Evidence for Large-Scale Gene-by-Smoking Interaction Effects on Pulmonary Function

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Abstract

Background: Smoking is the strongest environmental risk factor for reduced pulmonary function. The genetic component of various pulmonary traits has also been demonstrated, and at least 26 loci have been reproducibly associated with either FEV₁ (forced expiratory volume in 1 second) or FEV_1/FC (FEV₁/forced vital capacity). Although the main effects of smoking and genetic loci are well established, the question of potential gene-by-smoking interaction effect remains unanswered. The aim of the present study was to assess, using a genetic risk score approach, whether the effect of these 26 loci on pulmonary function is influenced by smoking.

Methods: We evaluated the interaction between smoking exposure, considered as either ever vs. never or pack-years, and a 26 SNPs genetic risk score in relation to $FEV₁$ or $FEV₁/FVC$ in 50 047 participants of European ancestry from the CHARGE and SpiroMeta consortia.

Results: We identified an interaction ($\beta_{int} = -0.036$, 95% confidence interval, -0.040 – -0.032, *P*=0.00057) between an unweighted 26 SNPs genetic risk score and smoking status (ever/never) on the FEV₁/FVC ratio. In interpreting this interaction, we showed that the genetic risk of falling below the FEV**1**/FVC threshold used to diagnose chronic obstructive pulmonary disease is higher among ever smokers than among never smokers.

Conclusions: This study highlights the benefit of using genetic risk scores for identifying interactions missed when studying individual SNPs, and shows for the first time that persons with the highest genetic risk for low FEV₁/FVC may be more susceptible to the deleterious effects of smoking.

Key words

FEV₁/FVC, smoking, gene-environment interaction, genetic risk score

Key messages

- Spirometric measures of pulmonary function are influenced by both smoking and genetics. This paper identified a genetic risk score-by-ever smoking interaction on FEV₁/FVC (forced expiratory volume in 1 second / forced vital capacity).
- \bullet In individuals of European ancestry, the reduction in FEV₁/FVC due to smoking was greater among individuals who are genetically predisposed to lower FEV₁/FVC ratio.
- Genetic risk score-by-ever smoking interaction can allow the identification of subgroups in the population whose genetic background makes them more susceptible to the deleterious effects of smoking.

Introduction

Spirometric measures of pulmonary function, such as the forced expiratory volume in one second (FEV₁) or its ratio with the forced vital capacity (FEV₁/FVC), form the basis of the diagnosis of chronic obstructive pulmonary disease (COPD). [1-3](#page-23-0) Pulmonary function measures are also used clinically to monitor severity and control of asthma and other respiratory diseases, and are independent risk factors for mortality.^{[1-3](#page-23-0)} Pulmonary function is strongly influenced by cigarette smoking and by multiple low-penetrance genetic variants. Indeed, genome-wide association studies (GWAS) of marginal genetic effects (i.e. not including interaction effects between genetic variants and smoking) have identified at least 26 loci associated with FEV₁ or FEV₁/FVC in the general population.^{[4](#page-23-1)} However, the interplay between genetic factors and environmental exposures has not been well established for pulmonary function or its associated traits. More broadly, while considerable efforts have been made to identify interaction effects between genetic variants and environmental exposures across the wide range of human traits and diseases,^{[5,](#page-23-2) [6](#page-23-3)} such investigations have been mostly unsuccessful in detecting robust gene-environment interactions.^{[5,](#page-23-2) [7](#page-23-4)} The well-established effect of cigarette smoking on numerous human health outcomes^{[8](#page-23-5)} makes it a serious candidate for identification of novel gene-environment interactions, especially for pulmonary traits.

Hypothesizing the presence of single nucleotide polymorphism (SNP)-by-smoking interaction, Hancock et al.^{[9](#page-23-6)} performed a genome-wide interaction study of pulmonary function, modeling single SNP main effects and their interactions with smoking in 50 047 participants of European ancestry across 19 studies within the Cohorts for Heart and Aging Research in

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Genomic Epidemiology (CHARGE)^{[10](#page-23-7)} and SpiroMeta consortia^{[11](#page-23-8)} -the largest genome-wide interaction study of pulmonary function as modified by smoking to date. However, rather than focusing on the interaction effects *per se*, they performed a meta-analysis of the joint test of SNP main effects and SNP-by-smoking interaction effects, in order to improve power for identifying genetic variants associated with pulmonary function.^{[12,](#page-23-9) [13](#page-23-10)} While they reported new candidate variants based on this joint test, the study did not identify any SNPs with genomewide significant interaction with smoking.

