# Supplementary Materials <br> Civil conflict and firm recovery: Evidence from Côte d'Ivoire 

## A. Additional Tables and Figures

Table A1: Number of firms per year

|  | Panel A: All firms |  |  |  | Panel B: Cohort 2009 |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| Year | Total | Entry | Exit |  | Total | Entry | Exit |
| 2006 | 3,392 | 266 | 491 |  | 2,123 | 91 | - |
| 2007 | 3,829 | 508 | 535 |  | 2,587 | 223 | - |
| 2008 | 4,654 | 634 | 693 |  | 3,556 | 319 | - |
| $\mathbf{2 0 0 9}$ | $\mathbf{5 , 0 7 1}$ | $\mathbf{8 4 4}$ | $\mathbf{9 5 9}$ |  | $\mathbf{5 , 0 7 1}$ | $\mathbf{6 6 3}$ | $\mathbf{8 1 9}$ |
| 2010 | 8,196 | 892 | 1116 |  | 3,387 | - | 287 |
| 2011 | 10,411 | 904 | 1515 |  | 3,524 | - | 312 |
| 2012 | 12,945 | 2650 | 2572 |  | 3,260 | - | 333 |
| 2013 | 14,136 | 2451 | 3551 |  | 2,994 | - | 428 |
| 2014 | 17,106 | 3694 | - |  | $\mathbf{2 , 8 9 6}$ | - | - |

Total reports the total number of firms; Entry the number of new firms; and, Exit the number of exit firms
Table A2: Input dependence by sector

|  | Staff Cost |  |  |  | Managers |  |  |  | Wage Average |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Mean | Median | SDev | \% | Mean | Median | SDev | \% | Mean | Median | SDev | \% |
| Agriculture | 0.14 | 0.10 | 0.13 | 0.21 | 0.59 | 0.02 | 1.84 | 0.25 | 3531 | 1640 | 4102 | 0.38 |
| Fishing | 0.31 | 0.31 | 0.02 | 0.50 | 0.08 | 0.08 | 0.01 | 0.50 | 4539 | 4539 | 3919 | 0.50 |
| Extraction | 0.21 | 0.19 | 0.16 | 0.50 | 0.18 | 0.17 | 0.15 | 0.50 | 10991 | 8146 | 11457 | 0.50 |
| Manufacturing | 0.19 | 0.13 | 0.22 | 0.37 | 0.18 | 0.05 | 0.48 | 0.21 | 3045 | 2057 | 3255 | 0.30 |
| Electricity, gaz and water | 0.15 | 0.10 | 0.14 | 0.10 | 0.28 | 0.26 | 0.21 | 0.33 | 3497 | 3709 | 1613 | 0.09 |
| Construction | 0.16 | 0.11 | 0.20 | 0.23 | 0.48 | 0.14 | 1.92 | 0.17 | 3373 | 2302 | 3469 | 0.38 |
| Trade | 0.09 | 0.05 | 0.13 | 0.26 | 0.31 | 0.10 | 0.75 | 0.23 | 3578 | 2297 | 3692 | 0.31 |
| Hotels and restaurants | 0.20 | 0.19 | 0.12 | 0.33 | 0.12 | 0.05 | 0.18 | 0.33 | 2411 | 1740 | 2321 | 0.35 |
| Transport and communication | 0.24 | 0.18 | 0.20 | 0.38 | 0.27 | 0.12 | 0.48 | 0.29 | 5114 | 3344 | 4718 | 0.34 |
| Services to enterprises | 0.28 | 0.23 | 0.22 | 0.35 | 0.36 | 0.15 | 0.86 | 0.30 | 4627 | 3085 | 4580 | 0.33 |
| Education | 0.31 | 0.28 | 0.23 | 0.36 | 0.28 | 0.03 | 0.76 | 0.23 | 1849 | 1372 | 1435 | 0.34 |
| Health and social | 0.17 | 0.14 | 0.14 | 0.38 | 0.24 | 0.05 | 0.40 | 0.29 | 2264 | 1751 | 1785 | 0.33 |
| TOTAL | 0.19 | 0.12 | 0.20 | 0.31 | 0.31 | 0.09 | 0.92 | 0.25 | 36949 | 2315 | 3862 | 0.33 |

Access to capital

|  | Debt ratio |  |  |  | Trade Credit |  |  |  | Financial cost |  |  |  | interest rate |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Mean | Median | SDev | \% | Mean | Median | SDev | \% | Mean | Median | SDev | \% | Mean | Median | SDev | \% |
| Agriculture | 0.80 | 0.77 | 0.53 | 0.25 | 0.14 | 0.18 | 0.58 | 0.38 | 0.001 | 0.012 | 0.958 | 0.95 | 0.022 | 0.001 | 0.045 | 0.88 |
| Fishing | 0.72 | 0.72 | 0.37 | 0.08 | 0.15 | 0.15 | 0.01 | 1.00 | 0.008 | 0.008 | 0.011 | 0.96 | 0.021 | 0.021 | 0.026 | 0.50 |
| Extraction | 0.59 | 0.41 | 0.63 | 0.15 | 0.14 | 0.03 | 0.24 | 0.75 | 0.000 | 0.000 | 0.000 | 0.93 | 0.003 | 0.000 | 0.006 | 1.00 |
| Manufacturing | 1.14 | 0.80 | 1.88 | 0.20 | 0.22 | 0.22 | 0.17 | 0.50 | 0.006 | 0.000 | 0.012 | 0.83 | 0.039 | 0.001 | 0.103 | 0.83 |
| Electricity, gaz and water | 0.67 | 0.75 | 0.35 | 0.09 | 0.37 | 0.33 | 0.06 | 0.00 | 0.023 | 0.014 | 0.019 | 0.67 | 0.040 | 0.012 | 0.052 | 0.67 |
| Construction | 1.15 | 0.79 | 1.94 | 0.12 | 0.19 | 0.16 | 0.17 | 0.52 | 0.007 | 0.000 | 0.017 | 0.79 | 0.029 | 0.000 | 0.097 | 0.82 |
| Trade | 1.05 | 0.78 | 1.57 | 0.16 | 0.27 | 0.32 | 0.18 | 0.43 | 0.007 | 0.001 | 0.016 | 0.78 | 0.043 | 0.002 | 0.109 | 0.75 |
| Hotels and restaurants | 1.49 | 0.92 | 2.11 | 0.23 | 0.13 | 0.04 | 0.16 | 0.68 | 0.008 | 0.000 | 0.021 | 0.83 | 0.006 | 0.000 | 0.017 | 0.83 |
| Transport and communication | 0.95 | 0.75 | 1.04 | 0.13 | 0.21 | 0.17 | 0.19 | 0.50 | 0.020 | 0.002 | 0.037 | 0.74 | 0.041 | 0.003 | 0.099 | 0.75 |
| Services to enterprises | 1.32 | 0.87 | 2.56 | 0.17 | 0.14 | 0.07 | 0.16 | 0.61 | 0.006 | 0.000 | 0.017 | 0.83 | 0.019 | 0.000 | 0.072 | 0.85 |
| Education | 0.94 | 0.81 | 1.06 | 0.35 | 0.12 | 0.01 | 0.16 | 0.63 | 0.004 | 0.000 | 0.018 | 0.88 | 0.012 | 0.000 | 0.044 | 0.89 |
| Health and social | 0.89 | 0.62 | 1.18 | 0.22 | 0.14 | 0.08 | 0.16 | 0.63 | 0.004 | 0.000 | 0.010 | 0.79 | 0.027 | 0.000 | 0.080 | 0.81 |
| TOTAL | 1.11 | 0.80 | 1.84 | 0.18 | 0.20 | 0.17 | 0.18 | 0.52 | 0.007 | 0.000 | 0.018 | 0.81 | 0.032 | 0.000 | 0.093 | 0.80 |


Table A3: Input usage, between-industry vs. within-industry variation

| Input | All <br> Observations |  | Cohort <br> (all obs.) |  | Cohort (in 2009) |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | W/out control | With control |  |
|  | $\mathrm{R}^{2}$ | Obs. |  |  | $\mathrm{R}^{2}$ | Obs. | $\mathrm{R}^{2}$ | Obs. | $\mathrm{R}^{2}$ | Obs. |
| Staff cost | 0.081 | 71296 | 0.083 | 26055 | 0.075 | 4687 | 0.081 | 4684 |
| Share of manager | 0.007 | 72345 | 0.004 | 25870 | 0.004 | 4818 | 0.006 | 4818 |
| Share of permanent workers | 0.008 | 72346 | 0.006 | 25870 | 0.010 | 4818 | 0.011 | 4818 |
| Average wage | 0.005 | 70901 | 0.049 | 25336 | 0.050 | 4732 | 0.129 | 4732 |
| Debt ratio | 0.009 | 80428 | 0.008 | 27846 | 0.007 | 5147 | 0.004 | 5144 |
| Trade credit | 0.007 | 81369 | 0.035 | 28154 | 0.111 | 5186 | 0.002 | 5183 |
| Financial cost | 0.018 | 71327 | 0.021 | 25938 | 0.024 | 4660 | 0.023 | 4657 |
| Interest rate | 0.007 | 74720 | 0.010 | 26995 | 0.007 | 4874 | 0.037 | 4871 |

This table reports $\mathrm{R}^{2}$ of the model explaining input usage (each row) in different specifications including industry dummies (and firm characteristics in the last specification). The first specification considers all observations available. The second specification considers all observations for firms operating in 2009 (coor not of firm level characteristics (nb. of employees, sales (in log), age (in log), foreign ownership, dummy for Abidjan and two dummies for legal status).

Table A4: The global impact of the crisis on productivity

| Panel A: Labor productivity (Value added per workers) |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\log (L P)$ |  |  |  | $\Delta[\log (L P)]$ |  |  |  |
|  | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| POST x CONFL | $\begin{array}{r} -0.0064^{* *} \\ (0.0008) \end{array}$ | $\begin{array}{r} -0.0025^{* *} \\ (0.0007) \end{array}$ | $\begin{gathered} -0.0031^{* *} \\ (0.0007) \end{gathered}$ | $\begin{array}{r} -0.0026^{* *} \\ (0.0006) \end{array}$ | $\begin{array}{r} -0.0105^{* *} \\ (0.0014) \end{array}$ | $\begin{array}{r} -0.0026^{* *} \\ (0.0010) \end{array}$ | $\begin{array}{r} -0.0048^{* *} \\ (0.0008) \end{array}$ | $\begin{array}{r} -0.0024^{* *} \\ (0.0008) \end{array}$ |

Panel B: Panel B: Total factor productivity

|  | $\log (T F P)$ |  |  |  | $\Delta[\log (T F P)]$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| POST x CONFL | $\begin{array}{r} \hline-0.0068^{* *} \\ (0.0015) \end{array}$ | $\begin{array}{r} \hline-0.0199^{* *} \\ (0.0009) \end{array}$ | $\begin{array}{r} \hline-0.0182^{* *} \\ (0.0009) \end{array}$ | $\begin{array}{r} -0.0134^{* *} \\ (0.0009) \end{array}$ | $\begin{array}{r} -0.0092^{* *} \\ (0.0030) \end{array}$ | $\begin{array}{r} \hline-0.0029 \\ (0.0030) \end{array}$ | $\begin{array}{r} \hline-0.0017 \\ (0.0034) \end{array}$ | $\begin{array}{r} \hline-0.0017 \\ (0.0035) \end{array}$ |

Panel C: Value added

|  | $\log (V A)$ |  |  |  | $\Delta[\log (V A)]$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| POST x CONFL | -0.0050** | -0.0006 | $0.0012^{\dagger}$ | 0.0018** | -0.0079** | -0.0004 | -0.0006 | -0.0008 |
|  | (0.0008) | (0.0007) | (0.0006) | (0.0006) | (0.0013) | (0.0009) | (0.0008) | (0.0007) |

