- 1 Appendix A: Experiments
- 2 Detailed methods for larvae temperature experiments
- 3 Experimental methodology

4 Source of Ostertagia gruehneri eggs

5 Two adult female reindeer housed at the Wildlife Research Station, Faculty of Veterinary 6 Medicine, University of Calgary, were experimentally infected with *O. gruehneri* third stage larvae 7 (L3) sourced from the Bathurst caribou herd, Northwest Territories and Nunavut, Canada (Hoar 8 2012). Fresh fecal samples were collected from the infected reindeer and refrigerated at 9 approximately 4°C. No more than 24 hours later, *O. gruehneri* eggs were isolated from the fecal 10 samples following the methods outlined in Hoar *et al.* (2012).

11 Data Collection

Following the isolation of O. gruehneri eggs from reindeer feces, forty Petri dishes of 3 cm 12 diameter were seeded with approximately 50 eggs each. These Petri dishes were then distributed 13 into incubators with 100% relative humidity and set at 5, 10, 15, 20, 25, 30, 35, and 40°C, 14 respectively, thus yielding five replicate cohorts for each temperature trial. Larval food (250 μ); 15 Hubert and Kerboeuf, 1984) was added to each Petri dish on the same day at least one egg had 16 hatched in all of the five replicates of a given temperature trial. Initially, each replicate was 17 examined daily at room temperature (~10-15 minutes per replicate) using a Leica L2 dissecting 18 microscope at 20-25x magnification, when the numbers of live individuals in each stage of 19 development (eggs, L1, L2, L3) were recorded. Larvae that were moving were hereby considered 20 to be alive, whereas immobile larvae were considered to be dead. Eggs were considered dead 21 when a breakdown of internal structure was noted or when the entire egg turned black. 22 Examination of all replicates within a given temperature treatment was reduced to every three 23 24 days once a minimum of one L3 was observed in each of the five replicates. The only exception to this sampling protocol was the 5°C treatment where, due to the slower development rate at 25 this temperature, replicates were examined every three days from the beginning of the trial. At 26 each sampling occasion, egg and larval counts were performed three times for every replicate 27

dish to minimize the effect of potential observer error in detecting eggs and larvae, in correctly
identifying the developmental stages of larvae, and in correctly assessing larval survival.
Replicates continued to be examined until no live larvae or eggs were observed, or for a
maximum of 120 days (a conservative estimate of the number of days with mean temperatures
>0°C in the Canadian low Arctic).

33 Model development

34 To extract development and mortality rates from the cohort data, we begin by describing a cohort model that tracks individuals through the pre-infective (egg, L1, L2) and infective (L3) life stages. 35 As we are primarily interested in the appearance of the infective L3, the model considers a 36 simplified life cycle where eggs, L1 and L2 are pooled into a "pre-infective" class (subscript 0), 37 and larvae in the L3 stage constitute the "infective" class (subscript 3). The model specifies the 38 expected number of individuals in the pre-infective and infective classes as a function of time and 39 temperature, and is specifically designed to account for variability in development rates among 40 individuals as well as for any mortality that may occur between the start of the experiment at 41 time $t_0 = 0$ and a given sampling day t. 42

We first determine the probability that an individual in experimental treatment *i* experiencing rearing temperature T_i is alive on day *t* but has not yet developed to the infective L3 stage. Régnière and Powell (2003) argue that in many cases the actual development time of an individual at temperature *T*, $t^*(T)$, will have a log-normal distribution with a mean that equals the cohort mean, $\tau(T)$. From this, it follows that development times to the infective stage are log-normally distributed:

$$t^*(T) \sim \log \left(-\frac{\sigma^2}{2} - \log(\rho_0(T)), \sigma^2\right), \tag{A1}$$

where the cohort mean development rate, $\rho_0(T) = 1/\tau(T)$, is the inverse of the cohort mean development time, σ is the scale parameter of the log-normal distribution, and $-\sigma^2/2$ is a correction to the mean of the lognormal so the expected development time is $1/\rho_0(T)$ (Hilborn

and Mangel 1997). Given a starting cohort of N_0 eggs, the expected number of live, pre-infective

