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People in pain make poorer decisions

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People in pain make poorer decisions

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Tables: 3

Figures: 3

People in pain make poorer decisions

Abstract

Chronic pain affects 1 in 5 people and has been shown to disrupt attention. Here, we investigated whether pain disrupts everyday decision making. In Study 1, 1322 participants completed two tasks online: a shopping decisions task and a measure of decision outcomes over the previous 10 years. Participants who were in pain during the study made more errors on the shopping task than those who were pain-free. Participants with a recurrent pain condition reported more negative outcomes from their past decisions than those without recurrent pain. In Study 2, 44 healthy participants completed the shopping decisions task with and without experimentally-induced pain. Participants made more errors while in pain than while pain-free. We suggest that the disruptive effect of pain on attending translates into poorer decisions in more complex and ecologically valid contexts, that the effect is causal, and that the consequences are not only attentional, but financial.

Keywords: Pain, cognitive disruption, decision making, finances, numeracy

Summary:

Pain had a disruptive effect on decision making in ecologically valid contexts. The effect is causal, and the consequences are not only cognitive, but financial.

1 Introduction

2 The disruptive effect of pain on attending has been demonstrated with
3 experimentally-induced pain[30,40], chronic pain[9,15] and transient pain such as
4 headache[4,24,31]. This field has predominantly focused on simple cognitive processes
5 (although sometimes using complex tasks combining multiple executive functions[25]). Few
6 studies have examined the effects of pain on higher-level cognition. One which did, found
7 that clinical pain was associated with less-abstract thinking[20], while another found no
8 evidence that experimentally-induced pain affected abstract thinking[2]. Here, we focus on
9 the potential impact of pain on higher-level real-world cognitive tasks requiring attention,
10 namely numerical reasoning and decision making, which have serious consequences if one
11 gets them wrong.

12 Reasoning and decision making are required in many areas of life and are influenced
13 by various cognitive and emotional factors. Here, we focus on numeracy as a domain that is
14 important in many areas of life, including budgeting, choosing a mortgage and choosing
15 insurance plans. Despite its importance, numeracy in adults is poor. In the quantitative
16 domain of the USA's 2003 National Assessment of Adult Literacy, 55% of adults performed
17 at a basic or below basic level (at best being able to locate easily-identifiable quantitative
18 information and solve one-step arithmetic problems when the operation was specified or
19 easily-inferred)[26]. These findings were echoed in a 2016 UK Money Advice Service
20 study[29] into the public's ability to choose the best supermarket deals for four products.
21 While 74% of participants chose the best deal for at least one product, only 2% chose
22 optimally for all four.

23 Attention is important for learning and performing numerical operations in both
24 children and adults[13,36]. Attention is also important in decision making, where we need

to consider various options, estimate their likely outcomes, then hold these in mind while choosing among them. Given that pain impairs attention, it may also influence numerical decision making. Indeed, there is some initial evidence that this is the case. Placing a hand into ice-cold water changed participants' risk-taking on a financial decision-making task[35].

We investigated the effect of pain on everyday decision making. In Study 1, a large general population sample recruited online reported whether they were currently in pain and whether they had any recurrent pain conditions. They completed two tasks: the shopping decision task used by the Money Advice Service[29] and the Decision Outcomes Inventory (DOI[10]), which measures real-world outcomes of everyday decisions made over the previous 10 years. We hypothesized that participants who were in pain would find the best shopping deal on fewer items than participants who were pain-free. If the effect of pain on attention does translate into poorer decision making, the outcomes of these poor decisions may accumulate in people with chronic pain. We therefore hypothesized that participants with pain that had lasted for three months or longer would report more negative decision outcomes on the DOI than other participants. In Study 2 we took an experimental approach to determine a causal relationship: participants completed an extended shopping-decisions task with their hand in warm or painfully-cold water.

Study 1 Method

Design and procedure

Participants (N = 1322) took part online and were recruited via Amazon's Mechanical Turk (N = 658) and Prolific.ac (N = 664). Research has shown data collected online for psychology studies is reliable[11,33] and that samples tend to be more diverse than traditional university-based samples[27]. The large sample size was decided based on

previous online pain studies[3,5,7], taking into account the likelihood of substantial non-completion, the noise associated with online data collection, and likely percentages of participants who would report current pain (estimated at 30%) or recurrent pain (estimated at 40%), to ensure sufficient numbers of participants in each cell (a minimum of 20) of the current pain x recurrent pain x sex interaction. We achieved this aim, with the smallest cell size in our primary analyses (after exclusions, see the “Data Cleaning” section) being 22 participants who were male with both current and recurrent pain.