Panel D: Number of workers

|  | Log(Workers) |  |  |  | $\Delta[\log ($ Workers $)]$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| POST x CONFL | $\begin{array}{r} 0.0007 \\ (0.0004) \end{array}$ | $\begin{gathered} 0.0016^{* *} \\ (0.0004) \end{gathered}$ | $\begin{gathered} 0.0042^{* *} \\ (0.0004) \end{gathered}$ | $\begin{gathered} 0.0043^{* *} \\ (0.0004) \end{gathered}$ | $\begin{gathered} 0.0014^{\dagger} \\ (0.0007) \end{gathered}$ | $\begin{gathered} \hline 0.0014^{* *} \\ (0.0005) \end{gathered}$ | $\begin{gathered} 0.0034^{* *} \\ (0.0005) \end{gathered}$ | $\begin{gathered} \hline 0.0009^{*} \\ (0.0004) \end{gathered}$ |
| Obs. | 6,535 | 8,762 | 10,913 | 13,380 | 5,093 | 6,954 | 8,813 | 10,779 |
| Firms | 2,877 | 2,877 | 2,877 | 2,877 | 2,288 | 2,475 | 2,568 | 2,644 |
| Obs. (Panel B: TFP) | 3,249 | 5,262 | 5,497 | 6333 | 808 | 1,369 | 1,441 | 1,525 |
| Firms (Panel B: TFP) | 2,225 | 2,587 | 2,644 | 2,689 | 657 | 960 | 1,007 | 1,051 |
| Year included |  |  |  |  |  |  |  |  |
| 2009 | x | x | x | x | x | x | x | x |
| 2010 | x | x | x | x | x | x | x | x |
| 2011 | x | x | x | x | x | x | x | x |
| 2012 |  | x | x | x |  | x | x | x |
| 2013 |  |  | x | x |  |  | x | x |
| 2014 |  |  |  | x |  |  |  | x |

The dependent variable is the logarithm of labor productivity (Panel A), total factor productivity (Panel B), value added (Panel C), the number of workers (Panel D). In columns (1) to (4), the dependent variable is expressed in logarithm and in difference in logarithm (growth) in columns (5) to (8). POST $\times C O N F L$ is a variable equal to zero before the crisis (e.g., 2011) and equals to the number of deaths per 100,000 inhabitants in the district after 2011. The years from 2011 to 2014 are included one by one as indicated at the bottom of the table. Firm-level and year fixed effects are included and standard errors are clustered at the firm-level. The number of observations and firms refers to the models in Panels A, C and D. Standard errors are clustered at the firm level, except in Panel B (bootstrapping with 500 replications because the dependent variable is a generated variable). ${ }^{\dagger},^{*}$, and ${ }^{* *}$ signal significance at the $10 \%, 5 \%$ and $1 \%$ levels, respectively.

Table A5: Heterogenous impact of the crisis, baseline results

|  | (0) | (1) | (2) | (3) | (4) | (5) | (6) | (7) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| POSTxCONF | $\begin{aligned} & -0.185^{* *} \\ & (0.010) \end{aligned}$ | $\begin{aligned} & -0.188^{* *} \\ & (0.010) \end{aligned}$ | $\begin{aligned} & -0.191^{* *} \\ & (0.010) \end{aligned}$ | $\begin{aligned} & -0.216^{* *} \\ & (0.011) \end{aligned}$ | $\begin{aligned} & -0.182^{* *} \\ & (0.010) \end{aligned}$ | $\begin{aligned} & -0.191^{* *} \\ & (0.010) \end{aligned}$ | $\begin{aligned} & -0.186^{* *} \\ & (0.010) \end{aligned}$ | $\begin{aligned} & -0.191^{* *} \\ & (0.010) \end{aligned}$ |
| (POSTxCONF) $\times \log ($ EMPL $)$ | $\begin{aligned} & -0.019^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.020^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.019^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.019^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.019^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.019^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.019^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.019 * * \\ & (0.001) \end{aligned}$ |
| (POSTxCONF) $\times \log$ (Sales) | $\begin{aligned} & 0.012^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.014^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.012^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.013^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.012^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.013^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.012^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.012^{* *} \\ & (0.001) \end{aligned}$ |
| (POSTxCONF) $\times \log$ (Age) | $\begin{aligned} & 0.001 \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.000 \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.002^{*} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.001) \end{aligned}$ |
| (POSTxCONF) x Abidjan | $\begin{aligned} & -0.002 \\ & (0.003) \end{aligned}$ | $\begin{aligned} & -0.011^{* *} \\ & (0.003) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.003) \end{aligned}$ | $\begin{aligned} & 0.005 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.004 \\ & (0.003) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.003) \end{aligned}$ | $\begin{aligned} & -0.002 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.005 \\ & (0.004) \end{aligned}$ |
| (POSTxCONF) x Foreign | $\begin{aligned} & -0.002 \\ & (0.001) \end{aligned}$ | $\begin{gathered} -0.002^{\dagger} \\ (0.001) \end{gathered}$ | $\begin{aligned} & -0.002 \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.002 \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.002 \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.002 \\ & (0.001) \end{aligned}$ |
| (POSTxCONF) x LimLiabilities | $\begin{aligned} & 0.002+ \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.003^{\dagger} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.005^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.002^{*} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.001) \end{aligned}$ |
| (POSTxCONF) x PublicCompany | $\begin{aligned} & -0.002 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.004^{*} \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.002 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.002 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.002 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.002 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.002 \\ & (0.002) \end{aligned}$ |
| (POSTxCONF) x Agriculture | $\begin{aligned} & 0.002 \\ & (0.007) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.007) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.007) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.007) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.007) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.007) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.007) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.007) \end{aligned}$ |
| (POSTxCONF) x Fishing | $\begin{aligned} & 0.007 \\ & (0.007) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.008) \end{aligned}$ | $\begin{aligned} & 0.008 \\ & (0.006) \end{aligned}$ | $\begin{aligned} & 0.008 \\ & (0.006) \end{aligned}$ | $\begin{aligned} & 0.008 \\ & (0.007) \end{aligned}$ | $\begin{aligned} & 0.007 \\ & (0.007) \end{aligned}$ | $\begin{aligned} & 0.007 \\ & (0.007) \end{aligned}$ | $\begin{aligned} & 0.007 \\ & (0.007) \end{aligned}$ |
| (POSTxCONF) x Extraction | $\begin{aligned} & -0.018^{\dagger} \\ & (0.011) \end{aligned}$ | $\begin{gathered} -0.018^{\dagger} \\ (0.010) \end{gathered}$ | $\begin{aligned} & -0.017^{\dagger} \\ & (0.010) \end{aligned}$ | $\begin{aligned} & -0.019^{\dagger} \\ & (0.010) \end{aligned}$ | $\begin{aligned} & -0.017 \\ & (0.011) \end{aligned}$ | $\begin{aligned} & -0.018^{\dagger} \\ & (0.011) \end{aligned}$ | $\begin{aligned} & -0.018^{\dagger} \\ & (0.011) \end{aligned}$ | $\begin{gathered} -0.018^{\dagger} \\ (0.011) \end{gathered}$ |
| (POSTxCONF) x Manufacturing | $\begin{aligned} & -0.000 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.000 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.000 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.000 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.000 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.000 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.000 \\ & (0.004) \end{aligned}$ |
| (POSTxCONF) x Electricity | $\begin{aligned} & 0.002 \\ & (0.007) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.007) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.007) \end{aligned}$ | $\begin{aligned} & -0.002 \\ & (0.007) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.007) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.007) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.007) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.007) \end{aligned}$ |
| (POSTxCONF) $\times$ Construction | $\begin{aligned} & -0.003 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.004 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.004 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.002 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.003 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.003 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.003 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.003 \\ & (0.004) \end{aligned}$ |
| (POSTxCONF) x Trade | $\begin{aligned} & -0.013^{* *} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.016^{* *} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.013^{* *} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.013^{* *} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.013^{* *} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.013^{* *} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.013^{* *} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.013^{* *} \\ & (0.004) \end{aligned}$ |
| (POSTxCONF) x Hotels | $\begin{aligned} & 0.012^{*} \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.011^{*} \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.013^{*} \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.013^{* *} \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.012^{*} \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.012^{*} \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.012^{*} \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.012^{*} \\ & (0.005) \end{aligned}$ |
| (POSTxCONF) x Transport | $\begin{aligned} & 0.001 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.000 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.004) \end{aligned}$ |
| (POSTxCONF) x ServicesEnt | $\begin{aligned} & 0.000 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.000 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.000 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.000 \\ & (0.004) \end{aligned}$ |
| (POSTxCONF) x Education | $\begin{aligned} & 0.010^{*} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.010^{*} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.010^{*} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.012^{*} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.010^{*} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.010^{*} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.010^{*} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.010^{*} \\ & (0.004) \end{aligned}$ |
| (POSTxCONF) x Social | $\begin{aligned} & 0.003 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.004 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.006 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.004 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.004 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.003 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.004 \\ & (0.004) \end{aligned}$ |
| (POSTxCONF) x StaffCost |  | $\begin{aligned} & 0.503^{* *} \\ & (0.040) \end{aligned}$ |  |  |  |  |  |  |
| (POSTxCONF) x Managers |  |  | $\begin{aligned} & -0.122^{* *} \\ & (0.040) \end{aligned}$ |  |  |  |  |  |
| (POSTxCONF) x AverWage |  |  |  | $\begin{aligned} & -0.322^{* *} \\ & (0.037) \end{aligned}$ |  |  |  |  |
| (POSTxCONF) x Debt |  |  |  |  | $\begin{aligned} & 0.142^{* *} \\ & (0.045) \end{aligned}$ |  |  |  |
| (POSTxCONF) x TradeCredit |  |  |  |  |  | $\begin{aligned} & 0.100^{* *} \\ & (0.037) \end{aligned}$ |  |  |
| (POSTxCONF) x FinCost |  |  |  |  |  |  | $\begin{aligned} & 0.006 \\ & (0.041) \end{aligned}$ |  |
| (POSTxCONF) x IntRate |  |  |  |  |  |  |  | $\begin{aligned} & 0.096^{*} \\ & (0.040) \end{aligned}$ |
| Combined effect | -0.185** | 0.315** | -0.313** | -0.538** | -0.040 | -0.091 | -0.179* | -0.095* |
| Obs. | 11178 | 11178 | $\mathrm{V}^{11178}$ | 11178 | 11178 | 11178 | 11178 | 11178 |
| \# firms | 2347 | 2347 | 2347 | 2347 | 2347 | 2347 | 2347 | 2347 |
| $R^{2}$ (within) | 0.178 | 0.202 | 0.179 | 0.187 | 0.179 | 0.179 | 0.178 | 0.178 |

The dependent variable is the logarithm of labor productivity. $P O S T \times C O N F L$ is a variable equal to zero before the crisis (e.g., 2011) and equals to the number of deaths per 100,000 inhabitants in the district after 2011. Combined effect measures the point estimates for firm with input dummy equals to one (and associated statistical significance). Within estimator (firm fixed effect) is used. Standard errors are clustered at the firm-level. ${ }^{\dagger},{ }^{*}$, and ${ }^{* *}$ signal significance at the $10 \%, 5 \%$ and $1 \%$ levels, respectively.

