systematic Biology*. (doi: 10.1093/sysbio/syu039)



# RevBayes Demo: Archery



The Rev language for calculating the probability of 1 data observation (observed\_shot) given a mean and variance

```
mean <- 60
var <- 3

alpha := (mean * mean) / var
beta := mean / var

observed_shot = 39.76

d ~ dnGamma(alpha, beta)
d.clamp(observed_shot)

d.lnProbability()

-90.03665
```



# Priors: Archery



What if we do not know  $m$  and  $v$ ?

We can use maximum likelihood or Bayesian methods to estimate their values

Maximum likelihood methods require us to find the values of  $m$  and  $v$  that maximize

$$f(d \mid m, v)$$

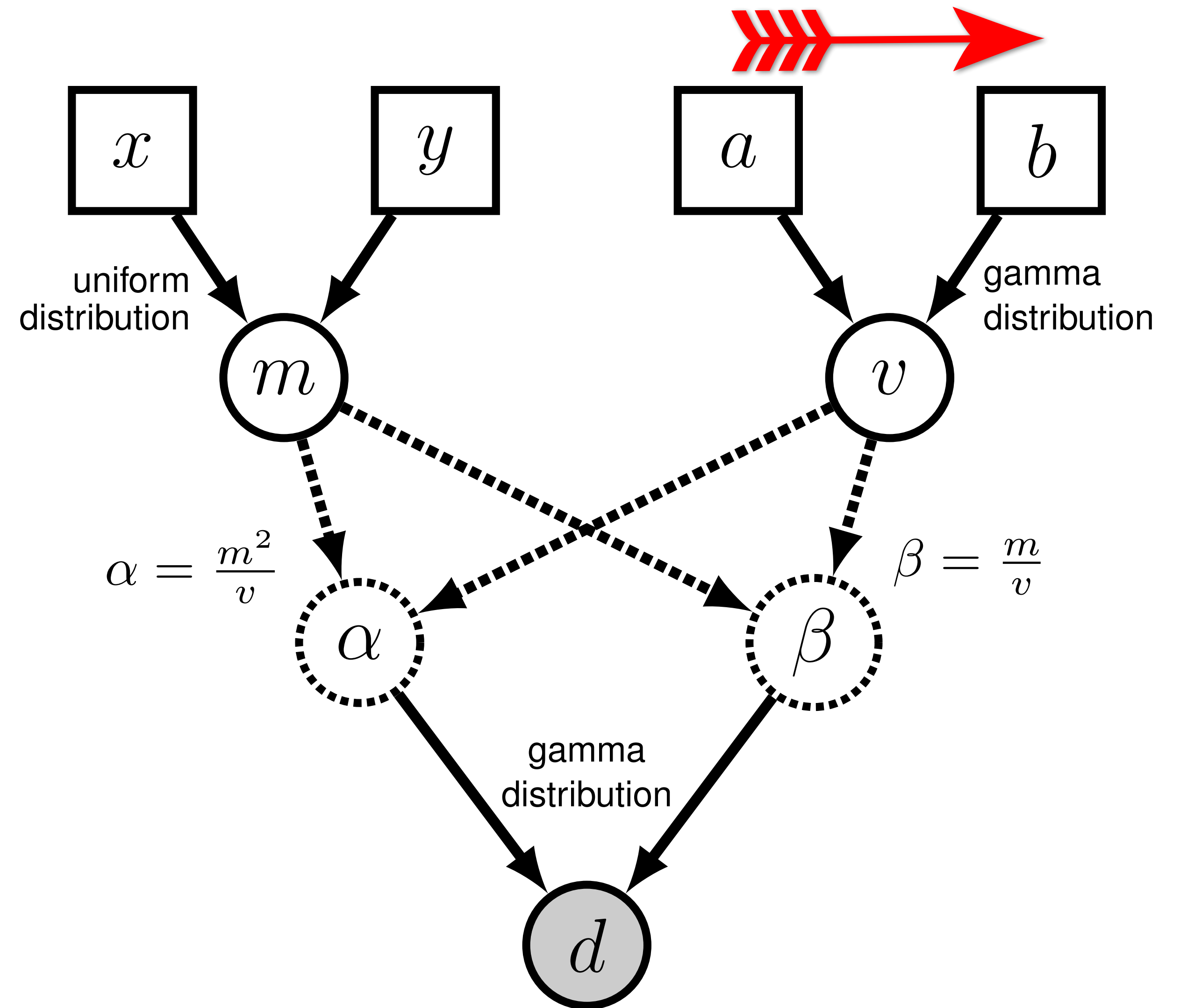
Bayesian methods use prior distributions to describe our uncertainty in  $m$  and  $v$  and estimate

$$f(m, v \mid d)$$



# Priors: Archery

We must define prior distributions for  $m$  and  $v$  to account for uncertainty and estimate the posterior densities of those parameters

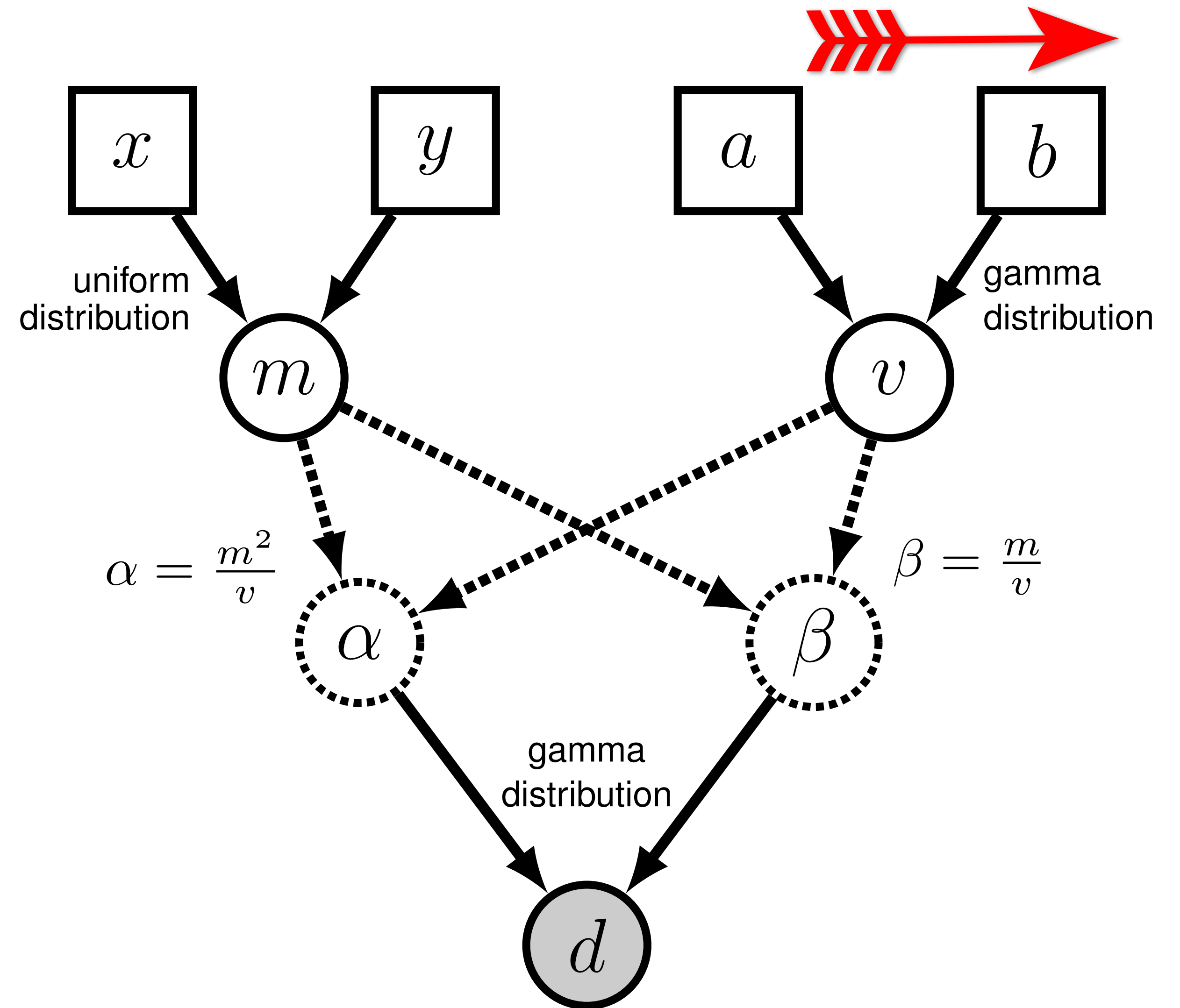




# Priors: Archery

Now  $x$  and  $y$  are the parameters of the uniform prior on  $m$

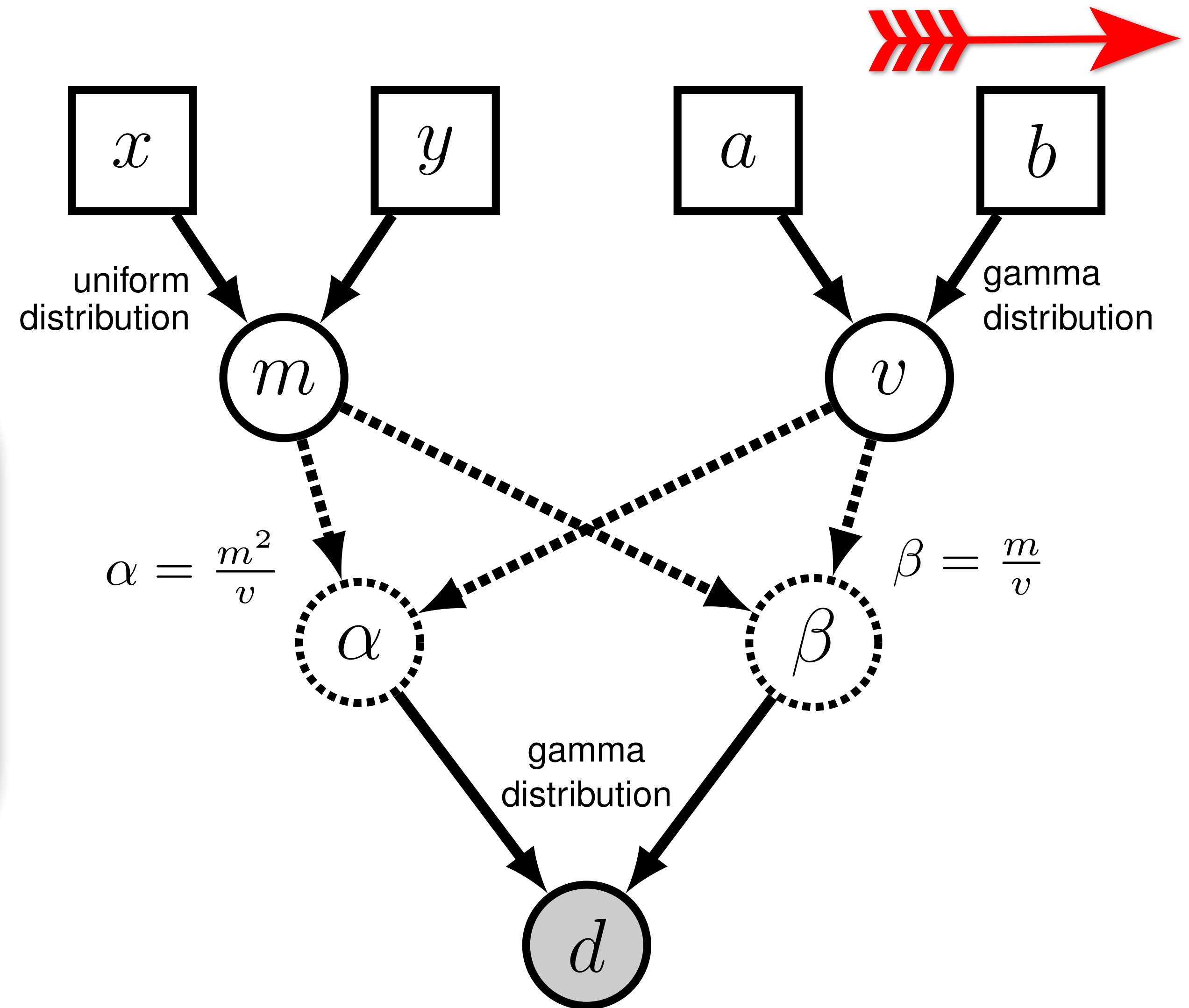
And  $a$  and  $b$  are the shape and rate parameters of the gamma prior on  $v$





# Priors: Archery

Stochastic nodes that are not observed are random variables that are unknown and estimated

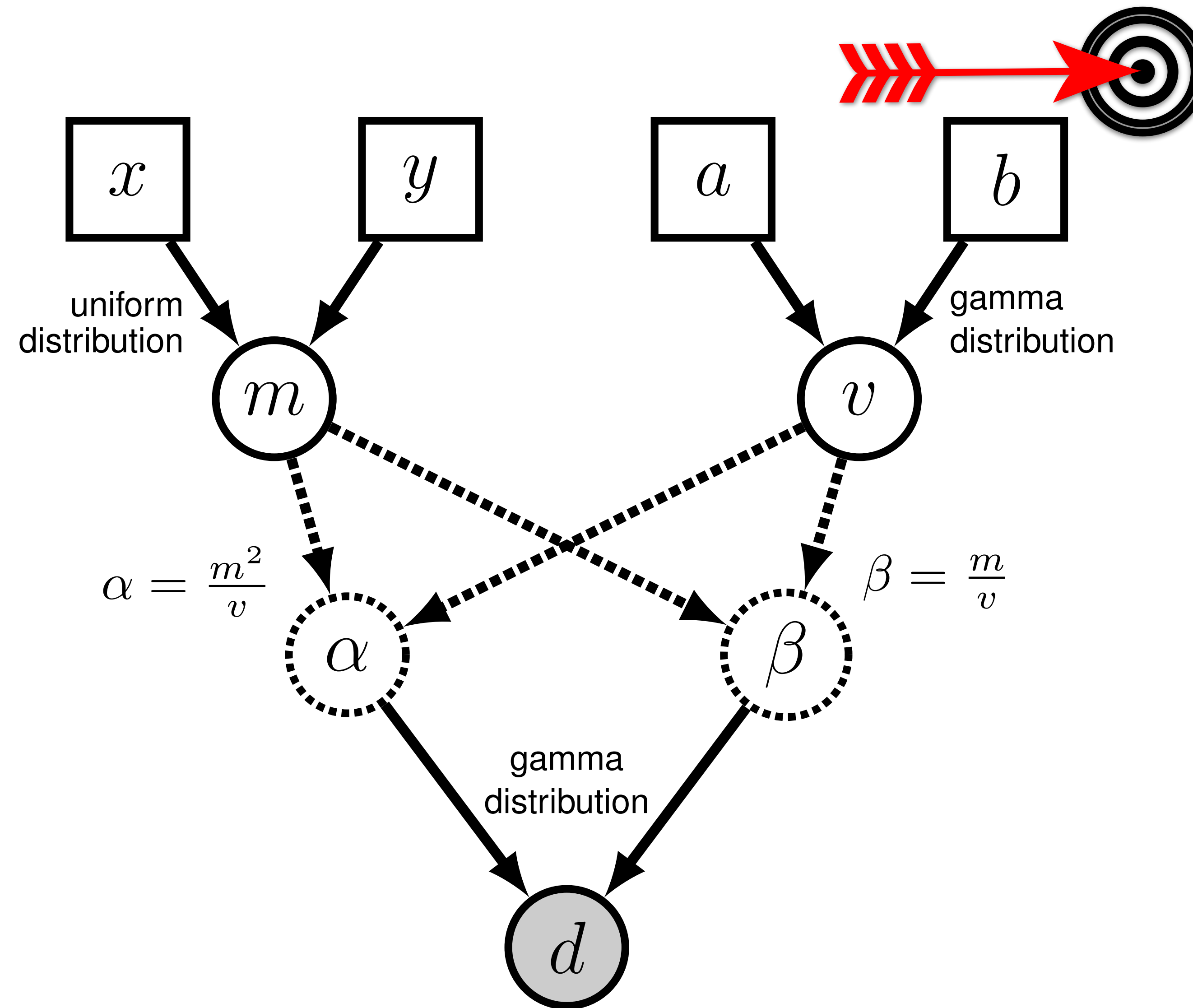




# Priors: Archery

The values we choose for the parameters of these prior distributions should reflect our prior knowledge