Participants were directed to the study pages in Qualtrics via the recruitment systems. After giving informed consent, participants then answered demographic questions, reported whether they had current and/or recurrent pain, completed the Decision Outcomes Inventory and completed a shopping decisions task (see below for question and task details). A ‘debrief’ page appeared at the end which gave information on the purpose of the study, how to contact the researchers with any questions, and how to receive payment on MTurk or Prolific (\$2 or £2 respectively, based on the recommended per-hour rates for each recruitment site).

Participants and demographics

Participants (N = 1322) were 699 males, 619 females, 3 trans females and 1 trans male. Ages ranged from 18 to 74 (M = 33.98, SD = 10.80). English was the first language for 1201 participants, with 16 reporting Hindi, 11 reporting Portuguese, and the remaining 94 reporting a wide variety of other languages. The majority of participants were based in the USA (N = 799), followed by the UK (N = 341), and India (N = 44), with the remaining 138 reporting a variety of other countries. Participants reported their highest educational qualification, and 54% reported a university degree or above.

Pain characteristics

Participants were asked whether they were in pain at the time of the study (“Are you in pain at this very moment, from any cause?” Yes: N = 337, 25%, No: N = 985, 75%). Those who reported a current pain condition answered follow-up questions about the type, intensity, and duration of their pain. Participants were then asked whether they had any recurrent pain conditions that were not causing them pain at the time (“Do you have a recurrent pain condition that is not causing you pain at this moment in time?” Yes = 407, 31%, No = 915, 69%). Those who reported a recurrent pain condition answered follow-up questions about the type, intensity, and duration of the pain. The type of pain was selected from a list of options, separately for current and recurrent pain (see Table 1; participants were able to select multiple types of pain). Pain intensity was indicated on a numerical rating scale from 0 “No pain at all” to 10 “Worst pain imaginable” separately for current and recurrent pain (current M = 3.47, SD = 1.79, 95% CI [3.28, 3.66]; recurrent M = 5.42, SD = 1.83, 95% CI [5.24, 5.60], participants were instructed to give an average pain rating if they had more than one recurrent painful condition). Participants were asked to indicate when their pain had started from a list of options (see Table 2), separately for current and recurrent pain. They were instructed to answer based on their longest standing painful condition if they had more than one type of current pain or more than one type of recurrent pain.

Insert Table 1 here

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Insert Table 2 here

Measures

Shopping-Decisions Task

The shopping-decision task was adapted from that used by the UK Money Advice Service[29] (see <https://figshare.com/s/17f3800d66375b92fb71> for the questions used here). We used the original four questions, based on actual deals available in supermarkets, and created two new ones to increase the potential variance in the data. Each question asked participants which of 4 or 5 options represented the best deal for buying a particular product. For example:

Of the following options for milk, which represents the best deal?:

- a) Six pints of milk for £1.80
- b) Four pints of milk for £1.40
- c) Two six-pint cartons of milk on offer for £3.50
- d) Two four-pint cartons of milk on offer for £2

We included a “don't know” option for each question. This was intended to serve two purposes: first, to reduce the risk of participants getting correct answers simply by guessing, and second, to reduce the risk of any participants with mathematics anxiety[21] feeling anxious about being asked to perform arithmetic, by making it clear that they could avoid it if they wished.

Answers were scored as either correct or incorrect and participants were given a total score ranging from 0 to 6.

Decision Outcomes Inventory (DOI)

The DOI asks participants to report whether or not they have experienced each of 41 negative decision outcomes, such as missing a flight, receiving parking tickets, or going bankrupt[10]. Thirty-five of these outcomes are preceded by filter questions asking whether participants had had the opportunity to experience the outcome. For example, the question “Have you had your electricity, cable, gas or water shut off because you didn't pay on time?” was preceded by “Have you been responsible for electricity, cable, gas or water payments?”.

Scores on the DOI are weighted by the proportion of the sample who reported experiencing each outcome, with less common outcomes being weighted more heavily than more common ones on the assumption that frequency reflects severity (e.g. many more people have been lost while driving than have spent a night in jail). For example, if 500 participants have bought their own food, and 470 of these have thrown away food that went bad, the score for experiencing this negative outcome would be $1 - 470/500 = .06$.