Table A6: Heterogenous impact of the crisis, total factor productivity

|  | (0) | (1) | (2) | (3) | (4) | (5) | (6) | (7) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| POSTxCONF | $\begin{aligned} & \hline-0.090^{* *} \\ & (0.015) \end{aligned}$ | $\begin{aligned} & \hline-0.101^{* *} \\ & (0.015) \end{aligned}$ | $\begin{aligned} & \hline-0.094 * * \\ & (0.016) \end{aligned}$ | $\begin{aligned} & \hline-0.115^{* *} \\ & (0.017) \end{aligned}$ | $\begin{aligned} & \hline-0.087^{* *} \\ & (0.015) \end{aligned}$ | $\begin{aligned} & \hline-0.095^{* *} \\ & (0.016) \end{aligned}$ | $\begin{aligned} & \hline-0.082^{* *} \\ & (0.015) \end{aligned}$ | $\begin{aligned} & \hline-0.092^{* *} \\ & (0.015) \end{aligned}$ |
| (POSTxCONF) $\times \log ($ EMPL $)$ | $\begin{aligned} & -0.007^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.008^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.007^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.008^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.007^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.007^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.007^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.007^{* *} \\ & (0.001) \end{aligned}$ |
| (POSTxCONF) $\times \log$ (Sales) | $\begin{aligned} & 0.006^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.007^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.006^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.006^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.006^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.006^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.006^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & 0.006^{* *} \\ & (0.001) \end{aligned}$ |
| (POSTxCONF) $\times \log$ (Age) | $\begin{aligned} & -0.004^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.005^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.004^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.003^{*} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.004^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.004^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.004^{* *} \\ & (0.001) \end{aligned}$ | $\begin{aligned} & -0.004^{* *} \\ & (0.001) \end{aligned}$ |
| (POSTxCONF) x Abidjan | $\begin{aligned} & -0.002 \\ & (0.006) \end{aligned}$ | $\begin{aligned} & -0.008 \\ & (0.006) \end{aligned}$ | $\begin{aligned} & -0.000 \\ & (0.006) \end{aligned}$ | $\begin{aligned} & 0.003 \\ & (0.006) \end{aligned}$ | $\begin{aligned} & -0.004 \\ & (0.006) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.006) \end{aligned}$ | $\begin{aligned} & -0.004 \\ & (0.006) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.006) \end{aligned}$ |
| (POSTxCONF) x Foreign | $\begin{aligned} & -0.001 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.002) \end{aligned}$ |
| (POSTxCONF) x LimLiabilities | $\begin{aligned} & -0.001 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.002 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.002) \end{aligned}$ | $\begin{aligned} & -0.001 \\ & (0.002) \end{aligned}$ |
| (POSTxCONF) x PublicCompany | $\begin{aligned} & -0.007^{* *} \\ & (0.003) \end{aligned}$ | $\begin{aligned} & -0.009^{* *} \\ & (0.003) \end{aligned}$ | $\begin{aligned} & -0.007^{* *} \\ & (0.003) \end{aligned}$ | $\begin{aligned} & -0.005^{\dagger} \\ & (0.003) \end{aligned}$ | $\begin{aligned} & -0.007^{* *} \\ & (0.003) \end{aligned}$ | $\begin{aligned} & -0.007^{* *} \\ & (0.003) \end{aligned}$ | $\begin{aligned} & -0.008^{* *} \\ & (0.003) \end{aligned}$ | $\begin{aligned} & -0.007^{* *} \\ & (0.003) \end{aligned}$ |
| (POSTxCONF) x Agriculture | $\begin{aligned} & 0.013^{*} \\ & (0.006) \end{aligned}$ | $\begin{aligned} & 0.014^{*} \\ & (0.006) \end{aligned}$ | $\begin{aligned} & 0.013^{*} \\ & (0.006) \end{aligned}$ | $\begin{aligned} & 0.013^{*} \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.013^{*} \\ & (0.006) \end{aligned}$ | $\begin{aligned} & 0.013^{*} \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.013^{*} \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.013^{*} \\ & (0.006) \end{aligned}$ |
| (POSTxCONF) x Fishing | $\begin{aligned} & 0.030^{* *} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.019^{* *} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.029^{* *} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.027^{* *} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.031^{* *} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.029^{* *} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.030^{* *} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.029^{* *} \\ & (0.004) \end{aligned}$ |
| (POSTxCONF) x Extraction | $\begin{aligned} & 0.024^{* *} \\ & (0.008) \end{aligned}$ | $\begin{aligned} & 0.021^{* *} \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.023^{* *} \\ & (0.008) \end{aligned}$ | $\begin{aligned} & 0.022^{* *} \\ & (0.008) \end{aligned}$ | $\begin{aligned} & 0.025^{* *} \\ & (0.008) \end{aligned}$ | $\begin{aligned} & 0.023^{* *} \\ & (0.008) \end{aligned}$ | $\begin{aligned} & 0.024^{* *} \\ & (0.008) \end{aligned}$ | $\begin{aligned} & 0.024^{* *} \\ & (0.008) \end{aligned}$ |
| (POSTxCONF) x Manufacturing | $\begin{aligned} & 0.003 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.003 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.003 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.003 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.004 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.003 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.003 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.003 \\ & (0.004) \end{aligned}$ |
| (POSTxCONF) x Electricity | $\begin{aligned} & 0.006 \\ & (0.008) \end{aligned}$ | $\begin{aligned} & 0.007 \\ & (0.008) \end{aligned}$ | $\begin{aligned} & 0.006 \\ & (0.008) \end{aligned}$ | $\begin{aligned} & 0.003 \\ & (0.007) \end{aligned}$ | $\begin{aligned} & 0.007 \\ & (0.008) \end{aligned}$ | $\begin{aligned} & 0.007 \\ & (0.008) \end{aligned}$ | $\begin{aligned} & 0.007 \\ & (0.008) \end{aligned}$ | $\begin{aligned} & 0.006 \\ & (0.008) \end{aligned}$ |
| (POSTxCONF) x Construction | $\begin{aligned} & 0.002 \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.004 \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.003 \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.003 \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.005) \end{aligned}$ |
| (POSTxCONF) x Trade | $\begin{aligned} & -0.008^{*} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.010^{*} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.008^{*} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.008^{*} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.007^{\dagger} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.008^{*} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & -0.009^{*} \\ & (0.004) \end{aligned}$ | $\begin{gathered} -0.008^{*} \\ (0.004) \end{gathered}$ |
| (POSTxCONF) x Hotels | $\begin{aligned} & 0.014^{* *} \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.013^{*} \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.014^{* *} \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.015^{* *} \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.014^{* *} \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.014^{* *} \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.014^{* *} \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.014^{* *} \\ & (0.005) \end{aligned}$ |
| (POSTxCONF) x Transport | $\begin{aligned} & 0.002 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.003 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.004) \end{aligned}$ |
| (POSTxCONF) x ServicesEnt | $\begin{aligned} & 0.002 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.003 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.003 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.001 \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.002 \\ & (0.004) \end{aligned}$ |
| (POSTxCONF) x Education | $\begin{aligned} & 0.012^{* *} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.013^{* *} \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.013^{* *} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.014^{* *} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.012^{* *} \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.012^{* *} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.012^{* *} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.012^{* *} \\ & (0.004) \end{aligned}$ |
| (POSTxCONF) x Social | $\begin{aligned} & 0.011^{*} \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.010^{*} \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.012^{* *} \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.013^{* *} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.012^{* *} \\ & (0.005) \end{aligned}$ | $\begin{aligned} & 0.012^{* *} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.011^{*} \\ & (0.004) \end{aligned}$ | $\begin{aligned} & 0.012^{*} \\ & (0.005) \end{aligned}$ |
| (POSTxCONF) x StaffCost |  | $\begin{aligned} & 0.408^{* *} \\ & (0.050) \end{aligned}$ |  |  |  |  |  |  |
| (POSTxCONF) x Managers |  |  | $\begin{aligned} & -0.079 \\ & (0.057) \end{aligned}$ |  |  |  |  |  |
| (POSTxCONF) x AverWage |  |  |  | $\begin{aligned} & -0.235^{* *} \\ & (0.052) \end{aligned}$ |  |  |  |  |
| (POSTxCONF) x Debt |  |  |  |  | $\begin{aligned} & 0.130^{*} \\ & (0.062) \end{aligned}$ |  |  |  |
| (POSTxCONF) x TradeCredit |  |  |  |  |  | $\begin{aligned} & 0.084^{\dagger} \\ & (0.047) \end{aligned}$ |  |  |
| (POSTxCONF) x FinCost |  |  |  |  |  |  | $\begin{aligned} & 0.150^{* *} \\ & (0.049) \end{aligned}$ |  |
| (POSTxCONF) x IntRate |  |  |  |  |  |  |  | $\begin{aligned} & -0.041 \\ & (0.048) \end{aligned}$ |
| Combined effect | -0.090** | 0.307** | $-0.173^{* *}$ | -0.350** | 0.043 | 0.011 | 0.068 | 0.133 |
| Obs. | 4640 | 4640 | Vi ${ }^{4640}$ | 4640 | 4640 | 4640 | 4640 | 4640 |
| \# firms | 2088 | 2088 | 2088 | 2088 | 2088 | 2088 | 2088 | 2088 |
| $R^{2}$ (within) | 0.138 | 0.161 | 0.139 | 0.146 | 0.140 | 0.139 | 0.141 | 0.138 |

The dependent variable is the logarithm of total factor productivity. $P O S T \times C O N F L$ is a variable equal to zero before the crisis (e.g., 2011) and equal to the number of deaths per 100,000 inhabitants in the region after 2011. Combined effect measures the point estimates for firm with input dummy equals to one (and associated statistical significance). Within estimator (firm fixed effect) is used. Standard errors are clustered at the firm-level. ${ }^{\dagger}$, ${ }^{*}$, and ${ }^{* *}$ signal significance at the $10 \%, 5 \%$ and $1 \%$ levels, respectively.

Table A7: Input reliance and its impact on value added and the number of workers

| Panel A: Value added (in log) |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Input $\rightarrow$ | Staff <br> (1) | Manager (2) | Avg wage (3) | Debt <br> (4) | Trade C. (5) | FinCost <br> (6) | IntRate (7) |
| (POSTxCONF) x Input | $\begin{aligned} & \hline 0.261^{* *} \\ & (0.037) \end{aligned}$ | $\begin{aligned} & 0.061 \\ & (0.038) \end{aligned}$ | $\begin{aligned} & \hline-0.207^{* *} \\ & (0.034) \end{aligned}$ | $\begin{aligned} & \hline 0.174^{* *} \\ & (0.045) \end{aligned}$ | $\begin{aligned} & \hline 0.128^{* *} \\ & (0.034) \end{aligned}$ | $\begin{aligned} & \hline 0.117^{* *} \\ & (0.038) \end{aligned}$ | $\begin{aligned} & \hline 0.189^{* *} \\ & (0.035) \end{aligned}$ |
| Obs. | 11178 | 11178 | 11178 | 11178 | 11178 | 11178 | 11178 |
| $R^{2}$ (within) | 0.083 | 0.076 | 0.080 | 0.078 | 0.077 | 0.077 | 0.079 |

Panel B: The number of workers (in log)

| Input $\rightarrow$ | Staff |  |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | $(1)$ | Manager <br> $(2)$ | Avg wage <br> $(3)$ | Debt <br> $(4)$ | Trade C. <br> $(5)$ | FinCost <br> $(6)$ | IntRate <br> $(7)$ |
| (POSTxCONF) x Input | $-0.241^{* *}$ | $0.182^{* *}$ | $0.119^{* *}$ | 0.023 | 0.029 | $0.115^{* *}$ | $0.092^{* *}$ |
|  | $(0.025)$ | $(0.025)$ | $(0.026)$ | $(0.028)$ | $(0.025)$ | $(0.028)$ | $(0.029)$ |
| Obs. |  |  |  |  |  |  |  |
| $R^{2}$ (within) | 11178 | 11178 | 11178 | 11178 | 11178 | 11178 | 11178 |

The dependent variable is the logarithm of value added in Panel A and the logarithm of workers in Panel B. $P O S T \times C O N F L$ is a variable equal to zero before the crisis (e.g., 2011) and equal to the number of deaths per 100,000 inhabitants in the district after 2011. Within estimator (firm fixed effect) is used. In each column, interactions with firms' characteristics are included (but unreported). Firm fixed effect as well as control interactions are included but unreported. Standard errors are clustered at the firm-level. * and ** indicate significance at the $5 \%$ and $1 \%$ levels, respectively.