Here, we explored gene-by-smoking interaction effects limited to genetic variants previously found to be associated with pulmonary function in standard marginal effects GWAS,^{[4](#page-23-1)} therefore not including the new variants reported by Hancock et al. [9](#page-23-6) based on the joint test of main effects plus interaction. Specifically, we aimed to determine whether smoking modifies the effect of established genetic variants, when considered singly or combined using a genetic risk score summarizing the genetic predisposition to abnormal pulmonary function. The primary motivation for using genetic risk score is statistical power^{[14,](#page-23-11) [15](#page-24-0)}. Indeed, several genetic risk score-by-exposure interactions have already been identified in cases where single SNPs did not show evidence for statistically significant interactions^{[16-21](#page-24-1)}. Genetic risk score-by-exposure interaction testing expands on the principle of omnibus test while leveraging the assumption that, for a given choice of coded alleles, most interaction effects will have the same direction. This is similar to burden tests that have been widely used for rare variant analysis^{[22](#page-24-2)} where a single parameter can accumulate evidence for association without increasing the number of degrees of freedom. When interaction effects are null on average (i.e. if interaction effects are both negative and positive so that the sum of interaction coefficients tend to zero), the single

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SNP approach will generally outperform the risk score-based approach. Conversely, if interaction effects tend to be in the same direction, the risk score-based approach can have dramatically higher power.^{[14](#page-23-11)}

Methods

Single SNP-by-smoking interaction

The present analysis relies on the Hancock et al.^{[9](#page-23-6)} genome-wide screening for main genetic effects plus interaction effects with smoking in relation to pulmonary function among 50 047 participants (56% women) of European ancestry from 19 studies. The mean age was 53 years at the time of pulmonary function testing. Approximately 15% were current-smokers and 56% were ever smokers. Among ever smokers, the average pack-years of smoking was 21. **Supplementary Table 3** (available as Supplementary data at IJE online) provides the main characteristics of the studies included, while complete details of study-specific pulmonary function testing protocols can be found in previous work^{[4](#page-23-1)}. For studies with spirometry at a single visit, we analyzed FEV₁/FVC and FEV₁ measured at that visit. For studies with spirometry at more than one visit, we analyzed measurements from the baseline visit or the most recent examination with spirometry data. Smoking history (current, former, and never smoking) was ascertained by questionnaire at the time of pulmonary function testing. Pack-years of smoking was calculated for current and past smokers by multiplying smoking amount (packs/day) and duration (years smoked). Approximately 2.5 million autosomal SNPs were tested for interaction with smoking status (ever smoking vs never smoking) and pack-years, for two outcomes, FEV**¹** and FEV**1**/FVC, using the following model (see **Supplementary Note**, available as Supplementary data at IJE online):

$$
Y \sim \beta_0 + \beta_G G + \beta_{GE_k} G E_k + \sum_{l=1 \dots 3} \beta_{E_l} E_l
$$
 (Equation 1)

where β_G and β_{E_l} are the main effect of the SNP G and exposure E_l , β_{GE_k} is the interaction effect between G and exposure E_k , and β_0 the intercept.

Multivariate interaction analysis overview

First, we considered an unweighted genetic risk score-by-smoking interaction where the risk score simply sums the number of risk alleles (i.e. alleles associated with a lower pulmonary function). This unweighted genetic risk score is most powerful when the interaction effects have the same direction as marginal SNP effects, i.e. the harmful effects of smoking are magnified in individuals with a genetic predisposition to reduced pulmonary function. Second, we used a weighted genetic risk score where SNPs were weighted by the absolute value of their marginal effect estimates obtained from stage 1 screening of FEV**¹** and FEV**1**/FVC from Soler Artigas et al. [4](#page-23-1) (**Supplementary Table 1** available as Supplementary data at IJE online). This weighting scheme is most powerful when the magnitude of interaction effects is proportional to the SNP marginal effects. Finally, for our third multivariate analysis, we derived a standard omnibus test of all interaction effects. This test will retain power in the presence of effects in both directions or of different magnitudes. Although there is strong correlation among the 12 tests performed (these three models, considering interaction with two different smoking

metrics, ever/never smoking or pack-years, for the two pulmonary function metrics FEV1 and FEV1/FVC), we used a stringent Bonferroni p-value correction threshold of 4×10⁻³ to account for multiple testing.