Participants received a positive weighted score if they had avoided a negative outcome (e.g. +0.06 if they had bought their own food and never had to throw any away because it had gone bad), a negative weighting if they had experienced the negative outcome (e.g. -0.06 if they had bought their own food and had to throw some away because it went bad), and no weighting if they had not had the opportunity to make the bad decision (e.g. 0 if they had never bought their own food). Therefore, lower scores indicated more and/or more severe negative outcomes from decisions.

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Ethics

The study was approved by the Loughborough University Ethical Approvals (Human Participants) Sub-Committee.

Hypotheses

We hypothesized that participants who were in pain at the time of the study would find the best shopping deal on fewer items than participants who were pain free. We did not expect participants with recurrent pain conditions that were not causing pain at the time to be affected on the shopping task: only current pain was hypothesized to have an effect. If the effect of pain on attention does translate into poorer decision making, the outcomes of these poor decisions may accumulate in people with recurrent or chronic pain. We therefore hypothesized that participants with chronic pain would report more negative outcomes from their decisions on the DOI than participants without chronic pain. This was because we expected pain to have disrupted the decision-making process at the time that some of the decisions were made. We did not expect current pain to affect DOI scores.

Study 1 Results

All analyses were two-tailed.

Data access

Data files and analysis scripts are openly available at <https://figshare.com/s/17f3800d66375b92fb71>

1 Data cleaning

2 A total of 1559 participants accessed the study pages. Of these, 1360 gave consent
3 and proceeded, and 1322 completed the study. Eight participants missed one or two items
4 on the DOI, but a score was calculated excluding the missing items and their data were
5 included.

6 Our analyses included participants' sex, given that females tend to report more pain
7 than males and at a higher intensity[18, 32], and they also tend to have higher self-reported
8 levels of cognitive intrusion from pain than males[3]. We therefore limited the analyses to
9 participants reporting their sex as male or female (N = 1318, excluding the four participants
10 who reported other sexes).

11 In the shopping-decision analysis, all available data (N = 1318) were included. In the
12 DOI analysis, however, only participants from the USA were included. This was because the
13 frequency and acceptability of the outcomes on the DOI may vary between countries, and
14 scores are weighted based on the relative frequency within the sample (see Method).
15 Therefore, including multiple countries may inappropriately skew the weightings given to
16 different decision outcomes (e.g. if alcohol-related outcomes are substantially more
17 common in Country A than in Country B, calculating weightings across the two countries
18 combined would overstate the severity in Country A by making drinking appear less
19 common that it is in that country, and understate the severity in Country B by making
20 drinking appear more common than it is in that country). The largest subgroup of
21 participants was from the USA (N = 795), so the scoring and analysis was limited to this
22 group. Eight negative outliers (more than 3 standard deviations below the group mean)
23 were removed from the DOI scores, bringing the sample size to 787.

We checked for repeated IP addresses within the dataset, and aside from cases of participants opening the link twice but only completing the study once, we identified 15 IP addresses which appeared with two complete submissions, and one which appeared with three complete submissions. It is not possible to know whether these participants completed the study multiple times or whether they were different individuals with the same IP address (e.g. because they were in the same household, public library, university student residence, or public or university computer room). We re-analysed the data excluding these cases, and the pattern of results remained the same as reported below where they are included.

We also checked for cases where participants completed the study very quickly, which might indicate that they were not responding seriously. We chose six minutes as a cut-off time based on a drop-off in the distribution of submission durations below this point. There were 71 participants with a study duration of less than six minutes, and when these participants were excluded the pattern of results remained the same as reported below where they are included.

Effect of pain on shopping deal scores

Scores on the shopping task (number of correct answers out of 6) were entered into a 2 (current pain: yes [N = 336], no [N = 982]) x 2 (recurrent pain: yes [N = 406], no [N = 912]) x 2 (sex: female [N = 619], male [N = 699]) ANCOVA with age as a covariate. Age was used as a covariate because it correlated with shopping scores, $r_s(1318) = .097$, $p < .001$, and differed by current (yes/no) and recurrent (yes/no) pain, both $ps < .001$. Highest qualification was correlated with shopping scores, $r_s(1310) = .109$, $p < .001$, but did not differ by current (yes/no) or recurrent (yes/no) pain, $ps > .406$. Our final sample of 1318

gave us 95.4% power to detect main effects and 2×2 interactions with a small effect size of $\eta_p^2 = 0.01$.