Table A8: Evolution of the value added and workers according to dependence to highskilled workers

|  | Log(Value added) |  |  |  | Log(Workers) |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Panel A: All firms |  |  |  |  |  |  |  |  |
| POST x CONFL | $\begin{array}{r} -0.00512^{* * *} \\ (0.000819) \end{array}$ | $\begin{gathered} -0.000668 \\ (0.000736) \end{gathered}$ | $\begin{array}{r} 0.00125^{*} \\ (0.000699) \end{array}$ | $\begin{array}{r} 0.00187^{* * *} \\ (0.000697) \end{array}$ | $\begin{array}{r} 0.000631 \\ (0.000386) \end{array}$ | $\begin{gathered} 0.00163^{* * *} \\ (0.000368) \end{gathered}$ | $\begin{gathered} 0.00351^{* * *} \\ (0.000392) \end{gathered}$ | $\begin{array}{r} 0.00386^{* * *} \\ (0.000399) \end{array}$ |

Panel B1: High dependence before the crisis (manager dummy=1)

| POST x CONFL | -0.00313 | 0.00188 | $0.00390^{* *}$ | $0.00526^{* * *}$ | $0.00369^{* * *}$ | $0.00479^{* * *}$ | $0.00675^{* * *}$ | $0.00758^{* * *}$ |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
|  | $(0.00192)$ | $(0.00175)$ | $(0.00168)$ | $(0.00159)$ | $(0.000738)$ | $(0.000683)$ | $(0.000718)$ | $(0.000744)$ |
| Panel B2: Low dependence before the crisis (manager dummy=0) |  |  |  |  |  |  |  |  |
| POST x CONFL | $-0.00579^{* * *}$ | $-0.00162^{* *}$ | 0.000257 | 0.000534 | -0.000582 | 0.000339 | $0.00218^{* * *}$ | $0.00232^{* * *}$ |
|  | $(0.000880)$ | $(0.000778)$ | $(0.000733)$ | $(0.000755)$ | $(0.000451)$ | $(0.000434)$ | $(0.000466)$ | $(0.000468)$ |

Panel C1: High dependence before the crisis (average wage dummy=1)

| POST x CONFL | $-0.00321^{* * *}$ | 0.000788 | $0.00247^{* *}$ | $0.00316^{* * *}$ | $0.00292^{* * *}$ | $0.00429^{* * *}$ | $0.00769^{* * *}$ | $0.00827^{* * *}$ |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
|  | $(0.00115)$ | $(0.00114)$ | $(0.00108)$ | $(0.00109)$ | $(0.000632)$ | $(0.000597)$ | $(0.000638)$ | $(0.000661)$ |

Panel C2: Low dependence before the crisis (average wage dummy=0)

|  |  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| POST x CONFL | $-0.00650^{* * *}$ | $-0.00171^{*}$ | 0.000304 | 0.000908 | $-0.000834^{*}$ | -0.0000912 | $0.000838^{*}$ | $0.000998^{* *}$ |
|  | $(0.00114)$ | $(0.000963)$ | $(0.000919)$ | $(0.000914)$ | $(0.000487)$ | $(0.000464)$ | $(0.000488)$ | $(0.000489)$ |


| Year included |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2009 | x | x | x | x | x | x | x | x |
| 2010 | x | x | x | x | x | x | x | x |
| 2011 | x | x | x | x | x | x | x | x |
| 2012 |  | x | x | x |  | x | x | x |
| 2013 |  |  | x | x |  |  | x | x |
| 2014 |  |  |  | x |  |  |  | x |

The specification is the same as that employed in Table 4, except dependent variables. The dependent variable is the logarithm of value added in columns (1) to (4) and the logarithm of total workers in columns (5) to (8) . In Panel A, we display results for all firms. In Panel B1/2, we display results for firms relying more (resp. less) on managers. In Panel C, we classify firms according to the value of average wage. Standard errors are clustered at the firm-level. *, **, and ${ }^{* * *}$ indicate significance at the $10 \%, 5 \%$ and $1 \%$ levels, respectively.

Table A9: Evolution of the share of managers and average wage

|  | Share of managers |  |  |  | Average wage |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Panel A: All firms POST x CONFL | $\begin{array}{r} -0.00469^{* * *} \\ (0.000348) \end{array}$ | $\begin{array}{r} -0.00342^{* * *} \\ (0.000303) \end{array}$ | $\begin{array}{r} -0.00312^{* * *} \\ (0.000292) \end{array}$ | $\begin{array}{r} -0.00292^{* * *} \\ (0.000305) \end{array}$ | $\begin{array}{r} -9.269^{* * *} \\ (1.270) \end{array}$ | $\begin{array}{r} 6.045^{* * *} \\ (1.257) \end{array}$ | $\begin{array}{r} -3.098^{* *} \\ (1.282) \end{array}$ | $\begin{array}{r} -3.090^{* *} \\ (1.251) \end{array}$ |
| Panel B1: High d POST x CONFL | $\begin{aligned} & \text { yendence befor } \\ & -0.0154^{* * *} \\ & (0.00120) \end{aligned}$ | $\begin{aligned} & \text { the crisis (ma } \\ & -0.0138^{* * *} \\ & (0.00101) \end{aligned}$ | $\begin{gathered} \text { ger dummy }= \\ -0.0134^{* * *} \\ (0.000954) \end{gathered}$ | $\begin{array}{r} -0.0134^{* * *} \\ (0.00101) \end{array}$ | $\begin{array}{r} -8.339^{* * *} \\ (2.916) \end{array}$ | $\begin{array}{r} -6.502^{* *} \\ (2.791) \end{array}$ | $\begin{gathered} -3.661 \\ (2.814) \end{gathered}$ | $\begin{gathered} -4.512 \\ (2.827) \end{gathered}$ |
| Panel B2: Low de POST x CONFL | endence befor $\begin{array}{r} -0.000990^{* * *} \\ (0.000125) \end{array}$ | $\begin{gathered} \text { he crisis (ma } \\ 0.000246^{* *} \\ (0.000115) \end{gathered}$ | $\begin{aligned} & \text { ager dummy= }= \\ & 0.000565^{* * *} \\ & (0.000118) \end{aligned}$ | $\begin{array}{r} 0.000822^{* * *} \\ (0.000118) \end{array}$ | $\begin{array}{r} 9.820^{* * *} \\ (1.381) \end{array}$ | $\begin{array}{r} 5.463^{* * *} \\ (1.384) \end{array}$ | $\begin{gathered} 2.934^{* *} \\ (1.420) \end{gathered}$ | $\begin{gathered} 2.576^{*} \\ (1.363) \end{gathered}$ |

Panel C1: High dependence before the crisis (average wage dummy=1)

|  |  |  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| POST x CONFL | $-0.00469^{* * *}$ | $-0.00265^{* * *}$ | $-0.00237^{* * *}$ | $-0.00189^{* * *}$ | $-23.96^{* * *}$ | $-24.66^{* * *}$ | $-25.87^{* * *}$ | $-26.74^{* * *}$ |
|  | $(0.000375)$ | $(0.000341)$ | $(0.000332)$ | $(0.000331)$ | $(3.108)$ | $(2.949)$ | $(3.010)$ | $(2.968)$ |
| Panel C2: Low dependence before the crisis (average wage dummy $=0)$ |  |  |  |  |  |  |  |  |
| POST x CONFL | $-0.00470^{* * *}$ | $-0.00386^{* * *}$ | $-0.00358^{* * *}$ | $-0.00352^{* * *}$ | -1.232 | $11.76^{* * *}$ | $9.349^{* * *}$ | $9.676^{* * *}$ |
|  | $(0.000510)$ | $(0.000439)$ | $(0.000421)$ | $(0.000444)$ | $(0.837)$ | $(1.045)$ | $(0.964)$ | $(0.909)$ |


| Year included |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 2009 | x | x | x | x | x | x | x | x |
| 2010 | x | x | x | x | x | x | x | x |
| 2011 | x | x | x | x | x | x | x | x |
| 2012 |  | x | x | x |  | x | x | x |
| 2013 |  |  | x | x |  |  | x | x |
| 2014 |  |  |  | x |  |  |  | x |

The specification is the same as that employed in Table 4, except dependent variables. The dependent variable is the share of managers in columns (1) to (4) and the average wage in columns (5) to (8). In Panel A, we display results for all firms. In Panel B1/2, we display results for firms relying more (resp. less) on managers. In Panel C, we classify firms according to the value of average wage. Standard errors are clustered at the firm-level. ${ }^{*},{ }^{* *}$, and ${ }^{* * *}$ indicate significance at the $10 \%$, $5 \%$ and $1 \%$ levels, respectively.

Table A10: Robustness checks (1) - Location

| Panel A: Firms outside Abidjan |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Input $\rightarrow$ | Staff <br> (1) | Manager <br> (2) | Avg wage <br> (3) | Debt <br> (4) | Trade C. <br> (5) | FinCost <br> (6) | IntRate <br> (7) |
| POST x CONFL | $\begin{aligned} & 0.376^{*} \\ & (0.112) \end{aligned}$ | $\begin{aligned} & -0.221+ \\ & (0.135) \end{aligned}$ | $\begin{aligned} & -0.196+ \\ & (0.107) \end{aligned}$ | $\begin{aligned} & \hline 0.143 \\ & (0.141) \end{aligned}$ | $\begin{aligned} & 0.062+ \\ & (0.063) \end{aligned}$ | $\begin{aligned} & -0.095 \\ & (0.115) \end{aligned}$ | $\begin{aligned} & 0.205 \\ & (0.167) \end{aligned}$ |
| Obs. | 1076 | 1076 | 1076 | 1076 | 1076 | 1076 | 1076 |
| $R^{2}$ (within) | 0.172 | 0.171 | 0.171 | 0.172 | 0.171 | 0.171 | 0.172 |

Panel B: Weigthed observations per number of firms in each locality

| Input $\rightarrow$ | Staff |  |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | $(1)$ | Manager |  |  |  |  |  |
| $(2)$ | Avg wage |  |  |  |  |  |  |
| $(3)$ | Debt <br> $(4)$ | Trade C. <br> $(5)$ | FinCost <br> $(6)$ | IntRate <br> $(7)$ |  |  |  |
| POST x CONFL | $0.596^{* *}$ | $-0.158^{* *}$ | $-0.408^{* *}$ | $0.148^{*}$ | $-0.090^{*}$ | -0.027 | $-0.142^{* *}$ |
|  | $(0.047)$ | $(0.050)$ | $(0.044)$ | $(0.059)$ | $(0.043)$ | $(0.050)$ | $(0.048)$ |
| Obs. |  |  |  |  |  |  |  |
| $R^{2}$ (within) | 11178 | 11178 | 11178 | 11178 | 11178 | 11178 | 11178 |

The dependent variable is the logarithm of labor productivity. $P O S T \times C O N F L$ is a variable equal to zero before the crisis (e.g., 2011) and equal to the number of deaths per 100,000 inhabitants in the district after 2011. Within estimator (firm fixed effect) is used. In each column, interactions with firms' characteristics are included (but unreported). Firm fixed effect as well as control interactions are included but unreported. Standard errors are clustered at the firm-level. $\dagger,{ }^{*}$ and ${ }^{* *}$ indicate significance at the $10 \%, 5 \%$ and $1 \%$ levels, respectively.

