

Supplement: Estimating latent baseline-by-treatment interactions in statistical mediation analysis

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1. Covariance expectations

Consider a two-wave mediation model representing a randomized pretest-posttest design.

$X \sim Bernoulli(p = 0.5)$, where $\mu_x = 0.5$, $\sigma_x^2 = 0.25$

$M_1 \sim N(0, 1)$

$Y_1 \sim N(0, 1)$, and $Cov(M_1, Y_1) = .575$

$$M_2 = i_3 + aX + s_m M_1 + s_{c1} Y_1 + h_1 X M_1 + h_2 X Y_1 + e_3$$

$$Y_2 = i_2 + c'X + bM_2 + s_y Y_1 + s_{c2} M_1 + h_3 X M_1 + h_4 X Y_1 + e_2$$

The implied covariance matrix of this model is:

	X	$M1$	$Y1$	$XM1$	$XY1$
X	σ_x^2				
$M1$	$\sigma_{x,m1}$	σ_{m1}^2			
$Y1$	$\sigma_{x,y1}$	$\sigma_{m1,y1}$	σ_{y1}^2		
$XM1$	$\sigma_x^2 \mu_{m1}$ + $\sigma_{x,m1} \mu_x$	$\sigma_{m1}^2 \mu_x$ + $\sigma_{x,m1} \mu_{m1}$	$\sigma_{m1,y1} \mu_x$ + $\sigma_{x,y1} \mu_{m1}$	$\sigma_{m1}^2 \mu_x^2 + \sigma_x^2 \mu_{m1}^2$ + $2\sigma_{x,m1} \mu_x \mu_{m1}$ + $\sigma_x^2 \sigma_{m1}^2 - \sigma_{x,m1}^2$	
$XY1$	$\sigma_x^2 \mu_{y1}$ + $\sigma_{x,y1} \mu_x$	$\sigma_{m1,y1} \mu_x$ + $\sigma_{x,y1} \mu_{m1}$	$\sigma_{x,y1} \mu_{y1}$ + $\sigma_{y1}^2 \mu_x$	$\sigma_x^2 \mu_{m1} \mu_{y1}$ + $\sigma_{x,y1} \mu_x \mu_{m1}$ + $\sigma_{x,m1} \mu_x \mu_{y1}$ + $\sigma_{m1,y1} \mu_x^2$ + $\sigma_x^2 \sigma_{m1,y1}$ + $\sigma_{x,y1} \sigma_{x,m1}$	$\sigma_{y1}^2 \mu_x^2 + \sigma_x^2 \mu_{y1}^2$ + $2\sigma_{x,y1} \mu_x \mu_{y1} + \sigma_x^2 \sigma_{y1}^2$ - $\sigma_{x,y1}^2$
$M2$	$a\sigma_x^2$ + $s_m \sigma_{x,m1}$ + $s_{c1} \sigma_{x,y1}$ + $h_1 \sigma_{x,xm1}$ + $h_2 \sigma_{x,xy1}$	$a\sigma_{x,m1}$ + $s_m \sigma_{m1}^2$ + $s_{c1} \sigma_{m1,y1}$ + $h_1 \sigma_{m1,xm1}$ + $h_2 \sigma_{m1,xy1}$	$a\sigma_{x,y1}$ + $s_m \sigma_{m1,y1}$ + $s_{c1} \sigma_{y1}^2$ + $h_1 \sigma_{y1,xm1}$ + $h_2 \sigma_{y1,xy1}$	$a\sigma_{x,xm1}$ + $s_m \sigma_{m1,xm1}$ + $s_{c1} \sigma_{y1,xm1}$ + $h_1 \sigma_{xm1}^2$ + $h_2 \sigma_{xm1,xy1}$	$a\sigma_{x,xy1} + s_m \sigma_{m1,xy1}$ + $s_{c1} \sigma_{y1,xy1} + h_1 \sigma_{xm1,xy1}$ + $h_2 \sigma_{xy1}^2$
$Y2$	$c' \sigma_x^2$ + $s_{c2} \sigma_{x,m1}$ + $s_y \sigma_{x,y1}$ + $b \sigma_{x,m2}$ + $h_3 \sigma_{x,xm1}$ + $h_4 \sigma_{x,xy1}$	$c' \sigma_{x,m1}$ + $s_{c2} \sigma_{m1}^2$ + $s_y \sigma_{m1,y1}$ + $b \sigma_{m1,m2}$ + $h_3 \sigma_{m1,xm1}$ + $h_4 \sigma_{m1,xy1}$	$c' \sigma_{x,y1}$ + $s_{c2} \sigma_{m1,y1}$ + $s_y \sigma_{y1}^2$ + $b \sigma_{y1,m2}$ + $h_3 \sigma_{y1,xm1}$ + $h_4 \sigma_{y1,xy1}$	$c' \sigma_{x,xm1}$ + $s_{c2} \sigma_{m1,xm1}$ + $s_y \sigma_{y1,xm1}$ + $b \sigma_{xm1,m2}$ + $h_3 \sigma_{xm1}^2$ + $h_4 \sigma_{xm1,y1}$	$c' \sigma_{x,xy1} + s_{c2} \sigma_{m1,xy1}$ + $s_y \sigma_{y1,xy1} + b \sigma_{xy1,m2}$ + $h_3 \sigma_{xm1,xy1} + h_4 \sigma_{xy1}^2$

	$M2$	$Y2$
$M2$	$a^2\sigma_x^2 + s_m^2\sigma_{m1}^2 + s_{c1}^2\sigma_{y1}^2 + h_1^2\sigma_{xm1}^2 + h_2^2\sigma_{xy1}^2$ $+ 2as_m\sigma_{x,m1} + 2as_{c1}\sigma_{x,y1}$ $+ 2ah_1\sigma_{x,xm1}$ $+ 2ah_2\sigma_{x,xy1}$ $+ 2s_ms_{c1}\sigma_{m1,y1}$ $+ 2s_mh_1\sigma_{m1,xm1}$ $+ 2s_mh_2\sigma_{m1,xy1}$ $+ 2s_{c1}h_1\sigma_{y1,xm1}$ $+ 2s_{c1}h_2\sigma_{y1,xy1}$ $+ 2h_1h_2\sigma_{xm1,xy1} + \sigma_{e_3}^2$	
$Y2$	$ac'\sigma_x^2 + ab\sigma_{x,m2} + as_y\sigma_{x,y1} + as_{c2}\sigma_{x,m1}$ $+ ah_3\sigma_{x,xm1} + ah_4\sigma_{x,xy1}$ $+ s_mc'\sigma_{m1,x} + s_mb\sigma_{m1,m2}$ $+ s_ms_y\sigma_{m1,y1} + s_ms_{c2}\sigma_{m1}^2$ $+ s_mh_3\sigma_{m1,xm1}$ $+ s_mh_4\sigma_{m1,xy1}$ $+ s_{c1}c'\sigma_{x,y1} + s_{c1}b\sigma_{y1,m2}$ $+ s_{c1}s_y\sigma_{y1}^2 + s_{c1}s_{c2}\sigma_{y1,m1}$ $+ s_{c1}h_3\sigma_{y1,xm1}$ $+ s_{c1}h_4\sigma_{y1,xy1}$ $+ h_1c'\sigma_{x,xm1} + h_1b\sigma_{m2,xm1}$ $+ h_1s_y\sigma_{y1,xm1}$ $+ h_1s_{c2}\sigma_{m1,xm1}$ $+ h_1h_3\sigma_{xm1}^2$ $+ h_1h_4\sigma_{xm1,xy1}$ $+ h_2c'\sigma_{x,xy1} + h_2b\sigma_{m2,xy1}$ $+ h_2s_y\sigma_{y1,xy1}$ $+ h_2s_{c2}\sigma_{m1,xy1}$ $+ h_2h_3\sigma_{xm1,xy1}$ $+ h_2h_4\sigma_{xy1}^2 + b\sigma_{e_3}^2$	$c'^2\sigma_x^2 + b^2\sigma_{m2}^2 + s_{y1}^2\sigma_{y1}^2$ $+ s_{c2}^2\sigma_{m1}^2 + h_3^2\sigma_{xm1}^2 + h_4^2\sigma_{xy1}^2$ $+ 2bc'\sigma_{x,m2} + 2c's_y\sigma_{x,y1}$ $+ 2c's_{c2}\sigma_{x,m1} + 2c'h_3\sigma_{x,xm1}$ $+ 2c'h_4\sigma_{x,xy1} + 2bs_y\sigma_{m2,y1}$ $+ 2bs_{c2}\sigma_{m1,m2} + 2bh_3\sigma_{m2,xm1}$ $+ 2bh_4\sigma_{m2,xy1} + 2s_y s_{c2}\sigma_{m1,y1}$ $+ 2s_yh_3\sigma_{y1,xm1} + 2s_yh_4\sigma_{y1,xy1}$ $+ 2s_{c2}h_3\sigma_{m1,xm1}$ $+ 2s_{c2}h_4\sigma_{m1,xy1}$ $+ 2h_3h_4\sigma_{xm1,xy1} + \sigma_{e_2}^2$

Recall that $E[XY] = \text{Cov}(X,Y) + E[X]*E[Y]$. For a situation in which we mean-center X, M1, and Y1 so that μ_x , μ_{m1} , and μ_{y1} are zero, and X is randomized so that $\text{Cov}(X, m1) = \text{Cov}(X, y1) = 0$, the covariance algebra of the model reduces to:

	X	$M1$	$Y1$	$XM1$	$XY1$
X	σ_x^2				
$M1$	0	σ_{m1}^2			
$Y1$	0	0	σ_{y1}^2		
$XM1$	0	0	0	$\sigma_x^2 \sigma_{m1}^2$	
$XY1$	0	0	0	$\sigma_x^2 \sigma_{m1,y1}$	$\sigma_x^2 \sigma_{y1}^2$
$M2$	$a\sigma_x^2$	$s_m \sigma_{m1}^2 + s_{c1} \sigma_{m1,y1}$	$s_m \sigma_{m1,y1} + s_{c1} \sigma_{y1}^2$	$h_1 \sigma_{xm1}^2 + h_2 \sigma_{xm1,xy1}$	$h_1 \sigma_{xm1,xy1} + h_2 \sigma_{xy1}^2$
$Y2$	$c' \sigma_x^2 + b \sigma_{x,m2}$	$s_{c2} \sigma_{m1}^2 + s_y \sigma_{m1,y1} + b \sigma_{m1,m2}$	$s_{c2} \sigma_{m1,y1} + s_y \sigma_{y1}^2 + b \sigma_{y1,m2}$	$b \sigma_{xm1,m2} + h_3 \sigma_{xm1}^2 + h_4 \sigma_{xm1,xy1}$	$b \sigma_{xy1,m2} + h_3 \sigma_{xm1,xy1} + h_4 \sigma_{xy1}^2$