The ANCOVA revealed a significant main effect of current pain, $F(1,1309) = 5.36$, $p = .021$, $\eta_p^2 = .004$, with participants who were in pain scoring lower ($M = 2.86$, $SD = 1.69$, 95% CI [2.68, 3.04]) than those who were not in pain ($M = 3.11$, $SD = 1.85$, 95% CI [3.00, 3.23], see Figure 1). There was no significant main effect of recurrent pain, $F(1,1309) = 1.19$, $p = .275$, $\eta_p^2 = .001$ (recurrent pain $M = 2.93$, $SD = 1.75$, 95% CI [2.76, 3.10]; no recurrent pain $M = 3.05$, $SD = 1.99$, 95% CI [2.92, 3.18]). There was no significant main effect of sex, $F(1,1309) = 2.77$, $p = .096$, $\eta_p^2 = .002$, (females $M = 3.08$, $SD = 1.94$, 95% CI [2.93, 3.23]; males $M = 2.90$, $SD = 2.04$, 95% CI [2.74, 3.05]). There were no significant interactions, all $ps > .149$.

Insert Figure 1 here

The data show that current but not recurrent pain was associated with more errors in finding the best shopping deal, with a small effect size. To explore the extent of this effect in terms of money spent, we calculated the average price per unit of the deals chosen by participants in the current pain and no current pain groups. These data are presented in Table 3 along with the best price available for each item. These data suggest that in general, participants would have spent more than necessary on all items, and that those in pain spent more than those not in pain on half of the items.

Insert Table 3 here

Effect of pain on DOI scores

For the DOI, we hypothesised that participants who had recurrent pain would score lower than those without recurrent pain, but that current pain would not affect scores. Therefore, participants who had current pain which had lasted for three months or longer were included in the recurrent pain group. DOI scores were analysed using a 2 (current pain: yes [N = 132], no [N = 655]) x 2 (recurrent pain: yes [N = 218], no [N = 569]) x 2 (sex: female [N = 352], male [N = 435]) ANCOVA. Age and highest qualification were both related to DOI scores, $ps < .001$, and age differed by recurrent pain, $p = .001$, so it was included as a covariate. Our final sample of 787 gave us 80.4% power to detect main effects and 2x2 interactions with a small effect size of $\eta_p^2 = 0.01$.

There was a main effect of recurrent pain on DOI scores, $F(1,778) = 4.73$, $p = .030$, $\eta_p^2 = .006$, where those with recurrent pain scored lower ($M = 13.80$, $SD = 7.60$, 95% CI [12.78, 14.81]) than those without recurrent pain ($M = 15.18$, $SD = 8.95$, 95% CI [14.45, 15.91], see Figure 2). There was no main effect of current pain on DOI scores, $F(1,778) = 2.39$, $p = .123$, $\eta_p^2 = .003$ (current pain $M = 13.99$, $SD = 6.56$, 95% CI [12.87, 15.12]; no current pain $M = 14.98$, $SD = 7.27$, 95% CI [14.43, 15.53]), and no main effect of sex, $F(1,778) = 2.13$, $p = .145$, $\eta_p^2 = .003$, (females $M = 14.95$, $SD = 8.46$; 95% CI [14.07, 15.84]; males $M = 14.02$, $SD = 9.39$, 95% CI [13.14, 14.91]). There were no significant interactions, all $ps > .252$.

Insert Figure 2 here

Interim discussion

Study 1 supported our hypotheses that current pain would be associated with lower scores on the shopping task while recurrent pain would be associated with lower scores on the DOI. However, due to the cross-sectional nature of Study 1, we cannot infer a causal relationship. The effect of pain on DOI scores cannot be tested experimentally because this would require sustained pain induction over the course of 10 years. However, the effect of pain on shopping decisions can be tested experimentally, and this was done in Study 2.

Study 2 Method

Design

The experiment followed a within-participants design where participants completed shopping deal questions in two conditions: with their non-dominant hand in painful cold water (8°C) and with their non-dominant hand in non-painful warm water (30°C).