Table A11: Robustness checks (2) - Dependent variable

| Panel A: Labor productivity |  |  |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Input $\rightarrow$ |  | Staff | Manager | Avg wage | Debt | Trade C. | FinCost |
|  | $(1)$ | $(2)$ | $(3)$ | $(4)$ | $(5)$ | $(6)$ | $(7)$ |
| POST x CONFL | $0.477^{* *}$ | $-0.093^{*}$ | $-0.283^{* *}$ | $0.142^{* *}$ | $0.092^{*}$ | 0.023 | $0.133^{* *}$ |
|  | $(0.039)$ | $(0.039)$ | $(0.036)$ | $(0.047)$ | $(0.036)$ | $(0.041)$ | $(0.037)$ |
|  |  |  |  |  |  |  |  |
| Obs. | 11171 | 11171 | 11171 | 11171 | 11171 | 11171 | 11171 |
| $R^{2}$ within | 0.139 | 0.116 | 0.123 | 0.117 | 0.116 | 0.115 | 0.117 |

Panel B: Labor productivity measured as value added per total payroll

| Input $\rightarrow$ | Staff <br> (1) | Manager (2) | Avg wage <br> (3) | Debt <br> (4) | Trade C. <br> (5) | FinCost <br> (6) | IntRate <br> (7) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| POST x CONFL | $\begin{aligned} & 0.240^{* *} \\ & (0.021) \end{aligned}$ | $\begin{aligned} & 0.016 \\ & (0.022) \end{aligned}$ | $\begin{aligned} & 0.021 \\ & (0.019) \end{aligned}$ | $\begin{aligned} & 0.110^{* *} \\ & (0.022) \end{aligned}$ | $\begin{aligned} & 0.026 \\ & (0.019) \end{aligned}$ | $\begin{aligned} & 0.017 \\ & (0.020) \end{aligned}$ | $\begin{aligned} & 0.044^{*} \\ & (0.019) \end{aligned}$ |
| Obs. <br> $R^{2}$ within | $\begin{aligned} & 11028 \\ & 0.053 \end{aligned}$ | $\begin{aligned} & 11028 \\ & 0.032 \end{aligned}$ | $\begin{aligned} & 11028 \\ & 0.033 \end{aligned}$ | $\begin{aligned} & 11028 \\ & 0.036 \end{aligned}$ | $\begin{aligned} & 11028 \\ & 0.033 \end{aligned}$ | $\begin{aligned} & 11028 \\ & 0.032 \end{aligned}$ | $\begin{aligned} & 11028 \\ & 0.033 \end{aligned}$ |
| Panel C: log of profit |  |  |  |  |  |  |  |
| Input $\rightarrow$ | Staff <br> (1) | Manager (2) | Avg wage <br> (3) | Debt <br> (4) | Trade C. <br> (5) | FinCost <br> (6) | IntRate (7) |
| POST x CONFL | $\begin{aligned} & 0.148^{*} \\ & (0.061) \end{aligned}$ | $\begin{aligned} & -0.022 \\ & (0.060) \end{aligned}$ | $\begin{aligned} & -0.220^{* *} \\ & (0.056) \end{aligned}$ | $\begin{aligned} & 0.051 \\ & (0.070) \end{aligned}$ | $\begin{aligned} & \hline 0.178^{* *} \\ & (0.056) \end{aligned}$ | $\begin{aligned} & \hline 0.281^{* *} \\ & (0.059) \end{aligned}$ | $\begin{aligned} & 0.158^{* *} \\ & (0.058) \end{aligned}$ |
| Obs. | 7922 | 7922 | 7922 | 7922 | 7922 | 7922 | 7922 |
| $R^{2}$ within | 0.053 | 0.032 | 0.033 | 0.036 | 0.033 | 0.032 | 0.033 |
| Panel D: Gross operating surplus divided by sales |  |  |  |  |  |  |  |


| Input $\rightarrow$ | Staff <br> $(1)$ | Manager <br> $(2)$ | Avg wage <br> $(3)$ | Debt <br> $(4)$ | Trade C. <br> $(5)$ | FinCost <br> $(6)$ | IntRate <br> $(7)$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| POST x CONFL | $0.093^{* *}$ | -0.002 | $-0.022^{*}$ | $0.042^{* *}$ | $0.019^{*}$ | 0.006 | $0.021^{* *}$ |
|  | $(0.013)$ | $(0.010)$ | $(0.009)$ | $(0.011)$ | $(0.009)$ | $(0.009)$ | $(0.007)$ |
|  |  |  |  |  |  |  |  |
| Obs. | 11112 | 11112 | 1112 | 11112 | 11112 | 11112 | 11112 |
| $R^{2}$ within | 0.040 | 0.028 | 0.029 | 0.030 | 0.029 | 0.028 | 0.029 |
|  |  |  |  |  |  |  |  |
| Panel E: Return on assets |  |  |  |  |  |  |  |
| Input $\rightarrow$ | Staff | Manager | Avg wage | Debt | Trade C. | FinCost | IntRate |
|  | $(1)$ | $(2)$ | $(3)$ | $(4)$ | $(5)$ | $(6)$ | $(7)$ |
| POST x CONFL | $0.094^{* *}$ | -0.002 | -0.009 | $0.348^{* *}$ | $0.098^{* *}$ | 0.019 | $0.063+$ |
|  | $(0.031)$ | $(0.035)$ | $(0.032)$ | $(0.053)$ | $(0.031)$ | $(0.029)$ | $(0.034)$ |
|  |  |  |  |  |  |  |  |
| Obs. | 11044 | 11044 | 11044 | 11044 | 11044 | 11044 | 11044 |
| $R^{2}$ within | 0.012 | 0.011 | 0.011 | 0.024 | 0.012 | 0.011 | 0.011 |

The dependent variable is defined in the heading of each panel. $P O S T \times C O N F L$ is a variable equal to zero before the crisis (e.g., 2011) and equal to the number of deaths per 100,000 inhabitants in the district after 2011. Within estimator (firm fixed effect) is used. In each column, interactions with firms' characteristics are included (but unreported). Firm fixed effect as well as control interactions are included but unreported. Standard errors are clustered at the firm-level. $\dagger,{ }^{*}$ and ${ }^{* *}$ indicate significance at the $10 \%, 5 \%$ and $1 \%$ levels, respectively.

Table A12: Robustness checks (3) - Interest variable; crisis definition and placebo test

| Panel A: Dummy based on median |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Input $\rightarrow$ | Staff <br> (1) | Manager (2) | Avg wage <br> (3) | Debt <br> (4) | Trade C. <br> (5) | FinCost <br> (6) | IntRate <br> (7) |
| POST x CONFL | $\begin{aligned} & 0.535^{* *} \\ & (0.037) \end{aligned}$ | $\begin{aligned} & -0.166^{* *} \\ & (0.037) \end{aligned}$ | $\begin{aligned} & -0.334^{* *} \\ & (0.041) \end{aligned}$ | $\begin{aligned} & 0.131^{* *} \\ & (0.037) \end{aligned}$ | $\begin{aligned} & 0.089 * * \\ & (0.038) \end{aligned}$ | $\begin{aligned} & 0.0404 \\ & (0.045) \end{aligned}$ | $\begin{aligned} & 0.111^{*} \\ & (0.048) \end{aligned}$ |
| Obs. <br> $R^{2}$ within | $\begin{aligned} & 11178 \\ & 0.207 \end{aligned}$ | $\begin{aligned} & 11178 \\ & 0.181 \end{aligned}$ | $\begin{aligned} & 11178 \\ & 0.188 \end{aligned}$ | $\begin{aligned} & 11178 \\ & 0.179 \end{aligned}$ | $\begin{aligned} & 11178 \\ & 0.178 \end{aligned}$ | $\begin{aligned} & 11178 \\ & 0.178 \end{aligned}$ | $\begin{aligned} & 11178 \\ & 0.178 \end{aligned}$ |
| Panel B: Continuous measure for input |  |  |  |  |  |  |  |
| Input $\rightarrow$ | Staff <br> (1) | Manager (2) | Avg wage <br> (3) | Debt <br> (4) | Trade C. <br> (5) | FinCost <br> (6) | IntRate <br> (7) |
| POST x CONFL | $\begin{aligned} & 0.0566^{* *} \\ & (0.0078) \end{aligned}$ | $\begin{gathered} -0.0017^{*} \\ (0.0007) \end{gathered}$ | $\begin{aligned} & -0.0000^{* *} \\ & (0.0000) \end{aligned}$ | $\begin{aligned} & \hline 0.0008^{*} \\ & (0.0004) \end{aligned}$ | $\begin{aligned} & 0.0132^{* *} \\ & (0.0037) \end{aligned}$ | $\begin{aligned} & 0.0196 \\ & (0.0312) \end{aligned}$ | $\begin{aligned} & \hline 0.0127+ \\ & (0.0072) \end{aligned}$ |
| Obs. | 11178 | 11178 | 11178 | 11178 | 11178 | 11178 | 11178 |
| $R^{2}$ within | 0.223 | 0.178 | 0.194 | 0.178 | 0.180 | 0.178 | 0.178 |

Panel C: Including 2008 in pre-crisis period

| Input $\rightarrow$ | Staff <br> $(1)$ | Manager <br> $(2)$ | Avg wage <br> $(3)$ | Debt <br> $(4)$ | Trade C. <br> $(5)$ | FinCost <br> $(6)$ | IntRate <br> $(7)$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| POST x CONFL | $0.428^{* *}$ | $-0.135^{* *}$ | $-0.330^{* *}$ | $0.154^{* *}$ | $0.075^{*}$ | 0.004 | $0.075^{*}$ |
|  | $(0.034)$ | $(0.036)$ | $(0.035)$ | $(0.041)$ | $(0.033)$ | $(0.036)$ | $(0.037)$ |
|  |  |  |  |  |  |  |  |
| Obs. | 13530 | 13530 | 13530 | 13530 | 13530 | 13530 | 13530 |
| $R^{2}$ within | 0.171 | 0.158 | 0.169 | 0.155 | 0.153 | 0.153 | 0.150 |

Panel D: Including 2010 as a crisis year

| Input $\rightarrow$ | Staff <br> $(1)$ | Manager <br> $(2)$ | Avg wage <br> $(3)$ | Debt <br> $(4)$ | Trade C. <br> $(5)$ | FinCost <br> $(6)$ | IntRate <br> $(7)$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| POST x CONFL | $0.576^{* *}$ | $-0.160^{* *}$ | $-0.420^{* *}$ | $0.190^{* *}$ | $0.165^{* *}$ | -0.009 | $0.138^{* *}$ |
|  | $(0.042)$ | $(0.045)$ | $(0.040)$ | $(0.047)$ | $(0.041)$ | $(0.043)$ | $(0.043)$ |
| Obs. | 11178 | 11178 | 11178 | 11178 | 11178 | 11178 | 11178 |
| $R^{2}$ within | 0.303 | 0.280 | 0.291 | 0.281 | 0.281 | 0.279 | 0.280 |
|  |  |  |  |  |  |  |  |
| Panel E: Placebo test |  |  |  |  |  |  |  |
| Input $\rightarrow$ | Staff | Manager | Avg wage | Debt | Trade C. | FinCost | IntRate |
|  | $(1)$ | $(2)$ | $(3)$ | $(4)$ | $(5)$ | $(6)$ | $(7)$ |
| POST x CONFL | -0.041 | -0.014 | -0.010 | 0.077 | 0.052 | 0.047 | -0.029 |
|  | $(0.046)$ | $(0.051)$ | $(0.044)$ | $(0.052)$ | $(0.040)$ | $(0.041)$ | $(0.043)$ |
|  |  |  |  |  |  |  |  |
| Obs. | 5852 | 5852 | 5852 | 5852 | 5852 | 5852 | 5852 |
| $R^{2}$ within | 0.041 | 0.038 | 0.038 | 0.043 | 0.041 | 0.041 | 0.043 |

The dependent variable is the logarithm of labor productivity defined as value added per worker in all specifications. In Panels A and B, the measure of input dependence is modified (dummy based on median value in the industry in Panel A Panels A and B, the measure of input dependence is modified (dummy based on median value in the industry in Panel A and continuous measure in Panel B). In Panel C, the pre-crisis period is extended to 2008. In Panel D, 2010 is considered as a crisis year. In Panel E, a placebo test is implemented (see Section 5.3). POST $\times C O N F L$ is a variable equal to zero
before the crisis (e.g., 2011) and equal to the number of deaths per 100,000 inhabitants in the district after 2011 . Within before the crisis (e.g., 2011) and equal to the number of deaths per 100,000 inhabitants in the district after 2011 . Within
estimator (firm fixed effect) is used. In each column, interactions with firms' characteristics are included (but unreported). estimator (firm fixed effect) is used. In each column, interactions with firms' characteristics are included (but unreported).
Firm fixed effect as well as control interactions are included but unreported. Standard errors are clustered at the firm-level. $\dagger,{ }^{*}$ and ${ }^{* *}$ indicate significance at the $10 \%, 5 \%$ and $1 \%$ levels, respectively.