When raw data are available, the weighted genetic risk score (GRS) is usually expressed as $GRS = \sum_m [w_i \times G_i]$ where m is the number of SNPs included in the genetic risk score and $w = (w_1, ..., w_m)$ are the weights attributed to each single SNP. Following previous notation, the test of interaction between the genetic risk score and the exposure E_k can be applied using the following model:

$$
Y \sim \gamma_0 + \gamma_{GRS} \times GRS + \gamma_{INT} \times GRS \times E_k + \sum_{l=1\ldots3} \gamma_{E_l} \times E_l
$$
 (Equation 2)

where γ_0 , γ_{GRS} , γ_{E_l} and γ_{INT} are the intercept, the main effect of the genetic risk score, the main effect of the exposure E_l and the interaction effect between E_k and the genetic risk score, respectively. However, as raw data were not directly available, we performed the test of γ_{INT} from summary statistics of interaction effects using an inverse-variance weighted sum as proposed by Aschard.^{[14](#page-23-11)} The chi-square for the interaction term γ_{INT} was derived as follows:

$$
\chi_{int}^{2} = \frac{\left(\sum_{i=1\ldots m} \frac{w_{i} \times \hat{\beta}_{G_{i}E_{k}}}{\hat{\sigma}_{\beta_{G_{i}E_{k}}}^{2}}\right)^{2}}{\sum_{i=1\ldots m} \frac{w_{i}^{2}}{\hat{\sigma}_{\beta_{G_{i}E_{k}}}^{2}}}
$$

where $\hat\beta_{G_t\times E_k}$ and $\hat\sigma^2_{\beta_{G_t\times E_k}}$ are the estimated effects and variance of the interaction between the exposure E_k and the SNP G_i obtained from *Equation 1*; and w_i is the weight applied to SNP

 G_i . Under the null hypothesis of no interaction effect, χ^2_{int} follows a chi-squared distribution with one degree of freedom.

The standard omnibus test of all interaction effects consisted in evaluating jointly $\boldsymbol{\alpha}_{\boldsymbol{G}\times\boldsymbol{E}_{\boldsymbol{k}}}=(\alpha_{G_{1}\times E_{k}},...,\alpha_{G_{m}\times E_{k}})$ from the model:

$$
Y \sim \alpha_0 + \sum_{i=1\ldots m} \left[\alpha_{G_i} \times G_i \right] + \sum_{i=1\ldots m} \left[\alpha_{G_i \times E_k} \times G_i \times E_k \right] + \sum_{l=1\ldots 3} \alpha_{E_l} \times E_l
$$

where α_0 , α_{G_i} , α_{E_l} and $\alpha_{G_i\times E_k}$ are the intercept, the main effects of SNP G_i and the exposure E_l , and the interaction effect between G_l and E_k . Leveraging the independence between the SNPs considered (a single SNP was selected for each independent locus), we also derived the omnibus test using summary statistics. Under this independence assumption, the $G_i \times E_k$ interaction terms would also be independents, [14](#page-23-11) so that it can be performed by summing the chi-square from each univariate interaction test to form a chi-square with m degree of freedom as follows:

$$
\chi_{joint}^{2} = \sum_{i=1\ldots m} \frac{\hat{\beta}_{G_{i} \times E_{k}}^{2}}{\hat{\sigma}_{\beta_{G_{i} \times E_{k}}}^{2}}
$$

where $\hat\beta_{G_l\times E_k}$ and $\hat\sigma^2_{\beta_{G_l}\times E_k}$ are the estimated effects and variance of the interaction between the exposure E_k and the SNP G_i obtained from *Equation 1*.