Participants

Participants were 44 students and members of staff at Loughborough University. There were 32 females and 12 males, who ranged in age from 18 to 47 ($M = 24.57$, $SD = 7.24$). Our sample size was pre-determined to give us 90% power to detect an effect size of $d = 0.5$ with a family-wise alpha level of 0.05 in a two-tailed paired-samples t -test. This effect size was chosen as the smallest effect size of interest.

Participants completed a health screening questionnaire to ensure that they were eligible for the study. Participants were ineligible for the study if they reported any of the following: current pain, chronic pain, uncorrected vision problems, heart problems, a history of fainting or seizures, serious cognitive impairment (e.g. dementia), asthma, a history of frostbite, arthritis in their hands or forearms, a fracture in their hands or wrists, open cuts or sores on their hands or wrists, a history of Reynaud's phenomenon, or having taken pain killers in the previous 12 hours. These criteria were intended to ensure participants' safety when taking part in the cold pressor task, and that the data would be of high quality.

Measures

Shopping decisions task

The shopping-decisions task was adapted for the cold pressor pain-induction paradigm used in this study. We created 90 items, which each posed a question with two answer options. For example:

Of the following options for buying grapes, which represents the best deal?

a) One 500g pack costing £2.00

b) Buy two get the third free deal on 500g packs costing £2.50 each

Reducing the number of answer options meant that participants would be able to complete more questions in the timeframe allowed by the pain induction procedure, which in turn increased the sensitivity of the measure. See

<https://figshare.com/s/17f3800d66375b92fb71> for the full set of questions.

Six 15-item blocks were created from the 90 items, and the blocks were presented in a fixed order. The starting condition (pain or control) was counterbalanced by participants, meaning that the items were counterbalanced by condition. The 15 items within each block

were presented in a random order for each participant. If participants responded to all 15 items within the 90-second time limit, the list restarted. This occurred for five participants in a total of 11 blocks (69 questions, 3.2% of the data). The analyses presented below include these repeated questions, but the results remained the same when they were excluded.

In each trial, the question was presented in the upper centre of the computer screen with the two answer options presented underneath on the left and right sides of the screen, all in 28-point black Arial font on a white background. The question and answers remained on screen until the participant responded by pressing the n or m key (to choose the left or right answer, respectively) on the computer keyboard with their dominant hand. Once the participant responded, the task moved immediately on to the next question.

Visual analogue scales (VASs)

After each block of the shopping task questions, participants responded to the question “How much pain were you in during the block?” by clicking on a visual analogue scale anchored by “No pain at all” on the left and “Worst pain imaginable” on the right.

Procedure

Participants took part individually in a quiet laboratory. They were sent an information sheet and health screening questionnaire prior to the study session for their reference. At the start of the session, they were asked to read the information sheet and complete a consent form, then complete the health screening questionnaire to ensure eligibility for the study (see “Participants” section for exclusion criteria).

The shopping task and VASs were presented on a desktop PC in E-Prime 3.0 with a 34cm x 27cm monitor. The study began with a practice VAS and a practice shopping task

question. Participants then completed six blocks of the shopping task, with each block followed by a VAS. Each shopping task block lasted for 90 seconds, and during each one participants had their non-dominant hand in either cold (8°C, pain condition) or warm (30°C, control condition) water. The pain and control blocks alternated in order, and the starting condition was counterbalanced. Due to experimenter error, one participant completed the wrong counterbalanced order, meaning that 23 participants started with a pain block and 21 with a control block. In between blocks the participants had the opportunity to warm their hand and only began the next block once they were pain-free. Participants were free to withdraw their hand from the water at any time during a block, but none did.

At the end of the study, participants were thanked, debriefed, and given £5 in cash.

Study registration

In between the completion of the first study and the start of the second we introduced into our labs the scientific practice of registering study protocols prior to study initiation, in line with current recommendations[12]. Our hypothesis, sample size and analysis plan for study 2 were registered at www.aspredicted.org (#8022; a pdf for peer review is available at <http://aspredicted.org/blind.php?x=rk2p33>). Our hypothesis was that “On a value-for-money shopping deal task, participants will give fewer correct answers AND/OR more incorrect answers while experiencing pain compared to when they are pain free.” Our analysis plan was to conduct two paired-samples *t*-tests to see whether (a) the number of correct responses and/or (b) the number of questions attempted but incorrect differed between the pain and no pain conditions. To maintain a family-wise alpha level of 0.05, a Bonferroni correction was applied, making the alpha level 0.025 for each *t*-test. Any a posteriori analyses that were not registered are labelled as ‘Exploratory’ below.