Table A13: Robustness checks (4) - Definition of credit exposure

| Panel A: ACLED data (Nb of deaths per 100,000 inhabitants) |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Input $\rightarrow$ | Staff <br> (1) | Manager <br> (2) | Avg wage <br> (3) | Debt <br> (4) | Trade C. <br> (5) | FinCost <br> (6) | IntRate <br> (7) |
| (POSTxCONF) x Input | $\begin{aligned} & 0.563^{* *} \\ & (0.044) \end{aligned}$ | $\begin{aligned} & -0.127^{* *} \\ & (0.045) \end{aligned}$ | $\begin{aligned} & -0.307^{* *} \\ & (0.044) \end{aligned}$ | $\begin{aligned} & 0.170^{* *} \\ & (0.054) \end{aligned}$ | $\begin{aligned} & 0.093^{*} \\ & (0.042) \end{aligned}$ | $\begin{aligned} & -0.006 \\ & (0.049) \end{aligned}$ | $\begin{aligned} & 0.123^{* *} \\ & (0.046) \end{aligned}$ |
| Obs. | 11178 | 11178 | 11178 | 11178 | 11178 | 11178 | 11178 |
| Panel B: Absolute number of deaths (NCI data) |  |  |  |  |  |  |  |
| Input $\rightarrow$ | Staff <br> (1) | Manager <br> (2) | Avg wage <br> (3) | Debt <br> (4) | Trade C. <br> (5) | FinCost <br> (6) | IntRate <br> (7) |
| POST x CONFL | 0.0566** | -0.0017* | -0.0000** | 0.0008* | 0.0132** | 0.0196 | 0.0127+ |
| (POSTxCONF) x Input | $\begin{aligned} & 0.448^{* *} \\ & (0.039) \end{aligned}$ | $\begin{aligned} & -0.095^{*} \\ & (0.039) \end{aligned}$ | $\begin{aligned} & -0.244^{* *} \\ & (0.038) \end{aligned}$ | $\begin{aligned} & 0.127^{* *} \\ & (0.044) \end{aligned}$ | $\begin{aligned} & 0.090^{*} \\ & (0.038) \end{aligned}$ | $\begin{aligned} & 0.003 \\ & (0.041) \end{aligned}$ | $\begin{aligned} & 0.100^{*} \\ & (0.039) \end{aligned}$ |
| Obs. | 11178 | 11178 | 11178 | 11178 | 11178 | 11178 | 11178 |

Panel C: Absolute number of deaths (ACLED data)

| Input $\rightarrow$ | Staff <br> $(1)$ | Manager <br> $(2)$ | Avg wage <br> $(3)$ | Debt <br> $(4)$ | Trade C. <br> $(5)$ | FinCost <br> $(6)$ | IntRate <br> $(7)$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| (POSTxCONF) x Input | $0.551^{* *}$ | $-0.124^{* *}$ | $-0.302^{* *}$ | $0.166^{* *}$ | $0.094^{*}$ | -0.004 | $0.121^{* *}$ |
|  | $(0.043)$ | $(0.044)$ | $(0.043)$ | $(0.053)$ | $(0.042)$ | $(0.048)$ | $(0.045)$ |
| Obs. | 11178 | 11178 | 11178 | 11178 | 11178 | 11178 | 11178 |

Panel D: Number of events (ACLED data)

| Input $\rightarrow$ | Staff |  |  |  |  |  |  |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | $(1)$ | Manager |  |  |  |  |  |
|  | $(2)$ | Avg wage <br> $(3)$ | Debt <br> $(4)$ | Trade C. <br> $(5)$ | FinCost <br> $(6)$ | IntRate <br> $(7)$ |  |
| (POSTxCONF) x Input | $0.468^{* *}$ | $-0.112^{* *}$ | $-0.258^{* *}$ | $0.131^{* *}$ | $0.100^{* *}$ | 0.013 | $0.117^{* *}$ |
|  | $(0.040)$ | $(0.040)$ | $(0.040)$ | $(0.046)$ | $(0.039)$ | $(0.042)$ | $(0.040)$ |
| Obs. | 11178 | 11178 | 11178 | 11178 | 11178 | 11178 | 11178 |

Panel E: Treatment dummy (NCI data)

| Input $\rightarrow$ |  | Staff <br> $(1)$ | Manager <br> $(2)$ | Avg wage <br> $(3)$ | Debt <br> $(4)$ | Trade C. <br> $(5)$ | FinCost <br> $(6)$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| (POSTxCONF) x Input | $0.583^{* *}$ | $-0.143^{* *}$ | $-0.348^{* *}$ | $0.165^{* *}$ | $0.107^{* *}$ | 0.012 | $0.157^{* *}$ |
|  | $(0.044)$ | $(0.045)$ | $(0.044)$ | $(0.055)$ | $(0.040)$ | $(0.049)$ | $(0.046)$ |
| Obs. | 11178 | 11178 | 11178 | 11178 | 11178 | 11178 | 11178 |

Panel F: Treatment dummy (ACLED data)

| Input $\rightarrow$ | Staff <br> $(1)$ | Manager <br> $(2)$ | Avg wage <br> $(3)$ | Debt <br> $(4)$ | Trade C. <br> $(5)$ | FinCost <br> $(6)$ | IntRate <br> $(7)$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| (POSTxCONF) x Input | $0.587^{* *}$ | $-0.153^{* *}$ | $-0.332^{* *}$ | $0.171^{* *}$ | $0.108^{* *}$ | 0.010 | $0.153^{*}$ |
|  | $(0.044)$ | $(0.046)$ | $(0.044)$ | $(0.056)$ | $(0.041)$ | $(0.049)$ | $(0.046)$ |
| Obs. | 11178 | 11178 | 11178 | 11178 | 11178 | 11178 | 11178 |

The dependent variable is the logarithm of labor productivity defined as value added per worker in all specifications. In Panel A, we employ ACLED data instead of NCI data. In Panels B and C, we exploit the absolute number of deaths per region (from NCI and ACLED respectively). In Panel D, we use the number of events (from ACLED). In Panels E and F, we create a dummy equals to one if a firm is located in region with at least one death (using NCI and ACLED data, respectively). Within dummy equals to one if a firm is located in region with at least one death (using NCI and ACLED data, respectively). Within
estimator (firm fixed effect) is used. In each column, interactions with firms' characteristics are included (but unreported). estimator (firm fixed effect) is used. In each column, interactions with firms' characteristics are included (but unreported).
Firm fixed effect as well as control interactions are included but unreported. Standard errors are clustered at the firm-level. $\dagger,{ }^{*}$ and ${ }^{* *}$ indicate significance at the $10 \%, 5 \%$ and $1 \%$ levels, respectively.

Table A14: Robustness checks (5) - Econometric issues

| Panel A: Value of labor productivity |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Input $\rightarrow$ | Staff <br> (1) | Manager (2) | Avg wage <br> (3) | Debt <br> (4) | Trade C. <br> (5) | FinCost <br> (6) | IntRate <br> (7) |
| POST x CONFL | $\begin{aligned} & 8088.8^{* *} \\ & (1823.7) \end{aligned}$ | $\begin{aligned} & \hline-4212.0^{* *} \\ & (1522.6) \end{aligned}$ | $\begin{aligned} & \hline-6420.1^{* *} \\ & (1638.0) \end{aligned}$ | $\begin{aligned} & \hline-867.0 \\ & (1173.7) \end{aligned}$ | $\begin{aligned} & \hline 862.9 \\ & (1302.1) \end{aligned}$ | $\begin{aligned} & \hline 2998.6^{*} \\ & (1433.7) \end{aligned}$ | $\begin{aligned} & \hline 224.4 \\ & (1398.3) \end{aligned}$ |
| Obs. <br> $R^{2}$ within | 11178 0.088 | 11178 0.085 | 11178 0.086 | 11178 0.083 | 11178 0.084 | 11178 0.084 | 11178 0.083 |

Panel B: Inverse hyperbolic sine transformation

| Input $\rightarrow$ | Staff <br> (1) | Manager (2) | Avg wage <br> (3) | Debt <br> (4) | Trade C. <br> (5) | FinCost <br> (6) | IntRate (7) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| POST x CONFL | $\begin{aligned} & 0.505^{* *} \\ & (0.039) \end{aligned}$ | $\begin{aligned} & -0.123^{* *} \\ & (0.040) \end{aligned}$ | $\begin{aligned} & -0.323^{* *} \\ & (0.037) \end{aligned}$ | $\begin{aligned} & 0.143^{* *} \\ & (0.045) \end{aligned}$ | $\begin{aligned} & 0.100^{* *} \\ & (0.038) \end{aligned}$ | $\begin{aligned} & 0.006 \\ & (0.042) \end{aligned}$ | $\begin{aligned} & 0.096^{*} \\ & (0.040) \end{aligned}$ |
| Obs. <br> $R^{2}$ within | 11178 0.088 | 11178 0.085 | 11178 0.086 | 11178 0.083 | 11178 0.084 | 11178 0.084 | 11178 0.083 |

Panel C: Inclusion of exiters

| Input $\rightarrow$ | Staff <br> (1) | Manager <br> (2) | Avg wage <br> (3) | Debt <br> (4) | Trade C. <br> (5) | FinCost (6) | IntRate (7) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| POST x CONFL | $\begin{aligned} & \hline 0.667^{* *} \\ & (0.075) \end{aligned}$ | $\begin{aligned} & \hline-0.083 \\ & (0.082) \end{aligned}$ | $\begin{aligned} & \hline-0.258^{* *} \\ & (0.076) \end{aligned}$ | $\begin{aligned} & \hline 0.229^{* *} \\ & (0.086) \end{aligned}$ | $\begin{aligned} & \hline 0.150^{*} \\ & (0.073) \end{aligned}$ | $\begin{aligned} & \hline 0.104 \\ & (0.083) \end{aligned}$ | $\begin{aligned} & \hline 0.085 \\ & (0.080) \end{aligned}$ |
| Obs. | 14749 | 14749 | 14749 | 14749 | 14749 | 14749 | 14749 |
| $R^{2}$ within | 0.090 | 0.083 | 0.084 | 0.083 | 0.083 | 0.083 | 0.083 |