Relative risk in ever smokers versus never smokers

Genetic risk score interaction effects can further be translated in terms of risk prediction. For pulmonary function, low FEV₁ or FEV₁/FVC increases the risk of death^{[23](#page-24-3)} and together form the basis for the diagnosis of COPD.^{[1-3](#page-23-0)} COPD stage 2 or higher are defined by the Global Initiative for Chronic Obstructive Lung Disease (GOLD) as having $FEV₁/FVC < 0.70$ and FEV₁ < 80% of the predicted value. According to recent studies,^{[2,](#page-23-12) [24](#page-24-4)} between 5% and 20% of European ancestry adults are expected to have FEV**1**/FVC <0.70, depending on smoking characteristics and age distribution. Moreover, several studies argue for a more stringent threshold to define COPD^{[24,](#page-24-4) [25](#page-24-5)} based on lower limit of normal predicted value, rather than a fixed absolute value, to prevent disease misclassification.

To explore the impact of interaction effect on the risk of disease, we derived the relative risk (RR) of having FEV**1**/FVC below a given threshold (1%, 5% and 20%) in ever smokers versus never smokers conditional on the unweighted genetic risk score. This quantity is defined as the joint probability of having both FEV₁/FVC in the interval $[-\infty, FEV₁/FVC_{un}]$ and the genetic risk score (GRS) in the interval $[GRS_{low}, GRS_{up}]$. This can be expressed as the following integral:

$$
\int_{-\infty}^{FEV_1/FVC_{up}} \int_{GRS_{low}}^{GRS_{up}} f_1(y|g, e) \times f_2(g|e) dy dg
$$

where y , e and g are FEV₁/FVC, smoking status and the genetic risk score, respectively, and f_1 and f_2 are the probability density function of y and g. The detailed derivation of the above integral is available as Supplementary data at IJE online.

Results

We selected 26 loci previously found to be associated with FEV₁ and/or FEV₁/FVC at genome-wide significance (P < 5×10⁻⁸) in marginal association tests^{[4,](#page-23-1) [11,](#page-23-8) [26](#page-24-6)} (i.e. not including interaction effects with smoking exposures), and replicated in the GWAS by Soler Artigas et al., [4](#page-23-1) the largest meta-analysis of marginal genetic effect conducted for these two traits in the general population. Additional loci for these two phenotypes have been identified in two recent studies.^{[27,](#page-24-7) [28](#page-24-8)} However, these new loci were not included in our analysis because both these studies used a large cohort ascertained through smoking status. For each of the 26 selected loci, we choose the SNP with the strongest evidence for association (i.e. smallest *p*-value) with each of these phenotypes. The final list included 26 SNPs per phenotype, with only two SNPs being different between FEV**¹** and FEV**1**/FVC as previously reported[4](#page-23-1) (**Supplementary Table 1** available as Supplementary data at IJE online). Estimated interaction effects of these SNPs were extracted from the meta-analysis summary statistics for the four tests performed in the Hancock et al.^{[9](#page-23-6)} analysis: SNP-by-smoking status (ever smoking vs never smoking) interaction effect on FEV**¹** and FEV**1**/FVC; and SNP-by-smoking pack-years interaction effect on FEV**1**/FVC and FEV**1**. As showed in **Supplementary Table 2** (available as Supplementary data at IJE online), nine SNPs showed nominal significance (*P* < 0.05) out of the 104 tests performed ; however, none remained significant after accounting for multiple testing (Bonferroni corrected p-value threshold of 5×10-4). The minimum *p*-value was observed for the interaction between rs993925, near the TGFβ2 gene, and smoking status on FEV₁ ($β_{int} = -0.036$, 95% confidence interval (CI), -0.009 – -0.032, *P*=0.007).

Next, using these data, we conducted three multivariate (as opposed to single SNP) interaction analyses, testing jointly for the interaction effects between those SNPs and either smoking status or pack-years on the two phenotypes (FEV**¹** and FEV**1**/FVC) for a total of 12 tests. As shown in **Table 1**, none of the multivariate interaction tests with pack-years were significant. However, four of the six multivariate interaction tests with smoking status (ever versus never) showed nominal significance, and two tests for FEV**1**/FVC had a *p*-value below the Bonferroni significance level (12 tests, P<4×10⁻³). . The strongest signal was observed for the unweighted genetic risk score by smoking status interaction effect on FEV₁/FVC ($\beta_{int} = -0.036$, 95% CI, -0.040 – -0.032, *P*=0.00057). The Cochran's Q test for heterogeneity of the interaction effect across studies was not significant (P=0.97), and the forest plot of study-specific results did not display any obvious outlier (**Supplementary Figure 1,** available as Supplementary data at IJE online).