	$M2$	$Y2$
$M2$	$a^2 \sigma_x^2 + s_m^2 \sigma_{m1}^2 + s_{c1}^2 \sigma_{y1}^2 + h_1^2 \sigma_{xm1}^2 + h_2^2 \sigma_{xy1}^2 + 2s_m s_{c1} \sigma_{m1,y1} + 2h_1 h_2 \sigma_{xm1,xy1} + \sigma_{e_3}^2$	
$Y2$	$ac' \sigma_x^2 + ab \sigma_{x,m2} + s_m b \sigma_{m1,m2} + s_m s_y \sigma_{m1,y1} + s_m s_{c2} \sigma_{m1}^2 + s_{c1} b \sigma_{y1,m2} + s_{c1} s_y \sigma_{y1}^2 + s_{c1} s_{c2} \sigma_{y1,m1} + h_1 b \sigma_{m2,xm1} + h_1 h_3 \sigma_{xm1}^2 + h_1 h_4 \sigma_{xm1,xy1} + h_2 b \sigma_{m2,xy1} + h_2 h_3 \sigma_{xm1,xy1} + h_2 h_4 \sigma_{xy1}^2 + b \sigma_{e_3}^2$	$c'^2 \sigma_x^2 + b^2 \sigma_{m2}^2 + s_{y1}^2 \sigma_{y1}^2 + s_{c2}^2 \sigma_{m1}^2 + h_3^2 \sigma_{xm1}^2 + h_4^2 \sigma_{xy1}^2 + 2bc' \sigma_{x,m2} + 2bs_y \sigma_{m2,y1} + 2bs_{c2} \sigma_{m1,m2} + 2bh_3 \sigma_{m2,xm1} + 2bh_4 \sigma_{m2,xy1} + 2s_y s_{c2} \sigma_{m1,y1} + 2h_3 h_4 \sigma_{xm1,xy1} + \sigma_{e_2}^2$

If we substitute everything in terms of exogenous variables, we get:

	X	$M1$	$Y1$	$XM1$	$XY1$
X	σ_x^2				
$M1$	0	σ_{m1}^2			
$Y1$	0	0	σ_{y1}^2		
$XM1$	0	0	0	$\sigma_x^2 \sigma_{m1}^2$	
$XY1$	0	0	0	$\sigma_x^2 \sigma_{m1,y1}$	$\sigma_x^2 \sigma_{y1}^2$
$M2$	$a\sigma_x^2$	$s_m \sigma_{m1}^2 + s_{c1} \sigma_{m1,y1}$	$s_m \sigma_{m1,y1} + s_{c1} \sigma_{y1}^2$	$h_1(\sigma_x^2 \sigma_{m1}^2) + h_2(\sigma_x^2 \sigma_{y1}^2)$	$h_1(\sigma_x^2 \sigma_{m1,y1}) + h_2(\sigma_x^2 \sigma_{y1}^2)$
$Y2$	$c'\sigma_x^2 + b(a\sigma_x^2)$	$s_{c2} \sigma_{m1}^2 + s_y \sigma_{m1,y1} + b(s_m \sigma_{m1}^2 + s_{c1} \sigma_{m1,y1})$	$s_{c2} \sigma_{m1,y1} + s_y \sigma_{y1}^2 + b(s_m \sigma_{m1,y1} + s_{c1} \sigma_{y1}^2)$	$b(h_1(\sigma_x^2 \sigma_{m1}^2) + h_2(\sigma_x^2 \sigma_{m1,y1})) + h_3(\sigma_x^2 \sigma_{m1}^2) + h_4(\sigma_x^2 \sigma_{m1,y1})$	$b(h_1(\sigma_x^2 \sigma_{m1,y1}) + h_2(\sigma_x^2 \sigma_{y1}^2)) + h_3(\sigma_x^2 \sigma_{m1,y1}) + h_4(\sigma_x^2 \sigma_{y1}^2)$

	$M2$	$Y2$
$M2$	$a^2 \sigma_x^2 + s_m^2 \sigma_{m1}^2 + s_{c1}^2 \sigma_{y1}^2 + h_1^2(\sigma_x^2 \sigma_{m1}^2) + h_2^2(\sigma_x^2 \sigma_{y1}^2) + 2s_m s_{c1} \sigma_{m1,y1} + 2h_1 h_2(\sigma_x^2 \sigma_{m1,y1}) + \sigma_{e_3}^2$	
$Y2$	$ac'\sigma_x^2 + ab(a\sigma_x^2) + s_m b(s_m \sigma_{m1}^2 + s_{c1} \sigma_{m1,y1}) + s_m s_y \sigma_{m1,y1} + s_m s_{c2} \sigma_{m1}^2 + s_{c1} b(s_m \sigma_{m1,y1} + s_{c1} \sigma_{y1}^2) + s_{c1} s_{c2} \sigma_{y1,m1} + h_1 b(h_1(\sigma_x^2 \sigma_{m1}^2) + h_2(\sigma_x^2 \sigma_{m1,y1})) + h_1 h_3(\sigma_x^2 \sigma_{m1}^2) + h_1 h_4(\sigma_x^2 \sigma_{m1,y1}) + h_2 b(h_1(\sigma_x^2 \sigma_{m1,y1}) + h_2(\sigma_x^2 \sigma_{y1}^2)) + h_2 h_3(\sigma_x^2 \sigma_{m1,y1}) + h_2 h_4(\sigma_x^2 \sigma_{y1}^2) + b\sigma_{e_3}^2$	$c'^2 \sigma_x^2 + b^2(a^2 \sigma_x^2 + s_m^2 \sigma_{m1}^2 + s_{c1}^2 \sigma_{y1}^2) + h_1^2(\sigma_x^2 \sigma_{m1}^2) + h_2^2(\sigma_x^2 \sigma_{y1}^2) + 2s_m s_{c1} \sigma_{m1,y1} + 2h_1 h_2(\sigma_x^2 \sigma_{m1,y1}) + \sigma_{e_3}^2 + s_{y1}^2 \sigma_{y1}^2 + s_{c2}^2 \sigma_{m1}^2 + h_3^2(\sigma_x^2 \sigma_{m1}^2) + h_4^2(\sigma_x^2 \sigma_{y1}^2) + 2bc'(a\sigma_x^2) + 2bs_y(s_m \sigma_{m1,y1} + s_{c1} \sigma_{y1}^2) + 2bs_{c2}(s_m \sigma_{m1}^2 + s_{c1} \sigma_{m1,y1}) + 2bh_3(h_1(\sigma_x^2 \sigma_{m1}^2)) + h_2(\sigma_x^2 \sigma_{m1,y1}) + 2bh_4(h_1(\sigma_x^2 \sigma_{m1}^2)) + h_2(\sigma_x^2 \sigma_{m1,y1}) + 2s_y s_{c2} \sigma_{m1,y1} + 2h_3 h_4(\sigma_x^2 \sigma_{m1,y1}) + \sigma_{e_2}^2$

Now, we if we assume no cross-lags ($s_{c1} = s_{c2} = 0$), the covariance matrix reduces to:

	X	$M1$	$Y1$	$XM1$	$XY1$
X	σ_x^2				
$M1$	0	σ_{m1}^2			
$Y1$	0	0	σ_{y1}^2		
$XM1$	0	0	0	$\sigma_x^2 \sigma_{m1}^2$	
$XY1$	0	0	0	$\sigma_x^2 \sigma_{m1,y1}$	$\sigma_x^2 \sigma_{y1}^2$
$M2$	$a\sigma_x^2$	$s_m \sigma_{m1}^2$	$s_m \sigma_{m1,y1}$	$h_1 \sigma_{xm1}^2 + h_2 \sigma_{xm1,xy1}$	$h_1 \sigma_{xm1,xy1} + h_2 \sigma_{xy1}^2$
$Y2$	$c' \sigma_x^2 + b \sigma_{x,m2}$	$s_y \sigma_{m1,y1} + b \sigma_{m1,m2}$	$s_y \sigma_{y1}^2 + b \sigma_{y1,m2}$	$b \sigma_{xm1,m2} + h_3 \sigma_{xm1,xy1} + h_4 \sigma_{xm1,xy1}$	

	$M2$	$Y2$
$M2$	$a^2 \sigma_x^2 + s_m^2 \sigma_{m1}^2 + h_1^2 \sigma_{xm1}^2 + h_2^2 \sigma_{xy1}^2 + 2h_1 h_2 \sigma_{xm1,xy1} + \sigma_{e_3}^2$	
$Y2$	$ac' \sigma_x^2 + ab \sigma_{x,m2} + s_m b \sigma_{m1,m2} + s_m s_y \sigma_{m1,y1} + h_1 b \sigma_{m2,xm1} + h_1 h_3 \sigma_{xm1}^2 + h_1 h_4 \sigma_{xm1,xy1} + h_2 b \sigma_{m2,xy1} + h_2 h_3 \sigma_{xm1,xy1} + h_2 h_4 \sigma_{xy1}^2 + b \sigma_{e_3}^2$	$c'^2 \sigma_x^2 + b^2 \sigma_{m2}^2 + s_y^2 \sigma_{y1}^2 + h_3^2 \sigma_{xm1}^2 + h_4^2 \sigma_{xy1}^2 + 2bc' \sigma_{x,m2} + 2bs_y \sigma_{m2,y1} + 2bh_3 \sigma_{m2,xm1} + 2bh_4 \sigma_{m2,xy1} + 2h_3 h_4 \sigma_{xm1,xy1} + \sigma_{e_2}^2$

If we substitute terms for only exogenous variables, we get:

	X	$M1$	$Y1$	$XM1$	$XY1$
X	σ_x^2				
$M1$	0	σ_{m1}^2			
$Y1$	0	0	σ_{y1}^2		
$XM1$	0	0	0	$\sigma_x^2 \sigma_{m1}^2$	
$XY1$	0	0	0	$\sigma_x^2 \sigma_{m1,y1}$	$\sigma_x^2 \sigma_{y1}^2$
$M2$	$a\sigma_x^2$	$s_m \sigma_{m1}^2$	$s_m \sigma_{m1,y1}$	$h_1(\sigma_x^2 \sigma_{m1}^2) + h_2(\sigma_x^2 \sigma_{m1,y1})$	$h_1(\sigma_x^2 \sigma_{m1,y1}) + h_2(\sigma_x^2 \sigma_{y1}^2)$
$Y2$	$c' \sigma_x^2 + b(a\sigma_x^2)$	$s_y \sigma_{m1,y1} + b(s_m \sigma_{m1,y1})$	$s_y \sigma_{y1}^2 + b(s_m \sigma_{m1,y1})$	$b(h_1(\sigma_x^2 \sigma_{m1}^2) + h_2(\sigma_x^2 \sigma_{m1,y1})) + h_3(\sigma_x^2 \sigma_{m1}^2) + h_4(\sigma_x^2 \sigma_{m1,y1})$	$b(h_1(\sigma_x^2 \sigma_{m1}^2) + h_2(\sigma_x^2 \sigma_{m1,y1})) + h_3(\sigma_x^2 \sigma_{m1,y1}) + h_4(\sigma_x^2 \sigma_{y1}^2)$

	M2	Y2
M2	$a^2\sigma_x^2 + s_m^2\sigma_{m1}^2 + h_1^2(\sigma_x^2\sigma_{m1}^2) + h_2^2(\sigma_x^2\sigma_{y1}^2)$ $+ 2h_1h_2(\sigma_x^2\sigma_{m1,y1}) + \sigma_{e_3}^2$	
Y2	$ac'\sigma_x^2 + ab(a\sigma_x^2) + s_m b(s_m\sigma_{m1}^2)$ $+ s_m s_y \sigma_{m1,y1}$ $+ h_1 b(h_1(\sigma_x^2\sigma_{m1}^2))$ $+ h_2(\sigma_x^2\sigma_{m1,y1})$ $+ h_1 h_3(\sigma_x^2\sigma_{m1}^2)$ $+ h_1 h_4(\sigma_x^2\sigma_{m1,y1})$ $+ h_2 b(h_1(\sigma_x^2\sigma_{m1}^2))$ $+ h_2(\sigma_x^2\sigma_{m1,y1})$ $+ h_2 h_3(\sigma_x^2\sigma_{m1,y1})$ $+ h_2 h_4(\sigma_x^2\sigma_{y1}^2) + b\sigma_{e_3}^2$	$c'^2\sigma_x^2$ $+ b^2(a^2\sigma_x^2 + s_m^2\sigma_{m1}^2)$ $+ h_1^2(\sigma_x^2\sigma_{m1}^2) + h_2^2(\sigma_x^2\sigma_{y1}^2)$ $+ 2h_1h_2(\sigma_x^2\sigma_{m1,y1}) + \sigma_{e_3}^2$ $+ s_{y1}^2\sigma_{y1}^2 + h_3^2(\sigma_x^2\sigma_{m1}^2)$ $+ h_4^2(\sigma_x^2\sigma_{y1}^2) + 2bc'(a\sigma_x^2)$ $+ 2bs_y(s_m\sigma_{m1,y1})$ $+ 2bh_3(h_1(\sigma_x^2\sigma_{m1}^2))$ $+ h_2(\sigma_x^2\sigma_{m1,y1})$ $+ 2bh_4(h_1(\sigma_x^2\sigma_{m1,y1}))$ $+ h_2(\sigma_x^2\sigma_{y1}^2)) + 2h_3h_4(\sigma_x^2\sigma_{m1,y1})$ $+ \sigma_{e_2}^2$