53 individuals on sampling day *t* is

$$L_{0}(t,T_{i}) = N_{0} \underbrace{\exp(-\mu_{0}(T_{i}) t)}_{\text{not yet dead}} \left(1 - \overline{\phi \left[\frac{\left[\ln(t \rho_{0}(T_{i})) + \frac{\sigma^{2}}{2} \right]}{\sigma} \right]}_{\text{not yet developed}} \right)$$
(A2)

where $\mu_0(T_i)$ is the per capita mortality rate of pre-infectives (eggs, L1, and L2) at temperature T_i and Φ is the cumulative distribution function of the standard normal distribution. To calculate the expected number of larvae in the L3 stage on day t, we first need to account for the fact that – due to the variability in development times – different individuals will start this life stage at different times. The expected number of individuals that start the L3 stage exactly at time t, $L_{3,\text{start}}(t,T_i)$, can be obtained from the probability density function of the lognormal development distribution, discounted for any mortality occurring in the pre-infective stage:

$$L_{3,\text{start}}(t,T_i) = \frac{N_0}{t\sqrt{2\pi\sigma^2}} \exp\left[-\frac{\left(\ln(t\rho_0(T_i)) + \frac{\sigma^2}{2}\right)^2}{2\sigma^2} - \mu_0(T_i)t\right]$$
(A3)

The expected number of larvae in the infective L3 stage on sampling day *t* is obtained by integrating over all larvae that have reached the L3 stage before day *t* while discounting for any mortality that has occurred since these larvae started the L3 stage:

$$L_{3}(t,T_{i}) = \int_{0}^{t} L_{3,\text{start}}(t',T_{i}) \exp(-\mu_{3}(T_{i})(t-t') dt'$$
(A4)

64 where $\mu_3(T_i)$ is the per capita mortality rate of infectious L3 at temperature T_i .

65 Metabolic Theory of Ecology relationships

- We assessed up to four different sub-models for the temperature dependencies of $\mu_0(T)$, $\mu_3(T)$, and $\rho_0(T)$: (1) no relationship among parameters at different temperatures, (2) constant parameter value across temperatures, (3) an Arrhenius relationship that describes increasing development and mortality rates with temperature (McCoy and Gillooly 2008), or (4) a Sharpe-Schoolfield relationship that builds on the Arrhenius relationship to include the
- possibility of upper and/or lower temperature thresholds (Schoolfield et al. 1981; Molnár et al.
- 72 2013).



Figure A1. Illustration of the possible MTE relationships for mortality and development parameters (written generally as y(T)), with corresponding equations on the right. The vertical grey lines indicate the lower and upper temperature thresholds in the Sharpe-Schoolfield relationships (green and blue dashed lines). The inverse Sharpe-Schoolfield relationship is tailored to the mortality rate, which tends to increase at extreme values. Parameters are described in Table A1, and $k = 8.62 \times 10^{-5}$ eV K⁻¹ is Boltzmann's Constant.

79

As described in the main text, testing all possible combinations of five MTE relationships (constant, Arrhenius, and Sharpe-Schoolfield with lower, upper, or both thresholds) for the three parameters ($n = 3^5 = 243$ models) was not practically feasible due to the computation time involved, and also would have increased the chance of spurious significant results. We chose nine models to test based on initial parameter estimates and limitations of the data (Table A2). Notably, this suite of 9 models does not include the Sharpe-Schoolfield model with a

86 lower temperature threshold, since the lower bound of temperatures tested in experiments

was 5 °C, which is well above the lower thermal limits of this parasite. We encountered

convergence issues when trying to estimate an upper thermal bound for L3 because of low

survival to this stage at high temperatures, but include results for one of those models from

90 which the estimates were applied in the population modelling.