Panel D: Sample selection

| Input $\rightarrow$ |  | Staff <br> $(1)$ | Manager <br> $(2)$ | Avg wage <br> $(3)$ | Debt <br> $(4)$ | Trade C. <br> $(5)$ | FinCost <br> $(6)$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| POST x CONFL | $0.461^{* *}$ | $-0.105^{* *}$ | $-0.297^{* *}$ | IntRate <br> $(7)$ |  |  |  |
|  | $(0.037)$ | $(0.039)$ | $(0.035)$ | $(0.044)$ | $(0.035)$ | $(0.038)$ | $(0.036)$ |
| $\hat{\lambda}$ | -0.073 | -0.085 | -0.052 | -0.094 | -0.113 | -0.096 | -0.086 |
|  | $(0.127)$ | $(0.129)$ | $(0.128)$ | $(0.129)$ | $(0.129)$ | $(0.129)$ | $(0.128)$ |
|  |  |  |  |  |  |  |  |
| Obs. | 10237 | 10237 | 10237 | 10237 | 10237 | 10237 | 10237 |
| $R^{2}$ within | 0.143 | 0.118 | 0.127 | 0.119 | 0.118 | 0.117 | 0.118 |

The dependent variable is the value of labor productivity defined as value added per worker in Panels A and C and the inverse hyperbolic sine in Panel B and D. In Panel A and B we add firms with a labor productivity which is null or negative. In Panel C and D, we consider exiters and set up their productivity to zero (or $\log (1)$ in Panel $C$ ) in the year of exit. In Panel E, a placebo test is implemented (see Section 5.3). POST $\times C O N F L$ is a variable equal to zero before the crisis (e.g., Panel E , a placebo test is implemented (see Section 5.3). POST $\times C O N F L$ is a variable equal to zero before the crisis (e.g.,
2011) and equal to the number of deaths per 100,000 inhabitants in the district after 2011. Within estimator (firm fixed $2011)$ and equal to the number of deaths per 100,000 inhabitants in the district after 2011. Within estimator (firm fixed
effect) is used. In each column, interactions with firms' characteristics are included (but unreported). Firm fixed effect as effect) is used. In each column, interactions with firms' characteristics are included (but unreported). Firm fixed effect as
well as control interactions are included but unreported. Standard errors are clustered at the firm level, except in Panel F well as control interactions are included but unreported. Standard errors are clustered at the firm level, except in Panel F
(bootstrapping with 500 replications is used). $\dagger, *$ and $* *$ indicate significance at the $10 \%$, $5 \%$ and $1 \%$ levels, respectively.


Figure A2: Conflict intensity per district
Panel (a): National Commission of Investigation


Figure displayed the absolute number of deaths (in left) and the relative number of deaths per 100,000 inhabitants (in right) per district in Cote d'Ivoire. Panel a) displays data from the NCI and Panel b) data from ACLED.

Figure A3: Testing parallel trends (base year $=2006$ )


The figure plots coefficients estimates and confidence intervals of the following model: $\log (P)_{i j k t}=\alpha_{i}+\mu_{t}+\beta_{t}\left(\mu_{t} \times T R E A T E D_{k}\right)+\varepsilon_{i j k t}$ where TREATED ${ }_{k}$ equals to one if a firm is located in a treated region $k$. A threshold of 10 deaths per 100,000 inhabitants is employed to discriminate between low intensity and high intensity. Base year is 2006. Dash line separates pre- and post-treatment periods.

Figure A4: Testing parallel trends (base year $=2010$ )


The figure plots coefficients estimates and confidence intervals of the following model: $\log (P)_{i j k t}=\alpha_{i}+\mu_{t}+\beta_{t}\left(\mu_{t} \times T R E A T E D_{k}\right)+\varepsilon_{i j k t}$ where $T R E A T E D_{k}$ equals to one if a firm is located in a treated region $k$. A threshold of 10 deaths per 100,000 inhabitants is employed to discriminate between low intensity and high intensity. Base year is 2010. Dash line separates pre- and post-treatment periods.

## B. Identifying false similar firms

To detect any possible irregularities, we consider six criteria: city, year of creation, sector, legal status, ownership structure and the time lag between two observations (inferior to two years). If two observations differ in at least four of the six criteria, we consider that the observations are indeed two different firms.

Let's consider the following firms (10001, 10002, 10003, and 10004) whose characteristics are shown in Table B1.

The first firm (id=10001) is a common observation in the dataset. In spite of a change in the ownership structure, we do not observe other changes that allow us to consider that the firm identified in 2010 is different from the firm operating in the following year.

The second identifier seems undoubtedly to refer to more than one different firms. We lack information in 2010 and 2011 and all characteristics have changed between 2009 and 2012. In our classification, we consider these to be two separate firms because more than 4 criteria have changed and we create a new identifier (20002) for the observations after 2012

The more complex case covers the last two situations (id=10003; id=10004). Between 2011 and 2012, many characteristics of firm 10003 changed. However, we consider that the firm referred to is the same because only three criteria of six are different (year of incorporation, ownership, and sector). For the same reason, we consider the observations of firm 10004 recover two different entities because four criteria have changed (year between two observations, year of incorporation, ownership structure and industry).

Table B1: Example of firms with a similar identifier

| id | year | year incorp. | city | ownership | legal | industry | final id |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| 10001 | 2009 | 2005 | Abidjan | foreign | Other | Trade | 10001 |
| 10001 | 2010 | 2005 | Abidjan | foreign | Other | Trade | 10001 |
| 10001 | 2011 | 2005 | Abidjan | local | Other | Trade | 10001 |
|  |  |  |  |  |  |  |  |
| 10002 | 2009 | 1995 | Bouake | local | Public company | Manufacturing | 10002 |
| 10002 | 2012 | 2011 | Abidjan | foreign | Limited L. | Construction | 20002 |
| 10002 | 2013 | 2011 | Abidjan | foreign | Limited L. | Construction | 20002 |
|  |  |  |  |  |  |  |  |
| 10003 | 2010 | 2008 | Abidjan | foreign | Limited L. | Manufacturing | 10003 |
| 10003 | 2011 | 2008 | Abidjan | foreign | Limited L. | Manufacturing | 10003 |
| 10003 | 2012 | 2011 | Abidjan | local | Limited L. | Construction | 10003 |
|  |  |  |  |  |  |  |  |
| 10004 | 2008 | 1998 | Abidjan | local | Limited L. | Manufacturing | 10004 |
| 10004 | 2011 | 2003 | Abidjan | foreign | Limited L. | Services | 20004 |
| 10004 | 2012 | 2003 | Abidjan | foreign | Limited L. | Services | 20004 |

## C. Estimation of the TFP

Suppose the production function is a Cobb-Douglas function in capital $K_{i t}$ and labor $L_{i t}$, the total factor productivity (TFP henceforth) can be estimated using the log transformation:

$$
\begin{equation*}
y_{i t}=\beta_{k} k_{i t}+\beta_{l} l_{i t}+\mu_{i t} \quad, \text { with } \quad \mu_{i t}=\Omega_{i t}+\eta_{i t} \tag{C1}
\end{equation*}
$$

with $y_{i t}$ representing the logarithm of the firm's output $i$ in period $t$, nd $l_{i t}$ and $k_{i t}$, respectively constitute the logarithm of labor and capital. The residual component is a mix of the productivity shock observed only by the firm affecting decision-making $\left(\Omega_{i t}\right)$ and the unexpected productivity shock that is by definition not observed by the firm $\left(\eta_{i t}\right)$. In this framework, we can estimate the TFP term if $\beta_{k}$ and $\beta_{l}$ are known.

Estimation of TFP with traditional methods raises several methodological problems (simultaneity and endogeneity problems) because the level of productivity and inputs are likely to be correlated (Olley and Pakes, 1996; Levinsohn and Petrin, 2003). Thus, the estimation by OLS poses a problem of simultaneity. In addition, the use of a balanced panel does not consider inputs and outputs, leading to selection bias, which results from the relationship between productivity shocks and the probability of bankruptcy or business interruption. In addition, these methodological challenges may be accentuated by the fact that the company's product choices may be related to their underlying productivity (Bernard et al., 2009). Also, most of the other traditional estimators (fixed effects, instrumental variables and generalized method of moments) used to overcome these endogeneity problems have not proved satisfactory in the case of production functions, particularly because of their underlying assumptions.

Faced with these methodological questions, several estimators (parametric and semiparametric) have emerged. Among the semi-parametric estimators, Olley and Pakes (1996) (OP) and Levinsohn and Petrin (2003) (LP) propose a semi-parametric estimator that considers simultaneity biases (and selection biases in the case of the OLS estimator). Indeed, Olley and Pakes (1996) are the first authors to propose an estimation method that explicitly considers the problem of selection and simultaneity by using a dynamic model that considers firm behavior and idiosyncratic productivity shocks. They propose a semi-parametric estimator that solves the simultaneity problem by using the company's
investment decision to replace unobserved productivity shocks. Under Levinsohn and Petrin (2003), the invertibility condition is likely to be invalidated in the presence of imperfect competition in the production markets, whereas it has no effect on the monotonicity condition under the OL method. We use the method of Olley and Pakes (1996) to estimate the overall factor productivity of the firms in our sample. Unfortunately, we cannot use the LP method because we do not have data on intermediate consumption and because of the methodological problems mentioned above.

We briefly describe the OP method used in this paper. Olley and Pakes (1996) assume that firms decide at the beginning of each period whether to continue or to stop production. If a firm decides to stop participating in the market, then it will receive a liquidation value equal to $\phi$. On the other hand, if the company chooses to remain in the market by continuing to produce, it will use its factors of production (labor, capital, etc.) and set its level of investment $I_{i t}$. Thus, the firm's results are conditioned by its stated variables at the beginning of the period, namely the capital stock $K_{i t}$, the level of productivity $\phi_{i t}$ and the age of the company $a_{i t}$. This model assumes that expected productivity is defined as a function of current productivity and capital, i. e., : $E\left[\Omega_{(i, t+1)} \mid \Omega_{i t}, K_{i t}\right]$ and the company's result depends on $\Omega_{i t}$ and $K_{i t}$.

This assumes that a firm will cease trading provided that its liquidation value $\phi$ is higher than its expected future returns. In other words, there is a threshold level of productivity $\left(\underline{\Omega_{i t}}\right)$ under which a firm decides to leave the market.

The semi-parametric estimation method proposed by Olley and Pakes (1996) allows for simultaneity and selection biases to be considered, unlike traditional methods. Its application involves using the investment decision function to control the correlation between the error term and the factors of production. This is based on the following underlying assumption: future productivity is strictly increasing ( $\Omega_{i t}$ follows a first-order Markov process) and firms that experience positive productivity shocks will invest more during this period, for any level of capital. The investment choice of the firm $I_{i t}$ also depends on productivity $\left(\Omega_{i t}\right)$, capital $\left(K_{i t}\right)$ and the age of the firm $\left(a_{i t}\right)$. Assuming positive investment, then the inverse function of the productivity shock is:

$$
\begin{equation*}
\Omega_{i t}=I^{-1}\left(I_{i t}, K_{i t}, a_{i t}\right)=h\left(I_{i t}, K_{i t}, a_{i t}\right) \quad, \text { with } \partial \Omega_{i t} / \partial I_{i t}>0 \tag{C2}
\end{equation*}
$$

The advantage of this function is control of the simultaneity bias. By substitution C2 in C1 we get :

$$
\begin{equation*}
y_{i t}=\beta_{l} l_{i t}+\phi\left(i_{i t}, k_{i t}\right)+\eta_{i t} \tag{C3}
\end{equation*}
$$

With $\phi\left(i_{i} t, k_{i t}\right)=\beta_{0}+\beta_{k} k_{i t}+h\left(i_{i t}, k_{i t}\right)$ and $\phi($.$) is approximated by the second-order$ polynomial series in capital and investment. We estimate Eq. C3 by OLS. The estimated coefficients of the variable production factor (labor) are therefore unbiased because $\phi$ (.) makes it possible to control unobserved productivity. As a result, the error term is no longer correlated with the factors of production. However, Eq. C3 does not identify $\beta_{k}$.