Ethics

The study received ethical approval from the Loughborough University Ethical Approvals (Human Participants) sub-committee, and the University of Bath's Psychology Ethics Committee and Research Ethics Approval Committee for Health.

Study 2 Results

All analyses were two-tailed.

Data access

Data files and analysis scripts are openly available at <https://figshare.com/s/17f3800d66375b92fb71>

Data cleaning

In line with our study registration, we ran several checks on the data before the main analysis was conducted. These checks resulted in four participants' data being removed and replaced. Two participants were removed because their average pain rating in the pain blocks was not significantly different to their average pain rating in the control blocks. One participant was removed because their number of attempted but incorrect questions was more than three standard deviations from the group mean in the control condition. Finally, one participant was removed because their mean response time was more than three standard deviations from the group mean response time in the control condition.

Effect of pain on shopping deal scores

Paired-samples t-tests were used to compare shopping-task performance between the pain and control conditions, in terms of (a) number of questions answered correctly, and (b) number of questions attempted and answered incorrectly (with the alpha level Bonferroni corrected to 0.025). There was no effect of pain condition on number of correct answers, $t(43) = 0.74$, $p = .464$, Hedges' $g_{av} = .095$ (Pain $M = 18.89$, $SD = 9.23$, 95% CI [16.08, 21.69]; Control $M = 18.14$, $SD = 6.00$, 95% CI [16.31, 19.96]). However, there were more incorrect answers in the pain condition ($M = 7.43$, $SD = 5.83$, 95% CI [5.66, 9.21]) than in the control condition ($M = 4.66$, $SD = 2.47$, 95% CI [3.91, 5.41]), $t(43) = 3.94$, $p < .001$, Hedges' $g_{av} = 0.608$ (see Figure 3).

Exploratory analyses

VAS ratings were compared in the pain and no-pain conditions as a manipulation check. VAS ratings were higher in the pain condition ($M = 63.11$, $SD = 17.08$, 95% CI [57.92, 68.31]) than in the control condition ($M = 3.23$, $SD = 3.57$, 95% CI [2.15, 4.32]), and the difference was significant, $t(43) = 23.27$, $p < .001$, Hedges' $g_{av} = 4.768$.

To see whether the effect of pain on incorrect answers on the shopping task differed between male and female participants, we ran a 2 (condition: pain, control) x 2 (sex: male, female) mixed ANOVA. There was no significant interaction, $F(1,42) = 0.06$, $p = .815$, $\eta_p^2 = .001$.

Insert Figure 3 here

In our main analysis we found that participants gave more incorrect answers in the pain condition than in the control condition, but no evidence that they gave fewer correct answers. This suggests that they may have answered the questions more quickly in the pain condition than in the control condition, thereby attempting more items but increasing their error rate. To test for this, we compared the total number of items attempted between the pain and control conditions. As expected, participants attempted more questions in the pain condition ($M = 26.32$, $SD = 13.60$, 95% CI [21.56, 31.08]) than in the control condition ($M = 22.80$, $SD = 7.14$, 95% CI [20.30, 25.29]), $t(43) = 2.35$, $p = .024$, Hedges' $g_{av} = 0.319$.

Discussion

Chronic pain patients often report difficulty in attending, and studies with both chronic and laboratory-induced pain have provided quantitative evidence that pain negatively affects performance on varieties of attention tasks. Here, we investigated for the first time whether pain disrupts performance on real-world and higher-level cognitive tasks that require attention, namely numerical and everyday decision making, which can have real and serious consequences for individuals.

In Study 1 we showed that people who were in pain at the time of completing an online shopping-decisions task made more errors than people who were not in pain at the time. Furthermore, people who reported a recurrent pain condition had experienced more negative events over the past 10 years of their lives, possibly resulting from poor decision making, than people who did not report a recurrent pain condition. In Study 2, we induced pain in healthy participants while they performed an extended version of the shopping task and found that they made more errors while they were in pain than while they were pain-free, supporting our hypothesis of a causal effect of pain on shopping decisions. This effect

was driven by participants attempting to answer more questions in the time allowed. We found no evidence of sex differences in any of these effects.