91 Parameter estimation

We fit the cohort model, incorporating temperature-dependent parameters $\mu_0(T)$, $\rho_0(T)$, and

 $\mu_3(T)$ in a Bayesian framework that could accommodate a hierarchical structure accounting for potential variability in parameters among replicates, latent variables, as well as the triplicate counts at each sampling occasion. We included a random effect for replicate (i.e., petri dish) on the parameters $\mu_0(T_i)$, $\rho_0(T_i)$, $\mu_3(T_i)$, and σ within each temperature trial *i*. Specifically, the parameter for replicate *k* was drawn from a log-normal distribution, e.g.,

98 $\mu_{0,i,k} \sim \log N(\text{mean } \log = \log(\mu_0(T_i)), \text{ sd} = 0.05)$. We initially tried to estimate the standard 99 deviation among replicates as a free parameter, but could not achieve convergence in that case 100 and so we assumed the standard deviation to have a fixed value of 0.05 which is relatively small 101 but nonetheless allowed the flexibility to capture observed differences in survival and mortality 102 among replicates (Figure A2).

Due to counting error, the initial number of eggs in each petri dish was not known and so it was treated as a latent variable to be estimated for each temperature, *i*, and replicate, *k*, with prior distribution $N_{0,i,k} \sim \log Normal(\log(50) - 0.3^2/2, 0.3)$, with an expected value of 50 eggs.

The likelihood of the pre-infective and infective counts from day 0 to day 120 was calculated assuming Poisson count error for the triplicate observations, with an expected count equal to the prediction from the cohort model for the given day, temperature, and replicate.

We used the MCMC software JAGS (Plummer 2003), and implemented the fitting via R (R Core
 Team 2021) using the packages rjags (Plummer 2019) and dclone (Sólymos 2010). R code for

- 111 the analysis is available at *https://github.com/sjpeacock/OsterBou_pop*. An example of a
- model function for Model 3 from Table 2 is provided below on pages 7-8.
- 113 We applied relatively uninformative priors on all MTE parameters, but constrained to
- 114 biologically reasonable values (Table A1).

115 Table A1. The priors on MTE hyperparameters used in the Bayesian parameter estimation (see Figure A1 for

116 equations).*

Hyperparameter	Description	Units	Prior distribution
${\mathcal{Y}}_0$	Parameter value at	d ⁻¹ (mortality	log-normal(log(0.5), 1)
	standardization	rate) or d	
	temperature T_0 = 15 °C	(development	
		time)	
Ε	Activation energy	eV	log-normal(log(0.65), 0.5)
E_H	Upper inactivation	eV	log-normal(log(3.25), 1)
	energy		
T_H	Upper temperature	°C	normal(23, 3)
	threshold		

* We report all temperatures in °C for ease of interpretation, with the conversion to K occurring in the model code.

In total, we had 2,520 unique sampling instances (i.e., temperature, petri dish, and day 118 combinations) with triplicate counts and two larval stages, yielding 15,120 data points. Due to 119 the large amount of data, the hierarchical (non-independent) structure of both the data and the 120 121 model, and the complexity of the model, we were concerned about overfitting. Thus, we evaluated models by holding out one of the five replicates when estimating parameters, using 122 the remaining four replicates as the training data, and then calculating how well the fitted 123 model predicted the hold-out replicate (i.e., the likelihood of the hold-out data given the model 124 fitted to the training data). This was repeated using each of the five replicates as the hold-out 125 data, and then taking the average likelihood over the five hold-out replicates. We took the 126 model with the lowest average negative log-likelihood among the validation sets as the best 127

- model and re-fitted that model to the entire dataset to yield the best parameter estimates for
- inference and prediction.