The contrast between this significant risk score interaction and the absence of strong single SNP interaction effects can be explained by looking at the distribution of the single SNP interaction effect estimates. **Figure 1** shows this distribution for the alleles associated with decreased FEV**1**/FVC. It highlights that while the 95% CI of most single SNP interaction effects encompass the null (and therefore the absence of significant single SNP interaction effect), there is an enrichment for negative interaction effects. Indeed, even a binomial test can be used to confirm the unbalanced direction of interaction effects (18 out of 26 interactions are negative leading to a *p*-value of 0.014 for a binomial test with an expected equiprobable distribution of 0.5). The genetic risk score-based interaction test exploits such enrichment by testing for the average interaction effect across all SNPs.^{[14](#page-23-11)} As with any multivariate approach

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based on a composite null hypothesis, this result indicates that at least a subset of these 26 SNPs interact with smoking status, but does not allow us to determine which SNP(s) or how many SNPs are driving the genetic risk score-by smoking interaction. The three other sets of single-SNP interaction tests showed a similar (but not significant after correction for multiple testing) trend with enrichment for negative interactions (**Supplementary Figs. 2-4** available as Supplementary data at IJE online). We summarized the contribution of the unweighted genetic risk score-by-smoking interaction on FEV**1**/FVC in **Table 2** and **Figure 2A**. This indicates that the deleterious effect of smoking is enhanced among carriers of the risk alleles or equivalently that the deleterious effect of smoking is reduced among subjects carrying the protective alleles.

We used two datasets of 8859 unrelated individuals and 9457 related individuals, respectively to test for independent replication of our results (**Supplementary Note** available as Supplementary data at IJE online). Both replication samples showed consistent negative GRSby-ever smoking interaction effect on FEV1/FVC ($\hat{\beta}_{int} = -0.0025$, 95% CI -0.0165, 0.0115, *P*=0.72 and $\hat{\beta}_{int}$ = -0.0030 , 95% CI -0.0214 0.0154, *P*=0.74, and overall interaction effect in the combined replication datasets $\hat{\beta}_{int} = -0.0027$, 95% CI -0.0136, 0.0082 P=0.63), and a Cochran's Q test for heterogeneity showed no significant difference in the three effect estimates (*P*=0.51). However, despite having total N=18,316 individuals, our combined replication sample remains underpowered (<50% power at nominal significance of 5%) and analyses including more participants would be necessary to strength the replication evidence for interaction effects.

To quantify the impact of this result from a public health perspective, we estimated the impact of the genetic risk score-by-smoking interaction on having FEV**1**/FVC below 1%, 5% and 20% in the lower tails of the distribution in the population. Specifically, we derived the relative risk (RR) of having FEV**1**/FVC below these cutoff points (1%, 5% and 20%) in ever smokers compared to never smokers. **Figure 2B** quantifies the excess RR (i.e. the RR minus one) of individuals across five genetic risk score quintiles. It highlights the higher risk associated with smoking among individuals carrying risk alleles (i.e. alleles associated with poorer pulmonary function) as compared to individuals carrying protective alleles (i.e. alleles associated with better pulmonary function). For example, among individuals with a genetic risk score above the 80 th percentile, smokers have on average a 26% excess RR of having FEV**1**/FVC in the lowest 1% of the population distribution, whereas ever smokers with a genetic risk score below the 20th percentile have on average a 18% excess RR of falling in that same FEV**1**/FVC category compared to never smokers. Applying the same approach for $FEV₁$, we observed a similar pattern (**Supplementary Figs. 5** available as Supplementary data at IJE online). However, as expected, the lower magnitude of the genetic risk score-by-ever smoking interaction on $FEV₁$ implied a lower difference in RR between ever smokers and never smokers.

Discussion

Using the largest dataset to date of European ancestry participants from the general population with pulmonary function (FEV₁/FVC and FEV₁), smoking, and genetic data, we identified a gene-by-smoking interaction effect on FEV $_1$ /FVC by using a genetic risk score composed of 26 SNPs identified and replicated in a prior GWAS meta-analysis of marginal genetic effects. Replication study showed interaction effect estimates in the same direction as the discovery study, though analyses including more participants would be necessary to strength the replication evidence for interaction effects. To our knowledge, our study is the first to report a synergistic action of genes and smoking on pulmonary function, i.e. the reduction in FEV**1**/FVC due to smoking is greater among individuals who are genetically predisposed to lower FEV**1**/FVC ratio. Our study also highlights the importance of developing and applying alternative strategies to evaluate interaction effects for lung phenotypes along with other complex traits and diseases. The genetic risk score-based approach enabled us to identify an interaction when the standard univariate test (i.e. evaluating each single genetic variant for interaction independently) failed to identify any interactions.