If we observed a previous shot at **39.76 cm**, then we can use this to parameterize our priors for analysis of future observations





# Priors: Archery

$$m \sim \text{Uniform}(x, y)$$

$$x = 10$$

$$y = 50$$

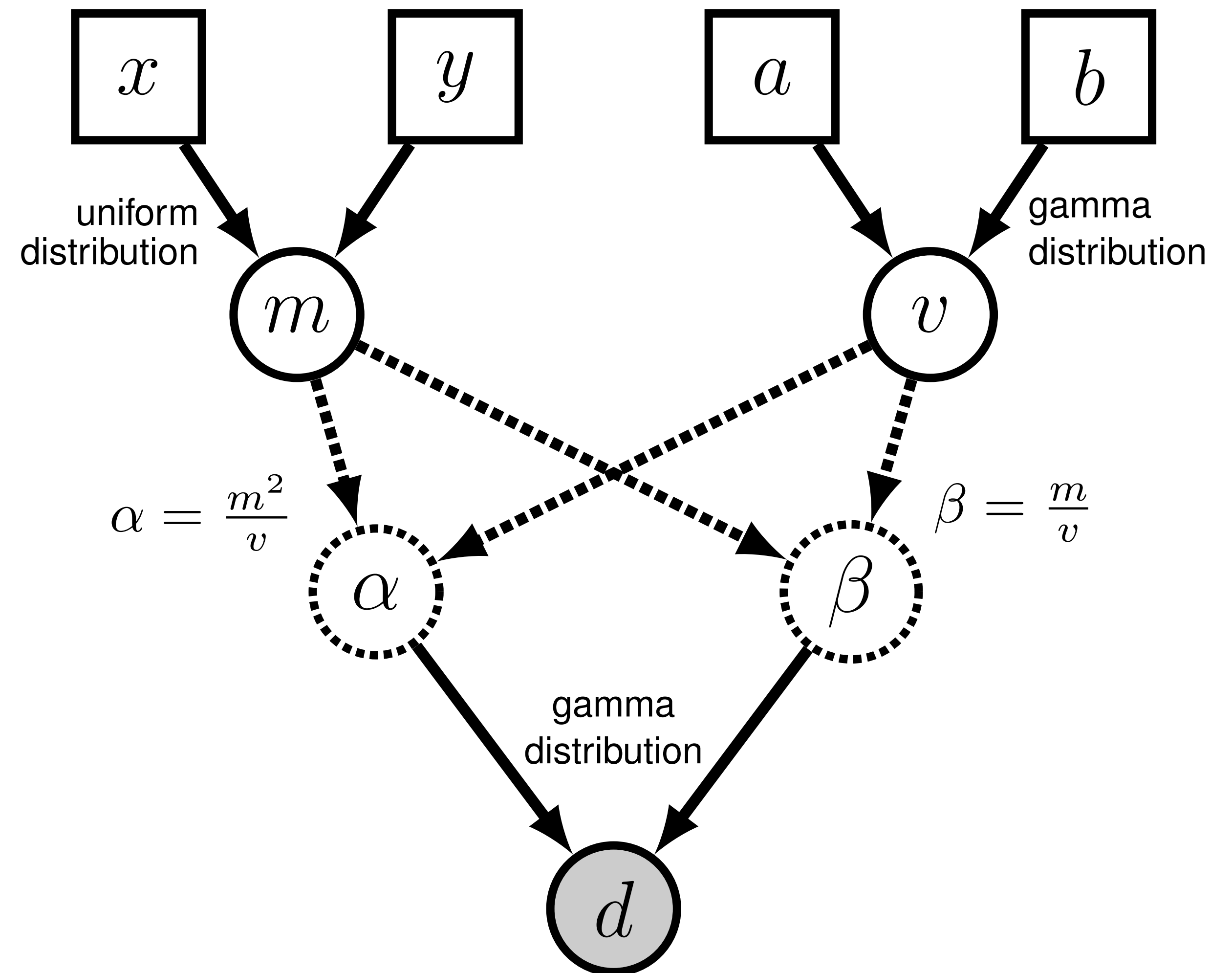
$$\mathbb{E}(m) = 30$$

$$v \sim \text{Gamma}(a, b)$$

$$a = 20$$

$$b = 2$$

$$\mathbb{E}(v) = 10$$





# RevBayes Demo: Archery



The Rev language specifying a hierarchical model on shot accuracy based on 1 new observation

```
mean ~ dnUnif(10, 50)
var ~ dnGamma(20, 2)

alpha := (mean * mean) / var
beta := mean / var

observed_shot = 35.21

d ~ dnGamma(alpha, beta)
d.clamp(observed_shot)

d.lnProbability()
```

depends on initial value of mean & var



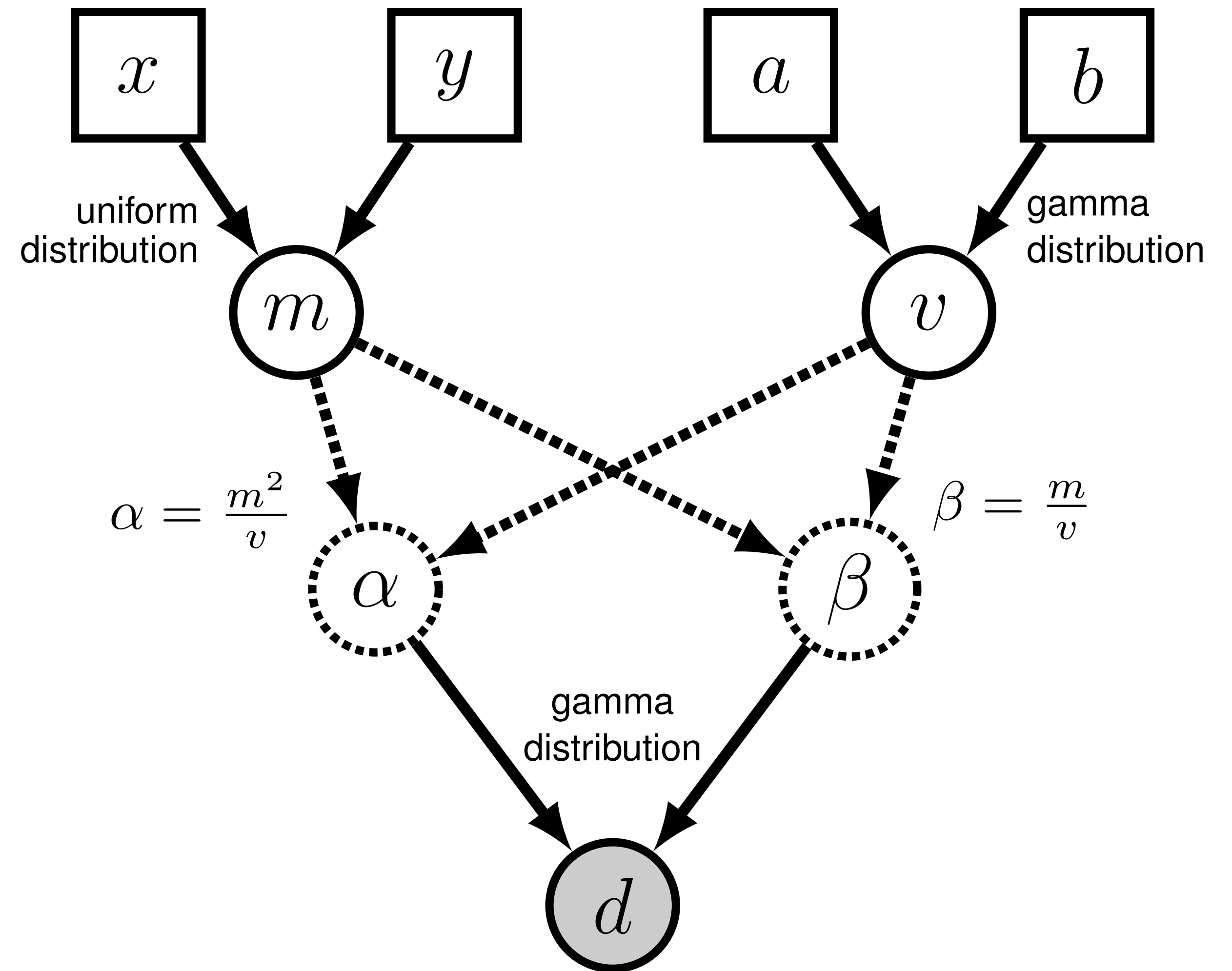
# Priors: Archery

Now that we have a defined model, how do we estimate the posterior probability density?

$$m \sim \text{Uniform}(x, y)$$

$$v \sim \text{Gamma}(a, b)$$

$$d \sim \text{Gamma}(\alpha, \beta)$$

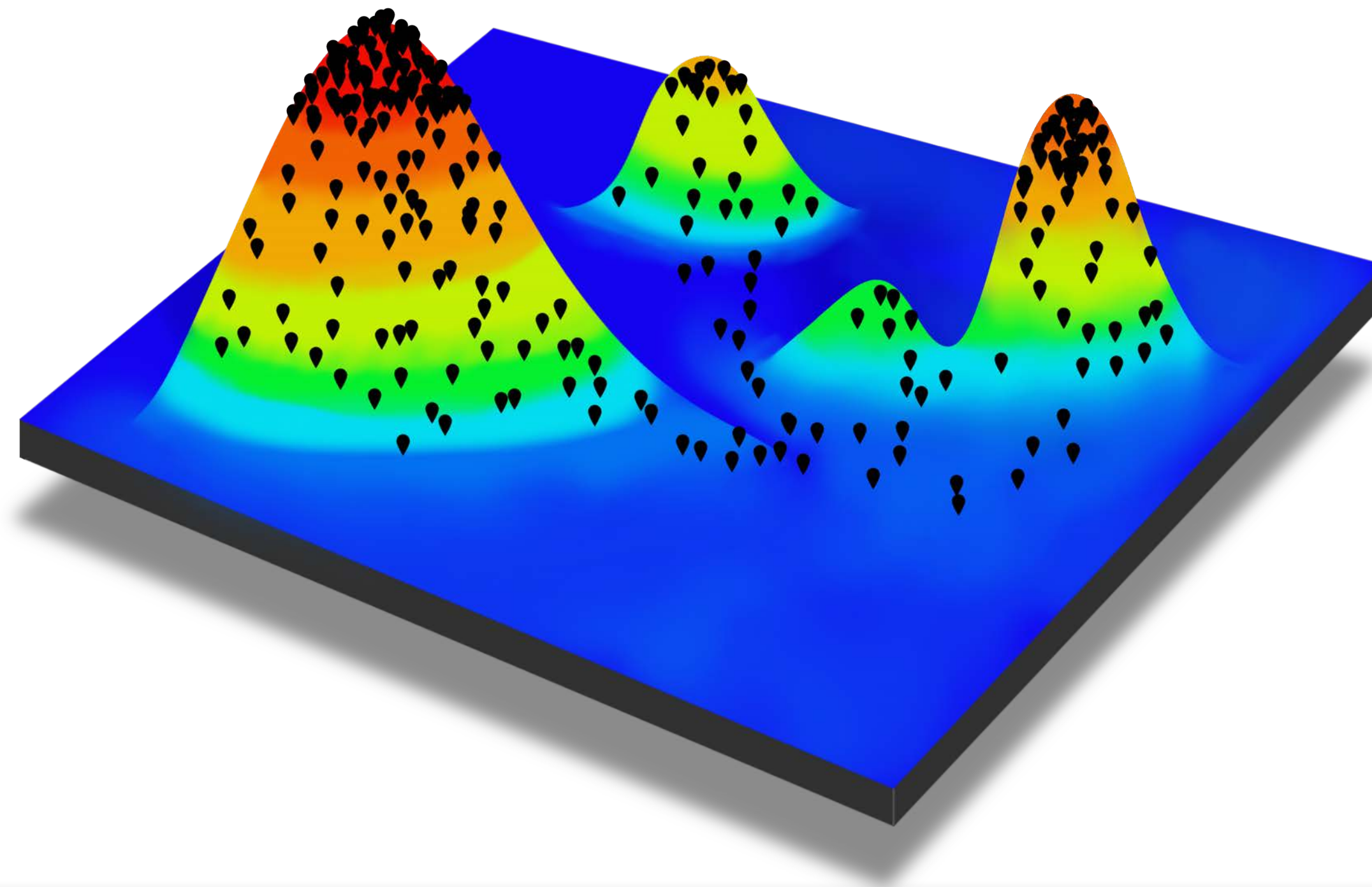


$$f(m, v \mid d, a, b, x, y) \propto f(d \mid , \alpha = \frac{m^2}{v}, \beta = \frac{m}{v}) f(m \mid x, y) f(v \mid a, b)$$



# Markov Chain Monte Carlo

An algorithm for approximating the posterior distribution



Metropolis, et al. 1953. Equations of state calculations by fast computing machines. J. Chem. Phys.

Hastings. 1970. Monte Carlo sampling methods using Markov chains and their applications. Biometrika.