130	***************************************
131	# Model 3: A A SSU
132	**********************
133	model3 <- function(){
134	#
135	# Parameter models
136	# Note: variance in prior distributions specified as precision = sd^{2}
137	#
138	$\frac{1}{2}$ For each parameter (10, 11, rbo), draw base value (a) and activation
139	
140	for(in 1.3)
1/1	$sii \simeq diporm(log(0.5), pow(12))$
1/12	$ \begin{aligned} \text{Efi} &\simeq \text{dim}(\log(0.5), \operatorname{pow}(4, -2)) \\ \text{Efi} &\simeq \text{dim}(\log(0.5), \operatorname{pow}(2, -2)) \end{aligned} $
1/2	د المالي الم
143	f Set up a threshold to zero for the two Arrhenius relationships
144	for (i = 1.2)
145	
140	
147	in[]] <- 0
148	}
149	# Draw upper thresholds for the SS relationship of ul
150	Eh[3] ~ dinorm(log(3.25), pow(1, -2))
151	Th[3] ~ dnorm(23, pow(3, -2))
152	# Calculate temperature-specific mean parameter value
153	for(i in 1:nT){ # for each temperature treatment
154	for(j in 1:2){ # for the two Arrhenius relationships (u0, u1)
155	p.mean[i,j] <- a[j]*exp(-E[j]/(8.62*10^-5)*(1/(T.obs[i]+273.15)-1/(15+273.15)))}
156	
157	# For the SS relationships (rho)
158	p.mean[i,3] <- a[3]*exp(-E[3]/(8.62*10^-5)*(1/(T.obs[i]+273.15)-1/(15+273.15)))*(1+exp(Eh[3]/(8.62*10^-
159	5)*(-1/(T.obs[i]+273.15)+1/(Th[3]+273.15))))^(-1)
160	}
161	
162	# Sigma - constant across temperature; no metabolic theory underpinning
163	sigma.log ~ dnorm(0, 0.01)
164	for(i in 1:nT){p.mean[i,4] <- exp(sigma.log)}
165	
166	# Draw random model parameters (u0, u1, rho, sigma) for each temp and replicate
167	for(i in 1:nT){ # for each temperature
168	for(k in 1:nk){ # for each replicate
169	for(j in 1:4){ # for each parameter
170	p.rep[i,j,k] ~ dlnorm(log(p.mean[i,j]), sigp^(-2))}}}
171	
172	# Draw starting number of eggs in each temp and replicate
173	for(i in 1:nT){ # for each temperature
174	<pre>for(k in 1:nk){ # for each replicate</pre>
175	N0[i,k] ~ dInorm(log(50)-1/2*(sigN0^2), 1/(sigN0^2))}}
176	
177	#
178	# Process model
179	#
180	<pre>for(i in 1:nT){ # for each temperature</pre>
181	for(k in 1:nk){ # for each replicate
182	
183	# Pre-infectives
184	for(j in 1:nt){
185	N[i,1,j,k]<-N0[i,k]*exp(-p.rep[i,1,k]*t[j])*(1-
186	phi((log(t[j]*p.rep[i,3,k])+p.rep[i,4,k]^2/2)/p.rep[i,4,k]))
187	}

Appendix A: Experiments

188	
189	# Infectives that develop at time "jj"
190	for(j in 1:nt){
191	N1_start[i,j,k]<-N0[i,k]/(t[j]*sqrt(2*3.141593*p.rep[i,4,k]^2))*exp(-
192	1/(2*p.rep[i,4,k]^2)*(log(t[i]*p.rep[i,3,k])+p.rep[i,4,k]^2/2)^2-p.rep[i,1,k]*t[j])
193	}
194	#and survive to time j
195	for(i in 1:nt){
196	for(jj in 1:j){
197	N1 surv[i,j,k,jj]<-N1 start[i,jj,k]*exp(-p.rep[i,2,k]*(t[i]-t[jj]))
198	}
199	# Summed from 1:i
200	N[i,2,j,k]<-sum(N1 surv[i,j,k,1:j]*dt)
201	}#end i
202	}}
203	
204	#
205	# Data model
206	#
207	<pre># n_obs[i,j,k,l,s] is the observed number of larvae alive in</pre>
208	# temperature i, time j, replicate k, count l, and stage s
209	
210	<pre>for(i in 1:nT){ #for each temperature</pre>
211	<pre>for(k in 1:nk){ # for each replicate</pre>
212	<pre>for(l in 1:3){ # for each of three counts</pre>
213	
214	# Likelihood of initial number of eggs
215	n_obs[i,1,k,l,1] ~ dpois(N0[i,k])
216	
217	# Likelihood of preinfective and infective stages
218	# at each timestep
219	for(s in 1:2){ # for each stage
220	<pre>for(j in 2:nt_obs[i,k]){ # for each timestep</pre>
221	n_obs[i,j,k,l,s] ~ dpois(max(10^-10, N[i,s,ind[i,j,k],k]))
222	}}}}
223	}# end model