Genetic risk score-by-exposure interaction can have higher clinical value than the identification of single SNP-by-exposure interaction as it can capture a wealth of information in a single measure to identify subgroups in the population whose genetic background makes them more susceptible to the deleterious effects of smoking.^{[19,](#page-24-9) [29,](#page-24-10) [30](#page-24-11)} Indeed, if single SNP-bysmoking interactions are distributed unconditionally on the marginal genetic effect (i.e. interaction effects have equal chances to be positive or negative given the coded alleles are the risk alleles), the genetic effect will be similar between ever and never smokers on average. The enrichment for negative interactions we identified through our genetic risk score approach reveals a stronger genetic component among the ever smoker subgroup in the population, and can allow the implementation of more efficient implementation of prevention strategies. For example, in the public health setting, programs targeting smoking cessation campaigns to

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individuals who are genetically predisposed to low pulmonary function may have a stronger impact in preventing COPD.

Our results may also elucidate biological mechanisms underlying the interplay between genes and smoking in pulmonary function. In particular, the higher statistical power for the genetic risk score-based interaction test points towards the potential presence of an unmeasured intermediate biomarker mediating the effect of the 26 loci on FEV₁/FVC. As shown in **Figure 3**, the most parsimonious model (i.e. the less complex following Occam's razor) that would explain multiple interactions going in the same direction (**Figure 1**) implies the genetic variants together influence an intermediate biomarker which itself interacts with smoking. Future studies with extended genomic data, including transcriptomic, proteomic, and/or metabolomic data, might be able to further assess such an hypothesis by evaluating i) the effect of the genetic risk score on those biomarkers, and ii) testing for interactions between smoking and the candidate biomarkers identified at step i).

This study has some limitations. The 26 selected variants together explain a relatively small proportion of the additive genetic variance in FEV₁/FVC and in FEV₁.^{[4](#page-23-1)} However, GWAS with increasing sample sizes will likely continue to provide additional associated genetic variants to further assess the role of SNP-by-smoking interaction effects on pulmonary phenotypes, and may increase the gap between smokers and never smokers to allow for a significant impact in the clinic or at the population level. Moreover, we focused on genetic variants previously found to be associated at genome-wide significance level, but future studies might consider less stringent criteria to select genetic variants, including those with only

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suggestive evidence, or alternatively candidate variants with functional annotation relevant to the outcomes and exposures in question. Obviously, the signal to noise ratio might decrease when relaxing the constraint on the SNP selection. However, as we recently showed, additional gain in statistical power might be achieved even if a substantial proportion of the variants do not interact with the exposure.^{[14](#page-23-11)} Finally, investigation of interaction effects with other environmental exposures such as secondhand smoke, air pollution, asbestos, or occupational risks may lead to a more comprehensive understanding of the biological and epidemiological significance of these variants.

In summary, the identification of interaction effects between genetic variants and environmental exposures in human traits is recognized as extremely challenging, and this quest has been mostly unsuccessful so far. In this study, we discovered novel gene-by-smoking interactions using risk scores, that were not observed at the level of individual genetic variants. This risk score analysis suggests that persons with a greater genetic predisposition to low pulmonary function are more susceptible to the deleterious effects of smoking. By extension, the use of a genetic risk score may help predict which smokers will fall below thresholds that establish the diagnosis of COPD.

Author Contributions

H.A., P.K., S.J.L., M.T., D.B.H. and A.J. were involved in designing the study. M.D.T., D.B.H., A.S., A.J., A.V.S., A.W.M., D.W.L., D.P.S., G.O.C., R.G.B., G.G.B., I.P.H., J.K.P., J.F.W., J.W., J.H.Z., K.d.J., L.V.W., M.S.A., H.M.B., M-R.J., M.F., P.A.C., S.A.G., S.R.H., V.G., W.T. S.J.L., I.R., O.P., J.E.H.,, C.H., A.C., D.J.P., S.E.H., I.J.D., S.E., U.G., LP.L., T.L., E.Z., B.P.P., and V.E.J. were involved in participant recruitment, sample collection or genotyping. H.A. performed analyses from the discovery study. H.A., V.E.J. and M.S.A performed the replication analysis. H.A. drafted the paper, with substantial editorial input from P.K., S.J.L., J.D., D.S., M.T., D.B.H. and A.J. All authors have reviewed and approved the final draft. This material has not been published previously in a substantively similar form.