2. Additional Tables

Table 1. Bias estimates and interval coverage of the parameters from the mediation model with latent BTIs with discrete, nonnormal indicators

	Relative Bias						Standardized Bias					
	a-path	b-path	c'-path	sm-path	sy-path	ab	cy-path	h1	h2	cm-path	h3	h4
Summed	-0.100	-0.041	-0.119	-0.118	-0.163	-0.138	-0.007	-0.043	-0.068	0.008	-0.094	-0.113
Factor	-0.071	-0.050	-0.088	-0.123	-0.167	-0.121	0.002	-0.059	-0.075	0.008	-0.104	-0.132
UPI	0.008	-0.103	0.028	-0.143	-0.179	-0.098	0.165	-0.020	-0.032	0.242	-0.073	-0.064
LMS	0.008	-0.105	0.028	-0.141	-0.178	-0.101	0.147	0.012	-0.111	0.240	-0.102	-0.109
Bayes	0.008	-0.108	0.022	-0.130	-0.173	-0.106	0.133	0.017	-0.099	0.202	-0.089	-0.094
<i>Raw Bias</i>												
Summed	-0.099	-0.026	-0.142	-0.116	-0.137	-0.062	-0.006	-0.006	-0.013	-0.009	-0.037	-0.022
Factor	-0.025	-0.019	-0.025	-0.122	-0.141	-0.013	-0.009	-0.010	-0.018	-0.005	-0.023	-0.025
UPI	0.002	-0.033	0.008	-0.141	-0.151	-0.011	0.012	-0.003	-0.007	0.026	-0.014	-0.011
LMS	0.002	-0.034	0.008	-0.139	-0.150	-0.011	0.010	0.004	-0.022	0.026	-0.018	-0.017
Bayes	0.002	-0.034	0.006	-0.128	-0.146	-0.012	0.008	0.006	-0.021	0.025	-0.018	-0.016
<i>Interval Coverage</i>												
Summed	0.939	0.947	0.937	0.853	0.731	0.933	0.947	0.954	0.951	0.950	0.951	0.948
Factor	0.942	0.947	0.942	0.830	0.702	0.934	0.949	0.954	0.953	0.952	0.952	0.947
UPI	0.944	0.907	0.939	0.721	0.552	0.932	0.944	0.946	0.947	0.939	0.945	0.943
LMS	0.944	0.906	0.939	0.735	0.561	0.931	0.946	0.944	0.946	0.941	0.945	0.939
Bayes	0.962	0.942	0.966	0.783	0.649	0.952	0.961	0.965	0.959	0.972	0.968	0.967

Note: There were seven categories with thresholds at -1.43, -.43, .38, .94, 1.44, 2.53. These were taken from Rhemtulla et al. (2012). True values of the a-path were .285 and .39, which are roughly both small effects.

Table 2. Statistical power of the mediation model parameters across methods for conditions with discrete, nonnormal indicators

	N = 250												
	a-path (s)	a-path (~s)	b-path (s)	b-path (m)	c'-path (s)	h1 (s)	h2 (s)	h3 (s)	h4 (s)	ab (s,s)	ab (m,s)	ab (s,m)	ab (m,m)
Summed	0.422	0.637	0.244	0.899	0.427	0.155	0.158	0.153	0.168	0.089	0.130	0.360	0.554
Factor	0.435	0.652	0.250	0.909	0.453	0.156	0.167	0.157	0.181	0.092	0.139	0.377	0.575
UPI	0.492	0.717	0.327	0.989	0.553	0.219	0.219	0.215	0.264	0.142	0.209	0.475	0.698
LMS	0.492	0.720	0.327	0.989	0.550	0.242	0.206	0.212	0.258	0.141	0.209	0.473	0.704
Bayes	0.419	0.659	0.252	0.978	0.457	0.192	0.184	0.155	0.190	0.082	0.134	0.396	0.628
	N = 500												
Summed	0.722	0.920	0.504	0.996	0.764	0.299	0.304	0.321	0.362	0.330	0.447	0.723	0.914
Factor	0.728	0.924	0.514	0.997	0.774	0.306	0.317	0.335	0.379	0.341	0.456	0.733	0.919
UPI	0.769	0.945	0.602	1.000	0.829	0.382	0.400	0.409	0.475	0.430	0.553	0.774	0.944
LMS	0.766	0.946	0.600	1.000	0.829	0.437	0.381	0.419	0.477	0.430	0.549	0.772	0.945
Bayes	0.726	0.932	0.528	1.000	0.775	0.367	0.350	0.333	0.388	0.343	0.471	0.736	0.931

Note: Power estimates for the stability of M and Y are not shown because the power was almost 1 across conditions. (s) is for small effect size and (m) is for medium effect size. There were seven categories with thresholds at -1.43, -.43, .38, .94, 1.44, 2.53. These were taken from Rhemtulla et al. (2012). True values of the a-path were .285 and .39, which are roughly both small effects.

Table 3. Type I error for parameters from the mediation model across methods in conditions with discrete, nonnormal indicators

	cy-path	cm-path	h1	h2	h3	h4
N = 250						
Summed	0.053	0.047	0.040	0.046	0.042	0.040
Factor	0.049	0.046	0.037	0.042	0.039	0.038
UPI	0.056	0.057	0.053	0.052	0.047	0.050
LMS	0.052	0.053	0.056	0.052	0.048	0.052
Bayes	0.038	0.025	0.036	0.040	0.028	0.028
N = 500						
Summed	0.053	0.053	0.050	0.047	0.050	0.056
Factor	0.052	0.051	0.050	0.047	0.049	0.055
UPI	0.056	0.065	0.056	0.053	0.053	0.063
LMS	0.055	0.063	0.068	0.050	0.052	0.067
Bayes	0.040	0.031	0.043	0.038	0.031	0.038

Note: There were seven categories with thresholds at -1.43, -.43, .38, .94, 1.44, 2.53. These were taken from Rhemtulla et al. (2012). True values of the a-path were .285 and .39, which are roughly both small effects.

3. Additional Figure

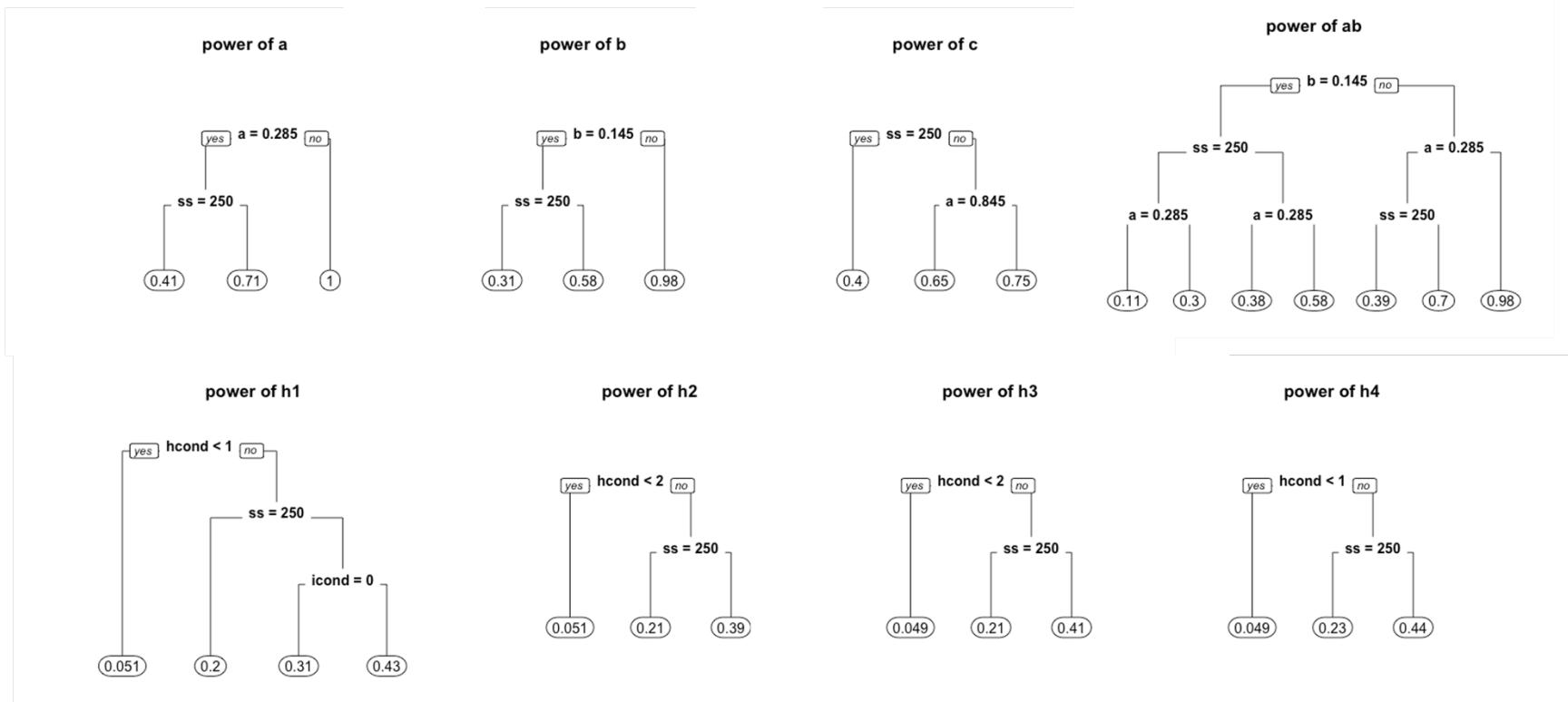
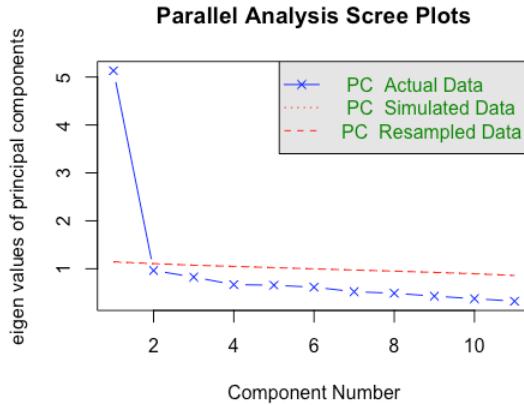
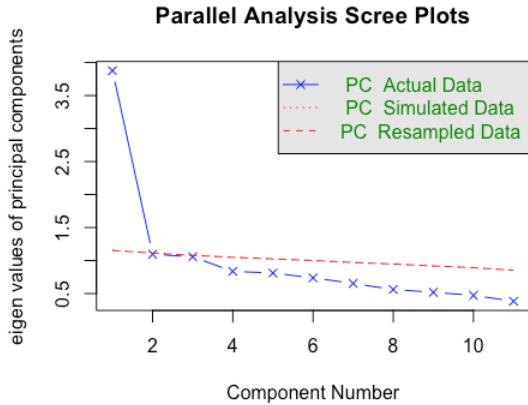
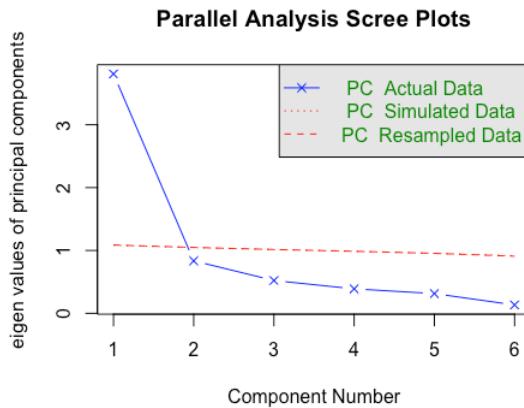
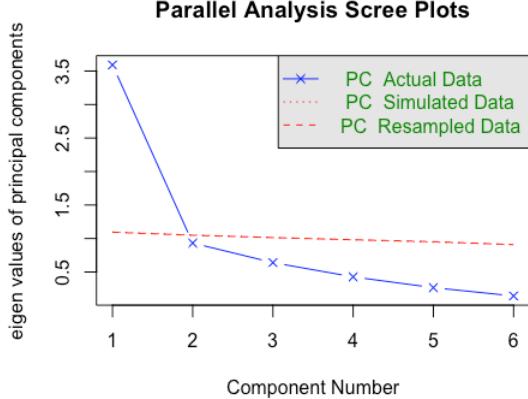