224

225 Supplemental results

Table A2. Model comparison statistics for the MTE models fit to experimental data of *O*.

gruehneri larval mortality and development at seven temperature treatments from 5 – 35 °C.

The negative log-likelihoods of the training data (four replicates) and of the validation data (one

replicate) were averaged across the five different hold-out replicates, and the overall negative

log-likelihood is from the given model fitted to all five replicates (which is why it is a larger

number than the sum of the other two values).

					Negative log-likelihood			
					mean	mean		-
Model #	μ_0^{1}	μ_3	$ ho_0$	np²	training	validation	overall	AIC
model1	I	I	I	22	15675.2	4396.5	20962.9	41970
model5 ³	Α	С	SSU	8	17495.7	4475.8	22663.3	45343
model3	А	А	SSU	9	17520.9	4486.1	22689.1	45396
model7	SSU	А	SSU	11	17519.8	4487.7	22687.1	45396
model10 ⁴	А	SSU	SSU	11	17507.7	4503.1	22777.6	45577
model4	А	С	А	6	19344	4897.4	25148.1	50308
model6	SSU	А	А	9	19370.1	4909.4	25187.7	50393
model2	А	А	А	7	19370.6	4909.6	25189.1	50392
model8	SSU	С	А	8	67626.2	11350.8	25143.3	50303
model9	SSU	С	SSU	10	69712.8	19447.9	440627.6	881275

232 1. Temperature relationships considered were: I = parameter estimated independently at each temperature (*n* = 7

233 parameters); C = constant parameter value across temperature (n = 1 hyperparameter); A = Arrhenius relationship

234 (*n* = 2 hyperparameters), SSU = Sharpe-Schoolfield relationship with an upper thermal bound (n = 4

235 hyperparameters).

- 236 2. The number of (hyper)parameters in the model. For each model, the standard deviation in the log-normal
- 237 distribution of development times, σ , was also estimated.
- 238 3. Model5 in bold was the best-fit MTE model, with the lowest average validation NLL.
- 239 4. There were some convergence issues for this model, see Table A5.
- 240
- Table A3. Parameter estimates¹ for pre-infective and infective mortality and pre-infective development of *O*.
- 242 gruehneri larvae, estimated separately for each of seven temperature treatments from 5 35 °C.

Temperature	μ_0	μ_3	$ ho_0$
5 °C	0.017 (0.016, 0.019)	0.011 (0.008, 0.014)	0.012 (0.012, 0.013)
10 °C	0.024 (0.023, 0.026)	0.046 (0.042, 0.049)	0.023 (0.022, 0.024)
15 °C	0.103 (0.097, 0.109)	0.024 (0.020, 0.028)	0.021 (0.020, 0.023)
20 °C	0.102 (0.095, 0.110)	0.018 (0.016, 0.019)	0.059 (0.056, 0.063)
25 °C	0.307 (0.288, 0.327)	0.008 (0.006, 0.010)	0.077 (0.072, 0.083)
30 °C	0.469 (0.436, 0.504)	0.072 (0.049, 0.099)	0.086 (0.077, 0.095)
35 °C	0.456 (0.425, 0.489)	1.739 (0.080, 51.570) ²	0.007 (0.001, 0.042) ²

243 1. The estimate for σ was 0.441 (0.425, 0.457).

244 2. Estimates in red did not converge, likely because there were no infective larvae counted in any replicate at 35
245 °C.