# MCMC Robot Rules





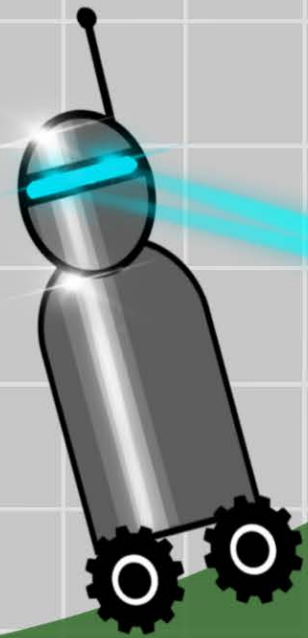
# MCMC Robot Rules





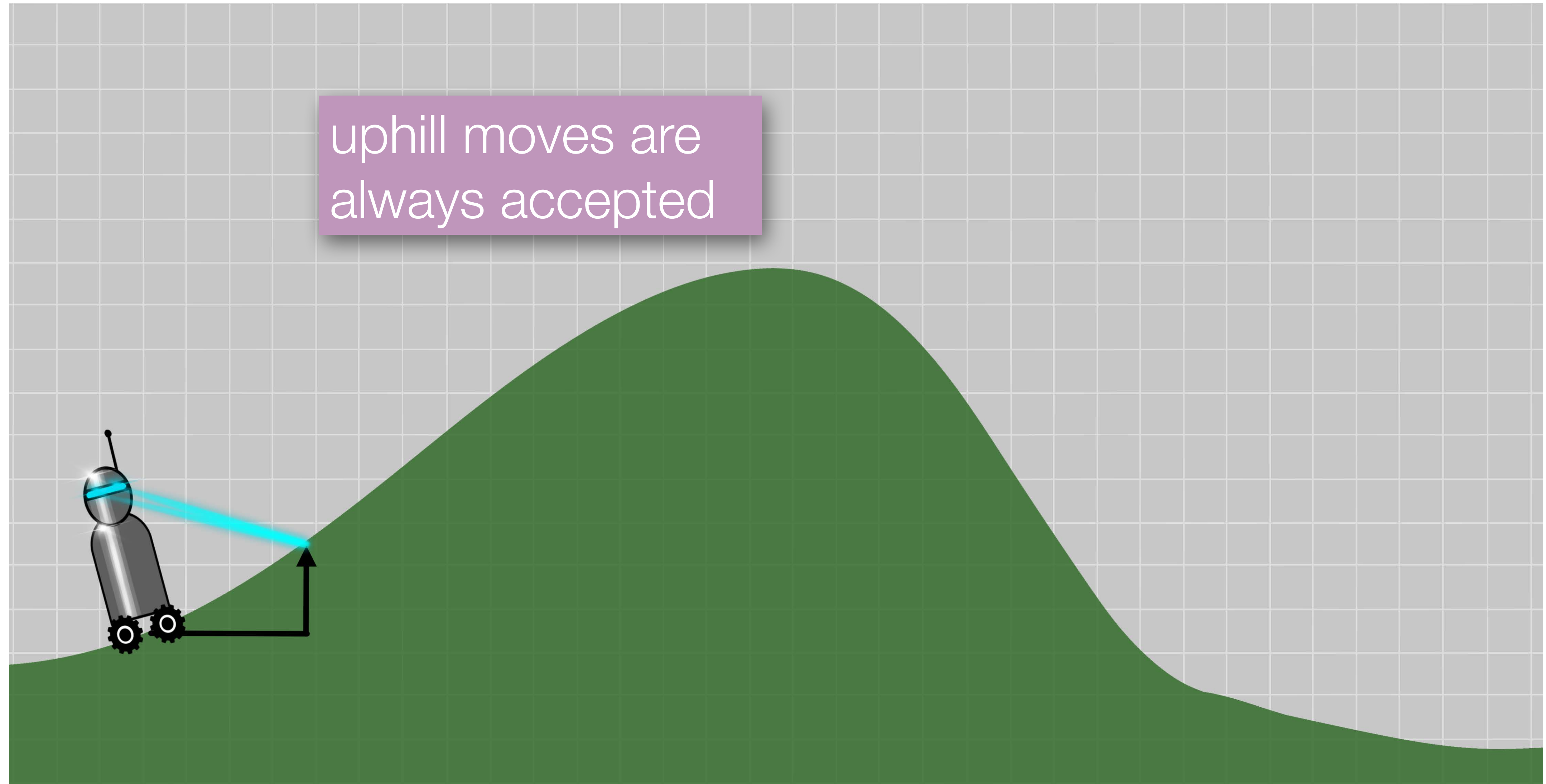
# MCMC Robot Rules

uphill moves are  
always accepted



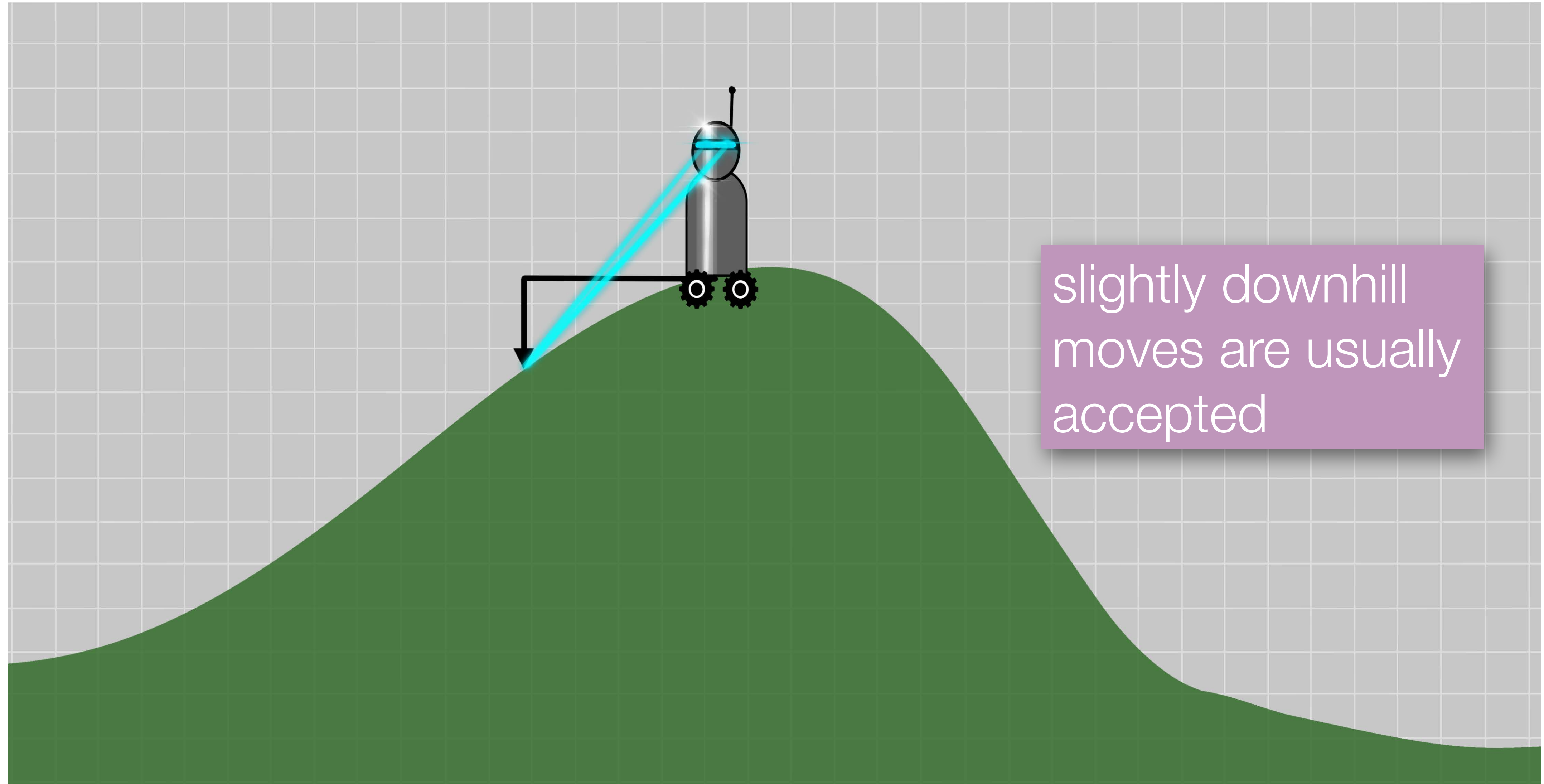


# MCMC Robot Rules



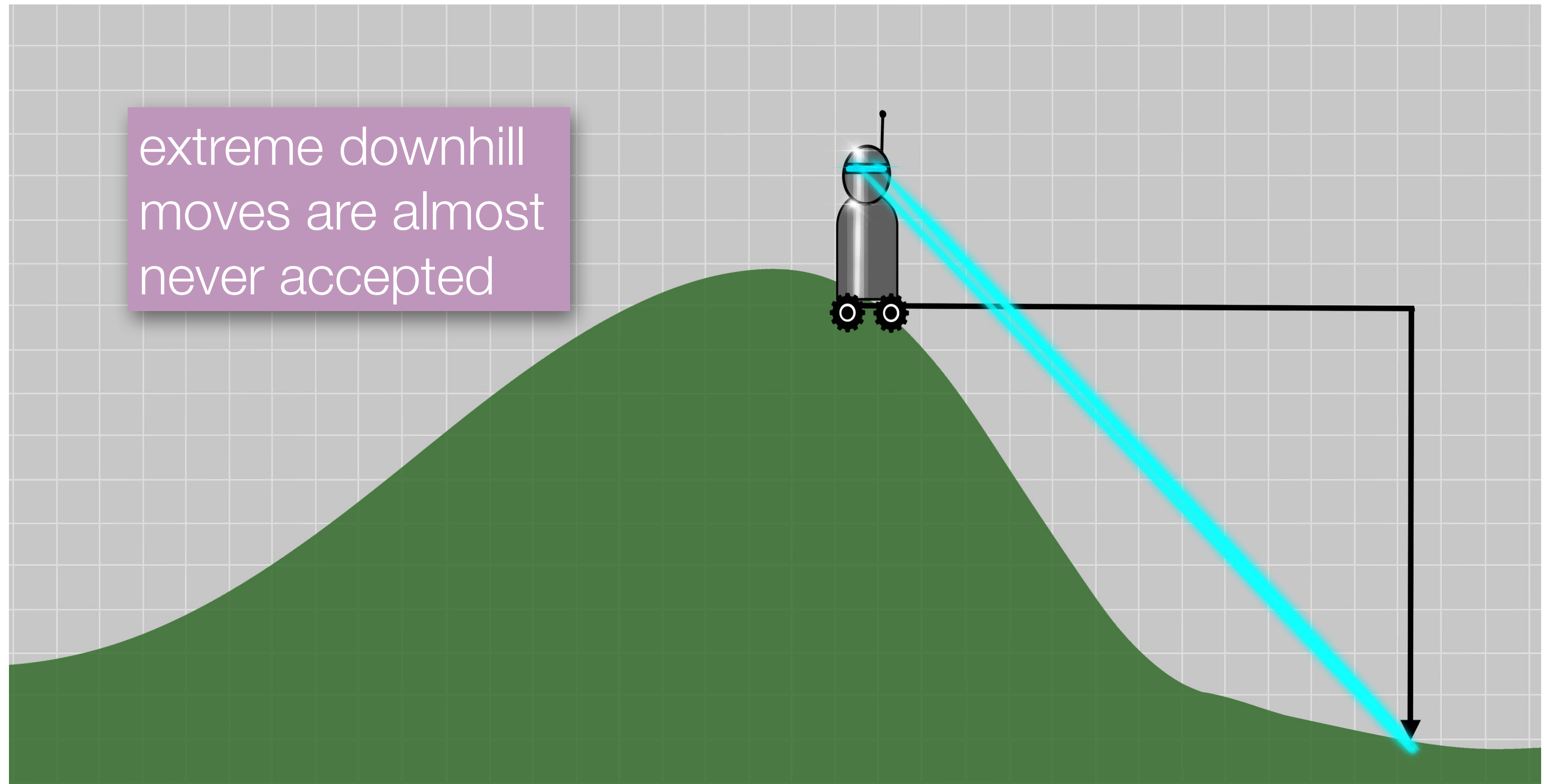


# MCMC Robot Rules





# MCMC Robot Rules





# Actual Rules (Metropolis Algorithm)

Metropolis et al. 1953. Equation of state calculations by fast computing machines. *J. Chem. Physics*.

when the robot makes a move  
it evaluates the new value  
compared to its current state

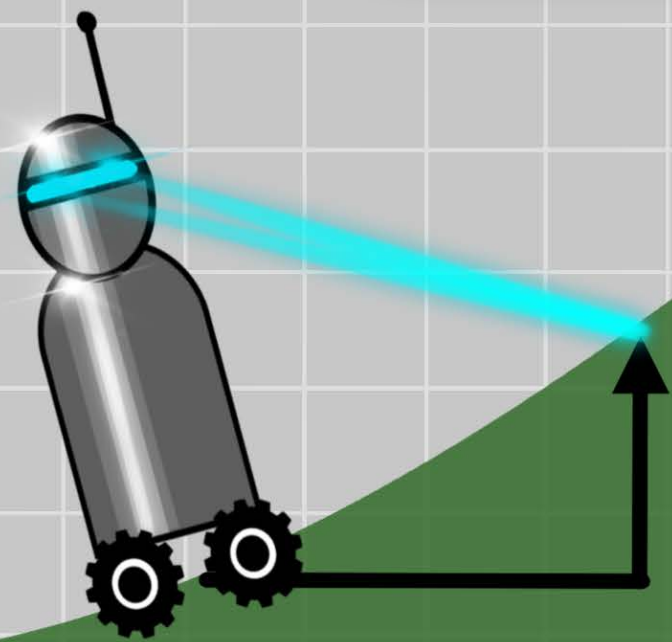
current = 3.5 m  
proposed = 5.5 m  
 $R = 5.5 / 3.5 = 2.0$

to decide if it should move to  
the proposed state it draws a  
variable ( $x$ ) from a uniform  
distribution between 0 and 1