Figure S1. Trees predicting power from simulation conditions for conditions with normally-distributed indicators. Note: $hcond = 0$ is $h1-h4 = 0$, $hcond = 1$ is $h1 \& h4 = .3$ and $h2 \& h3 = 0$, $hcond = 2$ is $h1-h4 = .3$, $icond = 0$ is 4 items for M and 8 items for Y, $icond = 1$ is 6 items for M and Y, ss is sample size.

4. Psychometric Analyses of the Illustration Data



Depression at baseline and posttest, respectively



Job search self-efficacy at baseline and posttest, respectively

Psychometric analyses for the factor analyses models per scale per timepoint, along with testing for longitudinal scalar invariance. Per our analyses, it seems like longitudinal scalar invariance models fit generally well to serve in our illustration.

DEP at T0 CFA

```
## lavaan 0.6-7 ended normally after 25 iterations
##
##    Estimator          DWLS
## Optimization method      NLMINB
## Number of free parameters        22
##
##                               Used     Total
## Number of observations       1170      1285
##
## Model Test User Model:
##
##    Test statistic           135.642
##    Degrees of freedom            44
##    P-value (Chi-square)        0.000
##
## Model Test Baseline Model:
##
##    Test statistic           4942.611
##    Degrees of freedom            55
##    P-value                      0.000
##
## User Model versus Baseline Model:
##
##    Comparative Fit Index (CFI)      0.981
##    Tucker-Lewis Index (TLI)          0.977
##
## Root Mean Square Error of Approximation:
##
##    RMSEA                     0.042
##    90 Percent confidence interval - lower  0.034
##    90 Percent confidence interval - upper  0.050
##    P-value RMSEA <= 0.05            0.941
##
## Standardized Root Mean Square Residual:
##
##    SRMR                      0.049
##
## Parameter Estimates:
##
##    Standard errors             Standard
##    Information                  Expected
##    Information saturated (h1) model Unstructured
##
```

```

## Latent Variables:
##          Estimate Std.Err z-value P(>|z|) Std.lv Std.all
## f1 =~
##   DEP01      0.378  0.018  21.440  0.000  0.378  0.369
##   DEP02      0.681  0.022  31.556  0.000  0.681  0.587
##   DEP03      0.737  0.021  35.467  0.000  0.737  0.667
##   DEP04      0.468  0.018  26.208  0.000  0.468  0.492
##   DEP05      0.632  0.021  30.205  0.000  0.632  0.565
##   DEP06      0.612  0.019  31.722  0.000  0.612  0.598
##   DEP07      0.608  0.018  33.882  0.000  0.608  0.661
##   DEP08      0.774  0.020  37.813  0.000  0.774  0.763
##   DEP09      0.076  0.008  9.596   0.000  0.076  0.176
##   DEP010     0.245  0.012  19.835  0.000  0.245  0.343
##   DEP011     0.379  0.016  23.191  0.000  0.379  0.430
##
## Variances:
##          Estimate Std.Err z-value P(>|z|) Std.lv Std.all
## .DEP01      0.905  0.043  21.226  0.000  0.905  0.864
## .DEP02      0.882  0.051  17.207  0.000  0.882  0.656
## .DEP03      0.679  0.051  13.214  0.000  0.679  0.555
## .DEP04      0.686  0.052  13.183  0.000  0.686  0.758
## .DEP05      0.853  0.055  15.643  0.000  0.853  0.681
## .DEP06      0.674  0.052  13.032  0.000  0.674  0.643
## .DEP07      0.478  0.041  11.652  0.000  0.478  0.564
## .DEP08      0.429  0.049  8.800   0.000  0.429  0.417
## .DEP09      0.181  0.030  5.949   0.000  0.181  0.969
## .DEP010     0.449  0.042  10.695  0.000  0.449  0.882
## .DEP011     0.633  0.050  12.638  0.000  0.633  0.815
## f1         1.000
##
## R-Square:
##          Estimate
## DEP01      0.136
## DEP02      0.344
## DEP03      0.445
## DEP04      0.242
## DEP05      0.319
## DEP06      0.357
## DEP07      0.436
## DEP08      0.583
## DEP09      0.031
## DEP010     0.118
## DEP011     0.185

modindices(fit.dep0, sort=TRUE)

##      lhs op    rhs    mi    epc sepc.lv sepc.all sepc.nox
## 24  DEP01 ~~  DEP02 25.447  0.201   0.201    0.225    0.225
## 78  DEP010 ~~ DEP011 19.190  0.122   0.122    0.229    0.229

```

```

## 40  DEP02 ~~  DEP09  8.793 -0.046  -0.046   -0.115   -0.115

DEP at T2 CFA

## lavaan 0.6-7 ended normally after 25 iterations
##
##    Estimator                      DWLS
##    Optimization method             NLMINB
##    Number of free parameters      22
##
##                               Used     Total
##    Number of observations        1254    1285
##
## Model Test User Model:
##
##    Test statistic                69.296
##    Degrees of freedom            44
##    P-value (Chi-square)          0.009
##
## Model Test Baseline Model:
##
##    Test statistic                6848.752
##    Degrees of freedom            55
##    P-value                        0.000
##
## User Model versus Baseline Model:
##
##    Comparative Fit Index (CFI)    0.996
##    Tucker-Lewis Index (TLI)        0.995
##
## Root Mean Square Error of Approximation:
##
##    RMSEA                         0.021
##    90 Percent confidence interval - lower 0.011
##    90 Percent confidence interval - upper 0.031
##    P-value RMSEA <= 0.05           1.000
##
## Standardized Root Mean Square Residual:
##
##    SRMR                          0.041
##
## Parameter Estimates:
##
##    Standard errors                 Standard
##    Information                     Expected
##    Information saturated (h1) model Unstructured
##
## Latent Variables:
##                  Estimate  Std.Err  z-value  P(>|z|)  Std.lv  Std.all

```

```

## f1 =~
##   DEP21      0.486    0.017  29.366    0.000    0.486    0.536
##   DEP22      0.772    0.020  39.049    0.000    0.772    0.713
##   DEP23      0.683    0.019  35.448    0.000    0.683    0.705
##   DEP24      0.555    0.019  29.262    0.000    0.555    0.583
##   DEP25      0.703    0.019  37.004    0.000    0.703    0.675
##   DEP26      0.796    0.021  37.961    0.000    0.796    0.690
##   DEP27      0.730    0.019  38.483    0.000    0.730    0.719
##   DEP28      0.870    0.021  42.136    0.000    0.870    0.814
##   DEP29      0.233    0.013  18.425    0.000    0.233    0.439
##   DEP210     0.413    0.016  26.198    0.000    0.413    0.517
##   DEP211     0.664    0.020  33.400    0.000    0.664    0.616
##
## Variances:
##             Estimate Std.Err z-value P(>|z|) Std.lv Std.all
##   .DEP21      0.585   0.048 12.304   0.000   0.585   0.712
##   .DEP22      0.578   0.055 10.461   0.000   0.578   0.492
##   .DEP23      0.472   0.055  8.539   0.000   0.472   0.503
##   .DEP24      0.599   0.058 10.375   0.000   0.599   0.660
##   .DEP25      0.591   0.052 11.272   0.000   0.591   0.544
##   .DEP26      0.697   0.063 11.117   0.000   0.697   0.524
##   .DEP27      0.498   0.054  9.301   0.000   0.498   0.483
##   .DEP28      0.386   0.058  6.673   0.000   0.386   0.338
##   .DEP29      0.228   0.041  5.618   0.000   0.228   0.807
##   .DEP210     0.467   0.048  9.754   0.000   0.467   0.732
##   .DEP211     0.723   0.061 11.791   0.000   0.723   0.621
##   f1         1.000
##
## R-Square:
##             Estimate
##   DEP21      0.288
##   DEP22      0.508
##   DEP23      0.497
##   DEP24      0.340
##   DEP25      0.456
##   DEP26      0.476
##   DEP27      0.517
##   DEP28      0.662
##   DEP29      0.193
##   DEP210     0.268
##   DEP211     0.379

```

SE at T1 CFA

```

## lavaan 0.6-7 ended normally after 19 iterations
## 
##   Estimator           DWLS
##   Optimization method NLMINB
##   Number of free parameters          12
## 

```

```

##                                         Used      Total
##   Number of observations                 1261     1285
##
## Model Test User Model:
##
##   Test statistic                      165.218
##   Degrees of freedom                   9
##   P-value (Chi-square)                0.000
##
## Model Test Baseline Model:
##
##   Test statistic                      3887.411
##   Degrees of freedom                  15
##   P-value                            0.000
##
## User Model versus Baseline Model:
##
##   Comparative Fit Index (CFI)        0.960
##   Tucker-Lewis Index (TLI)           0.933
##
## Root Mean Square Error of Approximation:
##
##   RMSEA                           0.117
##   90 Percent confidence interval - lower 0.102
##   90 Percent confidence interval - upper 0.133
##   P-value RMSEA <= 0.05            0.000
##
## Standardized Root Mean Square Residual:
##
##   SRMR                            0.094
##
## Parameter Estimates:
##
##   Standard errors                  Standard
##   Information                      Expected
##   Information saturated (h1) model Unstructured
##
## Latent Variables:
##             Estimate Std.Err z-value P(>|z|) Std.lv Std.all
##   f1 =~
##     SE11       0.616  0.020  30.989  0.000  0.616  0.633
##     SE12       0.784  0.023  34.448  0.000  0.784  0.745
##     SE13       0.809  0.024  34.006  0.000  0.809  0.736
##     SE14       0.724  0.023  31.144  0.000  0.724  0.695
##     SE15       0.851  0.025  34.713  0.000  0.851  0.749
##     SE16       0.814  0.024  33.370  0.000  0.814  0.739
##
## Variances:
##             Estimate Std.Err z-value P(>|z|) Std.lv Std.all
##   .SE11      0.569  0.047  11.998  0.000  0.569  0.600