246 Table A4. Hyperparameter estimates¹ for MTE relationships describing pre-infective and infective mortality and

247	pre-infective develo	nment of O	<i>aruehneri</i> larvae acro	oss seven temr	peratures (Table S2)	1
24/	pre-intective develo	prine in or o.	grachilen laivae acit	Jaa seven temp	Jeratures (Table 52	.,

	Parameter				
Hyperparameter	μ_0	μ_1	ρ		
y_0	0.068 (0.066, 0.070)	0.022 (0.021, 0.023)	0.032 (0.031, 0.033)		
Ε	0.884 (0.865, 0.902)		0.686 (0.659, 0.713)		
E_H			7.957 (3.781, 13.390)		
T_H			30.568 (30.202, 31.170)		

248 1. The estimate for σ was 0.436 (0.421, 0.451).

249 Estimation of lower temperature thresholds for population modelling

As described in the main text, we were unable to estimate freezing survival of pre-infective and 250 infective larvae from experiments because the minimum temperature tested was 5 °C. We 251 therefore estimated lower thresholds on mortality parameters by fixing other hyperparameters 252 in the Sharpe-Schoolfield relationship according to their estimated values in Table A5, and 253 estimated just the lower inactivation energy (E_L) and lower temperature threshold (T_L) using 254 data on freezing survival of Marshallagia marshalli larvae reported in Figure 3 of Aleuy et al. 255 256 (2020). Specifically, we fit exponential curves to the mean survival over time of eggs, L1, and L3 at -9 °C and -20 °C, yielding the mortality rates for *M. marshalli* at each of those temperatures 257 (Figure A2). We then estimated the two unknown parameters in the Sharpe-Schoolfield 258 relationship for pre-infective and infective stages from these two points for L1 and L3, 259 respectively, assuming a log-normal error distribution (Figure A3, Table A5). We used the 260 mortality rates for L1, rather than eggs, to estimate the lower bounds for the pre-infective 261 larvae because M. marshalli eggs are less sensitive to freezing and thus not the limiting stage 262 determining lower thermal tolerance (Figure A2). 263





Figure A2. The proportion of eggs, L1, and L3 stages of *Marshallagia marshalli* surviving over 80 days at (A) -9 °C,
(B) -20 °C, (C) -35 °C, taken from Aleuy et al. (2020). The solid lines are the fitted exponential curves, with the
estimated mean mortality rates indicated on the figure.

269

265

270

Table A5. Hyperparameters assumed in population modelling for MTE relationships describing pre-infective and
 infective mortality and pre-infective development of *O. gruehneri*.¹

	Parameter			
Hyperparameter	μ_0	μ_3	$ ho_0$	
<i>y</i> ₀	0.068 (0.066, 0.070)	0.0211 (0.020, 0.022)	0.032 (0.031, 0.033)	
Ε	0.884 (0.865, 0.902)	0.208 (0.171, 0.246)	0.686 (0.659, 0.713)	
E_H		3.554 (3.217, 3.970) ²	7.957 (3.781, 13.390)	
T_H		27.6 (26.7, 28.3) ²	30.568 (30.202, 31.170)	
E_L	-3.358 (-4.279, -2.406)	-19.318 (-20.596, - 18.084)		
T_L	2.928 (2.759, 3.088)	3.409 (1.961, 5.136)		

- 273 1. Hyperparameters in red were estimated from data on *Marshallagia marshalli* from Aleuy et al. (2020).
- 274 Hyperparameters in blue were estimated from Model 10. Note that the hyperparameters for
- 275 μ_0 and ρ_0 were not significantly different between Model 5 and Model 10 (Table A2).
- 276 2. Parameters showed some convergence issues with $\hat{R} > 1.1$.

277



279 Figure A3. Full MTE curves for the mortality rates (d⁻¹) of (A) pre-infective stage and (B) infective stage larvae,

280 including lower thermal bounds estimated from *M. marshalli* data (red points).