$x \sim \text{Uniform}(0,1)$

the robot moves if:  
 $x \leq R$

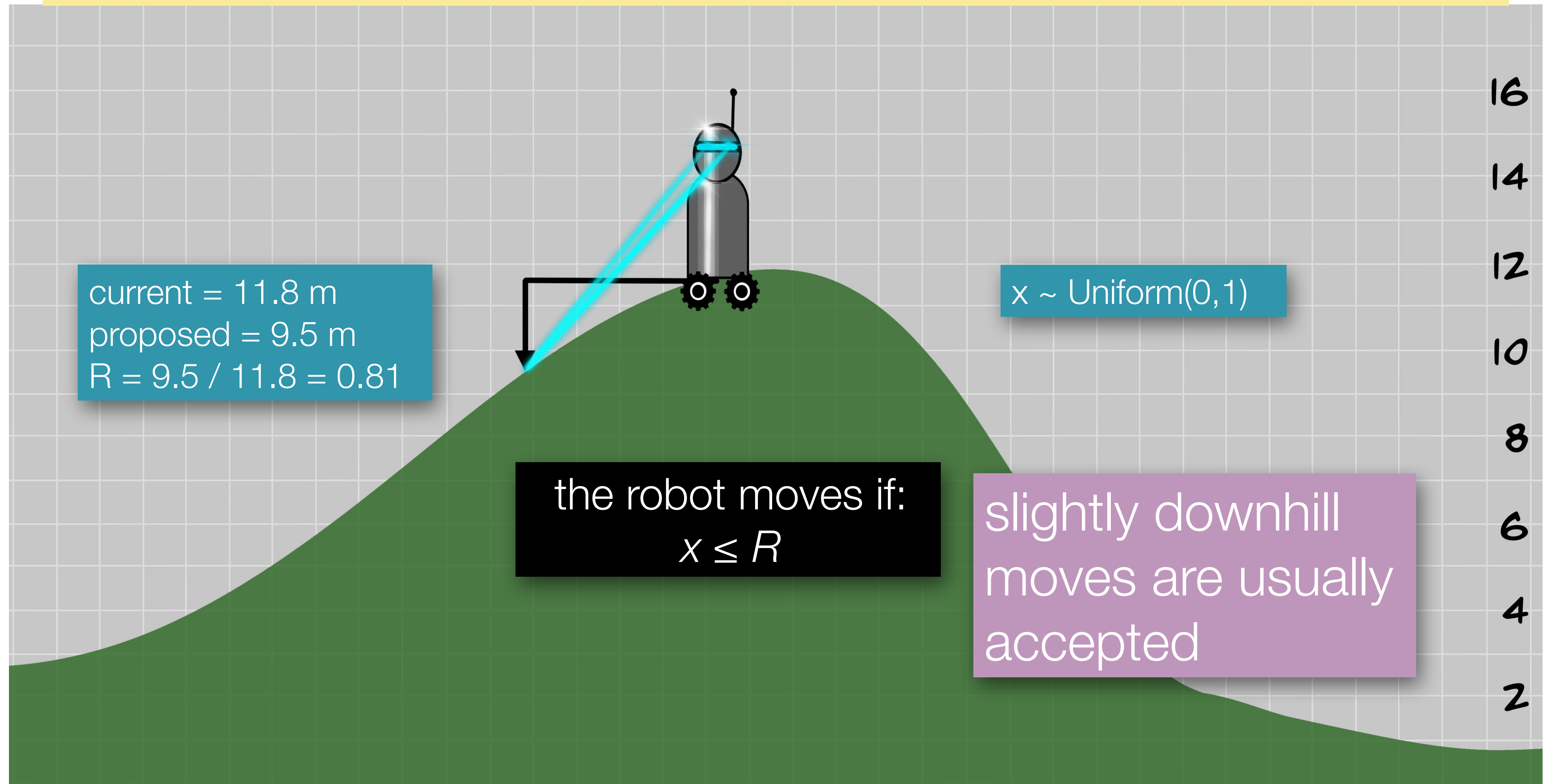
uphill moves are  
always accepted





# Actual Rules (Metropolis Algorithm)

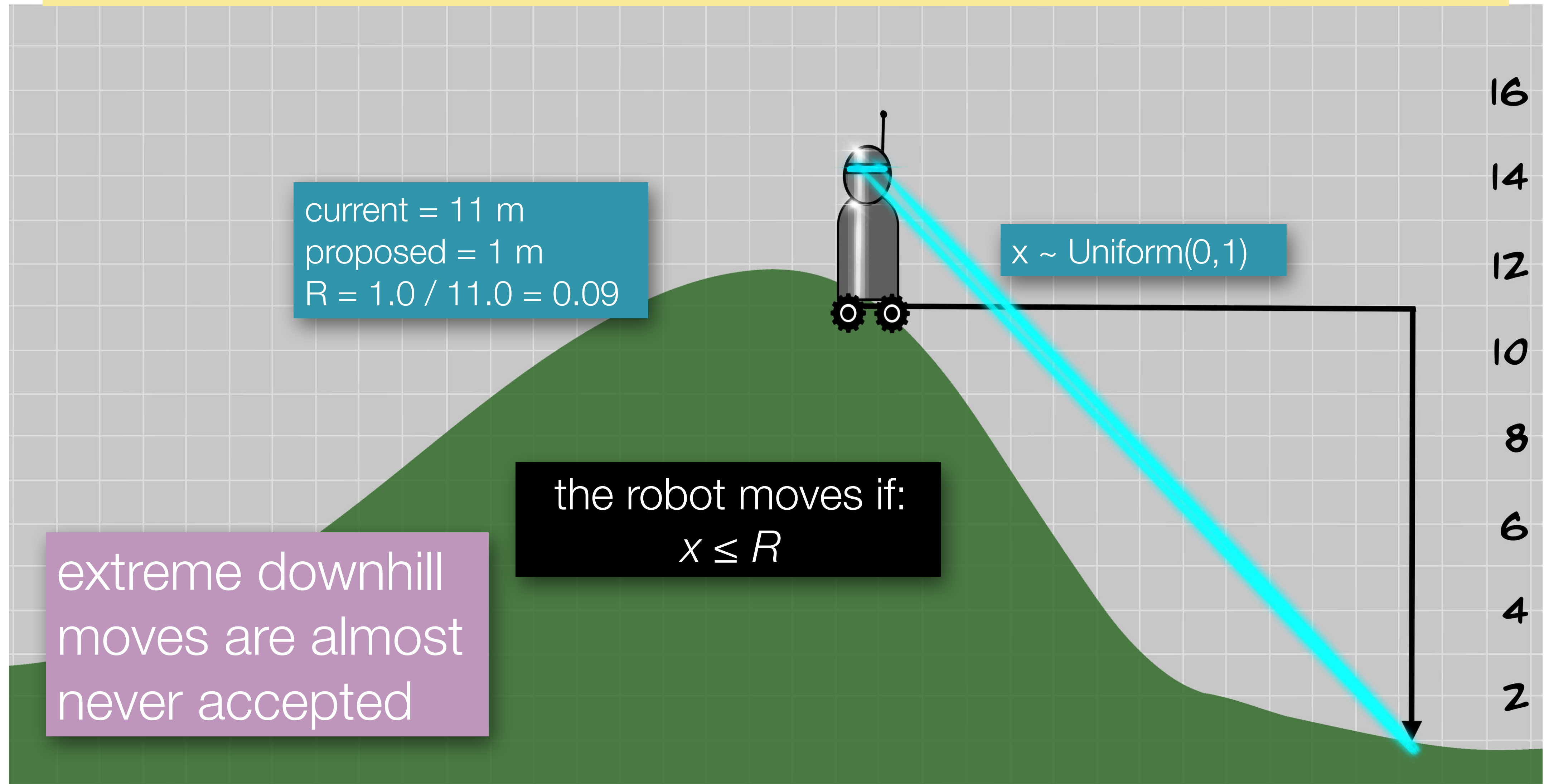
Metropolis et al. 1953. Equation of state calculations by fast computing machines. *J. Chem. Physics*.





# Actual Rules (Metropolis Algorithm)

Metropolis et al. 1953. Equation of state calculations by fast computing machines. *J. Chem. Physics*.





# Bayes Rule

POSTERIOR PROBABILITY

LIKELIHOOD

PRIOR PROBABILITY

$$\Pr(\theta \mid D) = \frac{\Pr(D \mid \theta) \Pr(\theta)}{\sum_{\theta} \Pr(D \mid \theta) \Pr(\theta)}$$

MARGINAL PROBABILITY OF THE DATA

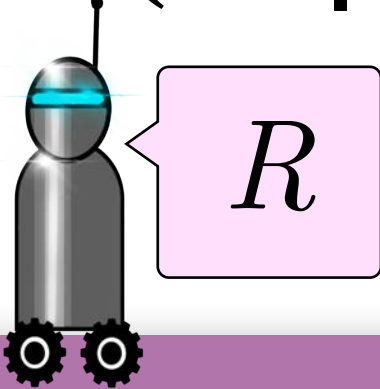
The diagram illustrates Bayes Rule with the following components and labels:

- POSTERIOR PROBABILITY:** Labeled with an arrow pointing to the yellow box containing  $\Pr(\theta \mid D)$ .
- LIKELIHOOD:** Labeled with an arrow pointing to the blue box containing  $\Pr(D \mid \theta)$ .
- PRIOR PROBABILITY:** Labeled with an arrow pointing to the pink box containing  $\Pr(\theta)$ .
- MARGINAL PROBABILITY OF THE DATA:** Labeled with an arrow pointing to the green box containing the denominator  $\sum_{\theta} \Pr(D \mid \theta) \Pr(\theta)$ .



# Canceling Out the Marginal Likelihood

$$\frac{\Pr(\theta^* \mid D)}{\Pr(\theta \mid D)} = \frac{\frac{\Pr(D \mid \theta^*) \Pr(\theta^*)}{\cancel{\Pr(D)}}}{\frac{\Pr(D \mid \theta) \Pr(\theta)}{\cancel{\Pr(D)}}} = \frac{\Pr(D \mid \theta^*) \Pr(\theta^*)}{\Pr(D \mid \theta) \Pr(\theta)}$$

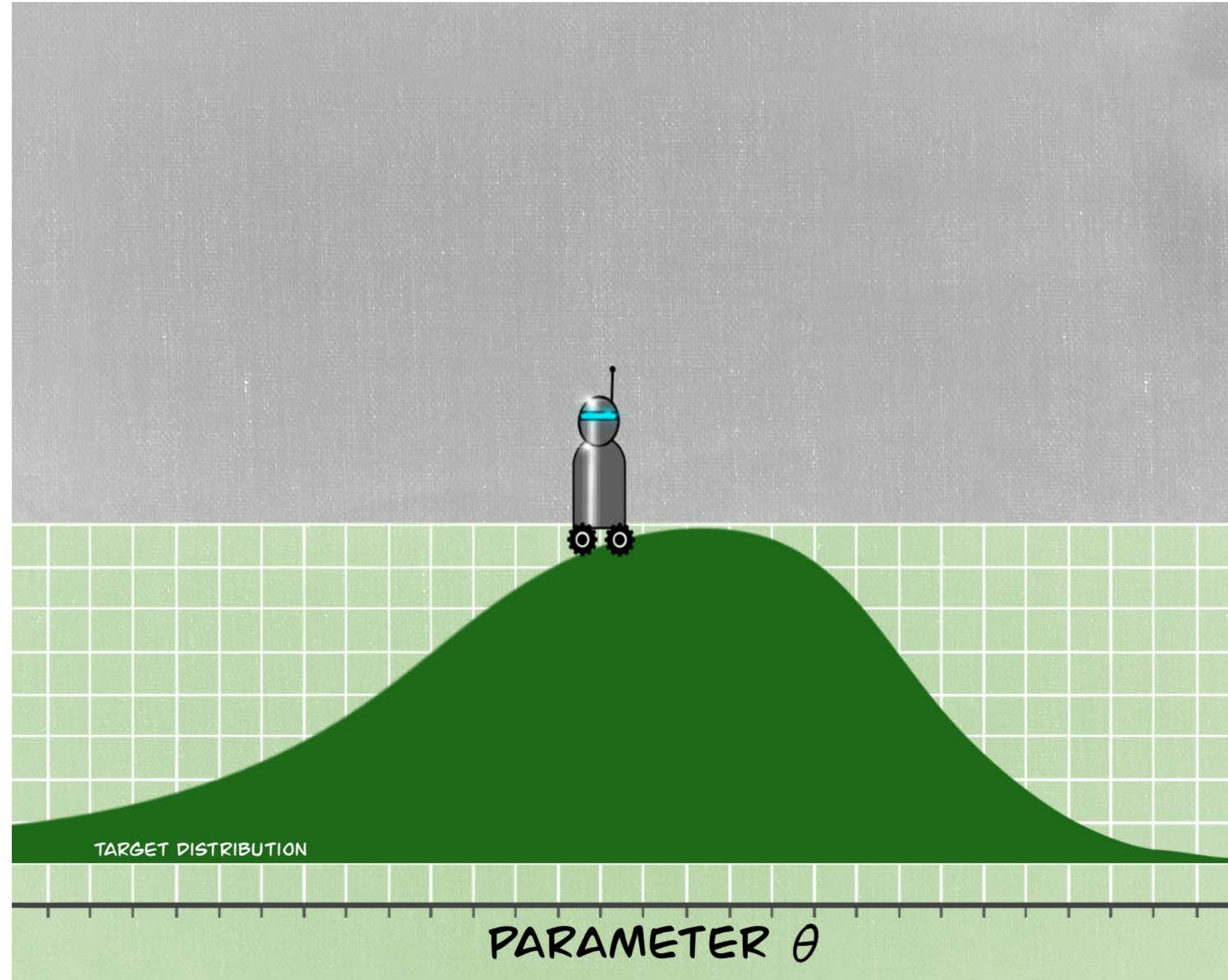
 Posterior odds

Bayes Rule!