```

```

##   .SE12          0.492  0.054  9.078  0.000  0.492  0.445
##   .SE13          0.554  0.058  9.553  0.000  0.554  0.458
##   .SE14          0.561  0.059  9.551  0.000  0.561  0.517
##   .SE15          0.565  0.061  9.290  0.000  0.565  0.438
##   .SE16          0.550  0.061  9.062  0.000  0.550  0.453
##   f1            1.000
##
## R-Square:
##             Estimate
##   SE11        0.400
##   SE12        0.555
##   SE13        0.542
##   SE14        0.483
##   SE15        0.562
##   SE16        0.547

modindices(fit.se1, sort=TRUE, minimum.value = 10)

##      lhs op rhs      mi     epc sepc.lv sepc.all sepc.nox
## 19 SE12 ~~ SE13 137.089  0.598  0.598    1.145    1.145
## 28 SE15 ~~ SE16  44.674  0.368  0.368    0.660    0.660
## 16 SE11 ~~ SE14  19.082  0.188  0.188    0.332    0.332
## 22 SE12 ~~ SE16  18.498 -0.214 -0.214   -0.411   -0.411
## 25 SE13 ~~ SE16  17.300 -0.216 -0.216   -0.391   -0.391

se1.cfa.cor<- 'f1 =~ SE11 + SE12 + SE13 + SE14+ SE15 + SE16
SE12~~SE13'

fit.se1.cor <- cfa(se1.cfa.cor, data=se1, std.lv=TRUE, estimator = 'DWLS')
summary(fit.se1.cor, fit.measures=TRUE, standardized=TRUE, rsquare=TRUE)

## lavaan 0.6-7 ended normally after 23 iterations
##
##   Estimator           DWLS
##   Optimization method NLMINB
##   Number of free parameters      13
##
##   Used      Total
##   Number of observations       1261      1285
##
## Model Test User Model:
##
##   Test statistic        27.852
##   Degrees of freedom          8
##   P-value (Chi-square)      0.001
##
## Model Test Baseline Model:
##
##   Test statistic        3887.411
##   Degrees of freedom         15
##   P-value                  0.000

```

```

## 
## User Model versus Baseline Model:
## 

## Comparative Fit Index (CFI)          0.995
## Tucker-Lewis Index (TLI)           0.990
## 

## Root Mean Square Error of Approximation:
## 

## RMSEA                                0.044
## 90 Percent confidence interval - lower 0.027
## 90 Percent confidence interval - upper 0.063
## P-value RMSEA <= 0.05                0.665
## 

## Standardized Root Mean Square Residual:
## 

## SRMR                               0.038
## 

## Parameter Estimates:
## 

## Standard errors                   Standard
## Information                         Expected
## Information saturated (h1) model   Unstructured
## 

## Latent Variables:
## 

## f1 =~                                Estimate Std.Err z-value P(>|z|) Std.lv Std.all
## SE11        0.644    0.021  30.727  0.000  0.644  0.661
## SE12        0.645    0.024  27.092  0.000  0.645  0.613
## SE13        0.662    0.025  26.471  0.000  0.662  0.602
## SE14        0.759    0.025  30.973  0.000  0.759  0.729
## SE15        0.898    0.026  34.283  0.000  0.898  0.791
## SE16        0.868    0.026  33.233  0.000  0.868  0.788
## 

## Covariances:
## 

## .SE12 ~~                                Estimate Std.Err z-value P(>|z|) Std.lv Std.all
## .SE13        0.564    0.046  12.258  0.000  0.564  0.773
## 

## Variances:
## 

## .SE11        0.534    0.049  10.953  0.000  0.534  0.563
## .SE12        0.691    0.051  13.529  0.000  0.691  0.624
## .SE13        0.771    0.055  14.144  0.000  0.771  0.638
## .SE14        0.508    0.061   8.349  0.000  0.508  0.469
## .SE15        0.482    0.065   7.465  0.000  0.482  0.374
## .SE16        0.459    0.064   7.125  0.000  0.459  0.379
## f1          1.000                               1.000  1.000
## 

## R-Square:
## 

## Estimate

```

```

##      SE11        0.437
##      SE12        0.376
##      SE13        0.362
##      SE14        0.531
##      SE15        0.626
##      SE16        0.621

lavTestLRT(fit.se1.cor, fit.se1)

## Chi-Squared Difference Test
##
##          Df AIC BIC   Chisq Chisq diff Df diff Pr(>Chisq)
## fit.se1.cor  8     27.852
## fit.se1      9    165.218     137.37      1 < 2.2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

modindices(fit.se1.cor, sort=TRUE, minimum.value = 10)

##      lhs op rhs mi epc sepc.lv sepc.all sepc.nox
## 28 SE15 ~~ SE16 19.02 0.264  0.264    0.562    0.562

se1.cfa.cor2<- 'f1 =~ SE11 + SE12 + SE13 + SE14+ SE15 + SE16
SE12~~SE13
SE15~~SE16'

fit.se1.cor2 <- cfa(se1.cfa.cor2, data=se1, std.lv=TRUE, estimator = 'DWLS')
summary(fit.se1.cor2, fit.measures=TRUE, standardized=TRUE, rsquare=TRUE)

## lavaan 0.6-7 ended normally after 27 iterations
##
##      Estimator           DWLS
##      Optimization method NLMINB
##      Number of free parameters       14
## 
##      Number of observations      Used      Total
##                               1261      1285
## 
## Model Test User Model:
##
##      Test statistic            8.865
##      Degrees of freedom                  7
##      P-value (Chi-square)            0.262
## 
## Model Test Baseline Model:
##
##      Test statistic            3887.411
##      Degrees of freedom                  15
##      P-value                         0.000
## 
## User Model versus Baseline Model:

```

```

##  

## Comparative Fit Index (CFI) 1.000  

## Tucker-Lewis Index (TLI) 0.999  

##  

## Root Mean Square Error of Approximation:  

##  

## RMSEA 0.015  

## 90 Percent confidence interval - lower 0.000  

## 90 Percent confidence interval - upper 0.039  

## P-value RMSEA <= 0.05 0.994  

##  

## Standardized Root Mean Square Residual:  

##  

## SRMR 0.022  

##  

## Parameter Estimates:  

##  

## Standard errors Standard  

## Information Expected  

## Information saturated (h1) model Unstructured  

##  

## Latent Variables:  

##  

## f1 =~  

## SE11 0.663 0.022 30.304 0.000 0.663 0.680  

## SE12 0.668 0.025 26.359 0.000 0.668 0.635  

## SE13 0.685 0.027 25.779 0.000 0.685 0.623  

## SE14 0.780 0.026 30.472 0.000 0.780 0.749  

## SE15 0.825 0.030 27.605 0.000 0.825 0.726  

## SE16 0.791 0.030 26.204 0.000 0.791 0.718  

##  

## Covariances:  

##  

## .SE12 ~~  

## .SE13 0.533 0.048 11.188 0.000 0.533 0.763  

## .SE15 ~~  

## .SE16 0.253 0.056 4.513 0.000 0.253 0.424  

##  

## Variances:  

##  

## .SE11 0.510 0.050 10.213 0.000 0.510 0.537  

## .SE12 0.660 0.053 12.432 0.000 0.660 0.596  

## .SE13 0.740 0.057 13.072 0.000 0.740 0.612  

## .SE14 0.476 0.063 7.612 0.000 0.476 0.439  

## .SE15 0.609 0.066 9.199 0.000 0.609 0.473  

## .SE16 0.587 0.066 8.861 0.000 0.587 0.484  

## f1 1.000  

##  

## R-Square:  

## Estimate

```

```

##    SE11      0.463
##    SE12      0.404
##    SE13      0.388
##    SE14      0.561
##    SE15      0.527
##    SE16      0.516

lavTestLRT(fit.se1.cor2, fit.se1.cor)

## Chi-Squared Difference Test
##
##          Df AIC BIC  Chisq Chisq diff Df diff Pr(>Chisq)
## fit.se1.cor2 7     8.865
## fit.se1.cor 8    27.852    18.987      1  1.316e-05 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

SE at T2 CFA

```

## lavaan 0.6-7 ended normally after 19 iterations
##
##   Estimator           DWLS
##   Optimization method NLMINB
##   Number of free parameters      12
##
##   Number of observations      Used      Total
##                               1261      1285
##
## Model Test User Model:
##
##   Test statistic        165.218
##   Degrees of freedom       9
##   P-value (Chi-square)  0.000
##
## Model Test Baseline Model:
##
##   Test statistic        3887.411
##   Degrees of freedom      15
##   P-value                 0.000
##
## User Model versus Baseline Model:
##
##   Comparative Fit Index (CFI)      0.960
##   Tucker-Lewis Index (TLI)        0.933
##
## Root Mean Square Error of Approximation:
##
##   RMSEA                  0.117
##   90 Percent confidence interval - lower  0.102
##   90 Percent confidence interval - upper  0.133

```

```

## P-value RMSEA <= 0.05                                0.000
##
## Standardized Root Mean Square Residual:
##
## SRMR                                         0.094
##
## Parameter Estimates:
##
## Standard errors                               Standard
## Information                                 Expected
## Information saturated (h1) model            Unstructured
##
## Latent Variables:
##                         Estimate Std.Err z-value P(>|z|) Std.lv Std.all
## f1 =~
##   SE11          0.616   0.020  30.989  0.000  0.616  0.633
##   SE12          0.784   0.023  34.448  0.000  0.784  0.745
##   SE13          0.809   0.024  34.006  0.000  0.809  0.736
##   SE14          0.724   0.023  31.144  0.000  0.724  0.695
##   SE15          0.851   0.025  34.713  0.000  0.851  0.749
##   SE16          0.814   0.024  33.370  0.000  0.814  0.739
##
## Variances:
##                         Estimate Std.Err z-value P(>|z|) Std.lv Std.all
## .SE11          0.569   0.047  11.998  0.000  0.569  0.600
## .SE12          0.492   0.054   9.078  0.000  0.492  0.445
## .SE13          0.554   0.058   9.553  0.000  0.554  0.458
## .SE14          0.561   0.059   9.551  0.000  0.561  0.517
## .SE15          0.565   0.061   9.290  0.000  0.565  0.438
## .SE16          0.550   0.061   9.062  0.000  0.550  0.453
##   f1           1.000
##
## R-Square:
##                         Estimate
##   SE11          0.400
##   SE12          0.555
##   SE13          0.542
##   SE14          0.483
##   SE15          0.562
##   SE16          0.547

modindices(fit.se2, sort=TRUE, minimum.value = 10)

##      lhs op rhs     mi     epc sepc.lv sepc.all sepc.nox
## 19 SE22 ~~ SE23 69.530  0.380   0.380    1.067   1.067
## 28 SE25 ~~ SE26 17.537  0.189   0.189    0.454   0.454
## 22 SE22 ~~ SE26 12.503 -0.140  -0.140   -0.373  -0.373

se2.cfa.cor<- 'f1 =~ SE21 + SE22 + SE23 + SE24+ SE25 + SE26
SE22~~SE23'

```

```

fit.se2.cor <- cfa(se2.cfa.cor, data=se2, std.lv=TRUE, estimator = 'DWLS')
summary(fit.se2.cor, fit.measures=TRUE, standardized=TRUE, rsquare=TRUE)