281

Figure A4. Instantaneous rates of (A) pre-infective mortality (μ_0) (B) development from pre-infective to infective (ρ_0), (C) infective mortality (μ_1), and (D) the proportion of eggs that survive to the infective stage ($\exp(-\mu_0 \rho_0)$) throughout the year at x = 660 km (approximate calving grounds). For each panel, rates are shown for current temperatures (black) and under RCP 2.6 (blue) and RCP 8.5 (red) climate change scenarios. These rates were calculated using the relationships from Fig. 5 and temperature curve from Fig. 4C.

288



- 289 Figures A5. Plot of experimental data showing the number of pre-infective (left) and infective (right) larvae alive
- from time 0 to 120 days at seven temperature treatments (rows) and 5 replicates within each temperature
- treatment (symbols). The lines show three different model predictions: the independent fits to each temperature
- treatment (model 1; red), the best-fitting MTE model (model 5; orange) and the assumed SSU model used in the
- 293 host-parasite population modelling (Assumed; turquoise). The following plots show the same data, but highlight
- 294 each replicate in turn (black points) and show the replicate-specific predictions using replicate-specific parameters
- 295 from the hierarchical model. The average model, corresponding to that shown above, is also shown in a lighter
- 296 dashed line.











302 References

Aleuy, O.A., Peacock, S., Hoberg, E.P., Ruckstuhl, K.E., Brooks, T., Aranas, M., and Kutz, S. 2020. 303 Phenotypic plasticity and local adaptation in freeze tolerance and its implications for 304 parasite dynamics in a changing world: the case of Marshallagia marshalli. Int. J. Parasitol. 305 **50**: 161–169. Australian Society for Parasitology Inc. doi:10.1016/j.ijpara.2019.12.004. 306 Hilborn, R., and Mangel, M. 1997. The Ecological Detective: Confronting Models with Data. 307 Princeton University Press, Princeton, New Jersey, USA. 308 Hoar, B.M. 2012. Ecology and Transmission Dynamics of Ostertagia gruehneri in Barrenground 309 Caribou. Available from https://prism.ucalgary.ca/handle/11023/289. 310 McCoy, M.W., and Gillooly, J.F. 2008. Predicting natural mortality rates of plants and animals. 311 Ecol. Lett. 11(7): 710–716. doi:10.1111/j.1461-0248.2008.01190.x. 312 Molnár, P.K., Kutz, S.J., Hoar, B.M., and Dobson, A.P. 2013. Metabolic approaches to 313 understanding climate change impacts on seasonal host-macroparasite dynamics. Ecol. 314 Lett. **16**: 9–21. doi:10.1111/ele.12022. 315 Plummer, M. 2003. JAGS: A Program for Analysis of Bayesian Graphical Models using Gibbs 316 Sampling. In Proceedings of the 3rd International Workshop on Distributed Statistical 317 Computing. Edited by F.L. Kurt Hornik and A. Zeileis. Vienna, Austria. pp. 1–10. Available 318 from http://www.ci.tuwien.ac.at/Conferences/DSC-2003/. 319 Plummer, M. 2019. rjags: Bayesian graphical models using MCMC. Version 4-10. Available from 320 http://cran.r-project.org/package=rjags. 321 R Core Team. 2021. R: A Language and Environment for Statistical Computing. Vienna, Austria. 322 323 Régnière, J., and Powell, J. 2003. Animal Life Cycle models. *In* Phenology: an integrative environmental science. Edited by M. Schwarz. Springer Berlin, Berlin, Germany. pp. 295-324 315. 325 Schoolfield, R.M., Sharpe, P.J.H., and Magnuson, C.E. 1981. Non-linear regression of biological 326 temperature-dependent rate models based on absolute reaction-rate theory. J. Theor. 327 Biol. 88(4): 719–731. doi:10.1016/0022-5193(81)90246-0. 328 Sólymos, P. 2010. dclone: Data Cloning in R. R J. 2(2): 29–37. Available from http://journal.r-329 330 project.org/.

331