Conflict of interest

Jaakko Kaprio consulted for Pfizer on nicotine dependence. Other authors declare no competing financial interest.

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Supplementary Data.

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Figure Legends

Figure 1. Distribution of interaction effects on FEV1/FVC.

Single SNP risk allele-by-smoking status (ever/never) interaction effect estimates (β_{int}) and 95% confidence intervals are plotted by increasing values. Negative and positive interactions are in dark blue and light blue, respectively. The unweighted genetic risk score-by-smoking status interaction is plotted in purple.

Figure 2. Overview of the unweighted genetic risk score-by-smoking interaction effect on FEV1/FVC.

Upper panel (A) presents the distribution of the unweighted genetic risk score (GRS, grey density plot) and the relationship between the unweighted GRS and standardized FEV**1**/FVC in ever smokers (red line) and never smokers (black line). Lower panel (B) shows the excess relative risk (RR) of having FEV**1**/FVC in the lowest 1%, 5% and 20% of the population for ever smokers compared to never smokers, as stratified by GRS quintiles.

Figure 3. Underlying causal model.

Potential causal diagrams underlying the gene and smoking interaction effects on $FEV₁/FVC$. Panel (A) presents a scenario where each genetic variant influences the outcome through a SNP-specific pathway, and interactions with the environmental exposure take place along these pathways. Panel B) presents an alternative (and simpler) model where multiple genetic variants influence an unmeasured intermediate biomarker, which effect on FEV₁/FVC depends on smoking. In scenario (A), the single SNP-by-smoking interaction test is the optimal approach, while in scenario (B), the single SNP-by-smoking interaction test can become inefficient and interaction would be easier to detect using a genetic risk score-by-smoking interaction test, as it summarizes all interaction effects in a single test.

Outcome	Exposure	Test	$\widehat{\beta}_{int}$ (CI)		p-value
FEV ₁	Smoking	uGRS		-0.0055 $(-0.011, 2.7x10^{-5})$	0.051
		wGRS		-0.21 $(-0.40,-0.033)$	0.020
	status [*]	CHISQ			0.49
FEV ₁	Pack-years	uGRS		$-1.6x10^{-5}$ $(-4.6x10^{-5}, 1.4x10^{-5})$	0.30
		wGRS		$-6.5x10^{-4}$ $(-1.6x10^{-3}, 3.3x10^{-4})$	0.19
		CHISQ			0.46
FEV_1/FVC	Smoking	uGRS		-0.0099 $(-0.016, -0.0043)$	0.00057 ⁺
	status	wGRS		-0.21 $(-0.33,-0.073)$	0.0022^+
		CHISQ			0.026
FEV ₁ /FVC	Pack-years	uGRS		$-4.4e-06$ $(-3.6x10^{-5}, 2.7x10^{-5})$	0.78
		wGRS		$-6.5x10^{-5}$ $(-8.0x10^{-4}, 6.6x10^{-4})$	0.85
		CHISQ			0.53

Table 1. Multivariate interaction tests of the 26 loci associated with pulmonary function.

uGRS is the genetic risk score using equal weights to all SNPs; wGRS is the genetic risk score weighted by effect *estimates from the marginal screening; CHISQ is the omnibus test of all interaction effects;* $\hat{\beta}_{int}$ is the estimated interaction effect between the GRS and the outcome; and CI is the confidence interval of that *estimate.*

Nominally significant tests are indicated in bold.

** Smoking status is defined as never smokers versus ever smokers.*

†*Significant* p*-value after Bonferroni correction.*

Table 2. Summary of effect estimates for genetic risk score-by-smoking status interaction on FEV1/FVC.

GRS is the unweighted genetic risk score ; beta is the effect estimates of each predictor; and SD the standard

deviation of the each beta.

** Smoking status was defined as never smokers versus ever smokers*