## lavaan 0.6-7 ended normally after 20 iterations
##
##   Estimator                      DWLS
##   Optimization method            NLMINB
##   Number of free parameters     13
## 
##                               Used      Total
##   Number of observations        1281      1285
## 
## Model Test User Model:
## 
##   Test statistic                17.259
##   Degrees of freedom             8
##   P-value (Chi-square)          0.028
## 
## Model Test Baseline Model:
## 
##   Test statistic                3352.659
##   Degrees of freedom              15
##   P-value                         0.000
## 
## User Model versus Baseline Model:
## 
##   Comparative Fit Index (CFI)    0.997
##   Tucker-Lewis Index (TLI)       0.995
## 
## Root Mean Square Error of Approximation:
## 
##   RMSEA                           0.030
##   90 Percent confidence interval - lower 0.010
##   90 Percent confidence interval - upper 0.050
##   P-value RMSEA <= 0.05           0.953
## 
## Standardized Root Mean Square Residual:
## 
##   SRMR                            0.034
## 
## Parameter Estimates:
## 
##   Standard errors                  Standard
##   Information                       Expected
##   Information saturated (h1) model Unstructured
## 
## Latent Variables:
## 
##   Estimate Std.Err z-value P(>|z|) Std.lv Std.all
##   f1 =~
```

```

##   SE21      0.658  0.022  30.085  0.000  0.658  0.763
##   SE22      0.618  0.024  25.922  0.000  0.618  0.669
##   SE23      0.625  0.024  26.046  0.000  0.625  0.662
##   SE24      0.643  0.022  28.841  0.000  0.643  0.730
##   SE25      0.812  0.025  32.179  0.000  0.812  0.801
##   SE26      0.689  0.023  30.120  0.000  0.689  0.751
##
## Covariances:
##          Estimate Std.Err z-value P(>|z|) Std.lv Std.all
## .SE22 ~~
## .SE23      0.365  0.042  8.655  0.000  0.365  0.751
##
## Variances:
##          Estimate Std.Err z-value P(>|z|) Std.lv Std.all
## .SE21      0.311  0.046  6.782  0.000  0.311  0.418
## .SE22      0.473  0.048  9.849  0.000  0.473  0.553
## .SE23      0.500  0.048 10.390  0.000  0.500  0.561
## .SE24      0.362  0.051  7.052  0.000  0.362  0.467
## .SE25      0.367  0.057  6.418  0.000  0.367  0.358
## .SE26      0.366  0.050  7.283  0.000  0.366  0.435
## f1        1.000                1.000  1.000
##
## R-Square:
##          Estimate
## SE21      0.582
## SE22      0.447
## SE23      0.439
## SE24      0.533
## SE25      0.642
## SE26      0.565

lavTestLRT(fit.se2.cor, fit.se2)

## Chi-Squared Difference Test
##
##          Df AIC BIC Chisq Chisq diff Df diff Pr(>Chisq)
## fit.se2.cor  8    17.259
## fit.se2     9    87.041    69.782      1 < 2.2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

modindices(fit.se2.cor, sort=TRUE, minimum.value = 10)

## [1] lhs      op      rhs      mi      epc      sepc.lv  sepc.all sepc.no
## <0 rows> (or 0-length row.names)

```

Longitudinal Invariance Modeling

		χ^2	df	CFI	RMSEA	SRMR	$\Delta \chi^2$	Δdf
Depression	Configural	327.577	197	0.991	0.024	0.038		
	Metric	549.159	207	0.978	0.038	0.050	221.59	10
	Scalar	845.993	217	0.959	0.050	0.059	296.82	10
Job Search Self-efficacy	Configural	320.873	47	0.975	0.068	0.066		
	Metric	341.341	52	0.973	0.067	0.068	20.468	5
	Scalar	345.132	57	0.973	0.063	0.068	3.792	5

5. Code to estimate the two-wave mediation model with latent BTIs

Estimating latent baseline-by-treatment interactions (BTIs) in statistical mediation analysis: R/Mplus Code

Below, we provide code to conduct five different approaches to estimate latent BTI effects in the two-wave mediation model. Standard errors for each of the paths could be estimated using the 95% percentile bootstrap confidence intervals, if desired.

First, we start by summarizing the dataset from the illustration, and then code is provided.

Jobs II dataset and preliminary analyses

Below is the jobs ii dataset that we use for this example. Note that cTx is the treatment/control indicator, which is centered. SE1 and SE2 is job search self-efficacy [the mediator] at time 1 and time 2, and DEP0 and DEP2 is depression symptoms [the outcome] at time 1 and time 2.

```
setwd('~/Downloads')
datt22=read.csv('jobsii_supp.csv')
head(datt22)

##          cTx SE11 SE12 SE13 SE14 SE15 SE16 DEP01 DEP02 DEP03 DEP04 DEP05 DEP06
## 1  0.3081705    5    5    5    5    4    4    1    4    2    2    2    2    1
## 2  0.3081705    5    4    4    5    5    5    2    1    1    1    1    1    1
## 3  0.3081705    3    4    3    3    3    3    1    3    2    1    2    1    1
## 4  0.3081705    2    3    4    4    3    4    4    4    2    2    2    2    2
## 5 -0.6918295    4    5    5    5    4    4    2    2    1    1    3    1    1
## 6  0.3081705    5    5    5    5    5    5    2    2    2    2    2    2    1
##          DEP07 DEP08 DEP09 DEP010 DEP011 SE21 SE22 S23 SE24 SE25 SE26 DEP21 DEP22
## 1      3     3     1     1     1     5     5     5     5     4     5     3     2
## 2      1     1     1     1     2     5     5     5     5     5     5     1     1
## 3      1     1     1     1     1     4     4     4     4     4     3     1     3
## 4      2     2     1     1     1     4     4     4     5     5     5     1     2
## 5      1     2     1     1     1     4     3     4     4     3     4     1     2
## 6      2     2     1     1     1     5     5     5     5     5     5     1     2
##          DEP23 DEP24 DEP25 DEP26 DEP27 DEP28 DEP2  DEP210 DEP211
## 1      1     1     2     1     3     2     1     1     2
## 2      1     1     1     1     1     1     1     1
## 3      2     1     3     1     2     2     1     4     2
## 4      2     3     3     5     2     2     2     1     1
## 5      1     1     3     2     2     2     1     1     1
## 6      1     1     1     1     1     1     1     1     1
```

As a preliminary analysis, we conduct longitudinal CFAs and use the output later.

```
library(lavaan)

#Longitudinal factor model for the mediator
m1='
m1=~NA*SE11+a*SE11+b*SE12+c*SE13+d*SE14+e*SE15+f*SE16
m2=~NA*SE21+a*SE21+b*SE22+c*S23+d*SE24+e*SE25+f*SE26
m2~m1
```

```

SE11~~SE21
SE12~~SE22
SE13~~S23
SE14~~SE24
SE15~~SE25
SE16~~SE26

SE12~~SE13 #correlated residual
SE22~~S23 #correlated residual
m1~~1*m1
m2~~m2
'

f1=cfa(m1,data=datt22,std.lv=T)

coef(f1)

##          a          b          c          d          e          f          a
##    0.661      0.675      0.686      0.739      0.910      0.854      0.661
##          b          c          d          e          f          m2~m1 SE11~~SE21
##    0.675      0.686      0.739      0.910      0.854      0.514      0.110
## SE12~~SE22  SE13~~S23  SE14~~SE24  SE15~~SE25  SE16~~SE26  SE12~~SE13  SE22~~S23
##    0.003      0.031      0.141      0.051      0.040      0.560      0.372
##          m2~m2 SE11~~SE11  SE12~~SE12  SE13~~SE13  SE14~~SE14  SE15~~SE15  SE16~~SE16
##    0.535      0.539      0.682      0.754      0.531      0.453      0.416
## SE21~~SE21  SE22~~SE22  S23~~S23  SE24~~SE24  SE25~~SE25  SE26~~SE26
##    0.355      0.480      0.499      0.359      0.373      0.310

#Longitudinal factor model for the outcome

d1='
de1=~a*DEP01+b*DEP02+c*DEP03+d*DEP04+e*DEP05+f*DEP06+g*DEP07+h*DEP08+i*DEP09+j*DEP010+k*DEP011
de2=~a*DEP21+b*DEP22+c*DEP23+d*DEP24+e*DEP25+f*DEP26+g*DEP27+h*DEP28+i*DEP29+j*DEP210+k*DEP211
de1~~de2
DEP01~~DEP21
DEP02~~DEP22
DEP03~~DEP23
DEP04~~DEP24
DEP05~~DEP25
DEP06~~DEP26
DEP07~~DEP27
DEP08~~DEP28
DEP09~~DEP29
DEP010~~DEP210
DEP011~~DEP211
'

f2=cfa(d1,data=datt22,std.lv=T)

```

```
coef(f2) #estimated coefficients
```

##	a	b	c	d	e
##	0.445	0.701	0.680	0.510	0.677
##	f	g	h	i	j
##	0.712	0.675	0.837	0.161	0.322
##	k	a	b	c	d
##	0.507	0.445	0.701	0.680	0.510
##	e	f	g	h	i
##	0.677	0.712	0.675	0.837	0.161
##	j	k	de1~~de2	DEP01~~DEP21	DEP02~~DEP22
##	0.322	0.507	0.473	0.178	0.118
##	DEP03~~DEP23	DEP04~~DEP24	DEP05~~DEP25	DEP06~~DEP26	DEP07~~DEP27
##	0.121	0.270	0.235	0.226	0.097
##	DEP08~~DEP28	DEP09~~DEP2	DEP010~~DEP210	DEP011~~DEP211	DEP01~~DEP01
##	0.066	0.045	0.122	0.218	0.899
##	DEP02~~DEP02	DEP03~~DEP03	DEP04~~DEP04	DEP05~~DEP05	DEP06~~DEP06
##	0.914	0.723	0.664	0.837	0.660
##	DEP07~~DEP07	DEP08~~DEP08	DEP09~~DEP09	DEP010~~DEP010	DEP011~~DEP011
##	0.456	0.413	0.182	0.442	0.628
##	DEP21~~DEP21	DEP22~~DEP22	DEP23~~DEP23	DEP24~~DEP24	DEP25~~DEP25
##	0.570	0.590	0.466	0.602	0.562
##	DEP26~~DEP26	DEP27~~DEP27	DEP28~~DEP28	DEP2~~DEP2	DEP210~~DEP210
##	0.677	0.511	0.379	0.227	0.458
## DEP211~~DEP211					
##	0.766				

Scoring M and estimating observed XM interaction

Summed scores, correcting for unreliability

```
library(lavaan)
library(psych)

#estimate summed scores, centered on baseline measure
cSE1=rowSums(datt22[,2:7])-mean(rowSums(datt22[,2:7]))
cDEP1=rowSums(datt22[,8:18])-mean(rowSums(datt22[,8:18]))
cSE2=rowSums(datt22[,19:24])-mean(rowSums(datt22[,2:7]))
cDEP2=rowSums(datt22[,25:35])-mean(rowSums(datt22[,8:18]))

#create dataset with only summed scores
datt3=cbind(
  datt22[,1],cSE1,cDEP1,cSE2,cDEP2,datt22[,1]*cSE1,datt22[,1]*cDEP1)
colnames(datt3)=c('cTx','cSE1','cDEP1','cSE2','cDEP2','XM1','XY1')

covsum=cov(datt3)
round(covsum,3)

##          cTx   cSE1   cDEP1   cSE2   cDEP2     XM1     XY1
## cTx    0.213 -0.171 -0.043  0.152 -0.177  0.066  0.016
## cSE1   -0.171 24.609 -5.516 11.323 -5.762  1.038  0.946
## cDEP1  -0.043 -5.516 37.883 -4.837 20.441  0.946 -0.599
## cSE2    0.152 11.323 -4.837 19.790 -7.616 -0.062  0.633
## cDEP2  -0.177 -5.762 20.441 -7.616 53.312  0.618 -0.977
## XM1     0.066  1.038  0.946 -0.062  0.618  4.819 -1.546
## XY1     0.016  0.946 -0.599  0.633 -0.977 -1.546  8.305