Our findings extend the literature showing that pain disrupts basic attention processes [4,14,30,31,40] by showing that this disruptive effect translates into poorer decisions in more complex and ecologically-valid contexts. This could have implications for many aspects of daily life for people in pain, such as school, work, and managing finances. The effects of pain on both tasks in Study 1 were small, but even a small effect of pain on each of the many decisions we need to make every day can add up, over the course of years, to substantially poorer quality of life in people with chronic pain. Indeed, in Study 2, with a more sensitive measure and controlled environment, we observed a medium-sized effect of pain on shopping decisions (Hedges $g_{av} = 0.608$).

The findings from Study 2 hint towards potential mechanisms of the effect of pain on shopping decisions. When participants were in pain, they attempted more questions and made more errors, but they did not give fewer correct answers. This is suggestive of three possible explanations. First, it could be that participants felt the need to rush through the questions when they were in pain in case it made the pain end sooner. Participants were told that every block would last for 90 seconds, but it is possible that they still consciously or subconsciously thought that going faster would make the pain end sooner. Second, it could be that pain reduces our confidence in our thinking skills so that participants were less willing to think through the arithmetic for items they found particularly difficult, and instead guessed. This would lead to more questions being attempted, and more errors being made. Third, it could be that pain increases confidence so that participants didn't check their answers whilst in pain, whereas they may have checked them if they were not in pain. This would also increase the number of items attempted in the time limit and lead to more

errors. Due to participants being aware that all blocks would last for exactly 90 seconds, we suggest that the second and third explanations are more likely than the first. These are both metacognitive explanations which could be tested using established paradigms from the metacognition literature. For example, participants could be asked to indicate their initial judgment of solvability for each item, and their interim and/or final confidence in each answer they give [1].

We suggest that our pattern of findings is not well-explained by a mechanism whereby pain disrupts participants' ability to perform arithmetic calculations. If that had been the case, we would have expected fewer correct answers in the pain condition than in the control condition. Nevertheless, in many studies, pain has been associated with lower accuracies and/or slower response times on cognitive tasks. Thinking about these more general patterns of findings, we can theorise mechanisms at the cognitive level as well as at the metacognitive level. At the cognitive level, executive functions may be useful for thinking about potential mechanisms. Executive functions are the processes by which we regulate our attention, thoughts and actions [28]. When we complete a task whilst in pain, we may repeatedly switch our attention between the pain and the task, focusing on each fully but for short periods, which could have engagement and disengagement costs in both time and accuracy: one 'loses track'. Alternatively, the pain and the task at hand may be processed simultaneously rather than alternately. In this scenario, rather than the task suffering from disengagement costs, it suffers from a continually lower working memory capacity being available. Inhibition skills may be important in controlling our attention deployment in either of these scenarios. Additionally, there are likely motivational [39] and attentional bias [23] elements to the deployment of attention between task- and pain-related goals.

1 If we can identify the mechanism(s) of the disruptive effect of pain on attention, and
2 these are found to be common across multiple tasks, then we may be able to develop
3 targeted interventions that mitigate its more extreme effects in everyday tasks that have
4 serious consequences of error [6]. Research is already making steps in this direction; Baker
5 et al [8] tested a broad programme of cognitive training as a way of improving cognition in
6 chronic pain patients, with some promising results.

7 In Study 1, we recruited participants reporting current pain and a history of pain.
8 Participants were blind to the task, and unaware of our hypotheses, but were not randomly
9 assigned to groups. This meant we could not test whether the relationships in question
10 were causal, and if so, in which direction. This shortcoming was remedied for the shopping
11 task in Study 2, where we found evidence for a causal effect of pain on shopping decisions
12 with a true experiment. Our finding that participants with a history of pain reported more
13 negative outcomes from real-life decisions should be conceptually replicated by giving the
14 DOI to a clinical chronic pain sample, with participants providing a more comprehensive
15 clinical history, and a matched control group. This would provide stronger evidence for a
16 claim that the effect is indeed due to the participants' pain and not due to a confound. Such
17 a study could consider how long a patient's pain has persisted, how often it occurs, and for
18 what proportion of time (e.g. per week) they generally experience pain. We would expect
19 that the more time a participant spends in pain per week, the greater the effect on their
20 decision making.