To control for selection bias, an estimate of survival probabilities is made. We know that the probability of a firm's survival at period $t$ therefore depends on productivity, age, and capital at t-1 (as well as to their squares and cross-products). Therefore, in our implementation, we estimate the probability of survival by fitting a probit model.

We use the method of Olley and Pakes (1996) using the method introduced by Yasar et al. (2008). This approach uses a bootstrap technique to group variables by treating all observations of an individual firm as a (sub)group.

The results obtained using Olley and Pakes (1996) and the OLS method are presented in Table C1.

Table C1: Production function parameters: OP and OLS estimations

| Variables | Olley and Pakes | OLS |
| :--- | ---: | ---: |
| Labor | $0.610^{* * *}$ | $0.630^{* * *}$ |
|  | $(0.130)$ | $(0.009)$ |
| Capital | $0.419^{* * *}$ | $0.338^{* * *}$ |
|  | $(0.428)$ | $(0.005)$ |
| Age | $0.012^{* * *}$ | $0.013^{* * *}$ |
|  | $(0.001)$ | $(0.001)$ |
| Trend | $-0.038^{* * *}$ | $-0.033^{* * *}$ |
|  | $(0.007)$ | $(0.006)$ |

Standard errors in parentheses.
In the OP model SEs are bootstrapped (250 rep)
*** Significant at the $1 \%$ level.

## D. Accounting for sample selection

In a first step, for each year we estimate a selection equation using a standard probit as follows:

$$
\begin{equation*}
\operatorname{Pr}\left(s_{i}=1\right)=\Phi\left(\delta X_{i j\left(t_{0}\right)}+\mu C_{i j\left(t_{0}\right)}\right) \quad(\forall t=0, \ldots, T) \tag{D1}
\end{equation*}
$$

where $s_{i}$ is a dummy equal to 1 if a firm survived in year $t$ and 0 if not. $X_{i j\left(t_{0}\right)}$ and $C_{i j\left(t_{0}\right)}$ are variables included in the baseline model (input usage and firm characteristics). Ideally, we should include a selection variable that affects only the selection process (i.e., exit) but not the outcome (performance of survivors). However, we fail to find a relevant selection variable in our case. Results using the first step as displayed below in Table D1. It should be noted that input usage does not tend to alter the likelihood probability to exit, at the weak exception of Debt variable (which reduced exit probability but before, during and after the crisis). Among other factors, only the age is a strong determinants of exit both before, during and after the crisis. Enterprises operating under the status of limited liabilities are more likely to exit, except in the year of the crisis. The role of size is stronger before the crisis.

In a second step, we compute the inverse of the Mills ratio for each firm $i$ for each year $t$ as follows:

$$
\begin{equation*}
\hat{\lambda}_{i}=\frac{\phi\left(\hat{\delta} X_{i j\left(t_{0}\right)}+\hat{\mu} C_{i j\left(t_{0}\right)}\right)}{\Phi\left(\hat{\delta} X_{i j\left(t_{0}\right)}+\hat{\mu} C_{i j\left(t_{0}\right)}\right)} \quad(\forall t=0, \ldots, T) \tag{D2}
\end{equation*}
$$

where $\Phi($.$) is the cumulative normal distribution function and \phi($.$) the normal density$ function.

Insofar as $\hat{\lambda}_{i}$ is computed for each period by running a probit model by period, we use a time-variant measure of the inverse of the Mills ratio $\left(\hat{\lambda}_{i t}\right)$ allowing us to include firm fixed effects as well as our crisis and post-crisis dummies. In a third step, we re-estimate the baseline model (Eq. 1) by adding the estimated inverse Mills ratio as covariates:

$$
\begin{equation*}
\log (P)_{i j k t}=\alpha_{i}+\mu_{t}+\beta_{1}\left(P O S T_{t} \times C O N F_{k}\right)+\beta_{2}\left(P O S T_{t} \times C O N F_{k}\right) \times X_{i j\left(t_{0}\right)}+\gamma \hat{\lambda}_{i t}+\varepsilon_{i j k t} \tag{D3}
\end{equation*}
$$

According to Wooldridge (1995), a simple test to detect sample selection is based on statistical significance of the inverse of the Mills ratio. Under the null hypothesis (absence
of bias) the coefficient is statistically equal to 0 . If not, we need to correct for sample selection bias. In this case, we cannot use standard errors because $\hat{\lambda}_{i t}$ is a generated variable. A simple way to get robust standard errors is by applying the bootstrapping method (Brownstone and Valletta, 2001).

Table D1: Determinants of exit per year

| Variables | Exit in |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | 2010 | 2011 | 2012 | 2013 | 2014 |
| Usual firms' characteristics |  |  |  |  |  |
| Log(EMPL) | $\begin{aligned} & 0.120^{* *} \\ & (0.037) \end{aligned}$ | $\begin{aligned} & 0.101^{*} \\ & (0.053) \end{aligned}$ | $\begin{aligned} & 0.089 \\ & (0.052) \end{aligned}$ | $\begin{aligned} & 0.032 \\ & (0.053) \end{aligned}$ | $\begin{aligned} & -0.018 \\ & (0.045) \end{aligned}$ |
| Log(Sales) | $\begin{aligned} & 0.023 \\ & (0.025) \end{aligned}$ | $\begin{aligned} & -0.044 \\ & (0.036) \end{aligned}$ | $\begin{aligned} & 0.043 \\ & (0.035) \end{aligned}$ | $\begin{aligned} & 0.081^{*} \\ & (0.039) \end{aligned}$ | $\begin{aligned} & 0.102^{* *} \\ & (0.035) \end{aligned}$ |
| Log(Age) | $\begin{aligned} & 0.186^{* *} \\ & (0.026) \end{aligned}$ | $\begin{aligned} & 0.264^{* *} \\ & (0.048) \end{aligned}$ | $\begin{aligned} & 0.143^{* *} \\ & (0.050) \end{aligned}$ | $\begin{aligned} & 0.240^{* *} \\ & (0.057) \end{aligned}$ | $\begin{aligned} & 0.090 \\ & (0.058) \end{aligned}$ |
| Abidjan | $\begin{aligned} & -0.034 \\ & (0.088) \end{aligned}$ | $\begin{aligned} & -0.121 \\ & (0.132) \end{aligned}$ | $\begin{aligned} & -0.142 \\ & (0.136) \end{aligned}$ | $\begin{aligned} & -0.140 \\ & (0.130) \end{aligned}$ | $\begin{aligned} & 0.057 \\ & (0.117) \end{aligned}$ |
| Foreign | $\begin{gathered} -0.131^{*} \\ (0.051) \end{gathered}$ | $\begin{aligned} & -0.085 \\ & (0.088) \end{aligned}$ | $\begin{aligned} & -0.055 \\ & (0.082) \end{aligned}$ | $\begin{aligned} & 0.088 \\ & (0.088) \end{aligned}$ | $\begin{aligned} & 0.113 \\ & (0.081) \end{aligned}$ |
| LimLiabilities | $\begin{aligned} & 0.152^{* *} \\ & (0.053) \end{aligned}$ | $\begin{aligned} & 0.220^{* *} \\ & (0.094) \end{aligned}$ | $\begin{aligned} & 0.038 \\ & (0.078) \end{aligned}$ | $\begin{aligned} & 0.147 \\ & (0.077) \end{aligned}$ | $\begin{aligned} & 0.243^{* *} \\ & (0.075) \end{aligned}$ |
| PublicCompany | $\begin{aligned} & -0.091 \\ & (0.093) \end{aligned}$ | $\begin{aligned} & 0.191 \\ & (0.140) \end{aligned}$ | $\begin{aligned} & -0.069 \\ & (0.131) \end{aligned}$ | $\begin{aligned} & -0.093 \\ & (0.132) \end{aligned}$ | $\begin{aligned} & 0.214 \\ & (0.125) \end{aligned}$ |
| Input mix |  |  |  |  |  |
| StaffCost | $\begin{aligned} & -0.232^{* *} \\ & (0.087) \end{aligned}$ | $\begin{aligned} & -0.177 \\ & (0.123) \end{aligned}$ | $\begin{aligned} & -0.144 \\ & (0.111) \end{aligned}$ | $\begin{aligned} & 0.152 \\ & (0.142) \end{aligned}$ | $\begin{aligned} & 0.115 \\ & (0.135) \end{aligned}$ |
| Managers | $\begin{aligned} & 0.025 \\ & (0.025) \end{aligned}$ | $\begin{aligned} & 0.008 \\ & (0.137) \end{aligned}$ | $\begin{aligned} & -0.159 \\ & (0.140) \end{aligned}$ | $\begin{aligned} & 0.027 \\ & (0.123) \end{aligned}$ | $\begin{aligned} & -0.012 \\ & (0.121) \end{aligned}$ |
| AverWage | $\begin{aligned} & 0.000 \\ & (0.000) \end{aligned}$ | $\begin{aligned} & 0.000 \\ & (0.000) \end{aligned}$ | $\begin{aligned} & 0.000 \\ & (0.000) \end{aligned}$ | $\begin{aligned} & 0.000 \\ & (0.000) \end{aligned}$ | $\begin{aligned} & -0.000 \\ & (0.000) \end{aligned}$ |
| Debt | $\begin{aligned} & -0.014 \\ & (0.010) \end{aligned}$ | $\begin{gathered} -0.031^{*} \\ (0.012) \end{gathered}$ | $\begin{aligned} & -0.024^{* *} \\ & (0.009) \end{aligned}$ | $\begin{aligned} & -0.021^{*} \\ & (0.010) \end{aligned}$ | $\begin{aligned} & -0.014 \\ & (0.009) \end{aligned}$ |
| TradeCredit | $\begin{aligned} & -0.223 \\ & (0.137) \end{aligned}$ | $\begin{aligned} & -0.420 \\ & (0.343) \end{aligned}$ | $\begin{aligned} & 0.169 \\ & (0.425) \end{aligned}$ | $\begin{aligned} & -0.458^{* *} \\ & (0.196) \end{aligned}$ | $\begin{aligned} & -0.321 \\ & (0.333) \end{aligned}$ |
| FinCost | $\begin{aligned} & 10.70 \\ & (10.47) \end{aligned}$ | $\begin{aligned} & -0.571 \\ & (10.89) \end{aligned}$ | $\begin{aligned} & -1.143 \\ & (1.685) \end{aligned}$ | $\begin{aligned} & 3.441 \\ & (2.333) \end{aligned}$ | $\begin{aligned} & -2.206 \\ & (1.641) \end{aligned}$ |
| IntRate | $\begin{aligned} & 0.441 \\ & (0.347) \end{aligned}$ | $\begin{aligned} & 0.211 \\ & (0.783) \end{aligned}$ | $\begin{aligned} & 0.598 \\ & (0.945) \end{aligned}$ | $\begin{aligned} & -0.821 \\ & (0.596) \end{aligned}$ | $\begin{aligned} & 0.180 \\ & (0.686) \end{aligned}$ |
| Sector dummies | Yes | Yes | Yes | Yes | Yes |
| Obs. | 4385 | 3941 | 3184 | 2987 | 2622 |
| $\text { pseudo-R }{ }^{2}$ | 0.05 | 0.07 | 0.05 | 0.06 | 0.04 |

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