Likelihood ratio      Prior odds



# Target and Proposal Distributions

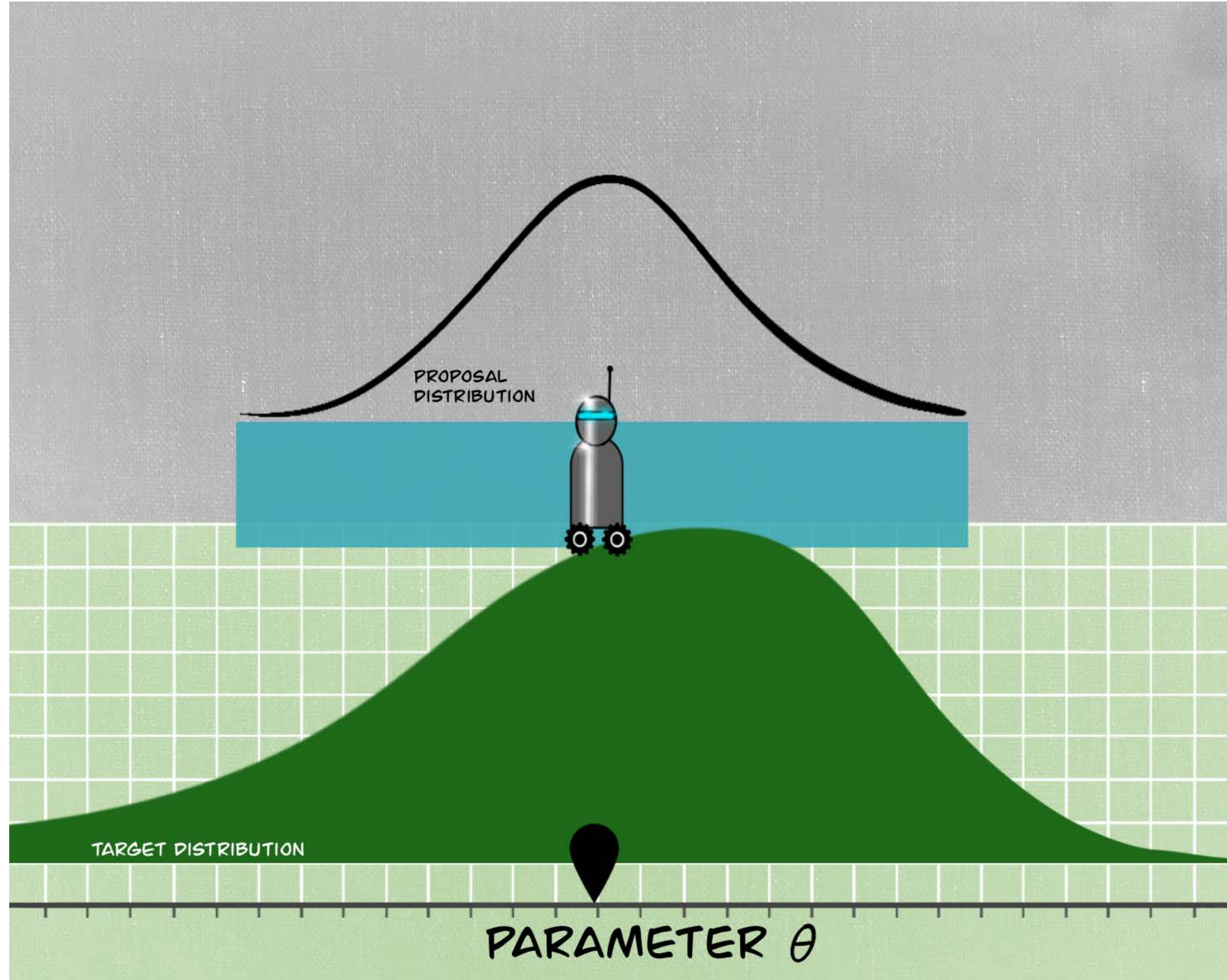


the target distribution is  
the landscape mapped by  
the robot

typically, this is the  
posterior distribution



# Target and Proposal Distributions

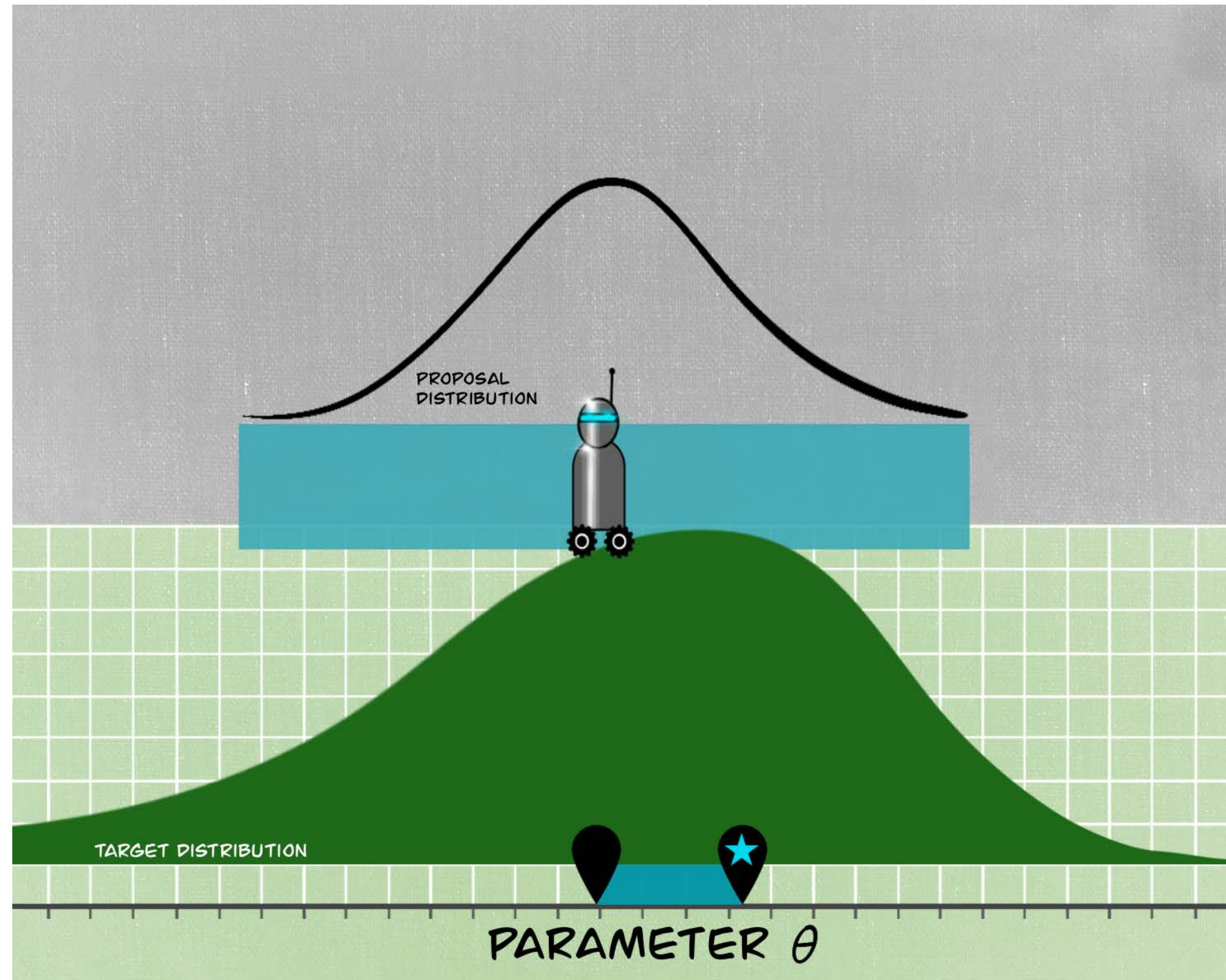


the proposal distribution is separate from the target distribution

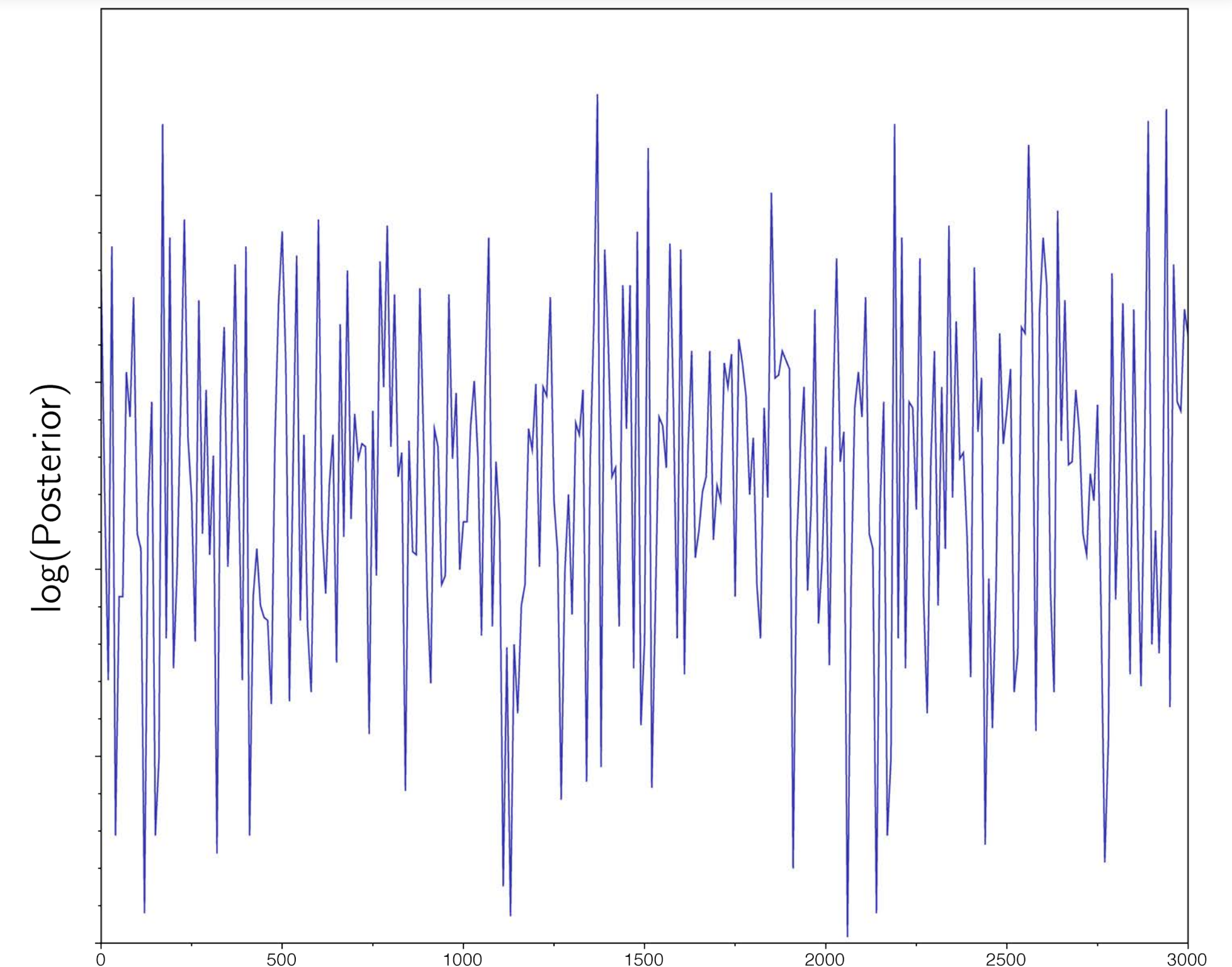
the robot uses the proposal distribution to choose the next spot to move



# Target and Proposal Distributions



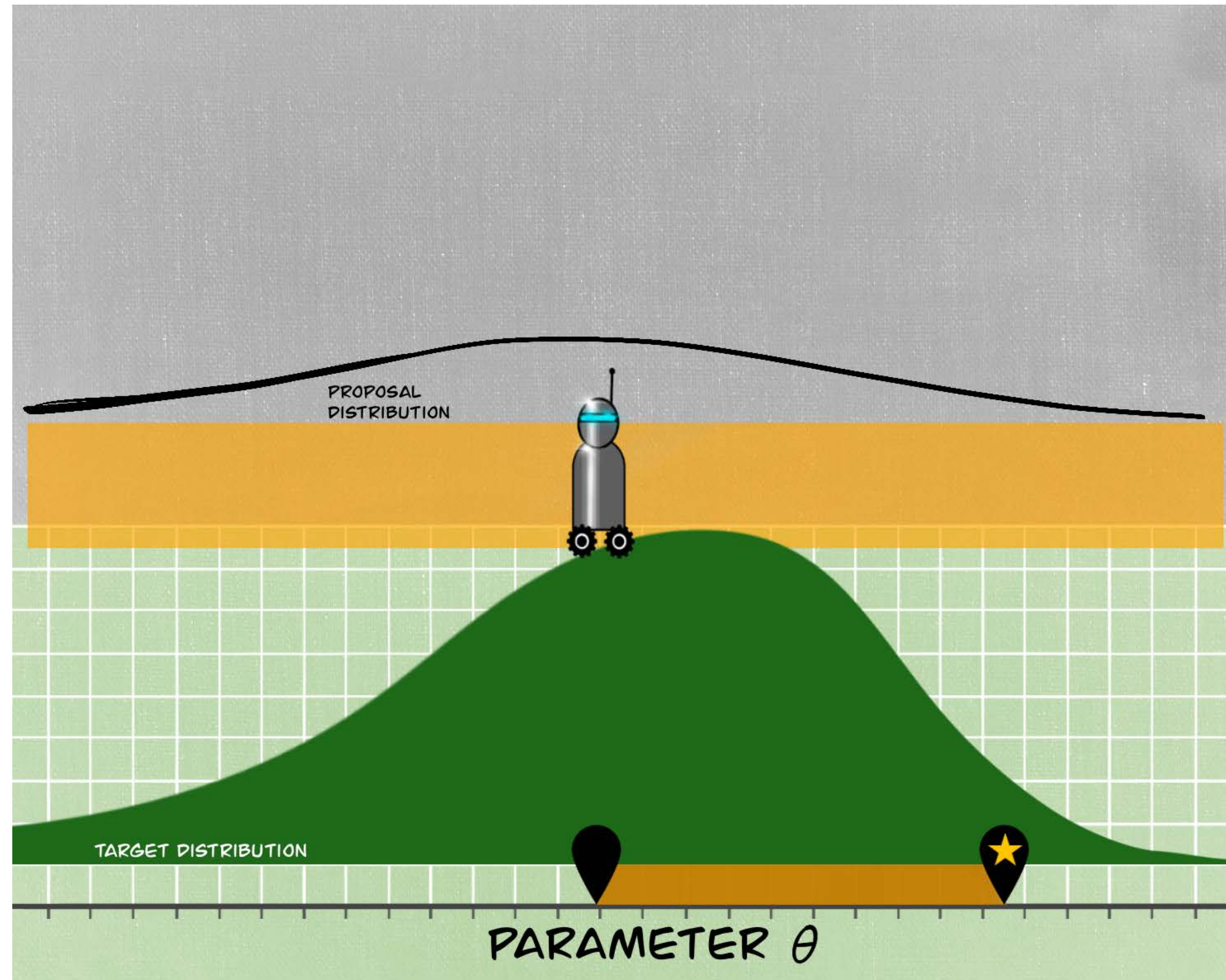
a good proposal distribution samples the target distribution effectively (i.e., "good mixing")



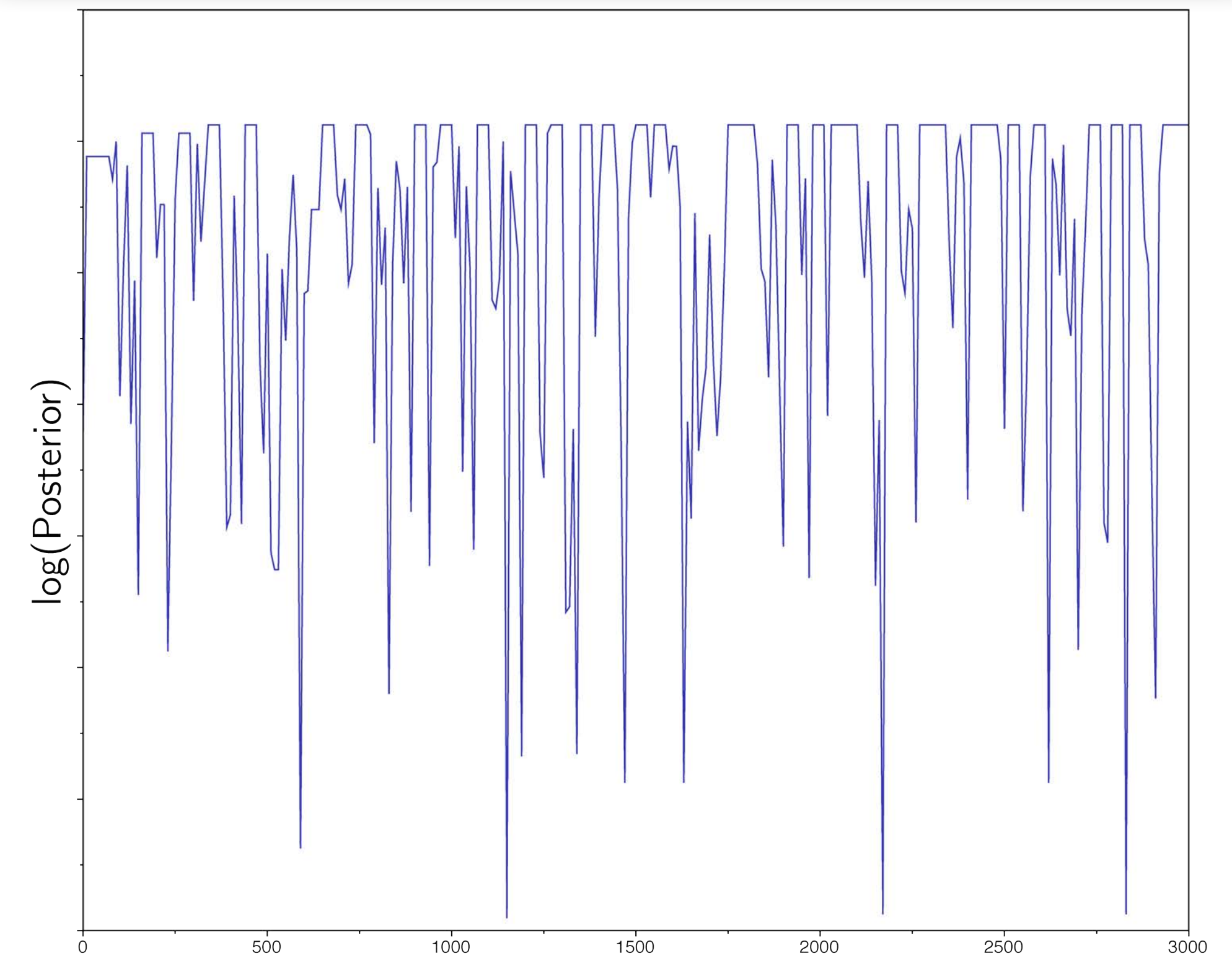
a trace plot of the sampled parameter values looks like white noise