#correct covariance matrix
covsum[2,2]=covsum[2,2]-sum(coef(f1)[23:28]) #correct varM1
covsum[4,4]=covsum[4,4]-sum(coef(f1)[29:34]) #correct varM2
covsum[3,3]=covsum[3,3]-sum(coef(f2)[35:45]) #correct varY1
covsum[5,5]=covsum[5,5]-sum(coef(f2)[46:56]) #correct varY2
covsum[6,6]=covsum[1,1]*covsum[2,2] #corrected varXM1
covsum[7,7]=covsum[1,1]*covsum[3,3] #corrected varXY1

#correct covM1M2
covsum[2,4]=covsum[4,2]=covsum[2,4]-sum(coef(f1)[14:19])
#correct covY1Y2
covsum[3,5]=covsum[3,5]=covsum[3,5]-sum(coef(f2)[24:34])

#fit model to covariance matrix
sc_model='
cSE2~cTx+cSE1+cDEP1+XM1+XY1
cDEP2~cTx+cSE2+cSE1+cDEP1+XM1+XY1'
```

```

scfit=cfa(sc_model, sample.cov=covsum, sample.nobs=1126)
summary(scfit)

## lavaan 0.6-12 ended normally after 1 iterations
##
##   Estimator                               ML
## Optimization method                      NLMINB
## Number of model parameters               13
##
##   Number of observations                  1126
##
## Model Test User Model:
##
##   Test statistic                          0.000
##   Degrees of freedom                     0
##
## Parameter Estimates:
##
##   Standard errors                        Standard
##   Information                            Expected
##   Information saturated (h1) model       Structured
##
## Regressions:
##             Estimate  Std.Err  z-value  P(>|z|)
## cSE2 ~
##   cTx          1.162    0.219   5.318   0.000
##   cSE1          0.518    0.023  22.816   0.000
##   cDEP1         -0.058    0.019  -3.143   0.002
##   XM1          -0.144    0.050  -2.879   0.004
##   XY1          -0.020    0.041  -0.492   0.623
## cDEP2 ~
##   cTx          -0.463    0.377  -1.227   0.220
##   cSE2         -0.281    0.051  -5.527   0.000
##   cSE1          0.035    0.047   0.745   0.456
##   cDEP1          0.619    0.032  19.541   0.000
##   XM1          -0.023    0.085  -0.270   0.787
##   XY1          -0.074    0.070  -1.061   0.289
##
## Variances:
##             Estimate  Std.Err  z-value  P(>|z|)
## .cSE2        11.282   0.475  23.728   0.000
## .cDEP2       32.743   1.380  23.728   0.000

```

Factor scores for M

```
library(lavaan)

#obtain factor scores from CFA output
k2=lavPredict(f1,method='bartlett',fsm=T)
tsmpM=attr(k2,'fsm')[[1]] #scoring matrix
head(k2) #factor scores

##          m1          m2
## [1,]  0.9838734  0.9372603
## [2,]  1.4733252  1.2291032
## [3,] -0.7934914 -0.3144488
## [4,] -0.4553417  0.8488736
## [5,]  0.7621510 -0.4138175
## [6,]  1.6454972  1.2510743

k3=lavPredict(f2,method='bartlett',fsm=T)
tsmpY=attr(k3,'fsm')[[1]] #scoring matrix
head(k3) #factor scores

##          de1          de2
## [1,]  0.32997209 -0.03963784
## [2,] -1.27753092 -1.33811621
## [3,] -0.96301956  0.31103609
## [4,]  0.06217877  0.69253732
## [5,] -0.73778490 -0.26187261
## [6,] -0.35835906 -1.17793331

#estimate covariance matrix of factor scores
facdat=cbind(
  datt22[,1],k2[,1],k3[,1],k2[,2],k3[,2],datt22[,1]*k2[,1],
  datt22[,1]*k3[,1])
covfac=cov(facdat)
colnames(covfac)=rownames(covfac)=c('cTx','cSE1','cDEP1',
                                     'cSE2','cDEP2','XM1','XY1')
round(covfac,3)

##          cTx      cSE1      cDEP1      cSE2      cDEP2      XM1      XY1
## cTx     0.213   -0.037   -0.010    0.035   -0.027    0.014    0.004
## cSE1   -0.037    1.162   -0.215    0.537   -0.220    0.043    0.036
## cDEP1  -0.010   -0.215    0.961   -0.172    0.506    0.036   -0.022
## cSE2    0.035    0.537   -0.172    0.917   -0.271   -0.015    0.022
## cDEP2  -0.027   -0.220    0.506   -0.271    1.314    0.027   -0.023
## XM1     0.014    0.043    0.036   -0.015    0.027    0.230   -0.060
## XY1     0.004    0.036   -0.022    0.022   -0.023   -0.060    0.213

#obtain error covariance matrix for M
errM=matrix(0,nrow=12,ncol=12)
diag(errM)=coef(f1)[23:34]
diag(errM[1:6,7:12])=coef(f1)[14:19]
diag(errM[7:12,1:6])=coef(f1)[14:19]
```

```

#obtain error covariance matrix for Y
errY=matrix(0,nrow=22,ncol=22)
diag(errY)=coef(f2)[35:56]
diag(errY[1:11,12:22])=coef(f2)[24:34]
diag(errY[12:22,1:11])=coef(f2)[24:34]

#correct covariance matrix among M1 and M2
covfac[c(2,4),c(2,4)] = covfac[c(2,4),c(2,4)]-tsmpM%*%errM%*%t(tsmpM)

#correct covariance matrix among Y1 and Y2
covfac[c(3,5),c(3,5)] = covfac[c(3,5),c(3,5)]-tsmpY%*%errY%*%t(tsmpY)

#corrected variances of XM1 and XM2
covfac[6,6]=covfac[1,1]*covfac[2,2]
covfac[7,7]=covfac[1,1]*covfac[3,3]

#estimate model
fc_model=
cSE2~cTx+cSE1+cDEP1+XM1+XY1
cDEP2~cTx+cSE2+cSE1+cDEP1+XM1+XY1

fcfit=cfa(fc_model,sample.cov=covfac,sample.nobs=1126)
summary(fcfit)

## lavaan 0.6-12 ended normally after 1 iterations
##
##    Estimator                      ML
##    Optimization method            NLMINB
##    Number of model parameters    13
##    Number of observations        1126
##    Model Test User Model:
##    Test statistic                 0.000
##    Degrees of freedom              0
##
## Parameter Estimates:
##    Standard errors                  Standard
##    Information                      Expected
##    Information saturated (h1) model Structured
##
## Regressions:
##                   Estimate Std.Err z-value P(>|z|)
##   cSE2 ~
##     cTx          0.267   0.047   5.701   0.000
##     cSE1          0.518   0.022  23.163   0.000

```

```

##   cDEP1      -0.065    0.025   -2.634    0.008
##   XM1       -0.197    0.049   -3.992    0.000
##   XY1       -0.063    0.055   -1.144    0.253
## cDEP2 ~
##   cTx       -0.063    0.062   -1.029    0.304
##   cSE2      -0.224    0.039   -5.782    0.000
##   cSE1      0.009    0.035    0.246    0.805
##   cDEP1      0.535    0.032   16.662    0.000
##   XM1       0.012    0.065    0.187    0.851
##   XY1       -0.032    0.071   -0.445    0.656
##
## Variances:
##             Estimate Std.Err z-value P(>|z|)
## .cSE2        0.516   0.022  23.728  0.000
## .cDEP2       0.869   0.037  23.728  0.000

```

Structural models to accommodate a latent XM interaction M

Unconstrained indicator approach

The UPI approach can be estimated in any structural equation modeling. Below we provide code to estimate the UPI approach in Mplus.

```
TITLE: Unconstrained product indicator for jobs ii data
DATA: FILE IS datt22.dat; !items centered at baseline
VARIABLE: NAMES ARE cTx se11-se16 depo1-depo11 se21-se26 dep1-dep11;
USEVARIABLE are cTx se11-se16 depo1-depo11 se21-se26 dep1-dep11
xm1-xm6 xy1-xy11;
DEFINE: !define product variables and include in USEVARIABLES above
xm1 = se11*cTx;
xm2 = se12*cTx;
xm3 = se13*cTx;
xm4 = se14*cTx;
xm5 = se15*cTx;
xm6 = se16*cTx;

xy1 = depo1*cTx;
xy2 = depo2*cTx;
xy3 = depo3*cTx;
xy4 = depo4*cTx;
xy5 = depo5*cTx;
xy6 = depo6*cTx;
xy7 = depo7*cTx;
xy8 = depo8*cTx;
xy9 = depo9*cTx;
xy10 = depo10*cTx;
xy11 = depo11*cTx;

ANALYSIS: estimator = ML; !bootstrap=500;
MODEL:

se1 by se11* (a1) !specify invariance across time
se12 (b1)
se13 (c1)
se14 (d1)
se15 (e1)
se16 (f1);
de1 by depo1* (a2)
depo2 (b2)
depo3 (c2)
depo4 (d2)
depo5 (e2)
depo6 (f2)
depo7 (g2)
depo8 (h2)
depo9 (i2)
```

```

depo10 (j2)
depo11 (k2);

se2 by se21* (a1)
se22 (b1)
se23 (c1)
se24 (d1)
se25 (e1)
se26 (f1);

de2 by dep1* (a2)
dep2 (b2)
dep3 (c2)
dep4 (d2)
dep5 (e2)
dep6 (f2)
dep7 (g2)
dep8 (h2)
dep9 (i2)
dep10 (j2)
dep11 (k2);

se11 with se21; !correlated residuals across time
se12 with se22;
se13 with se23;
se14 with se24;
se15 with se25;
se16 with se26;

se12 with se13; !correlated residual within time
se22 with se23; !correlated residual within time

depo1 with dep1; !correlated residuals across time
depo2 with dep2;
depo3 with dep3;
depo4 with dep4;
depo5 with dep5;
depo6 with dep6;
depo7 with dep7;
depo8 with dep8;
depo9 with dep9;
depo10 with dep10;
depo11 with dep11;

xm by xm1* xm2-xm6; !define latent interaction term XM1
xm2 with xm3; !correlated residual within time
xm@.213; !constrain to var(x)*var(m1)
xy by xy1* xy2-xy11; !define latent interaction term XY1
xy@.213; !constrain to var(x)*var(y1)

```

```

[se11] (a3); [se21] (a3); !longitudinal invariance on intercepts
[se12] (b3); [se22] (b3);
[se13] (c3); [se23] (c3);
[se14] (d3); [se24] (d3);
[se15] (e3); [se25] (e3);
[se16] (f3); [se26] (f3);

[depo1@0] (a4); [dep1@0] (a4); !longitudinal invariance intercepts
[depo2] (b4); [dep2] (b4);
[depo3] (c4); [dep3] (c4);
[depo4] (d4); [dep4] (d4);
[depo5] (e4); [dep5] (e4);
[depo6] (f4); [dep6] (f4);
[depo7] (g4); [dep7] (g4);
[depo8] (h4); [dep8] (h4);
[depo9] (i4); [dep9] (i4);
[depo10] (j4); [dep10] (j4);
[depo11] (k4); [dep11] (k4);

se2 on cTx (apath) !regression for M2
se1
de1
xm (h1path)
xy;

de2 on cTx !regression for Y2
se2 (bpath)
se1 de1 xm xy;

xm with se1@0; !constraints where interactions do not correlate
xm with de1@0;
xy with se1@0;
xy with de1@0;
xm with cTx@0;
xy with cTx@0;

[se1@0]; [de1@0]; [se2]; [de2]; !(conditional) means
se1@1; de1@1; se2; de2; !(residual) variances

OUTPUT: tech1 modindices(10); !cinterval(bootstrap);

```

Latent moderated structural equations

Below we provide Mplus code to estimate latent BTI effects in the two wave model.