21 It is possible that the effect we observed on the DOI reflects a memory or
22 interpretation bias. For example, compared to people without recurrent pain, people with
23 recurrent pain may remember past experiences in a more negative light, or have a better
24 recollection of negative events, or be more likely to mis-remember events as having bad

1 outcomes due to an expectation that their pain will have disrupted their decision making. A
2 longitudinal study regularly measuring decision outcomes in a chronic pain sample and
3 matched control group would help to differentiate these explanations from our proposal
4 that pain disrupts decision making.

5 These studies begin to address an important gap in our knowledge. Although
6 numeracy, reasoning and decision making are core skills embedded in almost all everyday
7 tasks, we know little about the effects of embodied distractions on them. We know that
8 time pressure [17,37], perceptual disfluency [38], and working memory load [19,22,34] can
9 detrimentally affect performance on reasoning and decision-making tasks, but the effects of
10 embodied (interoceptive or proprioceptive) distractions such as pain, itch, imbalance,
11 hunger, thirst and fatigue are less investigated [16].

12 In summary, we have shown that both naturally-occurring and experimentally-
13 induced pain is associated with poorer decisions when making grocery purchasing decisions,
14 and that the negative effect of pain on cognition may accumulate over time in people with
15 recurrent or chronic pain. People with pain frequently report that the cognitive
16 consequences are fatiguing, distressing, and disruptive. Here we provide data that suggest
17 that the consequences of that disruption are not only psychological, but also financial.

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Tables

Table 1. Type of current and recurrent pain conditions reported in the Study 1 sample.

Percentages are of the number of participants reporting current/recurrent pain. Participants were able to select multiple options so the total is greater than 100%.

Type of pain	Current	Recurrent
Arthritis	45 (12%)	61 (15%)
Back (non-muscular)	90 (27%)	103 (26%)
Lower back	59 (18%)	63 (16%)
Upper back	57 (17%)	42 (10%)
Hangover	7 (2%)	19 (5%)
Headache	90 (27%)	89 (22%)
Joint	69 (21%)	86 (21%)
Menstrual	24 (7%)	62 (15%)
Muscular (non-back)	27 (8%)	21 (5%)
Migraine	12 (4%)	60 (15%)
Nerve	24 (7%)	21 (5%)
Post-surgical	5 (1%)	9 (2%)
Sciatica	21 (6%)	33 (8%)
Stomach	33 (10%)	44 (11%)
Throat	11 (3%)	6 (1%)
Tooth	39 (12%)	39 (10%)
Other	37 (11%)	33 (8%)

Table 2. Duration of current and recurrent pain conditions reported in the Study 1 sample.

Duration of pain	Current	Recurrent
Up to one hour	39 (12%)	-
Up to 24 hours	71 (21%)	-
Up to a week	47 (14%)	18 (4%)
Up to a month	28 (8%)	21 (5%)
Up to 3 months	18 (5%)	22 (5%)
Up to 6 months	12 (4%)	23 (6%)
Up to a year	45 (13%)	78 (19%)
Up to a decade	52 (15%)	146 (36%)
Over a decade	24 (7%)	93 (23%)

Table 3. Mean amount of money spent per unit for each item in the shopping decision task in the current pain and no current pain groups in Study 1.

	Best price available	No current pain average spend	Current pain average spend	Pain/no pain spend difference
Milk (per pint)	25p	27p	27p	0p
Lemons (per kilo)	£2.10	£2.26	£2.27	1p
Ketchup (per kilo)	£2.45	£2.61	£2.65	4p
Eggs (each)	11p	12p	12p	0p
Cheese (per kilogram)	£9.20	£9.64	£9.67	3p
Cola (per litre)	75p	80p	80p	0p

Figures

Figure 1. Shopping decision scores in Study 1 by current (yes/no) and recurrent (yes/no) pain with age as a covariate. There was a significant main effect of current pain, but no effect of recurrent pain and no interaction. Error bars show ± 1 standard error of the mean.

Figure 2. Decision Outcome Inventory scores in Study 1 by current (yes/no) and recurrent (yes/no) pain with age as a covariate. There was a significant main effect of recurrent pain, but no effect of current pain and no interaction. Error bars show ± 1 standard error of the mean.

- 1 Figure 3. Violin plot showing distributions and means of number of errors and number of
- 2 correct answers on the shopping decisions task by pain condition in Study 2. There was a
- 3 significant main effect of pain condition on errors but not on number of correct answers.
- 4 Error bars show ± 1 standard error of the mean.

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