# Target and Proposal Distributions



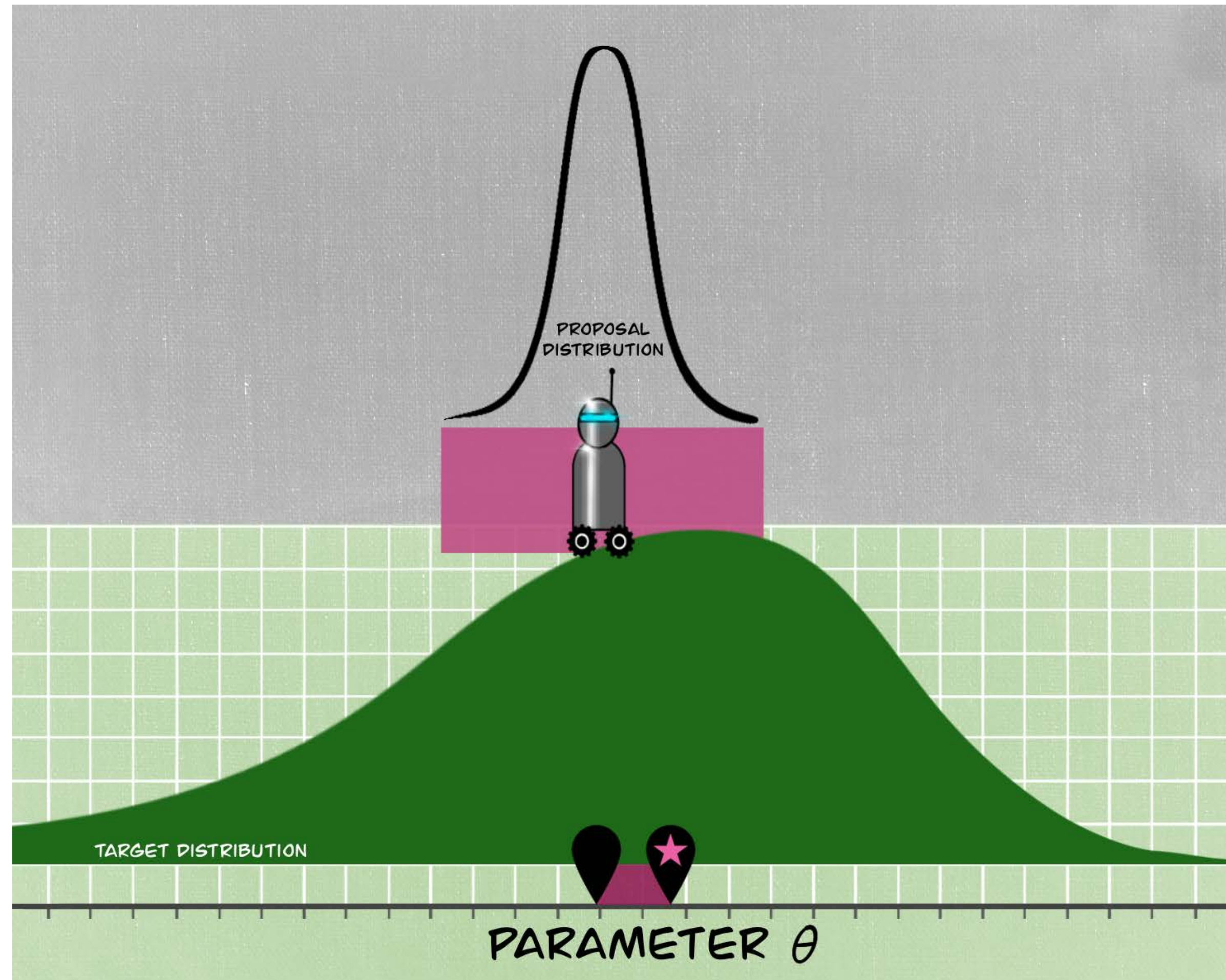
an overly bold proposal results in many rejected moves



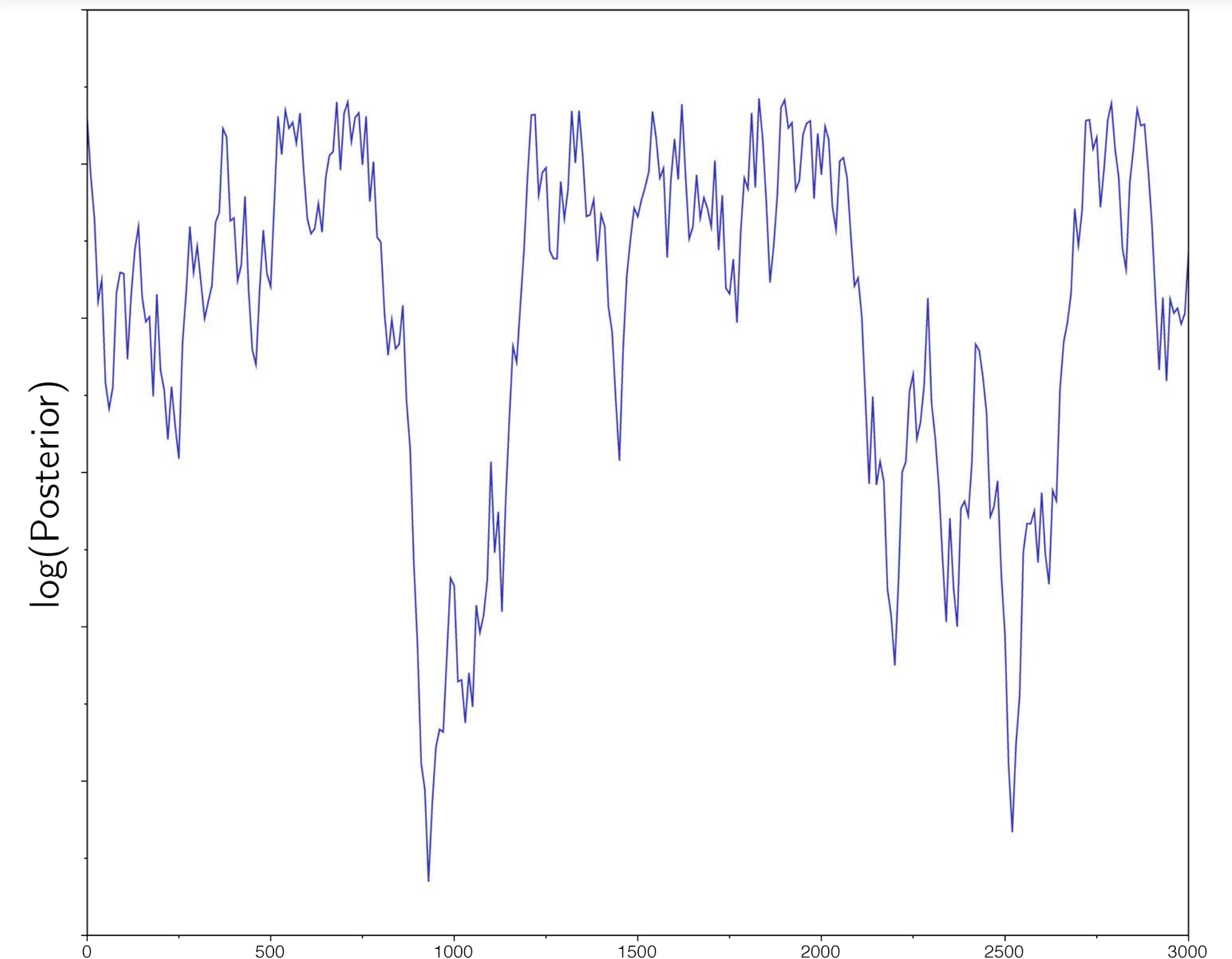
this causes the robot to get stuck, seen as plateaus in the trace plot



# Target and Proposal Distributions



a proposal distribution that only allows for baby steps results in lots of accepted moves



this causes big waves in the trace plot as the robot takes small incremental samples





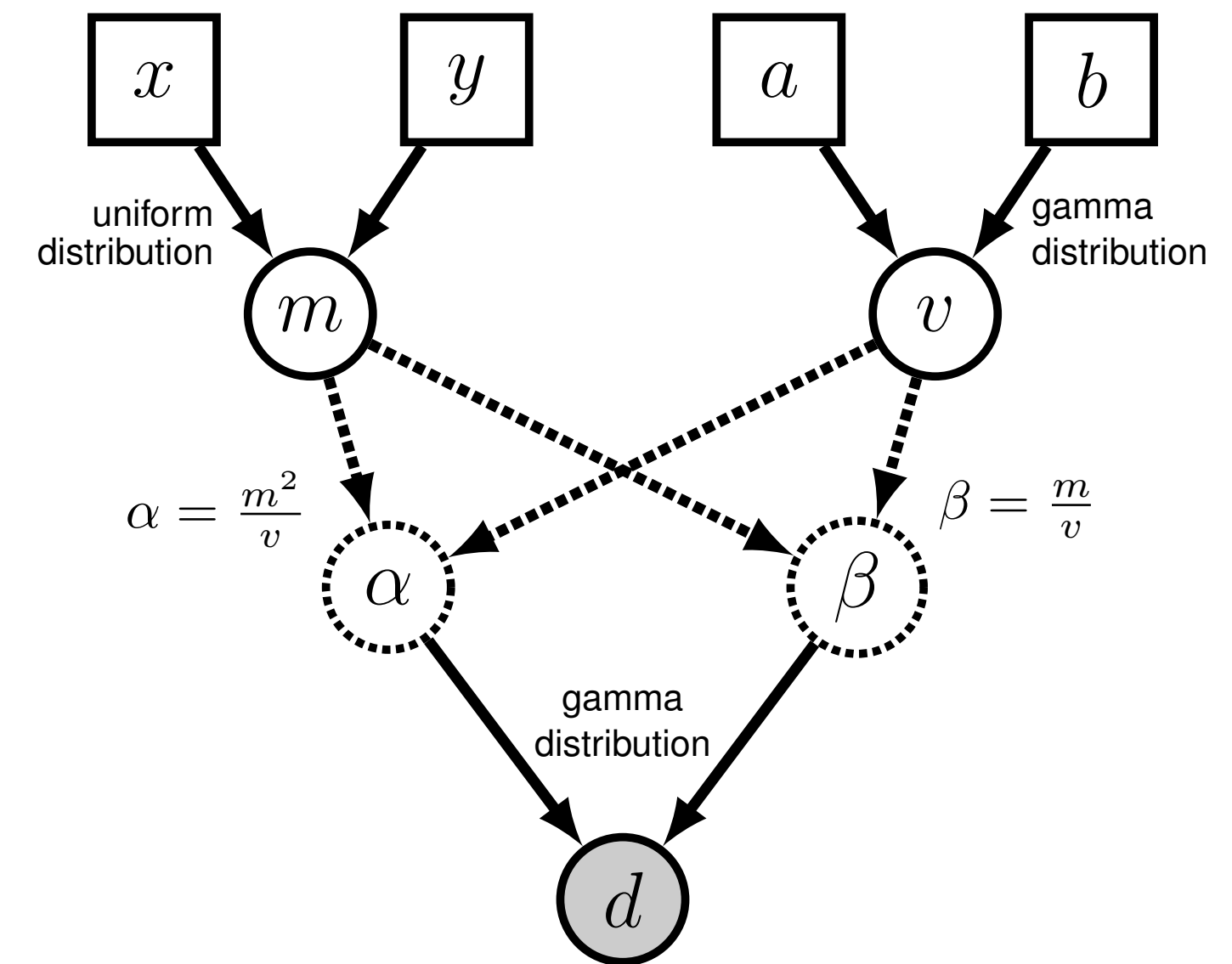
# RevBayes Demo: Archery

We can use MCMC to estimate  $m$  and  $v$

First, let's generate our observed data using simulation

```
true_accuracy = 35.0
true_variance = 4.0
true_alpha = (true_accuracy^2) / true_variance
true_beta = true_accuracy / true_variance

num_shots = 6
observed_shots = rgamma(num_shots, true_alpha, true_beta)
```



The values in `observed_shots` are data generated from the underlying distribution



# RevBayes Demo: Archery



Now we can specify the model for our new observations

```
mean ~ dnUnif(10,40)
var ~ dnGamma(20,2)

alpha := (mean * mean) / var
beta := mean / var

for(i in 1:num_shots){
  d[i] ~ dnGamma(alpha,beta)
  d[i].clamp(observed_shots[i])
}
```



# RevBayes Demo: Archery



The Rev language specifying the MCMC sampler for the hierarchical model on archery accuracy

```
mymodel = model(beta)

moves[1] = mvSlide(mean, delta=1.0, weight=3.0)
moves[2] = mvScale(var, lambda=1.0, weight=3.0)

monitors[1] = mnModel(filename="archery_mcmc_1.log", printgen=10)
monitors[2] = mnScreen(printgen=1000, mean, var)

mymcmc = mcmc(mymodel, monitors, moves)

mymcmc.burnin(generations=10000, tuningInterval=1000)

mymcmc.run(generations=30000)
```