```
TITLE: LMS examples for jobs ii data
DATA: FILE IS datt22.dat; !items centered at baseline
VARIABLE: NAMES ARE cTx se11-se16 depo1-depo11 se21-se26 dep1-dep11;
USEVARIABLE are cTx se11-se16 depo1-depo11 se21-se26 dep1-dep11;
ANALYSIS: Type = random; estimator = ML; algorithm = integration;

MODEL:
se1 by se11* (a1) !specify invariance across time
se12 (b1)
se13 (c1)
se14 (d1)
se15 (e1)
se16 (f1);
de1 by depo1* (a2)
depo2 (b2)
depo3 (c2)
depo4 (d2)
depo5 (e2)
depo6 (f2)
depo7 (g2)
depo8 (h2)
depo9 (i2)
depo10 (j2)
depo11 (k2);
depo11 (l2);
se2 by se21* (a1)
se22 (b1)
se23 (c1)
se24 (d1)
se25 (e1)
se26 (f1);
de2 by dep1* (a2)
dep2 (b2)
dep3 (c2)
dep4 (d2)
dep5 (e2)
dep6 (f2)
dep7 (g2)
dep8 (h2)
dep9 (i2)
dep10 (j2)
dep11 (k2);

[se11@0] (a3); [se21@0] (a3); !longitudinal invariance on intercepts
[se12] (b3); [se22] (b3);
[se13] (c3); [se23] (c3);
```

```

[se14] (d3); [se24] (d3);
[se15] (e3); [se25] (e3);
[se16] (f3); [se26] (f3);

[depo1@0] (a4); [dep1@0] (a4); !longitudinal invariance on intercepts
[depo2] (b4); [dep2] (b4);
[depo3] (c4); [dep3] (c4);
[depo4] (d4); [dep4] (d4);
[depo5] (e4); [dep5] (e4);
[depo6] (f4); [dep6] (f4);
[depo7] (g4); [dep7] (g4);
[depo8] (h4); [dep8] (h4);
[depo9] (i4); [dep9] (i4);
[depo10] (j4); [dep10] (j4);
[depo11] (k4); [dep11] (k4);

se11 with se21; !correlated residuals across time
se12 with se22;
se13 with se23;
se14 with se24;
se15 with se25;
se16 with se26;
se12 with se13; !correlated residual within time
se22 with se23; !correlated residual within time

depo1 with dep1; !correlated residuals across time
depo2 with dep2;
depo3 with dep3;
depo4 with dep4;
depo5 with dep5;
depo6 with dep6;
depo7 with dep7;
depo8 with dep8;
depo9 with dep9;
depo10 with dep10;
depo11 with dep11;

xm1 | se1 XWITH cTx; ! specify latent interaction
xy1 | de1 XWITH cTx; ! specify latent interaction

se2 on cTx (apath) !regression for M2
se1
de1
xm1 (h1path)
xy1;

de2 on cTx !regression for Y2
se2 (bpath)
se1 de1 xm1 xy1;

```

```
[se1]; [de1]; [se2]; [de2]; !(conditional) means  
se1@1; de1@1; se2@1; de2@1; !(residual) variances  
  
de1 with se1; !covariance for baseline measures  
  
OUTPUT: tech1;
```

Bayesian mediation

Below we provide R code that interacts with JAGS estimate latent BTI effects in the two-wave model using MCMC with a Gibbs sampler in JAGS.

```
library(R2jags) #JAGS should be installed in computer
bcfaa=function() { #specify model via function

#####
# Specify Latent var model for M1 and M2
#####
for (i in 1:n){
  #two Latent variables: construct + specific factor for corr residuals
  #invariance via using same loading for time 1 and time 2
  mu1[i,1] <- tau[1] + nu[i,9]*lambda[1] + nu[i,1]*beta[1]
  m1[i,1] ~ dnorm(mu1[i,1], inv.psi1[1]) #normal distribution
  mu2[i,1] <- tau[1] + ksi2[i]*lambda[1]+ nu[i,1]*beta[1]
  m2[i,1] ~ dnorm(mu2[i,1], inv.psi2[1])

  #additional Latent variable for item 2/3 for within time correlated residual
  mu1[i,2] <- tau[2] + nu[i,9]*lambda[2] + nu[i,2]*beta[2] + nu[i,7]*beta[7]
  m1[i,2] ~ dnorm(mu1[i,2], inv.psi1[2])
  mu2[i,2] <- tau[2] + ksi2[i]*lambda[2] + nu[i,2]*beta[2] + nu[i,8]*beta[8]
  m2[i,2] ~ dnorm(mu2[i,2], inv.psi2[2])

  mu1[i,3] <- tau[3] + nu[i,9]*lambda[3] + nu[i,3]*beta[3] + nu[i,7]*beta[7]
  m1[i,3] ~ dnorm(mu1[i,3], inv.psi1[3])
  mu2[i,3] <- tau[3] + ksi2[i]*lambda[3] + nu[i,3]*beta[3] + nu[i,8]*beta[8]
  m2[i,3] ~ dnorm(mu2[i,3], inv.psi2[3])

  mu1[i,4] <- tau[4] + nu[i,9]*lambda[4] + nu[i,4]*beta[4]
  m1[i,4] ~ dnorm(mu1[i,4], inv.psi1[4])
  mu2[i,4] <- tau[4] + ksi2[i]*lambda[4] + nu[i,4]*beta[4]
  m2[i,4] ~ dnorm(mu2[i,4], inv.psi2[4])

  mu1[i,5] <- tau[5] + nu[i,9]*lambda[5] + nu[i,5]*beta[5]
  m1[i,5] ~ dnorm(mu1[i,5], inv.psi1[5])
  mu2[i,5] <- tau[5] + ksi2[i]*lambda[5] + nu[i,5]*beta[5]
  m2[i,5] ~ dnorm(mu2[i,5], inv.psi2[5])

  mu1[i,6] <- tau[6] + nu[i,9]*lambda[6] + nu[i,6]*beta[6]
  m1[i,6] ~ dnorm(mu1[i,6], inv.psi1[6])
  mu2[i,6] <- tau[6] + ksi2[i]*lambda[6] + nu[i,6]*beta[6]
  m2[i,6] ~ dnorm(mu2[i,6], inv.psi2[6])
}

#####
# Specify Latent var model for Y1 and Y2
#####
for (i in 1:n){
  for(jj in 1:JJ){
```

```

    yu1[i,jj] <- tau.y[jj]+mu.y1[i]*lambda.y[jj] + nu.y[i,jj]*beta.y[jj]
    y1[i,jj] ~ dnorm(yu1[i,jj], inv.psi1.y[jj])
    yu2[i,jj] <-tau.y[jj]+ksi.y2[i]*lambda.y[jj] + nu.y[i,jj]*beta.y[jj]
    y2[i,jj] ~ dnorm(yu2[i,jj], inv.psi2.y[jj])
  }
}

#####
# Specify the (prior) distribution for M1 and M2
#####
for(j in 1:J){

  inv.psi1[j] ~ dgamma(1,1) # Precisions for observables
  inv.psi2[j] ~ dgamma(1,1) # Precisions for observables
  psi1[j] <- 1/inv.psi1[j]   # Variances for observables
  psi2[j] <- 1/inv.psi2[j]   # Variances for observables

}

for(j in 1:J){

  lambda[j] ~ dnorm(1,.5);T(0,) #half normal distribution for Loadings
}

tau[1] <- 0 #first intercept to 0
for(j in 2:J){
  tau[j] ~ dnorm(0,.5)
}

for(k in 1:8){
  beta[k] ~ dnorm(1,.5);T(0,) #half normal distribution for Loadings
}

mu.m ~ dnorm(0,1) #mean of m
nu.mean=c(0,0,0,0,0,0,0,mu.m) #mean vector
Imat=kk
for(i in 1:n){
  nu[i,1:9] ~ dmnorm.vcov(nu.mean,Imat) #multivariate prior
}
#####
# Specify the (prior) distribution for Y1 and Y2
#####

for(j in 1:JJ){
  inv.psi1.y[j] ~ dgamma(1,1) # Precisions for observables
  inv.psi2.y[j] ~ dgamma(1,1) # Precisions for observables
  psi1.y[j] <- 1/inv.psi1.y[j]   # Variances for observables
  psi2.y[j] <- 1/inv.psi2.y[j]   # Variances for observables
}

```

```

for(j in 1:JJ){
  lambda.y[j] ~ dnorm(1,.5);T(0,,) #half normal distribution for Loadings
  beta.y[j] ~ dnorm(1,.5);T(0,,)
}
tau.y[1] <- 0
for(j in 2:JJ){
  tau.y[j] ~ dnorm(0,.5) # normal distribution for intercepts
}

nu.mean.y=c(0,0,0,0,0,0,0,0,0,0,0)
Imat2=mm
for(i in 1:n){
  nu.y[i,1:11] ~ dmnorm.vcov(nu.mean.y,Imat2) #multivariate priors
}

#####
# Structural pars
#####
int1~dnorm(0,.001)
int2~dnorm(0,.001)
apath~dnorm(0,.001)
bpath~dnorm(0,.001)
cpath~dnorm(0,.001)
smpath~dnorm(0,.001)
sypath~dnorm(0,.001)
cmpath~dnorm(0,.001)
cypath~dnorm(0,.001)
h1path~dnorm(0,.001)
h2path~dnorm(0,.001)
h3path~dnorm(0,.001)
h4path~dnorm(0,.001)

int0~dnorm(0,.001) #intercept in regression Y1 on M1
cov1~dnorm(0,.001) #reg coefficient in regression Y1 on M1

for (i in 1:n){
  mu.y[i]<-int0+cov1*nu[i,9]
  mu.y1[i] ~ dnorm(mu.y[i],1)
  kappa[i]<-int1+apath*x[i]+smpath*nu[i,9]+cypath*mu.y1[i]+
    h1path*nu[i,9]*x[i]+h2path*mu.y1[i]*x[i]
 ksi2[i] ~ dnorm(kappa[i],1) # distribution for the Latent variables
  kappa.y[i]<-int2+cpath*x[i]+bpath*ksi2[i]+sypath*mu.y1[i]+cmpath*nu[i,9]+
    h3path*nu[i,9]*x[i]+h4path*mu.y1[i]*x[i]
  ksi.y2[i] ~ dnorm(kappa.y[i],1) # distribution for the Latent variables
}

} # closes the model

```

```

#data and specification of values for Loops
jagsdat=list(m1=datt22[,2:7],m2=datt22[,19:24],y1=datt22[,8:18],
y2=datt22[,25:35], n=nrow(datt22), J=6, JJ=11, kk=diag(9), mm=diag(11), x=datt22[,1])

#parameters to track
bpar <- c("tau", "lambda", 'beta', 'psi1', 'psi2', 'int1', 'int2', 'smpath', 'sypath',
'cmpath', 'cypath', 'tau.e', 'tau.e2', "tau.y", "lambda.y", 'beta.y', 'psi1.y',
'psi2.y', 'apath', 'bpath', 'cpath', 'h1path', 'h2path', 'h3path', 'h4path',
'mu.m', 'cov1', 'int0')

#run model with the following specifications
#parallelization with three chains
#total iterations of 12000
#burn first 2000 iterations per chain
jagcfa4 <- jags.parallel(data = jagsdat,
                           parameters.to.save = bpar, n.chains=3,
                           n.iter = 12000, n.burnin = 2000, model.file = bcafa)

```