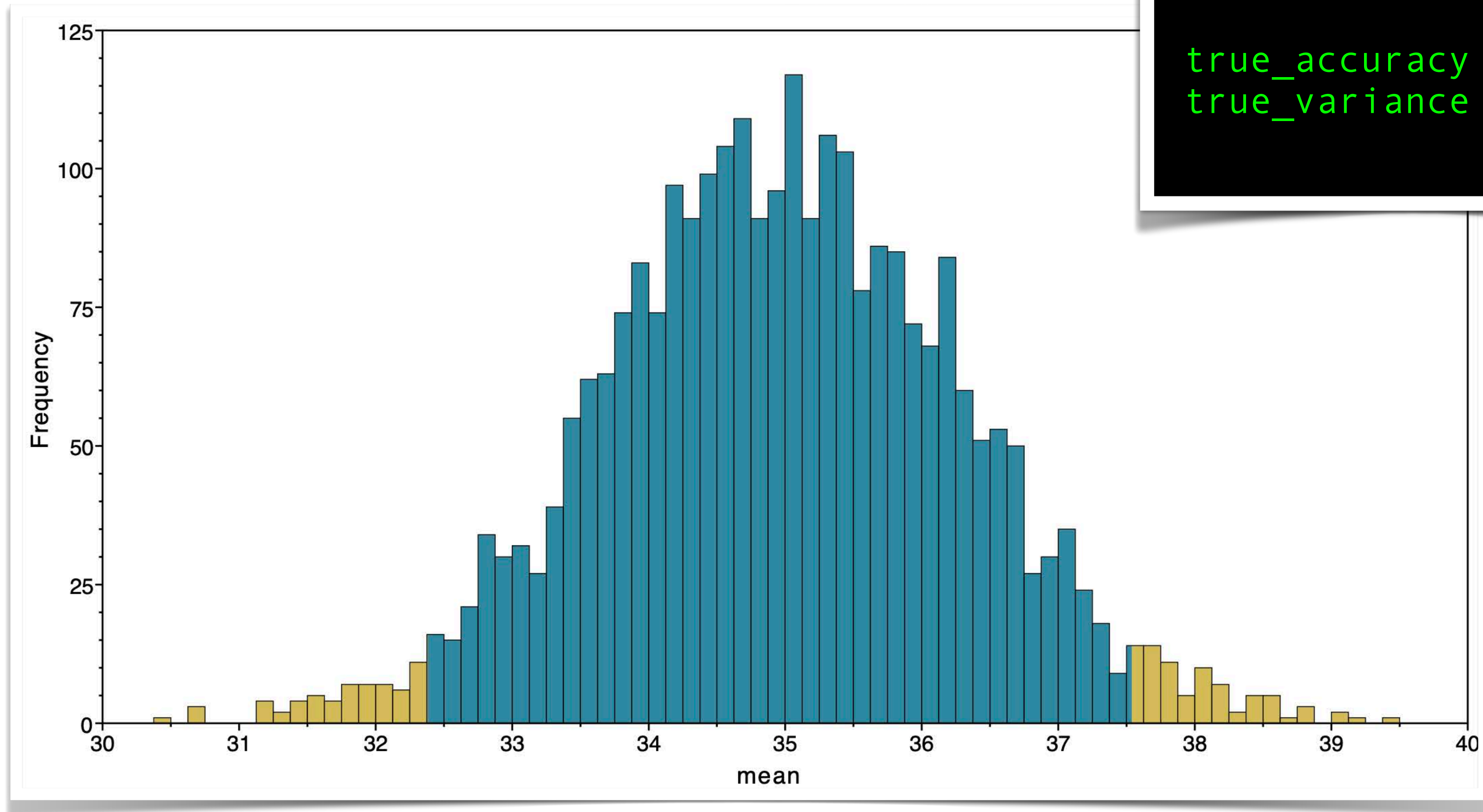
MCMC screen output



# RevBayes Demo: Archery



Summary of the MCMC sample for the mean distance from the target center



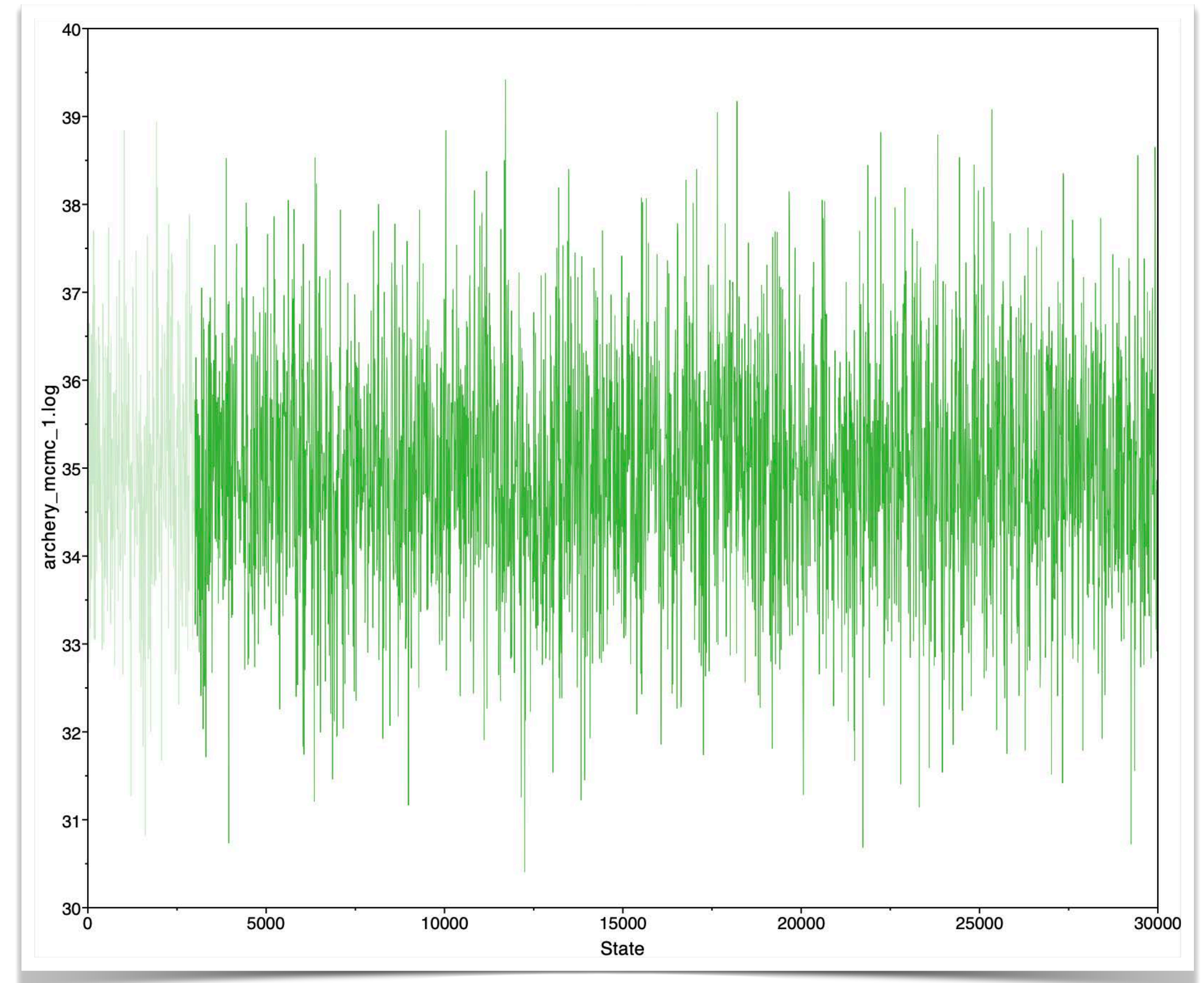
```
true_accuracy = 35.0  
true_variance = 4.0
```



# RevBayes Demo: Archery



The trace plot of the MCMC samples for the mean distance from the target center

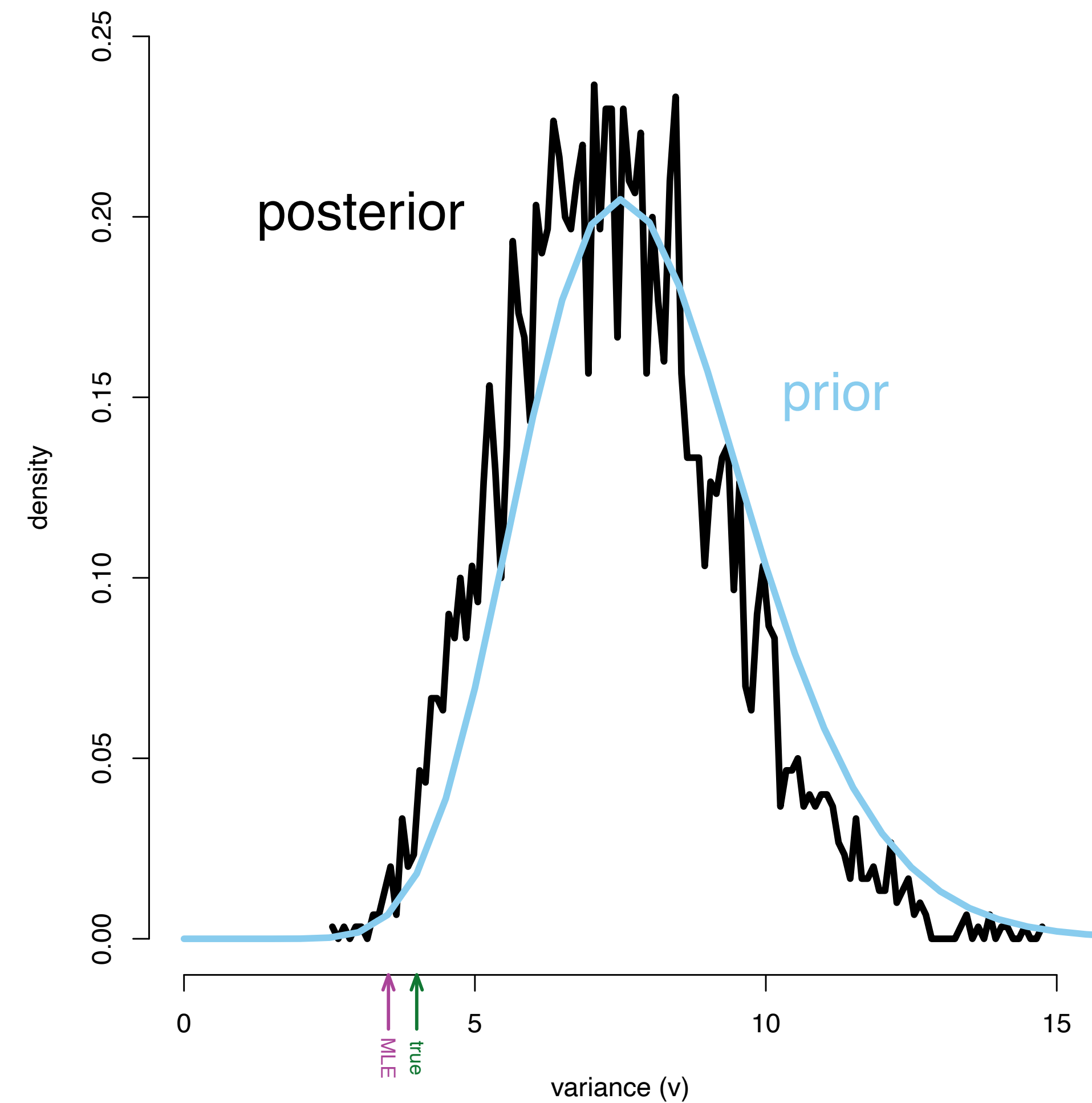




# RevBayes Demo: Archery



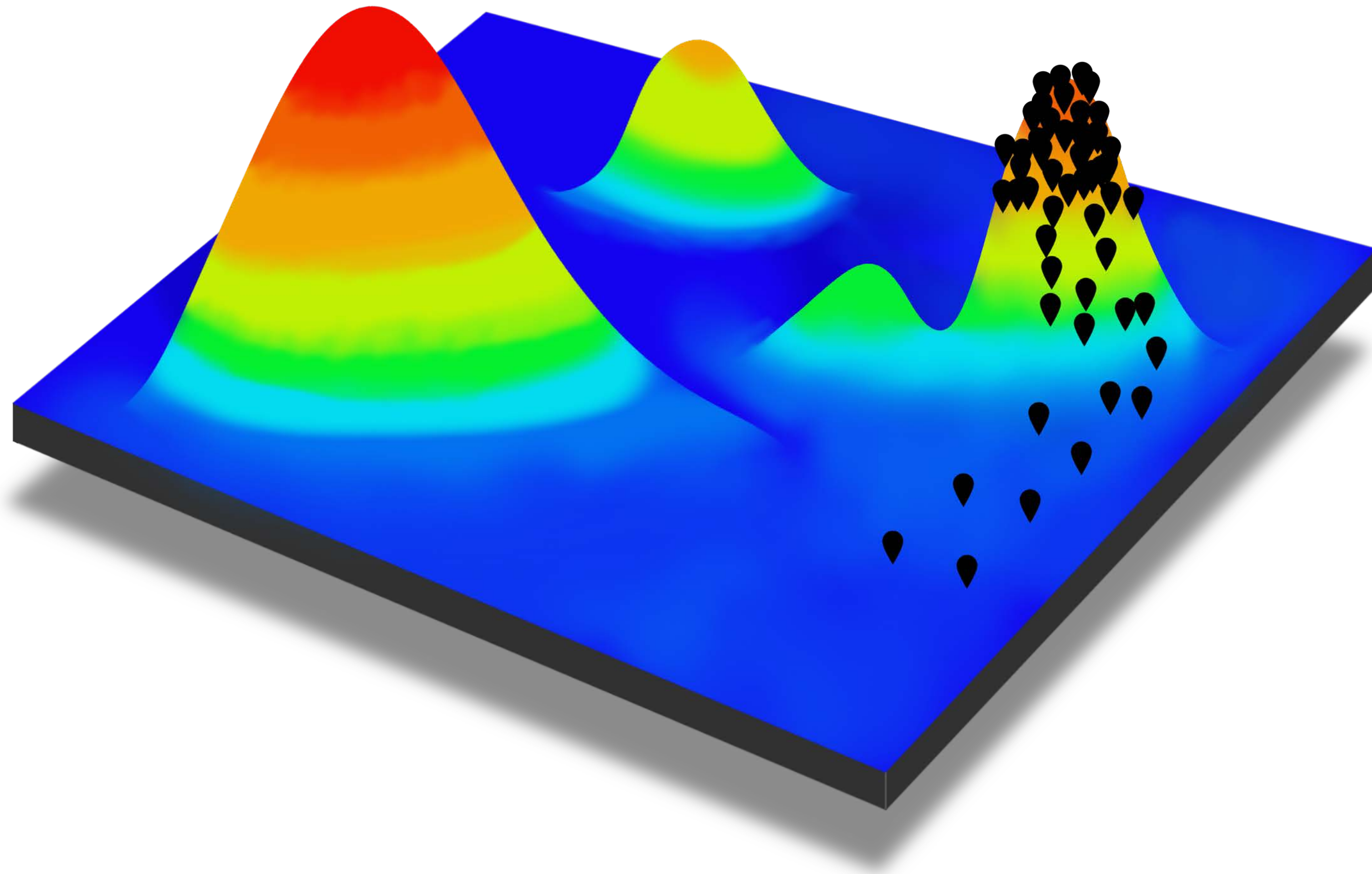
The posterior estimate for the variance ( $v$ ) is quite different from the true value (4.0) and from the highest likelihood value found by our MCMC (MLE = 3.51374)



This indicates that the prior is having a strong influence on the posterior. Why do you think that is?



# Metropolis Coupled MCMC



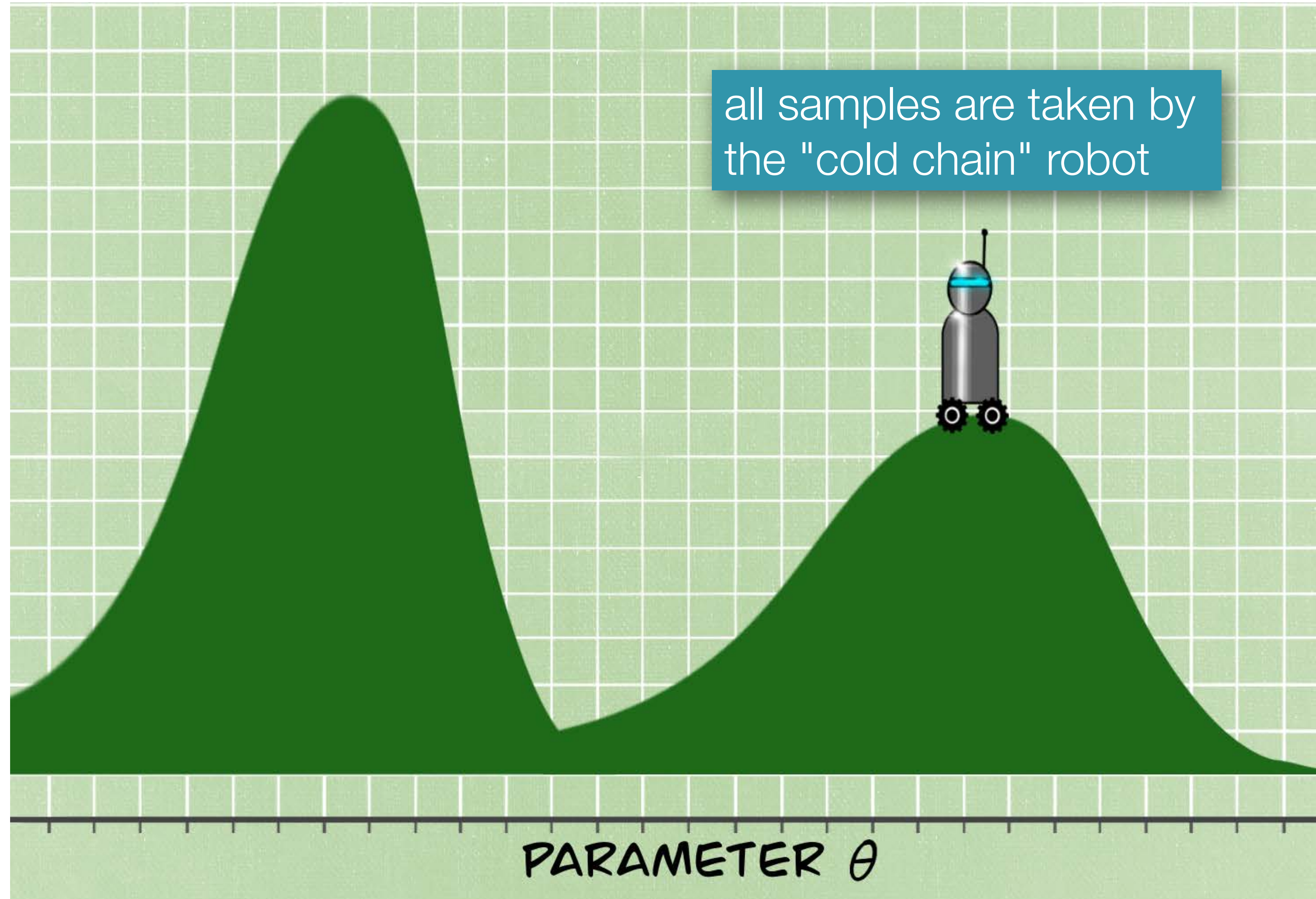
sometimes even good robots need help

MCMCMC introduces helper robots that act as scouts to explore more parameter space

Geyer, C. J. 1991. Markov chain Monte Carlo maximum likelihood for dependent data. Pages 156-163 in Computing Science and Statistics (E. Keramidas, ed.).

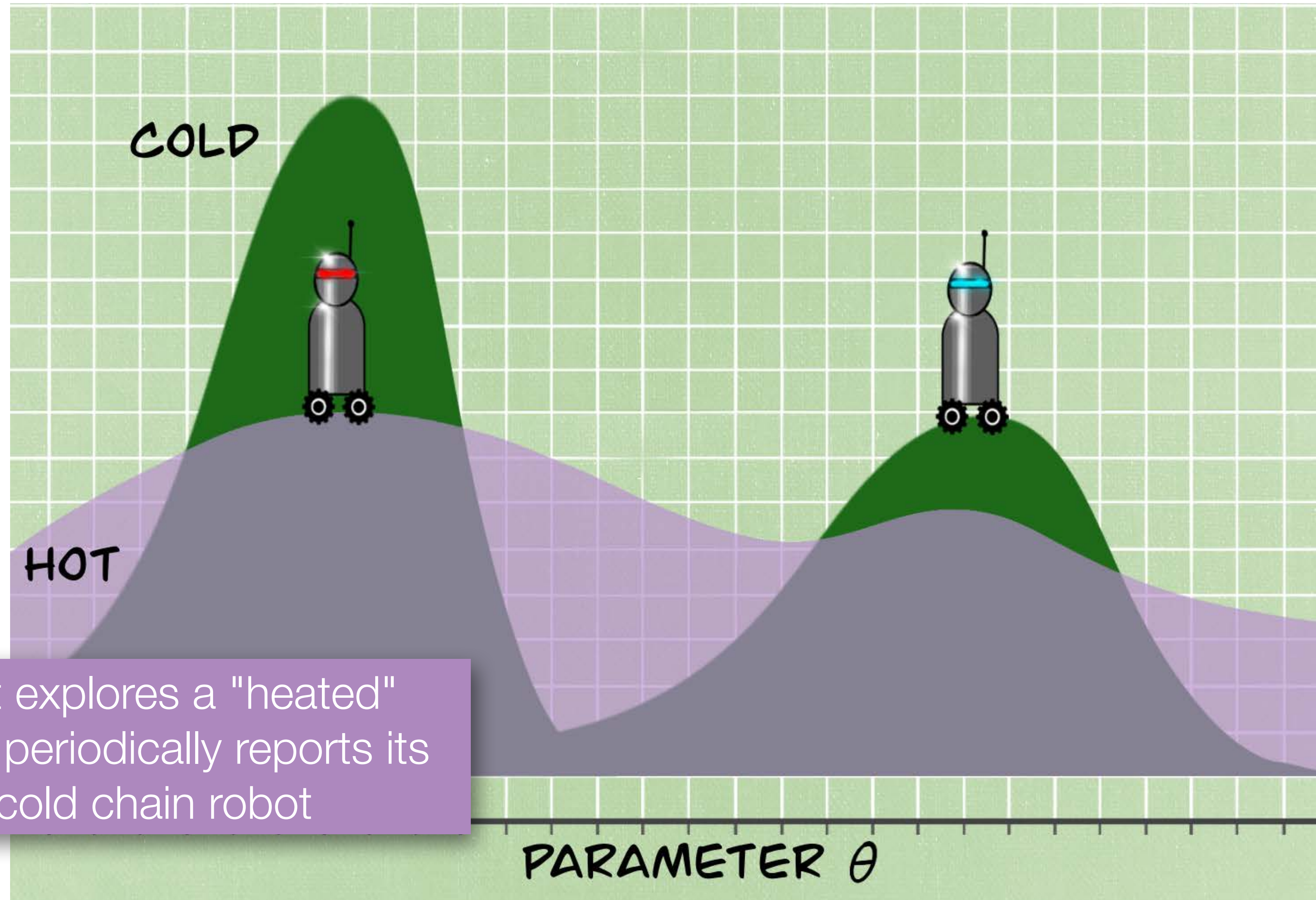


# Metropolis Coupled MCMC



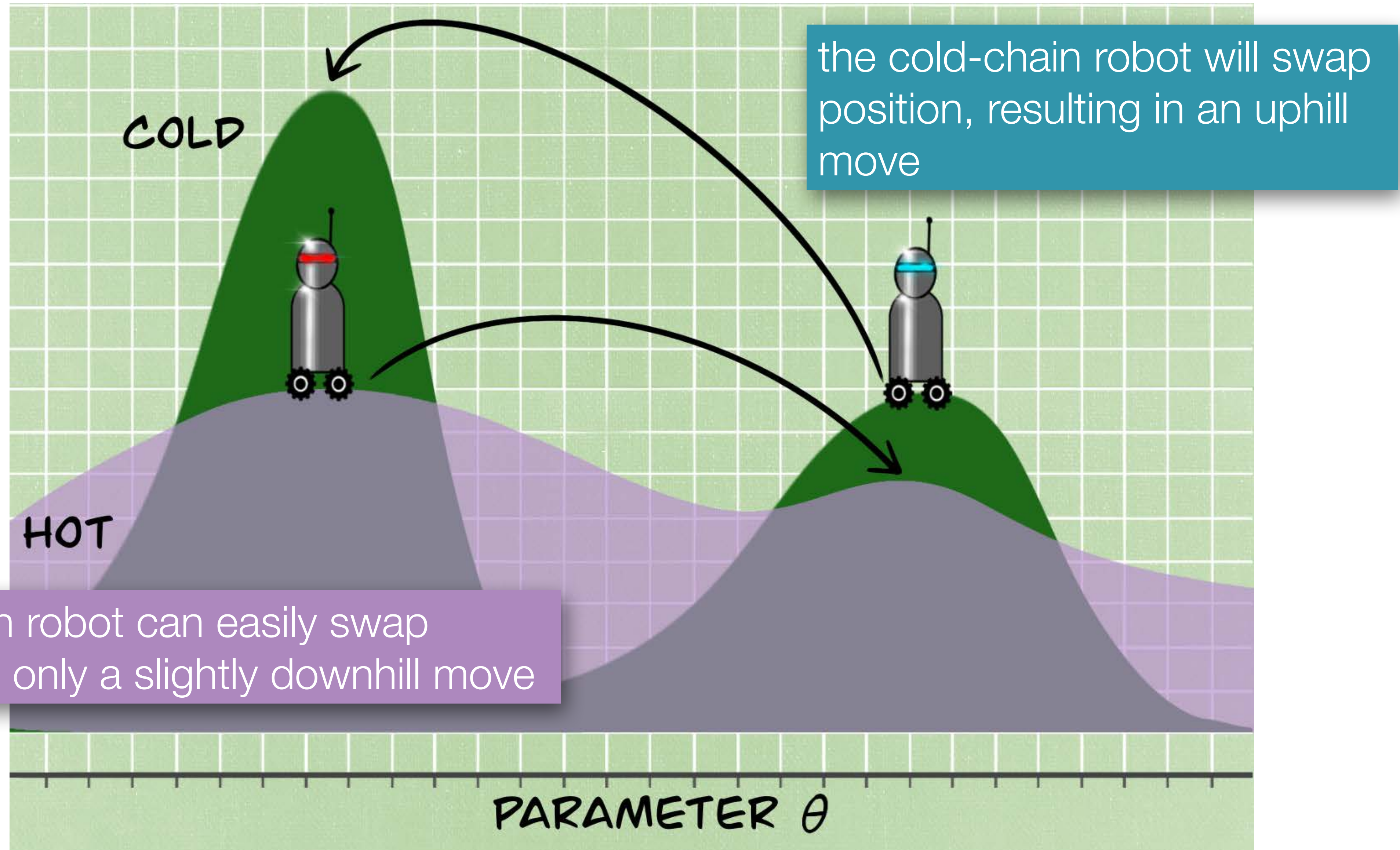


# Metropolis Coupled MCMC





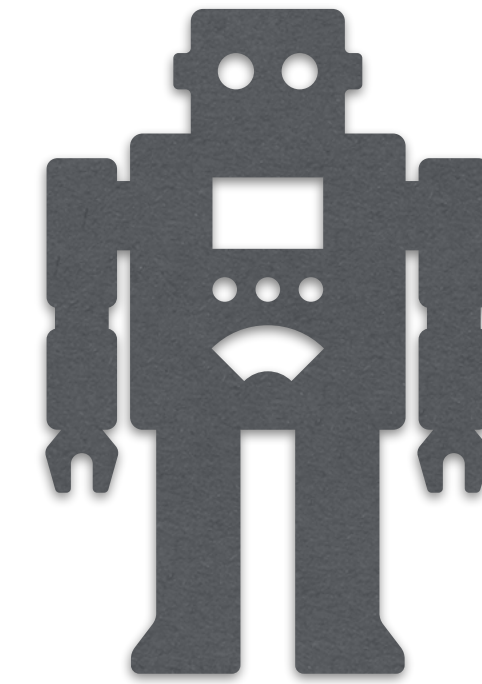
# Metropolis Coupled MCMC





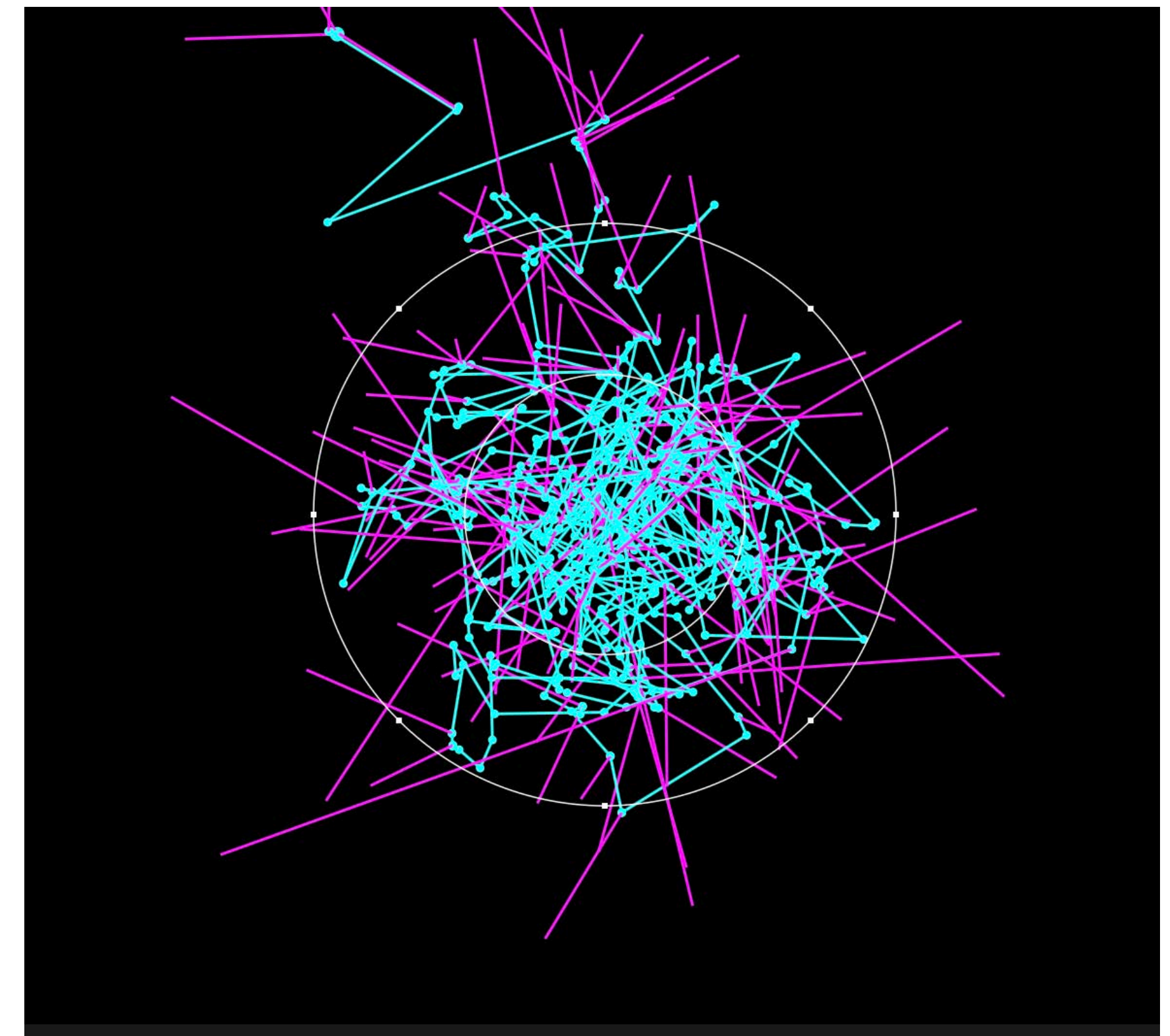
# Markov Chain Monte Carlo

Learn more about MCMC!



<https://phylogeny.uconn.edu/mcmc-robot/>

MCMCRobot, a helpful tool for learning MCMC by Paul Lewis





# Markov Chain Monte Carlo

Learn more about MCMC!

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DOI: 10.1038/s41559-017-0280-x

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## A biologist's guide to Bayesian phylogenetic analysis

Fabrícia F. Nascimento <sup>1,4\*</sup>, Mario dos Reis <sup>2</sup> and Ziheng Yang <sup>3\*</sup>

<https://thednainus.wordpress.com/2017/03/03/tutorial-bayesian-mcmc-phylogenetics-using-r/>